Vortioxetine treatment for generalised anxiety disorder: a meta-analysis of anxiety, quality of life and safety outcomes

Bin Qin,1 Guangsu Huang,2 Qian Yang,2 Mingjun Zhao,3 Hong Chen,2 Wen Gao,2 Mingxiu Yang2

ABSTRACT

Objectives The aim of this study was to investigate the efficacy, tolerability, safety, and impact on quality of life (QoL) and functional status of vortioxetine treatment for patients with generalised anxiety disorder (GAD) by performing a meta-analysis of randomised controlled trials (RCTs).

Design Systematic review and meta-analysis.

Data sources Data mining was conducted in January 2019 across PubMed, EMBASE, PsycINFO, Cochrane Central Register of Controlled Trials Cochrane Library, Web of science and ClinicalTrials.gov.

Eligibility criteria for selecting studies All published RCTs, which assessed the effect of vortioxetine treatment for patients with GAD when compared with a placebo group, were included.

Data extraction and synthesis Relevant data were extracted and synthesised narratively. Results were expressed as standardised mean differences or ORs with 95% CIs.

Results Our meta-analysis showed that multiple doses (2.5, 5 and 10 mg/day) of vortioxetine did not significantly improve the response rates, compared with placebo (OR 1.16, 95% CI 0.84 to 1.60, p=0.38; OR 1.41, 95% CI 0.82 to 2.41, p=0.21; and OR 1.05, 95% CI 0.76 to 1.46, p=0.75). Moreover, there was no statistically significant difference regarding the remission rates, discontinuation for any reason rates, discontinuation due to adverse events rates, Short-Form 36 Health Survey scores or Sheehan Disability Scale scores between administration of multiple doses (2.5, 5 and 10 mg/day) of vortioxetine and placebo.

Conclusions Although our results suggest that vortioxetine did not improve the GAD symptoms, QoL and functional status impairments of patients with GAD, it was safe and well tolerated. Clinicians should interpret and translate our data with caution, as the meta-analysis was based on a limited number of RCTs.

INTRODUCTION

Generalised anxiety disorder (GAD) is a common, chronic, costly and disabling mental disorder that is marked by persistent anxiety and worry, and multiple psychological and physical symptoms.1,2 It is also characterised by various psychological and somatic complaints, such as autonomic arousal, restlessness, fatigue, problems with concentration, irritability and sleeplessness.1 The estimated 1-year prevalence rate of GAD is between 1.2% and 1.9%, and the lifetime prevalence is between 4.3% and 5.9%.2,3 Since most patients are still affected for 6–12 years after diagnosis, GAD is usually considered a chronic disorder and a major burden on the individual, their family and healthcare services.2,4 Anxiety is a common comorbidity of chronic medical diseases, including atopic dermatitis, asthma, rheumatoid arthritis, lupus and stroke. Anxiety has a negative impact on the quality of life (QoL) of patients suffering from chronic diseases.2,5

Vortioxetine is a multimodal antidepressant that was approved for the treatment of major depressive disorder (MDD), by the US Food and Drug Administration (FDA) in September 2013. Vortioxetine’s mechanism of action is related to its multimodal activity, which combines two pharmacological properties: direct modulation of receptor activity and inhibition of the 5-HT transporter, and thereby with potential benefits in the treatment of major depressive episode and, probably, GAD and anxiety disorders.7 In
addition, vortioxetine causes a significant increase in the hippocampal brain-derived neurotrophic factors levels as compared with selective serotonin reuptake inhibitors. Several meta-analyses have proved the efficacy of vortioxetine in treatment of MDD. A recent systematic analysis reported that the popularity of vortioxetine is rising. Clinical trials evaluating its efficacy for the treatment of GAD, with doses up to 10 mg/day, have also yielded some interesting findings. Moreover, as vortioxetine has been proven to be efficient in the treatment of MDD comorbid with GAD, it is possible that it constitutes an effective treatment for GAD alone, as well.

