Does preoperative rehabilitation for patients planning to undergo joint replacement surgery improve outcomes? A systematic review and meta-analysis of randomised controlled trials

Li Wang,1,2,3 Myeongjong Lee,4 Zhe Zhang,5 Jessica Moodie,1 Davy Cheng,1,6 Janet Martin1,6,7

ABSTRACT

Objectives: The clinical impact of preoperative physiotherapy on recovery after joint replacement remains controversial. This systematic review aimed to assess the clinical impact of prehabilitation before joint replacement.

Design: We searched PubMed, Embase and Cochrane CENTRAL up to November 2015 for randomised controlled trials comparing prehabilitation versus no prehabilitation before joint replacement surgery. Postoperative pain and function scores were converted to Western Ontario and McMaster Universities Osteoarthritis Index (WOMAC) pain and function subscales (0–100, high scores indicate worse outcome). Random effects meta-analysis was performed to calculate weighted mean differences (WMD, 95% CI), subgrouped by hip and knee surgery.

Primary and secondary outcomes: Postoperative pain and function scores, time to resume activities of daily living, quality of life, length of hospital stay, total cost, patient satisfaction, postoperative complications, any adverse events and discontinuations.

Results: Of 22 studies (1492 patients), 18 had high risk of bias. Prehabilitation slightly reduced pain scores within 4 weeks postoperatively (WMD −6.1 points, 95% CI −10.6 to −1.6 points, on a scale of 0–100), but differences did not remain beyond 4 weeks. Prehabilitation slightly improved WOMAC function score at 6–8 and 12 weeks (WMD −4.0, 95% CI −7.5 to −0.5), and time to climbing stairs (WMD −1.4 days, 95% CI −1.9 to −0.8 days), toilet use (−0.9 days, 95% CI −1.3 to −0.5 days) and chair use (WMD −1.2 days, 95% CI −1.7 to −0.8 days). Effects were similar for knee and hip surgery. Differences were not found for SF-36 scores, length of stay and total cost. Other outcomes of interest were inadequately reported.

Conclusions: Existing evidence suggests that prehabilitation may slightly improve early postoperative pain and function among patients undergoing joint replacement; however, effects remain too small and short-term to be considered clinically-important, and did not affect key outcomes of interest (ie, length of stay, quality of life, costs).

INTRODUCTION

Total joint replacement surgery is considered one of the most successful medical interventions, with significant pain relief and improvement in physical function and quality of life for patients with severe osteoarthritis. However, the recovery for a significant proportion of patients remains difficult and prolonged, and many never gain optimal functionality post-operatively. Therefore, researchers, clinicians and policymakers are still looking for better ways to improve the timelines and extent of recovery for patients undergoing total joint replacement.

Strengths and limitations of this study

- The methodology was rigorous, and included a comprehensive systematic search—without limits by language, date or publication status—that identified seven randomised controlled trials (RCTs) not included in any previous systematic reviews.
- We went beyond previous published systematic reviews by analysing the effect of prehabilitation by converting to a standardised measurement of Western Ontario and McMaster Universities Osteoarthritis Index (WOMAC) pain and function scores, and used different presentation methods to enhance interpretability and to improve ability to find potential signals in effect size through meta-analysis.
- This meta-analysis addressed all available clinically relevant outcomes, while previous reviews addressed only a few selected outcomes. Application of GRADE for rating quality of evidence provides improved context for interpreting the findings in light of inherent strengths and limitations of the included studies.
- Compliance with prehabilitation was problematic in some studies, and was not reported in a number of studies.
Physiotherapy has been delivered to patients for rehabilitation, traditionally after total joint replacement. However, preoperative physiotherapy and exercise programmes (also known as ‘prehabilitation’) have been proposed as a potential way to expedite recovery times and improve overall extent of recovery in patients planning to undergo joint replacement. One recently published review recommended preoperative exercise to maintain or improve function and pain,\(^6\) however, this recommendation was based on only one narrative systematic review with indeterminate effects.\(^6\) Although it seems intuitive that prehabilitation should improve patient disposition at the time of surgery, and may prepare patients for a better recovery after surgery, significant uncertainties remain about the overall balance of benefits and risks (and costs) for prehabilitation.

