BMJ Open Trajectories of body mass index in adulthood and all-cause and cause-specific mortality in the Melbourne Collaborative Cohort Study

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ABSTRACT

Objective Limited research has assessed the association between patterns of body mass index (BMI) change across adulthood and mortality. We aimed to identify groups of individuals who followed specific group-based BMI trajectories across adulthood, using weight collected on three occasions and recalled data from early adulthood, and to examine associations with all-cause and cause-specific mortality.

Design Prospective cohort study.

Setting Melbourne, Australia.

Participants Adults (n=29 881) enrolled in the Melbourne Collaborative Cohort Study, who were aged from 40 to 70 years between 1990 and 1994, and had BMI data for at least three time points.

Outcome Deaths from any cause before 31 March 2017 and deaths from obesity-related cancers, cardiovascular diseases (CVDs) and other causes before 31 December 2013.

Results We identified six group-based BMI trajectories: lower-normal stable (TR1), higher-normal stable (TR2), normal to overweight (TR3), chronic borderline obesity (TR4), normal to class I obesity (TR5) and overweight to class II obesity (TR6). Generally, compared with maintaining lower-normal BMI throughout adulthood, the lowest mortality was experienced by participants who maintained higher-normal BMI (HR 0.90; 95% CI 0.84 to 0.97); obesity during midlife was associated with higher all-cause mortality even when BMI was normal in early adulthood (HR 1.09; 95% CI 0.98 to 1.21) and prolonged borderline obesity from early adulthood was also associated with elevated mortality (HR 1.16; 95% CI 1.01 to 1.33). These associations were stronger for never-smokers and for death due to obesity-related cancers. Being overweight in early adulthood and becoming class II obese was associated with higher CVD mortality relative to maintaining lower-normal BMI (HR 2.27; 95% CI 1.34 to 3.87).

Conclusion Our findings highlight the importance of weight management throughout adulthood to reduce mortality.

INTRODUCTION

Obesity is associated with increased risk of mortality,^{1–5} cardiovascular disease (CVD)

Strengths and limitations of this study

- Multiple exposure assessments were used to derive trajectories in body mass index (BMI) across adulthood.
- The analytic approach identified specific higher-risk BMI trajectories that would not have been identifiable using the conventional approach of grouping participants into categories of change.
- We did not have information on weight between age 18 and 21 years and baseline, which limited us from identifying trajectories of how people gained weight between the two time points.

and several cancers.^{6 7} Most of this evidence is based on anthropometric data collected at one time, typically in participants' mid to late life at recruitment to a cohort. The importance of considering body mass as a time-varying exposure to account for changes over the lifespan has been highlighted.⁸ Some studies have assessed changes in body mass using predefined categories (eg, weight gain, weight stable, weight loss).^{9–14} This approach can introduce statistical bias by creating groups that reflect predominantly random variation and can fail to identify more complex patterns of change.¹⁵ One study used pictograms to assess body shape throughout life to create body shape trajectories.¹⁶ Stable lean body shape from age 5 to 50 years had the lowest mortality.¹⁶ Two studies using body mass index (BMI) found the lowest mortality for individuals consistently overweight during their mid to late life; the highest mortality was for obese individuals who gained weight¹⁷ and for thin individuals who lost weight.¹⁸ As these studies only used data from mid-adulthood onwards, they could not assess the effects of long-term obesity versus short-term obesity and its potential impact on mortality.

In this study, we used data from the Melbourne Collaborative Cohort Study (MCCS). We have previously reported on mortality in relation to baseline BMI and change in BMI over approximately 12 years.^{12 19} Here, we aimed to identify groups of individuals who followed specific BMI trajectories across adulthood, using weight collected on three occasions and recalled data from early adulthood, and to examine associations with all-cause and cause-specific mortality.

METHODS

Study population

The MCCS is a prospective cohort study of 41 513 residents of Melbourne Australia with a mean age of 55 years at recruitment (99% aged between 40 and 69 years) from 1990 to 1994 (baseline). Southern European migrants were deliberately recruited to widen the range of social and lifestyle exposures. At baseline, data were obtained by physical measurements, a self-administered food frequency questionnaire²⁰ and an interviewer-administered questionnaire on demographics, lifestyle and medical conditions. Further waves of data collection occurred from 1995 to 1998 (88% responded by mail or phone) and 2003–2007 (68% attended a study centre). Details about the cohort have previously been described.²¹

Participant and public involvement

Participants had no involvement in the design of the study. Results are disseminated to participants in annual newsletters.²¹

BMI assessment

Weight was recorded for four times (age 18–21 years, years 1990–1994, 1995–1998 and 2003–2007). At baseline, weight was measured following standard protocols²² and participants reported their weight at age 18–21 years. Weight was self-reported in 1995–1998 and measured in 2003–2007. To adjust for measurement error arising from using self-reported weight in 1995–1998, a correction was applied to data based on sex-specific and age-specific comparison of self-reported and measured weight in the Australian National Nutrition Survey.^{23 24} Recalled weight at age 18–21 years was not adjusted, because the correction formula might not be applicable to recalled weight. Height was measured at baseline. BMI was calculated as weight (kg) divided by square of height (m²).