Interestingly, the efficacy of vortioxetine therapy in reducing anxiety symptom severity in GAD is summarised in two previous meta-analyses. Both reviews analysed its efficacy only in terms of symptom severity on the underlying continuous rating scales, and did not assess dichotomous outcomes of response and remission rates. However, a recent meta-analysis examined the efficacy of multiple doses of vortioxetine in terms of dichotomous response outcomes, and the results showed no significant improvement in the outcomes of treating GAD with vortioxetine compared with treating GAD with placebo. The efficacy was only assessed using continuous rating scales or dichotomous response; thus, the authors of these meta-analyses and of a relevant narrative review noted that a comprehensive summary of efficacy data is missing. Further, both these reviews only provided an assessment of efficacy and safety outcomes, and did not include important outcomes of patient-focused assessments, such as assessment of functional impairment and QoL. Currently, there is growing interest in assessing the QoL and functional status impairment in patients with psychological disorders. In addition, the importance of including such assessments in evaluations of the influence of psychological disorders and their treatment is widely recognised. Our previous network meta-analysis concluded that risperidone and aripiprazole improved the QoL of patients with treatment-resistant depression. Despite the growing interest in the field, studies addressing the impairment of QoL and functional status caused by anxiety disorders have progressed slowly. Moreover, GAD is an important public mental health problem that causes poor QoL and functional status impairment, with substantial impact on work and social roles. Thus, the outcome of post-treatment QoL assessments is recognised as an important measure of treatment efficacy for patients with GAD.

The assessment of antianxiety therapy benefits on QoL and functional status impairment in patients with GAD is becoming increasingly common in clinical studies, mainly because, both aspects are important for the patients’ overall well-being and recovery. Currently, the direct effect of vortioxetine treatment on QoL and functional status impairment in patients with GAD is unclear. Therefore, this meta-analysis was conducted to provide a comprehensive estimate of the efficacy, safety, and improving QoL and functional status impairment profiles of vortioxetine in treatment of GAD.

METHODS

All steps of this review were performed in strict accordance with the Cochrane Handbook for Systematic Reviews of Interventions. The Preferred Reporting Items for Systematic Reviews and Meta-Analyses statement guidelines were followed during the meta-analysis and preparation of this review.

Search strategy

As of January 2019, we searched PubMed, EMBASE, PsycINFO, Cochrane Central Register of Controlled Trials (Cochrane Library), Web of science and ClinicalTrials.gov (www.clinicaltrials.gov). Search terms included “vortioxetine OR Lu AA21004” OR Brintellix” AND “anxiety OR anxiety disorder OR anxiety disorders OR mood disorder OR mood disorders” (online supplementary table 1). No language or time restrictions were applied. Titles and abstracts were screened by two independent reviewers, before full texts of potentially relevant articles were retrieved for further evaluation. The decision to include a study was then made by two independent reviewers (QB and WG), after full-text review. The reference lists of included articles were further hand-searched to identify additional relevant articles.

Eligibility criteria and study selection

We included all clinical trials meeting the following criteria: (1) randomised controlled trials (RCTs) involving patients (≥18 years old) primarily diagnosed with GAD, according to the criteria of the Diagnostic and Statistical Manual of Mental Disorders, fourth edition, text revision and (2) RCTs comparing outcomes in efficacy, QoL and functional impairment between vortioxetine and placebo. We excluded (1) retrospective and observational studies; (2) non-human studies; (3) theses and conference abstracts; and (4) studies including patients who had any concurrent psychiatric disorders with GAD or any prior history of psychiatric disorders, such as manic or hypomanic episodes, schizophrenia or substance use disorders. Eligibility screening was performed in two steps, each by two independent reviewers (QB and WG): (1) title and abstract screening for relevance to the study objective and (2) full-text screening for eligibility for a meta-analysis. Conflicts were resolved by the opinion of a third reviewer (MY).

Outcomes

Efficacy measures

Response was defined across studies as a 50% improvement of the Hamilton Anxiety Rating Scale (HAM-A) total score from baseline to the end point. Remission was defined as an HAM-A total score of ≤7 at the study end point. The continuous measure of efficacy was the mean change from baseline in total scores on the HAM-A, as defined by the individual study.

Safety and tolerability measures

Data on the discontinuation for any reason (tolerability) and the numbers of discontinuation due to adverse events (AEs) (safety) were included in the analysis.
QoL and functioning measures

QoL can be assessed by study-designed questionnaires and disease-specific or generic instruments. These instruments assess an individual’s physical, emotional, psychological and social health. We selected the Short-Form 36 Health Survey (SF-36) scores as the outcome indicator for QoL to preserve sufficient homogeneity for meta-analysis, because this instrument is used to measure QoL for the GAD population in many studies. Studies were excluded if the QoL outcome was reported by the other rating scales. The Sheehan Disability Scale (SDS), a reliable, valid, brief, self-report scale that assesses disability in work, social and family life, is the only measure of functional impairment employed by the trials included in this meta-analysis.