A number of related systematic reviews and meta-analyses have been published in the recent decade, with inconsistent methods and varied conclusions.\(^6\)–\(^13\) Two papers suggested that prehabilitation reduced pain for patients undergoing joint replacement,\(^8\)–\(^11\) and improved physical function for patients undergoing hip replacement surgery, but not knee replacement surgery,\(^6\) while the remainder suggested prehabilitation did not clearly demonstrate beneficial effects or were unable to provide definitive conclusions.\(^6\)–\(^9\),\(^10\)–\(^12\),\(^13\) Furthermore, significant methodological limitations or errors have been identified among the existing systematic reviews. Some only qualitatively summarised the results,\(^6\)–\(^9\),\(^11\)–\(^13\) two other meta-analyses\(^8\),\(^10\) are outdated, or mistakenly included some trials in which postoperative outcomes were not reported. Thus, we conducted an updated methodologically rigorous systematic review with meta-analysis to clarify whether evidence supports prehabilitation for patients planning to undergo joint replacement.

METHODS
Eligibility criteria
We systematically searched three databases up to 10 November 2015, including PubMed, Embase and Cochrane Central Register of Controlled Trials (CENTRAL). Eligible studies had to be randomised controlled trials comparing preoperative rehabilitation programmes (ie, prescribed and supervised exercises or physiotherapy with or without co-interventions such as education, nutritional counselling, acupuncture, transcutaneous electrical nerve stimulation, etc) versus no formal preoperative rehabilitation programmes, reporting at least one clinically-relevant outcome of interest during the postoperative period. Clinical outcomes of interest included postoperative pain scores (Visual Analogue Scale (VAS), or pain subcomponents of Western Ontario and McMaster Universities osteoarthritis index (WOMAC) or pain-related subdomains of other instruments), patient functionality (WOMAC function score, SF-36 physical functioning subdomain or other function-related instruments), time to resume activities of daily living (ADL), quality of life, patient satisfaction, infection, transfusions, stroke and death or overall postoperative complications. Resource-related outcomes of interest included hospital length of stay, readmissions and total hospital costs or total health system costs. Timeframes of relevance included in-hospital outcomes, as well as clinical or resource-related outcomes over the longer term, postoperatively.

Search terms included MeSH and keyword terms for exercise, prehabilitation, physiotherapy, physical therapy, activity, weight training, weight lifting, aquatic, swimming, strength training, endurance training, cycling, biking, kinesiotherapy, hydrotherapy, fitness, orthopaedic surgery and joint replacement and ‘random’\(^*\). Limitations were placed neither on date of publication nor on language. Detailed search strategies are provided in the Appendix.

Literature screening and data extraction
Two reviewers (ML, ZZ) independently screened the articles by title and abstract using the predetermined eligibility criteria. Any disagreements were resolved by a third reviewer (IW). The third reviewer (IW) also checked all the reference lists of existing systematic reviews or meta-analyses and other reviews for potentially additional eligible articles.

Two reviewers (ML, JM) independently assessed the risk of bias of the included trials using the methods recommended by Cochrane Collaboration,\(^14\) including random sequence generation, allocation concealment, missing or incomplete outcome data, and blinding of patients, study personnel and outcome assessors. Any discrepancies were resolved by the third reviewer (IW).

Standardised data extraction forms were developed to specify the study characteristics, patient characteristics and outcomes. Three reviewers (ML, ZZ and IW) extracted the data. Data were verified by a fourth reviewer (JM).

Statistical analysis
Meta-analysis was performed using the random effects model. For discrete outcomes, relative risk (RR) and 95% CIs (RR, 95% CI) were calculated. For continuous outcomes, for example, pain score and function score, weighted mean differences (WMD, 95% CI) were calculated after conversion to the Western Ontario and McMaster Universities osteoarthritis index (WOMAC) pain score (0–100) and WOMAC function score (0–100), in which a higher score indicates worse outcome. Sensitivity analysis was conducted by calculating standardised mean differences (SMD) and ratio of means (RoM).

If different pain scores were reported in one article (eg, WOMAC pain, SF-36 pain score), the WOMAC pain score was preferentially used. If WOMAC pain score was not reported, the pain score reported in the study was converted to a WOMAC pain score to allow for
comparison across studies, and to allow for estimation of overall effect size.\(^\text{15}\) If pain scores were reported at rest and during activity, the pain score during activity was preferentially used for analysis. If pain scores were reported during different types of activities, the largest change of pain score during the most active movement was used preferentially. If different function scores were reported, the WOMAC function score was used preferentially for analysis. In the absence of WOMAC function scores, the alternate function score provided in the study was converted to a WOMAC function score.\(^\text{15}\) Two studies\(^\text{16}-\text{17}\) only reported total scores of Hospital for Special Surgery Knee Rating System (HSSK) and WOMAC, respectively. Given the function score accounting for most of the total score and with similar trends of change over time as total score, we used the total score to replace the function score. To test whether this changed the effect size, sensitivity analysis was performed after removing the total scores from function measures to recalculate effect size.