Ascertainment of deaths

Deaths were ascertained through record linkage to the Victorian Registry of Births, Deaths and Marriages, and the National Death Index at the Australian Institute of Health and Welfare. Underlying cause of death was defined according to the International Classification of Diseases codes versions (ICD-10). When the cause was unknown, it was classified as 'other'. Obesity-related cancer deaths were defined as those for which the International Agency for Research on Cancer classified the

evidence for causality as sufficient: oesophagus adenocarcinoma, gastric cardia, colon and rectum, liver, gall-bladder, pancreas, postmenopausal breast, corpus uteri, ovary, renal cell, meningioma, thyroid, multiple myeloma; or as limited: male breast cancer, fatal prostate cancer, diffuse large B-cell lymphoma.⁶

Participants selected for this analysis

We excluded participants if they were aged <40 or >70 years at baseline, were underweight (BMI <18.5 kg/m²) at baseline, were diagnosed with cancer before their last BMI measurement, had missing data for any confounding variables, had extreme BMI values at any time (BMI <15 kg/m² or >45 kg/m²) or if they had BMI for fewer than three times to allow estimation of potentially non-linear trajectories. These exclusions left 29 881 participants (online supplementary figure 1).

Statistical analysis

BMI trajectories

We used a semiparametric group-based trajectory model^{15 25} to identify latent homogeneity in participants' BMI trajectories from early (age 18–21 years) to late adulthood. The model is an application of finite mixture modelling which assumes the population is composed of a mixture of groups following homogenous developmental courses.¹⁵ Longitudinal BMI data were fitted as a mixture of several latent trajectories in a censored normal model (allowing for the lower and upper BMI limits) with a polynomial function of age.²⁵ We used the Bayesian information criterion (BIC) and the log Bayes factor to select optimal shapes and number of trajectory groups through a two-stage approach.¹⁵ First, the number of groups was determined assuming all trajectory groups were quadratic functions of age. Second, the preferred order of the polynomial (ie, linear or quadratic) for each trajectory was determined. Participants were assigned to the group for which their posterior predicted probability calculated from the final model was highest. The adequacy of the final model was evaluated using recommended diagnostic measures: average posterior probability of assignment for each group of 0.7 or higher; odds of correct classification of 5.0 or higher; the proportion of a sample assigned to a certain group close to the proportion estimated from the model and a reasonably narrow confidence interval (CI) around each trajectory.¹⁵

Association of BMI trajectories with mortality

Cox regression with age as the time scale²⁶ was used to estimate HRs and 95% CIs for mortality for each trajectory group. Follow-up began at the date of last BMI assessment and ended at date of death, date when mortality data were complete (31 March 2017 for all-cause mortality and 31 December 2013 for cause-specific mortality), whichever came first. Models included birth cohort (<1930, 1930–1939, 1940–1949, ≥1950), height (continuous in cm), sex (male, female), country of birth (Australia/New Zealand/Northern Europe, Southern Europe), an

Figure 1 Body mass index (BMI) trajectories in the Melbourne Collaborative Cohort Study.

area-based measure of socioeconomic position (Socioeconomic Indexes For Areas (SEIFA)) derived from the Australian Census²⁷ (quintiles), highest level of education attained (primary school or lower, some high or technical school, completed high or technical school, completed tertiary education), smoking status at last BMI assessment (never, former, current) and these baseline covariates: alcohol drinking status (lifetime abstainer, former, current lower (1–39 g/day for male; 1–19 g/day for female), current higher (≥40 g/day for male; ≥20 g/day for female)), physical activity score (0, >0 and <4, ≥4 and <6, ≥6; computed from frequency of walking, vigorous and less vigorous exercise over the last 6 months²⁸) and Mediterranean Diet Score (MDS; 0–3, 4–6, 7–9; detailed calculation is described elsewhere²⁹).

Graphical checks based on log–log plots of survival and tests based on Schoenfeld residuals were used to examine the proportional hazards assumption. We stratified the Cox models on covariates that did not satisfy the proportional hazards assumption and fitted interactions between attained age and BMI trajectory for each trajectory that did not meet the assumption. The analyses were repeated after restriction to never smokers at last BMI assessment, because adjustment for smoking status may not be sufficient to control for confounding by smoking.^{5 30 31}

We tested interactions between BMI trajectories and these factors using likelihood ratio tests: sex (male, female), country of birth (Australian/New Zealand/northern Europe, southern Europe), alcohol consumption (lifetime abstainer, former, current lower, current higher), history of diabetes (yes, no), history of heart diseases (yes, no), follow-up duration (<5 years, 5 to <10 years, ≥10 years) and baseline age (<60 years, ≥60 years).

We repeated the analysis excluding (1) the first 2 years of follow-up and (2) non-respondents for the 2003–2007 data collection. Finally, we compared the fit of the models with BMI trajectories and models with last BMI assessment to assess whether trajectory models had better fit

than models using the conventional approach based on one BMI assessment. Last BMI measurements were categorised into six groups: underweight (BMI <18.5 kg/m²), lower normal (BMI ≥18.5 kg/m² and <23 kg/m²), higher normal (BMI ≥23 kg/m² and <25 kg/m²), overweight (BMI ≥25 kg/m² and <30 kg/m²), class I obesity (BMI ≥30 kg/m² and <35 kg/m²) and class II obesity (BMI ≥35 kg/m²).

Most participants excluded due to insufficient BMI records had baseline and recalled weight at age 18–21 years, but no BMI data in 1995–1998 and 2003–2007. To see whether selection of participants into the analysis was related to both BMI and death, we compared mean BMI (at age 18–21 years and at baseline) of participants who had three or more records with those who had fewer than three records; we also used unadjusted logistic regression to estimate the odds ratio (OR) of death for these two groups.

Statistical analyses were performed using Stata V.14.2. (Stata Corp., CollegeStation, Texas, USA).

RESULTS

The analyses included 29 881 participants (17 475 never-smokers at the latest BMI assessment) (online supplementary figure 1). Six trajectories (TR) of BMI were identified and named according to their shapes in relation to the WHO classification of BMI³² (figure 1). Almost all participants (94%) had normal BMI at age 18–21 years (TR1, TR2, TR3 and TR5). Two-thirds had an increase in BMI to higher normal (TR2) or overweight (TR3) later in life. A smaller proportion maintained a lower-normal BMI (TR1, 19%) throughout adulthood. Few participants (13%) became obese in midlife (TR5 and TR6) or were consistently borderline obese (TR4; 3%). Model adequacy diagnostics are presented in the online supplementary table 1.