Data extraction

Two independent reviewers (QB and WG) extracted the following data from the included studies: (1) baseline characteristics of enrolled patients, (2) general characteristics of the study design, (3) information on efficacy, safety, tolerability, QoL and functioning outcome. Data were summarised by one investigator and checked by a second reviewer. Any discrepant data were, again, examined by a third reviewer (MY), to ensure accurate data were obtained.

Risk of bias assessment

The risk of bias within each study was assessed by two independent reviewers (QB and WG) using the Cochrane Risk of Bias assessment tool, as described in the Cochrane Handbook for Systematic Reviews of Interventions 5.1.0. This tool classifies the studies as having low, unclear or high risk of bias across six domains: sequence generation, allocation concealment, blinding, missing data, selective reporting and other biases.

Data analysis

The meta-analysis was conducted using RevMan V.5.3 software (Cochrane Collaboration, London, UK) and Stata V.13.0 software (StataCorp). ORs and 95% CIs were used to assess binary outcomes, such as response, remission rates, as well as discontinuation for any reason rates. In addition, we converted continuous data to standardised mean differences (SMDs) and 95% CIs. The statistical heterogeneity among trials was measured by Q statistics and the I² test. Higher I² values indicate greater heterogeneity, with I² values of 25, 50 and 75% signifying mild, moderate and high heterogeneity, respectively. Based on heterogeneity, data were pooled to estimate the overall effect of all the interventions by random-effect modelling. Fixed-effect models assume that the population effect sizes are the same for all studies. In contrast, random-effects model attempted to generalise findings beyond the included studies by assuming that the selected studies are random samples from a larger population. Sensitivity analyses were performed to test the impact robustness of every single study on the overall results. Publication bias could be assessed by visual inspection of a funnel plot, and the Egger test was also used to evaluate publication biases. However, according to Egger et al, assessing publication bias using the funnel plot-based methods is not reliable when fewer than 10 pooled studies are used in the direct comparison.

Patient and public involvement

Patients or the public were not involved in the conceptualisation or carrying out of this research.

RESULTS

Search results

We identified 94 references from the electronic literature search. After screening the titles and abstracts, 85 were excluded because they failed to meet the inclusion criteria. By reading the full text of the remaining nine articles, five more were excluded: one study included patients with depression, one study focused on vortioxetine in the prevention of relapse of GAD, another one constituted a conference abstract and did not provide treatment outcomes, and two studies were eliminated because they were review articles. Ultimately, only four studies that fully satisfied the pre-established inclusion criteria of this meta-analysis were included (see online supplementary figure 1).

Study characteristics

Four included studies were published between 2012 and 2014 (table 1). The collective patient population comprised 1074 individuals in the vortioxetine group and 613 individuals in the placebo group. The administered doses of vortioxetine were 2.5, 5 and 10 mg/day. The mean age of participants ranged from 36.8 to 45.3 years. All studies were characterised by a preponderance of female subjects, with proportions ranging from 60% to 70%. The main characteristics of these studies are presented in table 1.

Study quality

The risks of bias in each study is summarised in online supplementary figure 2. All studies claimed randomisation and three articles described the method of random sequence generation (random number table, computer generated). Three trials provided information that allowed us to assess whether an adequate concealment of the allocation procedure was used. All studies reported the blinding of participants. Therefore, all trials were deemed to have a mild-to-moderate risk of bias.

Efficacy

In terms of response, a total of four studies were included in the analysis; the overall ORs observed for groups treated with multiple doses (2.5, 5, and 10 mg/day) of vortioxetine compared with placebo were 1.16 (95% CI 0.84 to 1.60, Z=0.89, p=0.38), 1.41 (95% CI 0.82 to 2.41, Z=1.25, p=0.21) and 1.05 (95% CI 0.76 to 1.46, Z=0.32, p=0.75), respectively (figure 1). The results showed that there was no statistically significant difference in the response rates among the vortioxetine and placebo groups. In addition,
there was no statistically significant difference for the remission rates in multiple doses (2.5, 5 and 10 mg/day) of vortioxetine compared with placebo (figure 2).

Pooled effect sizes for the mean change from baseline in total scores on the HAM-A are provided in figure 3. The overall SMDs observed for the groups treated with multiple doses (2.5, 5 and 10 mg/day) of vortioxetine compared with the placebo were −0.13 (95% CI −0.29 to 0.03, Z=1.56, p=0.12), −0.15 (95% CI −0.48 to 0.18, Z=0.87, p=0.38) and −0.08 (95% CI −0.24 to 0.08, Z=1.01, p=0.31), respectively. The results also showed that there was no statistically significant difference in the mean change from baseline in total scores on the HAM-A among the vortioxetine and placebo groups.