To improve clinical relevance and interpretation of the results for postoperative pain and function improvement, we also converted continuous data from WOMAC pain score and WOMAC function score to a RR for achieving a ‘patient acceptable symptom state’ (defined as the number of patients achieving the threshold pain score or function score at which patients consider themselves ‘well’ or ‘satisfied’) derived from previous research.\(^\text{18}-\text{21}\) To calculate the RR, we assumed a normal distribution of WOMAC pain or function scores for the intervention and control groups, and used a threshold of 30 on the WOMAC 0–100 scale to represent the threshold for the patient acceptable symptom state. The proportion of patients in the intervention and control groups with WOMAC pain or function ≤30 was then calculated and combined across studies to derive a pooled RR.\(^\text{15}-\text{22}\) Finally, to further add to clinical applicability of the patient reported outcomes, we calculated the risk difference for the number of patients achieving this threshold of ≤30 per 100 patients, using the RR and median risk among the control groups in the included studies.\(^\text{23}\) Subsequently, sensitivity analyses were performed to explore whether using different thresholds (20 and 40) changed the conclusions, since our threshold of 30 represents a compromise of 20–40, suggested in previous studies of hip or knee surgery over the short term or long term. Heterogeneity was estimated using the \(\chi^2\) test and \(I^2\) statistic. Predefined subgroup analyses included separate analysis for hip and knee surgery patients, to test the existing hypothesis from a previous systematic review that prehabilitation improves postoperative pain and function more among patients undergoing hip replacement than among patients with knee replacement.\(^\text{8}\) Publication bias was explored using both visual inspection of funnel plots and Egger’s test only when there were at least 10 studies included in the meta-analysis.\(^\text{14}\)

GRADE methodology was used to summarise certainty in estimates of effect (quality of evidence) in the critically important outcomes for decision-making,\(^\text{25–29}\) including WOMAC pain scores and function scores from early follow-up to 24 weeks after surgery.

RESULTS

Studies identified

Figure 1 outlines study inclusion and exclusion. A total of 399 titles and abstracts were screened for inclusion, of which 110 studies were collected in full-text for review. Of these, 88 were excluded for the following reasons: no prehabilitation arm (ie, education only or postoperative rehabilitation only, n=46), not randomised (n=21), duplicate studies (n=4), no postoperative outcomes data (n=9), no outcome of interest (n=2), conference abstracts (n=3) and protocol only (n=3). In total, 22 randomised controlled trials (1492 patients) of prehabilitation versus no prehabilitation met the inclusion criteria. Twenty studies provided usable data for the meta-analysis, and 22 studies contributed qualitative or quantitative data.

Description of included studies

Among the 22 included studies, 8 studies were of patients undergoing total hip replacement;\(^\text{17}-\text{30–36}\) 12 studies included patients undergoing total knee replacement\(^\text{16,37–47}\) and two studies included either hip or knee replacement.\(^\text{48,49}\) Most studies were conducted in developed countries (North American and Europe), except for three in developing countries (Serbia,\(^\text{34}\) Thailand\(^\text{45}\) and Turkey\(^\text{32}\)). The median sample size of included studies was 54, ranging from 21 to 165 patients. Mean age ranged from 51 to 76 years (table 1).

Nine studies compared physiotherapist-supervised exercise plus home exercise versus no intervention, or usual care.\(^\text{17,35–36,38,40–42,44,46}\) Five compared physiotherapist-supervised exercise versus no intervention, or usual care.\(^\text{16,35,37,39,40}\) Two studies compared home exercise only versus no intervention.\(^\text{30,45}\) Three studies compared physiotherapist-supervised exercise plus education versus no intervention.\(^\text{31,32,34}\) One each compared physiotherapist-supervised exercise plus education versus education,\(^\text{48}\) kinesiologist-supervised exercise versus placebo (kinesiologist supervised upper body exercise)\(^\text{45}\) and physiotherapist-supervised exercise plus home exercise versus education plus home exercise,\(^\text{47}\) respectively, (see online supplementary table S1).