The proportion of participants in the highest SEIFA quintile (ie, most advantaged) and with the highest education decreased across TR1–TR6 (table 1). Compared with other trajectory groups, participants in the lower-normal stable trajectory (TR1) tended to follow a more Mediterranean diet, be more physically active and to drink alcohol. MDS, physical activity level and proportion of current alcohol drinkers gradually decreased across TR1–TR6. A higher proportion of people in the borderline obesity trajectory (TR4) had prevalent heart disease and diabetes at baseline.

The table 2 presents HR between trajectories and mortality using stable lower-normal BMI trajectory (TR1) as the reference. All-cause mortality was slightly lower for the stable higher-normal BMI trajectory (TR2: HR 0.90, 95% CI 0.84 to 0.97); trajectories of chronic borderline obesity (TR4) and normal to class I obesity (TR5) had higher all-cause mortality (HR 1.16, 95% CI 1.01 to 1.33, and HR 1.09, 95% CI 0.98 to 1.21, respectively). As the most obese trajectory (TR6) showed evidence of non-proportional hazards, we fitted an interaction between

Table 1 Characteristics of participants by BMI trajectory groups in the Melbourne Collaborative Cohort Study

	BMI trajectory groups					
	TR1 lower-normal stable	TR2 higher-normal stable	TR3 normal to overweight	TR4 chronic borderline obesity	TR5 normal to class I obesity	TR6 overweight to class II obesity
	N=5748	N=11 229	N=8130	N=933	N=2893	N=948
Baseline age, mean (SD)	54.3 (8.9)	55.3 (8.7)	54.7 (8.3)	55.2 (8.5)	53.8 (8.1)	52.3 (8.0)
BMI age 18–21 kg/m ² , mean (SD)	20.1 (2.0)	21.5 (2.3)	22.2 (2.3)	28.8 (2.6)	22.8 (2.5)	25.3 (3.9)
BMI 1990–1994, kg/m ² , mean (SD)	21.8 (1.4)	25.1 (1.5)	28.5 (1.8)	30.5 (2.4)	32.6 (2.2)	37.5 (2.7)
BMI 1995–1998, kg/m ² , mean (SD)	22.4 (1.4)	25.8 (1.5)	29.3 (1.7)	31.0 (2.2)	33.6 (2.1)	38.8 (2.4)
BMI 2003–2007, kg/m ² , mean (SD)	22.0 (1.7)	25.7 (1.6)	29.5 (1.8)	31.0 (2.3)	34.1 (2.2)	39.0 (2.7)
Sex, N (%)						
Male	1443 (25.1)	5027 (44.8)	3936 (48.4)	535 (57.3)	995 (34.4)	218 (23.0)
Female	4305 (74.9)	6202 (55.2)	4194 (51.6)	398 (42.7)	1898 (65.6)	730 (77.0)
Country of birth, N (%)						
Australia/New Zealand/Northern Europe	5388 (93.7)	9282 (82.7)	5640 (69.4)	564 (60.5)	1832 (63.3)	605 (63.8)
Southern Europe	360 (6.3)	1947 (17.3)	2490 (30.6)	369 (39.5)	1061 (36.7)	343 (36.2)
SEIFA quintiles N (%)						
1 (east advantaged)	724 (12.6)	1734 (15.4)	1554 (19.1)	205 (22.0)	672 (23.2)	235 (24.8)
2	912 (15.9)	2143 (19.1)	1814 (22.3)	242 (25.9)	690 (23.9)	233 (24.6)
3	843 (14.7)	1749 (15.6)	1354 (16.7)	141 (15.1)	480 (16.6)	181 (19.1)
4	1177 (20.5)	2211 (19.7)	1515 (18.6)	151 (16.2)	525 (18.1)	143 (15.1)
5 (most advantaged)	2092 (36.4)	3392 (30.2)	1893 (23.3)	194 (20.8)	526 (18.2)	156 (16.5)
Level of education, N (%)						
Primary or lower	290 (5.0)	1374 (12.2)	1844 (22.7)	281 (30.1)	791 (27.3)	261 (27.5)
Some high/technical school	2245 (39.1)	4341 (38.7)	3024 (37.2)	326 (34.9)	1142 (39.5)	405 (42.7)
Completed high/technical school	1335 (23.2)	2536 (22.6)	1654 (20.3)	173 (18.5)	549 (19.0)	163 (17.2)
Completed tertiary education	1878 (32.7)	2978 (26.5)	1608 (19.8)	153 (16.4)	411 (14.2)	119 (12.6)
Mediterranean Diet Score 1990–1994, N (%)						
0–3	1753 (30.5)	3645 (32.5)	2763 (34.0)	358 (38.4)	1024 (35.4)	318 (33.5)
4–6	3238 (56.3)	6090 (54.2)	4373 (53.8)	481 (51.6)	1578 (54.5)	541 (57.1)
7–9	757 (13.2)	1494 (13.3)	994 (12.2)	94 (10.1)	291 (10.1)	89 (9.4)
Mediterranean Diet Score 2003–2007, N (%)						
0–3	956 (23.1)	1732 (23.1)	1241 (24.3)	125 (24.2)	425 (24.4)	151 (26.6)
4–6	2515 (60.9)	4626 (61.8)	3181 (62.3)	328 (63.4)	1134 (65.1)	360 (63.5)
7–9	660 (16.0)	1128 (15.1)	681 (13.3)	64 (12.4)	182 (10.5)	56 (9.9)
Physical Activity Score 1990–1994, N (%)						
0	866 (15.1)	2061 (18.4)	1985 (24.4)	243 (26.0)	859 (29.7)	318 (33.5)