Tolerability and safety

No significant difference was observed between the vortioxetine and placebo groups in terms of the likelihood of
**Figure 2**  ORs and 95% CIs of the individual studies and the pooled data, comparing the remission rates between the vortioxetine and placebo groups. (A) 2.5 mg/day vortioxetine versus placebo, (B) 5 mg/day vortioxetine versus placebo and (C) 10 mg/day vortioxetine versus placebo. M-H, Mantel-Haenszel; IV, Inverse Variance.

**QoL and functional status results**

Three studies in this analysis reported SF-36 scores as the outcome measure of QoL. The overall SMDs of groups treated with multiple doses (2.5, 5 and 10 mg/day) of vortioxetine compared with the placebo-treated group were 1.03 (95% CI 0.71 to 1.47, Z=0.14, p=0.89), 0.87 (95% CI 0.63 to 1.20, Z=0.87, p=0.38) and 1.06 (95% CI 0.74 to 1.52, Z=0.32, p=0.75), respectively (online supplementary figure 3). Additionally, there was no statistically significant difference in the discontinuation due to AEs (safety) between the groups treated with multiple doses (2.5, 5 and 10 mg/day) of vortioxetine and the group treated with placebo (online supplementary figure 4).

**DISCUSSION**

In this meta-analysis of four randomised trials studying vortioxetine as a treatment for GAD, we found that vortioxetine (2.5, 5 and 10 mg once-daily doses) did not significantly improve GAD symptoms, QoL and functional status compared with a placebo treatment. However, vortioxetine might be safe and well tolerated in this patient population. Our findings have some clinical implications for comprehensively understanding the risk–benefit profiles of vortioxetine treatment for GAD.

Our results are not consistent with those of the previous meta-analysis conducted by Pae et al, as that study found that vortioxetine was significantly more effective than the placebo. In their study, they only performed the analysis of mean change from baseline in total scores on the HAM-A, and included all randomised subjects. However,
Figure 3  Standardised mean differences and 95% CIs of the individual studies and the pooled data, comparing the mean change from baseline in total scores on the HAM-A, between the vortioxetine and placebo groups. (A) 2.5 mg/day vortioxetine versus placebo, (B) 5 mg/day vortioxetine versus placebo and (C) 10 mg/day vortioxetine versus placebo. HAM-A, Hamilton Anxiety Rating Scale; M-H, Mantel-Haenszel; IV, Inverse Variance.

Figure 4  Standardised mean differences and 95% CIs of the individual studies and the pooled data, comparing the Short-Form 36 Health Survey scores between the vortioxetine and placebo groups. (A) 2.5 mg/day vortioxetine versus placebo, (B) 5 mg/day vortioxetine versus placebo and (C) 10 mg/day vortioxetine versus placebo. M-H, Mantel-Haenszel; IV, Inverse Variance.
significant difference regarding the response rates among groups treated with either multiple doses (2.5, 5 and 10 mg/day) of vortioxetine or a placebo.

Furthermore, the results of our meta-analysis demonstrated no statistically significant difference in the mean change from baseline in total scores on the HAM-A, remission rates, QoL and functional status among the groups. Thus, the results of our meta-analysis were more comprehensive.

Several reasons for these outcomes may have contributed to the negative results. A previous analysis of the FDA database concluded that negative results were commonly seen in anxiolytic agents administered for the treatment of anxiety disorders, including GAD, where less than one-half (48%) of the treatment arms were statistically superior to the placebo.

In this case, all anxiolytic agents included in the study are approved for GAD treatment in the USA, but only three out of seven treatment arms were separated from the placebo. Moreover, the results some studies have found that negative results had a higher placebo response rate than those with positive results.

Although this correlation has not been established in GAD, it is reasonable to speculate that an elevated placebo response could reduce the treatment effect. Unfortunately, because no positive control was included in our meta-analysis, it is impossible to determine the lack-of-treatment effect. In addition, the racial diversity may have introduced differences in response and remission rates; for example, the studied population of the trial that showed negative results was racially diverse, whereas the population of the trial that showed positive results was almost entirely Caucasian. Hence, the results of the STAR*D study demonstrate that non-Caucasians were significantly less likely to achieve remission.