Risk of bias

Among the 22 trials, adequate sequence generation was reported in 17 trials\(^\text{16,30–33,36–43,45–48}\) and allocation concealment was reported in 8 trials.\(^\text{30,35,37–39,42,43,47,49}\) The patients were blinded in one study,\(^\text{39}\) healthcare providers were blinded in three studies\(^\text{38,39,44}\) and outcome assessors were blinded in 12 studies.\(^\text{38,39,44}\) Seventeen studies\(^\text{17,30–34,36–40,42,43,46–49}\) reported loss to follow-up, ranging from 1.7% to 65.3%; among which the
proportion of loss to follow-up was more than 15% in 10 studies.\textsuperscript{30 34 36–39 42 43 47 48} Ten of 17 studies with incomplete data used intention to treat analysis.\textsuperscript{30 33 36 37 39 42 43 47–49} Overall, 4 of the 22 included trials were rated as having a low risk of bias\textsuperscript{30 33 37 39} and 18 trials were rated as having a high risk of bias (table 2).

Online supplementary table S2 qualitatively summarises the major findings of included studies. In total, 22 studies described at least one clinical or resource-related benefit for prehabilitation versus control, and 18 studies\textsuperscript{16 30–37 39 40 42–48} described no significant improvement for prehabilitation versus control. We conducted meta-analysis for pain scores, function scores, SF-36 PCS and MCS, hospital length of stay and total costs, based on data availability.

\textbf{Postoperative pain}

Fifteen trials with 18 comparisons and 1046 patients reported postoperative pain scores using different instruments, namely, WOMAC,\textsuperscript{31 37 40 42 43 45 48} VAS,\textsuperscript{32 34 36 44 47} Knee injury and Osteoarthritis Outcome Score (KOOS)/Hip disability and Osteoarthritis Outcome Score (HOOS)\textsuperscript{49} and 10-graded scale.\textsuperscript{46} Only two trials\textsuperscript{31 49} reported significant improvement in pain at early follow-up (≤3 months), including VAS pain at 3 months,\textsuperscript{31} and KOOS/HOOS pain at 6 weeks postoperatively, but not significant at 3 months.\textsuperscript{49}

After converting to WOMAC pain 0–100, prehabilitation significantly reduced postoperative pain at 4 weeks or less; however, the reduction of pain was clinically nominal (4 trials, 213 patients, WMD $-6.1$, 95\% CI $-10.6$ to $-1.6, \text{figure 2}$, GRADE: low certainty in estimates, tables 3 and 4). Differences in WOMAC pain scores after 4 weeks were no longer statistically significant for prehabilitation versus control (WOMAC pain score at 6–8 weeks, 5 trials, 488 patients, WMD $-1.4$, 95\% CI $-5.5$ to $+2.6$; at 12 weeks, 10 trials, 806 patients, WMD $-2.9$, 95\% CI $-6.2$ to $+0.3$; at 24 weeks, 3 trials, 247 patients, $-2.5$, 95\% CI $-5.6$ to $+0.6$; at 1 year, 1 trial, 109 patients, WMD $-2.0$, 95\% CI $-7.5$ to $+3.5$; GRADE: low to moderate certainty in estimates, tables 3 and 4).

When expressed as a RR, patients undergoing prehabilitation were more likely to achieve the acceptable pain state (WOMAC pain score ≤30) with RR 1.09. When expressed as an absolute risk difference, 3.9\% more patients with prehabilitation achieved the acceptable pain state (WOMAC pain score ≤30) than patients without prehabilitation at 4 weeks (online supplementary table S3). However, this small difference would be considered clinically nominal.\textsuperscript{18–21}

\textbf{Postoperative function}

Of 16 trials reporting on postoperative function, only four reported significant improvement in
function, including higher hip external rotation \(^{31,49}\) or higher flexion range of motion scores, WOMAC physical function and total score, \(^{17}\) and less time to reach 90\(^\circ\) of knee flexion \(^{41}\) and greater improvement in ADL \(^{49}\) after surgery.

Sixteen trials (1118 patients) reported postoperative function scores, using different instruments, namely, WOMAC, \(^{17,31,41,49}\) Harris hip score, \(^{32,34}\) SF-36 physical component summary (PCS), \(^{30}\) SF-36 physical functioning score, \(^{38}\) HSSK score, \(^{16}\) HOOS function in daily living, \(^{36}\) and KOOS/HOOS ADL. \(^{49}\) After converting function scores to WOMAC function score (0–100), the difference was slightly improved (but numerically small on a scale of 0–100) with prehabilitation versus no prehabilitation at early follow-up (WOMAC function score at 6–8 weeks; 5 trials, 488 patients, WMD −3.9, 95% CI −7.6 to −0.3, RR=1.10, figure 3, GRADE: moderate certainty in estimates, tables 3 and 4), and at 12 weeks (12 trials, 836 patients, WMD −4.0, 95% CI −7.5 to −0.5, RR=1.02, figure 4, GRADE: very low certainty in estimates, tables 3 and 4). No significant difference for WOMAC function score was found after 12 weeks (at 24 weeks, 5 trials, 345 patients, WMD −0.5, 95% CI −5.8 to +4.7; at 1 year, 6 trials, 296 patients, WMD −0.6, 95% CI −2.6 to +1.5, GRADE: low certainty in estimates, tables 3 and 4).