Continued

Table 1 Continued

	BMI trajectory groups					
	TR1 lower-normal stable	TR2 higher-normal stable	TR3 normal to overweight	TR4 chronic borderline obesity	TR5 normal to class I obesity	TR6 overweight to class II obesity
	N=5748	N=11 229	N=8130	N=933	N=2893	N=948
>0 and <4	1065 (18.5)	2193 (19.5)	1647 (20.3)	172 (18.4)	643 (22.2)	239 (25.2)
≥4 and <6	2079 (36.2)	4026 (35.9)	2844 (35.0)	338 (36.2)	975 (33.7)	280 (29.5)
≥6	1738 (30.2)	2949 (26.3)	1654 (20.3)	180 (19.3)	416 (14.4)	111 (11.7)
METs 2003–2007, median (IQR)	21.5 (9.9–40.0)	19.3 (8.3–36.7)	15 (6.2–29.8)	14 (5.0–31.4)	11 (3.7–24.4)	9.3 (2.8–19.8)
Drinking status 1990–1994, N (%)						
Lifetime abstainers	1437 (25.0)	2691 (24.0)	2275 (28.0)	287 (30.8)	1060 (36.6)	429 (45.3)
Ex-drinkers	195 (3.4)	376 (3.3)	315 (3.9)	58 (6.2)	121 (4.2)	43 (4.5)
Current (1–39 g (male) 1–19 (female))	3460 (60.2)	6995 (62.3)	4734 (58.2)	516 (55.3)	1498 (51.8)	429 (45.3)
Current (≥40 (male), ≥20 (female))	656 (11.4)	1167 (10.4)	806 (9.9)	72 (7.7)	214 (7.4)	47 (5.0)
Smoking status at last BMI measurement, N (%)						
Never	3585 (62.4)	6547 (58.3)	4504 (55.4)	488 (52.3)	1749 (60.5)	602 (63.5)
Former	1710 (29.7)	3905 (34.8)	3004 (36.9)	359 (38.5)	941 (32.5)	274 (28.9)
Current	453 (7.9)	777 (6.9)	622 (7.7)	86 (9.2)	203 (7.0)	72 (7.6)
Had angina before 1990–1994, N (%)	141 (2.5)	391 (3.5)	308 (3.8)	38 (4.1)	106 (3.7)	48 (5.1)
Had heart attack before 1990–1994, N (%)	103 (1.8)	289 (2.6)	228 (2.8)	30 (3.2)	75 (2.6)	21 (2.2)
Had stroke before 1990–1994, N (%)	37 (0.6)	124 (1.1)	94 (1.2)	13 (1.4)	25 (0.9)	12 (1.3)
Had diabetes before 1990–1994, N (%)	75 (1.3)	253 (2.3)	326 (4.0)	69 (7.4)	147 (5.1)	58 (6.1)
Birth cohort, N (%)						
Before 1930	1250 (21.7)	2725 (24.3)	1659 (20.4)	233 (25.0)	479 (16.6)	117 (12.3)
After 1930, before 1940	1674 (29.1)	3662 (32.6)	2908 (35.8)	307 (32.9)	1028 (35.5)	310 (32.7)
After 1940, before 1950	1999 (34.8)	3513 (31.3)	2639 (32.5)	298 (31.9)	1033 (35.7)	370 (39.0)
After 1950	825 (14.4)	1329 (11.8)	924 (11.4)	95 (10.2)	353 (12.2)	151 (15.9)

BMI, body mass index; MET, metabolic equivalent; SEIFA, Socio-Economic Indexes for Areas.

this trajectory and attained age and presented HRs for age 74, 80 and 85 (25th, 50th and 75th percentiles of age at death). For this trajectory, the HR was greater at younger attained age than at older attained age (from HR 1.31, 95% CI 1.11 to 1.54 for age 74 years to HR 1.10, 95% CI 0.88 to 1.38 for age 85 years). In never-smokers, while the proportional hazards assumption was not violated, the same pattern (ie, weaker HRs with increasing age)

was seen for TR6. The associations between obesity across adulthood (TR4, TR5 and TR6) and all-cause mortality were stronger than for the whole cohort (HR 1.25, 95% CI 1.02 to 1.53, HR 1.19, 95% CI 1.03 to 1.37, and HR 1.42, 95% CI 1.13 to 1.78 for age 74 years to HR 1.16, 95% CI 0.88 to 1.54 for age 85 years, respectively).

Being obese across adulthood (TR4, TR5 and TR6) showed stronger associations with mortality from

Table 2 HRs for all-cause mortality according to BMI trajectory groups in the Melbourne Collaborative Cohort Study, overall and in never-smokers

BMI trajectory groups	Overall n=(29 881)*						Never-smokers n=(17 475)†					
	Person-years	Deaths	HR	95% CI	Person-years	Deaths	HR	95% CI	Person-years	Deaths	HR	95% CI
All-cause mortality												
TR1 Lower-normal stable	72 077	1162	1.00		45 432	604	1.00					
TR2 Higher-normal stable	141 989	2507	0.90	0.84 to 0.97	83 881	1220	0.96	0.87 to 1.07				
TR3 Normal to overweight	103 976	1826	0.96	0.89 to 1.04	58 515	820	1.00	0.90 to 1.12				
TR4 Borderline obesity chronic	11 711	280	1.16	1.01 to 1.33	6210	120	1.25	1.02 to 1.53				
TR5 Normal to class I obesity	37 235	600	1.09	0.98 to 1.21	22 762	315	1.19	1.03 to 1.37				
TR6 Overweight to class II obesity‡	12 357	187			7838	107						
Age 74 years			1.31	1.11 to 1.54			1.42	1.13 to 1.78				
Age 80 years			1.19	1.00 to 1.42			1.27	1.02 to 1.59				
Age 85 years			1.10	0.88 to 1.38			1.16	0.88 to 1.54				

*Adjusted for age, birth cohort, height, sex, country of birth, SEIFA, highest level of education, smoking status at latest BMI assessment, baseline MDS, alcohol drinking status and physical activity score.