Furthermore, the mean baseline HAM-A total scores in most of the included studies were relatively high (ranging from 24.5 to 27); inflated baseline HAM-A total scores are a possible consequence of less stringent screening practices. Relative to placebo, vortioxetine treatment, in doses of 10 and 20 mg, demonstrated significant improvement in depressive symptoms; however, the RCTs that were studied, in accordance with those included in our meta-analysis, used the optimal doses of 2.5–10 mg. Higher dosages, such as 20 mg or even more, would be more beneficial in anxiety disorders, whereas lower dosages may unaffect anxiety symptoms. However, it is unlikely that any single reason can explain the inconsistent results observed in the vortioxetine for GAD.

Although our meta-analysis did not demonstrate a statistically significant anxiolytic effect of vortioxetine, it did provide the information regarding drug tolerability. Our study found that there was no statistically significant difference for the discontinuation for any reason rates and discontinuation due to AEs rates among groups receiving either multiple doses (2.5, 5 and 10 mg/day) of vortioxetine compared with placebo, which is similar to the findings of a previous meta-analysis.

Thus, the vortioxetine doses were well tolerated, and were associated with similar discontinuation for any reason rates and discontinuation due to AEs rates compared with placebo.
discontinuation due to AEs rates when compared with the placebo.

GAD is associated with significant functional impairment in many areas, including social, occupational and mental consequences; and when combined with physical impairment, together they influence QoL. The effectiveness of GAD treatment for improving QoL and functional outcomes is potentially confounded by the bidirectional relationship of anxiety symptoms and QoL/functional impairment. This is the first meta-analysis to report prospective assessment of QoL/functional status impairment for vortioxetine in treatment of patients with GAD. Unfortunately, our meta-analysis of RCTs with patients with GAD showed no significant improvement in the aforementioned aspects after vortioxetine treatment compared with after treatment with the placebo. Our results are not consistent with those of a previous meta-analysis of the effect of vortioxetine treatment on overall functioning in patients with MDD. The meta-analysis, conducted by Florea et al, demonstrated that vortioxetine, in doses of 5–20mg for 6/8 weeks, improved overall functioning in patients with MDD. Relative to placebo, vortioxetine treatment, in doses of 10 and 20mg, demonstrated significant improvement in SDS total score and functional remission. However, the RCTs that were studied, in accordance with those included in our meta-analysis, used the optimal doses of 2.5–10mg. Thus, the reason for the lack of congruence between these two meta-analyses may be the difference in the optimal doses. Meanwhile, a recent meta-analysis of vortioxetine in working patients with GAD showed that vortioxetine benefits adult patients who are working and/or pursuing an education. Thus, future research should be directed to provide more RCTs, specifically targeted to individuals with GAD, in order to assess the efficacy of vortioxetine in a larger sample, as well as to define the best therapeutic dosage.

The results of our meta-analysis should be interpreted in light of the following potential limitations. First, we included only four RCTs, which may have influenced the reliability of the results. Second, the duration of each trial included in our meta-analysis was 8 weeks; this is an important issue because patients with GAD typically require long-term pharmacological treatment. We found only one study focusing on long-term relapse prevention, which showed no significant improvement of relapse prevention effect after long-term (maintenance) vortioxetine treatment for GAD compared with placebo. Furthermore, the systematic review, conducted by Perna et al, indicated that although some recent data support the long-term efficacy of vortioxetine for GAD and showed a favourable tolerability profile, the conflicting short-term studies and limited clinical experience recommend its use only as second-line therapy. In addition, owing to a limited number of studies included in our meta-analysis, we did not compare the onset time between the groups treated with multiple doses of vortioxetine and placebo. Finally, all the included trials were supported by the Takeda Pharmaceutical Company as part of a joint clinical development programme with H. Lundbeck A/S, which may have also influenced the results.

CONCLUSIONS

In summary, GAD is an illness that is characterised by severe anxiety symptoms, and by diminished functioning and QoL. The challenge for interventions is to achieve improvement of symptoms, but also to enhance patients’ functioning ability and QoL. Our comprehensive evaluation of efficacy, safety and impact on QoL provides a critical insight that may be useful for clinicians attempting to thoroughly understand the risk–benefit profiles of vortioxetine treatment for GAD. Vortioxetine did not significantly improve GAD symptoms and QoL as compared with the placebo; nevertheless, it was safe and well tolerated in this patient population. However, our results should be interpreted and translated into clinical practice with caution, owing to the limited number of RCTs included in the present meta-analysis.

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