When expressed as a RR for achieving the acceptable threshold, the relative increases were small (RR 1.10 at 6–8 weeks; 1.02 at 12 weeks). When expressed as an absolute difference in likelihood of achieving the acceptable threshold, the differences ranged from 1.3% to 5.4% more patients achieving a WOMAC function score ≤30 at 6–8 weeks and 12 weeks, respectively, (online supplementary table S3).

### Resumption of ADL

Resumption of ADL was rarely reported. In the two studies (99 patients) \(^{32,34}\) that reported ADL, meta-analysis suggested significantly earlier resumption of activities, including climbing stairs (2 trials, 99 patients, WMD −1.4 days, 95% CI −1.9 to −0.8 days), use of toilet (2 trials, 99 patients, −0.9-day, 95% CI −1.3 to −0.5 days) and use of chair (2 trials, 99 patients, −1.2 days, 95% CI −1.7 to −0.8 days), but not for time to first day of walking (2 trials, 99 patients, −0.2-day, 95% CI −0.4 to +0.0-day, (table 3). However, based on the total time-course of recovery, the difference was small.

### Quality of life

In nine studies, significant differences were not found for quality of life as reported in the SF-36, \(^{30,37,39,42,43}\) Quality of Well Being instrument, \(^{16}\) HOOS Hip-related quality of life, \(^{36}\) KOOS/HOOS Quality of Life subscale, \(^{49}\) and Patient Specific Complaints (PSC) questionnaire. \(^{33,36}\) However, three trials reported a significant difference in quality of life scores within the first month, \(^{31,33,36}\) including improved physical function scores or physical composite scores as reported in the SF-36 \(^{31}\) or EuroQol five Dimension Health Questionnaire (EQ5D). \(^{49}\)
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LTFU, lost to follow-up.
Nevertheless, the numeric differences were small and the significance disappeared at 3 months. Three studies including 149 patients reported SF-36 Physical Component Summary (SF-36 PCS) and Mental Component Summary (SF-36 MCS). Meta-analysis of SF-36 PCS and MCS did not detect significant differences at any time-point (from 6 weeks to 1 year, table 3).

Length of hospital stay and total cost
Only one out of 10 studies found a significant reduction in hospital length of stay (mean difference = -0.8-day). Meta-analysis of these studies did not detect significant differences in hospital length of stay for prehabilitation versus control (7 studies, 507 patients, WMD = -0.3 days, 95% CI = -0.8 to +0.1 days, figure 5).

Of the few studies that reported on costs, none reported significant reduction of overall costs with prehabilitation, but one described significantly increased physiotherapy costs with prehabilitation (mean difference = £136.5). Even when total costs were converted to Canadian dollars and combined through meta-analysis, the results did not differ for prehabilitation versus none (2 studies, 242 patients, WMD + $0.5, 95% CI = $384 to +$393).

Other outcomes
Other outcomes of interest, including patient satisfaction, stroke, cardiovascular events and readmissions, were inadequately reported for meta-analysis. Adverse events and discontinuations were rarely reported within the studies; however, in at least one study, there was concern about increased cardiovascular events and stroke, and poorer SF-36 general health, energy and mental health among the withdrawn patients although the author stated no evidence that study withdrawal varied by group. In some studies, there were reports of patient withdrawals due to adverse events. Some studies reported no significant postoperative complications between groups, no serious adverse events or no adverse events.

Subgroup analysis and sensitivity analysis
Effect sizes were similar between hip and knee replacement subgroups for WOMAC pain and function scores (see online supplementary table S3), as indicated by non-significant p values for interaction.

Sensitivity analysis using SMD (instead of WMD), RoM and different thresholds for defining patient acceptable symptom state (20 and 40, instead of 30), and replacing function sub-score with total score, did not significantly change the results (see online supplementary tables S4 and S5).

While publication bias was not indicated for pain score, asymmetric funnel plots indicated the possibility of publication bias for function scores (see online supplementary figures S1 and S2, table 2).