†Adjusted for age, birth cohort, height, sex, country of birth, SEIFA, highest level of education, baseline MDS, alcohol drinking status and physical activity score.

‡The p refers interaction between TR6 and attained age=0.04, tested using likelihood ratio test.

BMI, body mass index; MDS, Mediterranean Diet Score; SEIFA, Socio-Economic Indexes For Areas.

Obesity-related cancer mortality

Cardiovascular mortality

Other-cause mortality

Figure 2 Forest plots of HRs for cause-specific mortality according to body mass index (BMI) trajectory groups in the Melbourne Collaborative Cohort Study, overall and in never-smokers.

obesity-related cancers than with all-cause mortality. Chronic borderline obesity (TR4) was associated with higher risk of death, especially from obesity-related cancers (HR 1.82, 95% CI 1.24 to 2.68 for all participants and HR 1.92, 95% CI 1.11 to 3.32 for never-smokers). The overweight to class II obesity trajectory (TR6) had increased risk of CVD mortality (HR 2.27, 95% CI 1.34 to 3.87 for the whole cohort; HR 1.62, 95% CI 1.02 to 2.56 for never-smokers). In general, BMI trajectories had weak associations with other-cause mortality (figure 2 and online supplementary table 2).

For the analyses of all-cause mortality, there was no evidence of interaction between BMI trajectory and sex ($p=0.2$), country of birth ($p=0.3$), alcohol consumption ($p=0.7$), history of diabetes ($p=0.2$), history of heart diseases ($p=0.3$), follow-up duration ($p=0.1$) or baseline age ($p=0.2$). HRs were similar after excluding the first 2 years of follow-up (results not shown). After excluding participants who had no 2003–2007 data, HRs for cause-specific mortality were generally similar (online supplementary table 3), but with wider CIs; the most obese

trajectory (TR6) interacted with attained age overall and in never-smokers for all-cause mortality (table 3).

Models including BMI trajectories fitted slightly better than those with last-measured BMI according to the Akaike information criterion and Bayesian information criterion (online supplementary table 4).

Participants with fewer than three BMI records had higher mean BMI at age 18–21 years (22.4 kg/m^2 and 21.9 kg/m^2 , respectively; $p<0.0001$) and at baseline (28.3 kg/m^2 and 26.7 kg/m^2 , respectively; $p<0.0001$) than those who had at least three records. They also had higher odds of all-cause mortality (OR 2.06, 95% CI 1.93 to 2.21).

DISCUSSION

Most participants had normal BMI at the age of 18–21 years. We identified six BMI trajectories. Most participants maintained a higher-normal BMI throughout adulthood or gained weight from higher-normal to become overweight. Only a small proportion of people developed class II obesity in mid to late life. Generally, the lowest mortality was experienced by participants who maintained higher-normal BMI throughout adulthood. Participants who were obese across the midlife had elevated all-cause mortality, even if they had normal BMI at younger age as did those who had prolonged exposure to borderline obesity. These associations were stronger for never-smokers and for obesity-related cancer deaths. Being overweight in early adulthood and becoming class II obese was associated with higher CVD mortality.

The key strength of our study is availability of weight at multiple times, which was used to derive trajectories in BMI across adulthood. This approach allowed us to categorise BMI over time into homogeneous groups and to identify a specific higher-risk group (ie, prolonged borderline obesity trajectory) that would not have been identifiable using the conventional approach of grouping participants into categories of change.

Several limitations should be noted. For some trajectories, there were few cause-specific deaths, with corresponding wide CIs. Recalled weight at age 18–21 years may be subject to measurement error. However, previous studies have generally found moderate to strong correlation between measured weight in young adulthood and recalled weight for the same age 32–56 years later.^{33 34} We did not have information on potential confounders at age 18–21. We also had no information on weight between age 18–21 years and baseline (20–50 years later), which limited us from identifying trajectories of how people gained weight after age 18–21 years and before entering the study in midlife. Participants excluded due to insufficient BMI records had higher BMI at age 18–21 and baseline. Having insufficient BMI records was also associated with higher odds of dying. Therefore, we may have underestimated the effect of obese trajectories on all-cause mortality.

Table 3 HRs for all-cause mortality according to BMI trajectory groups in the Melbourne Collaborative Cohort Study, overall and in never-smokers, excluding participants who did not attend 2003–2007 data collection

BMI trajectory groups	Overall* n=(21 903)					Never-smokers† n=(13 253)				
	Person-years	Deaths	HR	95% CI		Person-years	Deaths	HR	95% CI	
TR1 Lower-normal stable	52 878	655	1.00			34 655	358	1.00		
TR2 Higher-normal stable	96 021	1283	0.88	0.80 to 0.97		58 420	632	0.91	0.80 to 1.03	
TR3 Normal to overweight	65 825	902	0.98	0.88 to 1.09		38 118	438	1.04	0.90 to 1.20	
TR4 Borderline obesity chronic	6624	124	1.21	0.99 to 1.47		3643	65	1.36	1.04 to 1.79	
TR5 Normal to class I obesity	22 598	280	1.13	0.98 to 1.30		14 114	164	1.22	1.01 to 1.48	
TR6 Overweight to class II obesity‡	7125	83				4768	50			
Age 74 years			1.43	1.13 to 1.83				1.61	1.17 to 2.20	
Age 80 years			1.08	0.84 to 1.39				1.19	0.86 to 1.65	
Age 85 years			0.85	0.62 to 1.17				0.93	0.62 to 1.39	

*Adjusted for age, birth cohort, height, sex, country of birth, SEIFA, highest level of education, smoking status at latest BMI assessment, baseline MDS, alcohol drinking status and physical activity score.