INTERPRETATION

Main findings
Existing evidence from 22 randomised controlled trials suggests that prehabilitation for patients planning to...
undergo joint replacement does not affect postoperative pain and function to a degree that would be considered clinically relevant; however, this is based on studies with significant limitations, providing very low certainty in estimates. While some differences reached statistical significance, the effects are too small to be considered clinically important (ie, an improvement of a few points on a scale of 0–100 is likely clinically irrelevant, and undetectable to patients). Our analysis shows that prehabilitation reduced WOMAC pain score by 6 with 95% CI (−10.6 to −1.6) within 4 weeks, and with no difference remaining beyond 4 weeks, which is generally smaller than the minimal clinically important improvement of at least 9.7 at 6 weeks,51,52 even when the most optimistic extremes of the CIs are considered in our analysis. Even when the ‘patient accepted pain state’ was defined as achieving 30 in WOMAC pain subscale 0–100, there was only an absolute increase of 3.9% of patients achieving this threshold. Similarly for function improvement, prehabilitation improved early function by 3.9–4.0 points on the WOMAC function subscale 0–100, which is much smaller than the threshold of minimally important difference, ranged from 7.9 to 25.9,51–56 and only 1.3% to 5.4% more patients reached a WOMAC function score ≤30. Although prehabilitation allowed patients to resume ADL 0.9–1.4 days earlier than no formal prehabilitation, the difference is trivial, and, importantly, very few studies reported on this time point (ie, 2 of 22 studies), which prevents definitive conclusions. Similarly, for the outcome of length of stay, there was no difference between groups, and if statistical significance had been achieved, the difference would have been only 0.3 days, which is a minimal difference. Jurisdictions considering implementation (or continuation) of prehabilitation services should consider whether resources could be better spent elsewhere on interventions of proven clinical benefit. Until sufficient evidence accrues to definitively conclude that prehabilitation provides meaningful benefit, investment in prehabilitation does not represent the best use of limited resources in a healthcare system where other opportunities with proven benefits could be funded instead.

Relation to prior reviews

Similar to this meta-analysis, most previous meta-analyses10,11 and systematic reviews7,9,12 suggested that the impact of prehabilitation has not been proven by the existing evidence. In contrast to our analysis, Gill and McBurney8 suggested that exercise-based interventions reduce pain and improve physical function for people awaiting hip replacement surgery, but not those awaiting knee replacement surgery. It is notable that there were some limitations in the analysis by Gill and McBurney,8 wherein some included trials did not report...
Table 4  GRADE evidence profile: prehabilitation versus no formal prehabilitation for total joint replacement

<table>
<thead>
<tr>
<th>Participants (studies)</th>
<th>Follow-up</th>
<th>Risk of bias</th>
<th>Inconsistency</th>
<th>Indirectness</th>
<th>Imprecision</th>
<th>Publication bias</th>
<th>Overall quality of evidence</th>
<th>Relative effect or WMD (95% CI)</th>
<th>Anticipated absolute effects</th>
<th>Risk difference with prehabilitation (95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Pain score at 4 weeks or less, measured with: WOMAC pain subscale 0–100; lower values indicate less pain</td>
<td>213 (4 studies)</td>
<td>Serious risk of bias¹</td>
<td>Serious inconsistency; p value on test for heterogeneity 0.08, I²=55%</td>
<td>No serious indirectness</td>
<td>No serious imprecision</td>
<td>Uncertain (only 4 studies)</td>
<td>⊕⊕⊝⊝ LOW due to risk of bias and inconsistency</td>
<td>WMD −6.1 (−10.6 to −1.6)</td>
<td>43.8% patients achieved acceptable pain state of ≤30 on WOMAC pain (0–100) scale</td>
<td>3.9% more patients achieving acceptable pain state of ≤30 on WOMAC pain (0–100) scale</td>
</tr>
<tr>
<td>Pain score at 6–8 weeks, measured with: WOMAC pain subscale 0–100; lower values indicate less pain</td>
<td>488 (5 studies)</td>
<td>Serious risk of bias*</td>
<td>No serious inconsistency; p value on test for heterogeneity 0.31, I²=16%</td>
<td>No serious indirectness</td>
<td>No serious imprecision†</td>
<td>Uncertain (only 5 studies)</td>
<td>⊕⊕⊕ MODERATE due to risk of bias</td>
<td>WMD −1.4 (−5.5 to +2.6)</td>
<td>62.2% patients achieved acceptable pain state of ≤30 on WOMAC pain (0–100) scale</td>
<td>0% more patients achieved acceptable pain state of ≤30 on WOMAC pain (0–100) scale</td>
</tr>
<tr>
<td>Pain score at 12 weeks, measured with: WOMAC pain subscale 0–100; lower values indicate less pain</td>
<td>806 (10 studies)</td>
<td>Serious risk of bias*</td>
<td>Serious inconsistency; p value on test for heterogeneity 0.05, I²=46%</td>
<td>No serious indirectness</td>
<td>No serious imprecision</td>
<td>Undetected; Egger’s test p=0.35</td>
<td>⊕⊕⊝⊝ LOW due to risk of bias and inconsistency</td>
<td>WMD −2.9 (−6.2 to +2.8)</td>
<td>60.9% patients achieved acceptable pain state of ≤30 on WOMAC pain scale (0–100)</td>
<td>1.2% more patients achieved acceptable pain state of ≤30 on WOMAC pain scale (0–100)</td>
</tr>
<tr>
<td>Pain score at 24 weeks, measured with: WOMAC pain subscale 0–100; lower values indicate less pain</td>
<td>247 (3 studies)</td>
<td>Serious risk of bias*</td>
<td>No serious inconsistency; p value on test for heterogeneity 0.22, I²=33%</td>
<td>No serious indirectness</td>
<td>No serious imprecision</td>
<td>Uncertain (only 3 studies)</td>
<td>⊕⊕ MODERATE due to risk of bias</td>
<td>WMD −2.5 (−6.6 to +0.6)</td>
<td>98% patients achieved acceptable pain state of ≤30 on WOMAC pain scale (0–100)</td>
<td>0% patients achieved acceptable pain state of ≤30 on WOMAC pain scale (0–100)</td>
</tr>
<tr>
<td>Function score at 4 weeks or less, measured with: WOMAC function subscale 0–100; lower values indicate better function</td>
<td>257 (5 studies)</td>
<td>Serious risk of bias*</td>
<td>Serious inconsistency; p value on test for heterogeneity &lt;0.001, I²=79%</td>
<td>No serious indirectness</td>
<td>No serious imprecision†</td>
<td>Uncertain (only 5 studies)</td>
<td>⊕⊕⊝⊝ LOW due to risk of bias and inconsistency</td>
<td>WMD −3.6 (−7.7 to +0.5)</td>
<td>26.8% patients achieved acceptable function state ≤30 on WOMAC function scale (0–100)</td>
<td>6.2% more patients achieved acceptable function state ≤30 on WOMAC function scale (0–100)</td>
</tr>
</tbody>
</table>