†Adjusted for age, birth cohort, height, sex, country of birth, SEIFA, highest level of education, baseline MDS, alcohol drinking status and physical activity score.

‡The p refers interaction between TR6 and attained age=0.0002 for overall and 0.003 for never-smokers, tested using likelihood ratio test. BMI, body mass index; MDS, Mediterranean Diet Score; SEIFA, Socio-Economic Indexes For Areas.

Participants were mostly born before 1950. Few people (2%) had BMI of 30 or higher at age 18–21 years. Hence, we were not able to estimate the effect of sustained obesity from young adulthood. The prevalence of overweight in Australia has increased in cohorts born since 1960; the mean BMI has been higher for each successive birth cohort at any given age.³⁵ The prevalence of childhood overweight and obesity has also increased globally.^{36–38} In Australia, the proportion of obese individuals at age 18–21 years was almost double in people born in 1994–1997 (15.2%) compared with their counterparts born 20 years earlier (8%).³⁹ The projected increase of obesity prevalence in young adults⁴⁰ also suggests that potential higher-risk trajectories of chronic and severe obesity may emerge. Therefore, it is expected that more people in future generations will have higher BMI for longer durations. Conventional analytic approaches using a single BMI assessment or an a priori categorisation of body size changes could overlook these potential high-risk groups. Trajectories models may help differentiate people who have been obese long term from those who become obese later in life.

Some studies have reported lower all-cause mortality in overweight individuals than for normal weight individuals.^{2,5} However, we did not observe lower risk of death in participants who were overweight in midlife (TR3). The US Health and Retirement Study (HRS) modelled BMI trajectories from 51 to 77 years for adults aged 51–61 years at baseline and found that people with sustained overweight had the lowest all-cause mortality.¹⁷ The National Survey of the Japanese Elderly identified no obesity in trajectories in a Japanese population aged 60 years and older, with only 5% ‘overweight stable’ individuals.¹⁸ The lowest risk of death was observed for participants with sustained overweight overall, but not in non-current smokers.¹⁸ However, these two studies only used BMI data in mid to late life. A study of health professionals in the USA using earlier life (age 5–50 years) body shape pictograms found that people remaining lean had the lowest mortality; people who experienced a marked increase from lean to heavy had a higher risk; whereas people who had been heavy since childhood were at highest risk.¹⁶ However, many participants in the study were not obese according to their BMI, even when they were classified as ‘heavy’ based on pictogram. Possible measurement errors in body shape data (ie, combining recalled somatotypes with BMI-predicted somatotypes) may have affected their trajectory shapes. Nevertheless, their results showed that larger body shapes from childhood are likely to persist into adulthood, that being consistently heavy is associated with increased risk of death from all-cause, CVD and cancer, and the increased risk was stronger for never-smokers.¹⁶ Another study of 70 000 men showed that both ‘normal to obese’ and ‘overweight to obese’ adulthood BMI trajectories were associated with higher mortality from fatal prostate cancer compared with stable normal BMI trajectory.⁴¹

In our study, participants who had long-term borderline obesity also had increased risk of death compared

with lower-normal BMI trajectory. Long-term obesity gives longer exposure to proinflammatory adipocytokines released by adipose tissue, which can lead to pathophysiological alterations such as dyslipidaemia, hypertension, fibrinolysis, inflammation, insulin resistance and immune system alterations.^{7 42 43} Obesity also increases bioavailable estrogens (via aromatisation, the conversion of adrenal androgens to estrone, in adipose tissue),^{44 45} a contributing factor to breast and gynaecological cancers. Long-term obesity may also be associated with partial exhaustion of the beta cell and insulinopenia which might cause depressed glucose oxidation and impaired glucose tolerance.⁴⁶ Longer duration of obesity is associated with several adverse health consequences that could contribute to increased mortality, including unfavourable biomarker profiles as indicated by higher C-peptide and lower insulin-like growth factor-binding protein 1 concentrations,⁴⁷ type 2 diabetes,^{48 49} cancer^{41 50–52} and CVD.⁵³

A systematic review of studies using baseline BMI showed that all-cause mortality was higher for older people who had a BMI <23.0 kg/m².⁵⁴ In older populations such as the National Survey of the Japanese Elderly, most people followed trajectories with a gradual decline in BMI over time.¹⁸ Relative to the 'mid-normal decreasing' group, all-cause mortality was higher in the 'low-normal decreasing' group (HR 1.17; 95% CI 1.02 to 1.33) and lower in the 'high-normal decreasing' group (HR 0.82; 95% CI 0.72 to 0.93). This suggests that for older people experiencing weight loss, a higher initial BMI may be protective. In our study, the estimated effect of being overweight in early adulthood and gaining weight in midlife (TR6) declined with age. This is consistent with the results of a study investigating weight change with different initial BMI levels in relation to all-cause mortality in older Americans.¹¹ The authors found that compared with having a stable weight, even a small weight loss in the elderly was associated with increased mortality in normal weight, overweight and even mildly obese individuals.¹¹ This association was stronger when initial BMI was lower.¹¹ Similarly, in a previous analysis of weight change and mortality using the MCCS, weight loss from 1990 to 1994 to 2003–2007 (ie, in later adulthood) was associated with higher risk of death.¹² Weight loss in older people can be an indicator of undesirable health conditions such as sarcopenia,⁵⁵ or a more severe disease state.⁵⁶ Individuals with higher body weight may tolerate weight loss better than individuals with a normal body weight because of higher metabolic reserves and body fat.⁵⁶ Measures of body composition and fat distribution are needed to capture more adequately the relationship of adiposity and lean mass to health risks in older populations.⁵⁷

CONCLUSION

Midlife obesity was associated with higher mortality even when BMI was normal in early adulthood. This highlights the importance of weight management throughout

adulthood. Prolonged borderline obesity was also associated with elevated risk of death from all-causes and obesity-related cancers relative to maintaining lower-normal BMI across adulthood. Since chronic obesity will likely be a major health issue for future generations, policies and prevention programmes are needed to target obesity starting early in life.