Continued
<table>
<thead>
<tr>
<th>Participants assessment</th>
<th>Summary of findings</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Function score at 6–8 weeks, measured with: WOMAC function subscale 0–100; lower values indicate better function</strong></td>
<td></td>
</tr>
<tr>
<td>Function score</td>
<td>Relative effect or WMD (95% CI)</td>
</tr>
<tr>
<td>Follow-up</td>
<td>Risk of bias</td>
</tr>
<tr>
<td>488 (5 studies)</td>
<td>Serious risk of bias*</td>
</tr>
<tr>
<td>836 (12 studies)</td>
<td>Serious risk of bias*</td>
</tr>
<tr>
<td>345 (7 studies)</td>
<td>Serious risk of bias*</td>
</tr>
</tbody>
</table>

*None of the studies in the meta-analyses blinded patients and only 2 studies blinded the care providers.
†We did not rate down due to imprecision although 95% CI includes no effect because either extreme of the 95% CI is too small to be clinically important difference.

WMD, weighted mean differences; WOMAC, Western Ontario and McMaster Universities Osteoarthritis Index.
**Figure 3**  Function score at 6–8 weeks (converted to WOMAC function subscale 0–100) for prehabilitation versus no prehabilitation in joint replacement surgery. WOMAC, Western Ontario and McMaster Universities Osteoarthritis Index.

**Figure 4**  Function score at 12 weeks (converted to WOMAC function subscale 0–100) for prehabilitation versus no prehabilitation in joint replacement surgery. ADL, activities of daily living; WOMAC, Western Ontario and McMaster Universities Osteoarthritis Index. KOOS, Knee injury and Osteoarthritis Outcome Score; HOOS, Hip disability and Osteoarthritis Outcome Score.
if the patients underwent surgery after the intervention, and/or failed to report postoperative outcomes, and one included trial allocated patients based on the geographic availability, which may have introduced selection bias and unit of measurement errors. Furthermore, a total of nine relevant trials were not included in Gill (2013).

**Strengths and limitations**

Strengths of this review include rigorous methodology, including the comprehensive systematic search without limits by language, date or publication status, which identified seven randomised controlled trials not included in any previous systematic reviews. Furthermore, we analysed the effect of prehabilitation by converting to a standardised measurement of WOMAC pain and function scores, and used different presentation methods to enhance interpretability and to improve ability to find potential signals in effect size through meta-analysis, which is beyond what other systematic reviews published. In addition, this meta-analysis addressed all available clinically relevant outcomes, while previous reviews addressed only a few selected outcomes. Application of GRADE for rating quality of evidence provides improved context for interpreting the findings in light of inherent strengths and limitations of the included studies.