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REFERENCES

1. Flegal KM, Kit BK, Orpana H, *et al.* Association of all-cause mortality with overweight and obesity using standard body mass index categories: a systematic review and meta-analysis. *JAMA* 2013;309:71–82.
2. Aune D, Sen A, Prasad M, *et al.* BMJ and all cause mortality: systematic review and non-linear dose-response meta-analysis of 230 cohort studies with 3.74 million deaths among 30.3 million participants. *Br Med J* 2016;353.
3. Calle EE, Rodriguez C, Walker-Thurmond K, *et al.* Overweight, obesity, and mortality from cancer in a prospectively studied cohort of U.S. adults. *N Engl J Med* 2003;348:1625–38.
4. Di Angelantonio E, Bhupathiraju SN, Wormser D, *et al.* Body-Mass index and all-cause mortality: individual-participant-data meta-analysis of 239 prospective studies in four continents. *The Lancet* 2016;388:776–86.
5. Berrington de Gonzalez A, Hartge P, Cerhan JR, *et al.* Body-Mass index and mortality among 1.46 million white adults. *N Engl J Med* 2010;363:2211–9.
6. Lauby-Secretan B, Scoccianti C, Loomis D, *et al.* Body fatness and cancer — viewpoint of the IARC Working Group. *N Engl J Med* 2016;375:794–8.
7. Van Gaal LF, Mertens IL, De Block CE. Mechanisms linking obesity with cardiovascular disease. *Nature* 2006;444:875–80.
8. Ferreira I, Stehouwer CDA. Obesity paradox or inappropriate study designs? Time for life-course epidemiology. *J Hypertens* 2012;30:2271–5.
9. Klenk J, Rapp K, Ulmer H, *et al.* Changes of body mass index in relation to mortality: results of a cohort of 42,099 adults. *PLoS One* 2014;9:e84817.

10. Zheng Y, Manson JE, Yuan C, *et al.* Associations of weight gain from early to middle adulthood with major health outcomes later in life. *JAMA* 2017;318:255–69.
11. Myrskylä M, Chang VW, Initial BMI. Weight change, initial BMI, and mortality among middle- and Older-aged adults. *Epidemiology* 2009;20:840–8.
12. Karahalios A, Simpson JA, Baglietto L, *et al.* Change in body size and mortality: results from the Melbourne Collaborative cohort study. *PLoS One* 2014;9:e99672.
13. Hirko KA, Kantor ED, Cohen SS, *et al.* Body mass index in young adulthood, obesity trajectory, and premature mortality. *Am J Epidemiol* 2015;182:441–50.
14. Karahalios A, English DR, Simpson JA. Change in body size and mortality: a systematic review and meta-analysis. *Int J Epidemiol* 2017;46:526–46.
15. Nagin D. *Group-Based modeling of development*. Cambridge: Harvard University Press, 2005.
16. Song M, Hu FB, Wu K, *et al.* Trajectory of body shape in early and middle life and all cause and cause specific mortality: results from two prospective US cohort studies. *Br Med J* 2016;353.
17. Zheng H, Tumin D, Qian Z. Obesity and mortality risk: new findings from body mass index trajectories. *Am J Epidemiol* 2013;178:1591–9.
18. Murayama H, Liang J, Bennett JM, *et al.* Trajectories of body mass index and their associations with mortality among older Japanese: do they differ from those of Western populations? *Am J Epidemiol* 2015;182:597–605.
19. Simpson JA, MacInnis RJ, Peeters A, *et al.* A comparison of adiposity measures as predictors of all-cause mortality: the Melbourne Collaborative cohort Study. *Obesity* 2007;15:994–1003.
20. Ireland P, Jolley D, Giles G, *et al.* Development of the Melbourne FFC: a food frequency questionnaire for use in an Australian prospective study involving an ethnically diverse cohort. *Asia Pac J Clin Nutr* 1994;3:19–31.
21. Milne RL, Fletcher AS, MacInnis RJ, *et al.* Cohort profile: the Melbourne Collaborative cohort study (health 2020). *Int J Epidemiol* 2017;46:1757.
22. Lohman TG, Roche AF, Martorell R. *Anthropometric standardization reference manual*. Champaign, IL: Human Kinetics Books, 1988.
23. Ball K, Crawford D, Ireland P, *et al.* Patterns and demographic predictors of 5-year weight change in a multi-ethnic cohort of men and women in Australia. *Public Health Nutr* 2003;6:269–80.
24. McLennan W, Podger AS. *National nutrition survey users' guide*, 1995. Canberra: Australian Bureau of Statistics [and] Commonwealth Department of Health and Family Services, 1998.
25. Jones BL, Nagin DS. A note on a Stata plugin for estimating group-based trajectory models. *Social Methods Res* 2013;42:608–13.
26. Thiébaud ACM, Bénichou J. Choice of time-scale in COX's model analysis of epidemiologic cohort data: a simulation study. *Stat Med* 2004;23:3803–20.
27. Pink B. *Socio-Economic Indexes for Areas (SEIFA) 2011*. Canberra: Australian Bureau of Statistics, 2013.
28. Ainsworth BE, Haskell WL, Leon AS, *et al.* Compendium of physical activities: classification of energy costs of human physical activities. *Med Sci Sports Exerc* 1993;25:71–80.
29. Hodge AM, English DR, Itsiopoulos C, *et al.* Does a Mediterranean diet reduce the mortality risk associated with diabetes: evidence from the Melbourne Collaborative cohort study. *Nutr Metab Cardiovasc Dis* 2011;21:733–9.
30. Tobias DK, Hu FB. Does being overweight really reduce mortality? *Obesity* 2013;21:1746–9.
31. Manson JE, Stampfer MJ, Hennekens CH, *et al.* Body weight and longevity. A reassessment. *JAMA* 1987;257:353–8.
32. World Health Organization. *Obesity: preventing and managing the global epidemic*. Geneva: World Health Organization, 2000.
33. Casey VA, Dwyer JT, Berkey CS, *et al.* Long-Term memory of body weight and past weight satisfaction: a longitudinal follow-up study. *Am J Clin Nutr* 1991;53:1493–8.
34. Must A, Willett WC, Dietz WH. Remote recall of childhood height, weight, and body build by elderly subjects. *Am J Epidemiol* 1993;138:56–64.
35. Allman-Farinelli MA, Chey T, Bauman AE, *et al.* Age, period and birth cohort effects on prevalence of overweight and obesity in Australian adults from 1990 to 2000. *Eur J Clin Nutr* 2008;62:898–907.
36. Australian Institute of Health and Welfare. Childhood overweight and obesity. In: *Australia's health 2014*. Australia's health series no. 14. Cat. no. AUS 178. Canberra: AIHW, 2014.
37. Lobstein T, Jackson-Leach R, Moodie ML, *et al.* Child and adolescent obesity: part of a bigger picture. *The Lancet* 2015;385:2510–20.
38. Ogden CL, Carroll MD, Lawman HG, *et al.* Trends in obesity prevalence among children and adolescents in the United States, 1988–1994 through 2013–2014. *JAMA* 2016;315:2292–9.
39. Australian Institute of Health and Welfare. *Overweight and obesity in Australia: a birth cohort analysis*. Cat. no. PHE 215. Canberra: AIHW, 2017.
40. Hayes AJ, Lung TWC, Bauman A, *et al.* Modelling obesity trends in Australia: unravelling the past and predicting the future. *Int J Obes* 2017;41:178–85.
41. Kelly SP, Graubard BI, Andreotti G, *et al.* Prediagnostic body mass index trajectories in relation to prostate cancer incidence and mortality in the PLCO cancer screening trial. *J Natl Cancer Inst* 2017;109:djw225.
42. Font-Burgada J, Sun B, Karin M, Obesity KM. Obesity and cancer: the oil that feeds the flame. *Cell Metab* 2016;23:48–62.
43. Kolb R, Sutterwala FS, Zhang W. Obesity and cancer: inflammation bridges the two. *Curr Opin Pharmacol* 2016;29:77–89.
44. Kendall A, Folkert EJ, Dowsett M. Influences on circulating oestrogens in postmenopausal women: relationship with breast cancer. *J Steroid Biochem Mol Biol* 2007;103:99–109.
45. Lukanova A, Kaaks R. Endogenous hormones and ovarian cancer: epidemiology and current hypotheses. *Cancer Epidemiol Biomarkers Prev* 2005;14:98–107.
46. Felber JP, Golay A, Jéquier E, *et al.* The metabolic consequences of long-term human obesity. *Int J Obes* 1988;12:377–89.
47. Kværner AS, Hang D, Giovannucci EL, *et al.* Trajectories of body fatness from age 5 to 60 Y and plasma biomarker concentrations of the insulin–insulin-like growth factor system. *Am J Clin Nutr* 2018;108:388–97.
48. Abdullah A, Stoelwinder J, Shortreed S, *et al.* The duration of obesity and the risk of type 2 diabetes. *Public Health Nutr* 2011;14:119–26.
49. Everhart JE, Pettitt DJ, Bennett PH, *et al.* Duration of obesity increases the incidence of NIDDM. *Diabetes* 1992;41:235–40.
50. Arnold M, Jiang L, Stefanick ML, *et al.* Duration of adulthood overweight, obesity, and cancer risk in the Women's Health Initiative: a longitudinal study from the United States. *PLoS Med* 2016;13:e1002081.
51. Yang B, Petrick JL, Kelly SP, *et al.* Adiposity across the adult life course and incidence of primary liver cancer: the NIH-AARP cohort. *Int J Cancer* 2017;141:271–8.
52. Petrick JL, Kelly SP, Liao LM, *et al.* Body weight trajectories and risk of oesophageal and gastric cardia adenocarcinomas: a pooled analysis of NIH-AARP and PLCO studies. *Br J Cancer* 2017;116:951–9.
53. Abdullah A, Amin FA, Stoelwinder J, *et al.* Estimating the risk of cardiovascular disease using an obese-years metric. *BMJ Open* 2014;4:e005629.
54. Winter JE, MacInnis RJ, Wattanapenpaiboon N, *et al.* Bmi and all-cause mortality in older adults: a meta-analysis. *Am J Clin Nutr* 2014;99:875–90.
55. Bales CW, Ritchie CS. Sarcopenia, weight loss, and nutritional frailty in the elderly. *Annu Rev Nutr* 2002;22:309–23.
56. Banack HR, Kaufman JS. The obesity paradox: understanding the effect of obesity on mortality among individuals with cardiovascular disease. *Prev Med* 2014;62(Supplement C):96–102.
57. Michels KB, Greenland S, Rosner BA. Does body mass index adequately capture the relation of body composition and body size to health outcomes? *Am J Epidemiol* 1998;147:167–72.