There were a number of specific limitations in the existing clinical trials comparing prehabilitation with control. The most significant limitation is the lack of large randomised controlled trials conducted in this area. Included studies were small (median 81 patient, ranging from 21 to 165), of relatively short duration of follow-up (median 3 months, ranging from 4 weeks to 1 year), and many of them provided inadequate description of the frequency, intensity and duration of prehabilitation provided. Definitions for prehabilitation and for outcomes measurements were heterogeneous across studies. Patient compliance with prehabilitation was reported as 75% in three studies, and 90% or greater in seven studies, and was not reported in the remainder of the studies. Most studies provided an inadequate description of the components of the prehabilitation programmes provided, and few described the fidelity of programme implementation. Future studies in this area should follow current guidelines for intervention description (TIDieR checklist) to enable transparent evaluation and replication of programmes. In a number of studies, cointerventions were provided in the prehabilitation (eg, education) and, in some cases, these cointerventions were not provided in the control group. Nevertheless, this would likely provide an overestimate of the potential benefit for prehabilitation; despite this potential positive bias, still no differences were found for prehabilitation. Considered...

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**Figure 5** Hospital length of stay (days) for prehabilitation versus no prehabilitation in joint replacement surgery.

<table>
<thead>
<tr>
<th>Study name</th>
<th>N, mean</th>
<th>N, mean</th>
<th>%</th>
<th>WMD (95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Beaupre 2004</td>
<td>-1.50 (-3.27, 0.27)</td>
<td>55, 10.2 (4.5)</td>
<td>60, 11.7 (5.2)</td>
<td>5.26</td>
</tr>
<tr>
<td>Bitterli 2011</td>
<td>0.00 (-1.12, 1.12)</td>
<td>41, 14.6 (2.5)</td>
<td>39, 14.6 (2.5)</td>
<td>13.22</td>
</tr>
<tr>
<td>D’Lima 1996</td>
<td>0.21 (-0.86, 1.28)</td>
<td>10, 6.29 (1)</td>
<td>5, 6.08 (1)</td>
<td>14.35</td>
</tr>
<tr>
<td>D’Lima 1996</td>
<td>0.02 (-1.05, 1.09)</td>
<td>10, 6.1 (1)</td>
<td>5, 6.08 (1)</td>
<td>14.35</td>
</tr>
<tr>
<td>Hoogeboom 2010</td>
<td>0.00 (-2.67, 2.67)</td>
<td>10, 6 (4.25)</td>
<td>10, 6.75 (2.3)</td>
<td>2.31</td>
</tr>
<tr>
<td>Matassi 2012</td>
<td>-0.80 (-1.58, -0.02)</td>
<td>61, 9.1 (2.1)</td>
<td>61, 9.9 (2.3)</td>
<td>27.08</td>
</tr>
<tr>
<td>Mitchell 2005</td>
<td>-0.40 (-1.59, 0.79)</td>
<td>56, 9.2 (3.4)</td>
<td>55, 9.6 (3)</td>
<td>11.64</td>
</tr>
<tr>
<td>Oosting 2012</td>
<td>-0.30 (-1.48, 0.88)</td>
<td>14, 5.1 (1)</td>
<td>15, 5.4 (2.1)</td>
<td>11.78</td>
</tr>
<tr>
<td>Overall (I-squared = 0.0%, p = 0.879)</td>
<td>-0.34 (-0.75, 0.06)</td>
<td>257</td>
<td>250</td>
<td>100.00</td>
</tr>
</tbody>
</table>
together, the heterogeneity of the included studies in types of prehabilitation programmes, control group interventions, compliance and fidelity within the programmes, and systematic differences in the study population, likely impacted the ability to detect differences, if any exist. Although we performed subgroup analysis for hip versus knee replacement surgery, this failed to explain the heterogeneity across studies. Owing to the limited numbers of studies, meta-analysis was not performed for the effect of different types of prehabilitation (eg, exercise only vs exercise plus education). Publication bias was not detected; however, the methodological quality of included studies is very low, which was the major reason that we downgraded the overall quality of evidence. The high risk of bias, combined with the selective reporting of important outcomes across the studies (eg, only two studies reported time to return to ADL and total costs) precludes definitive conclusions, despite at least 22 randomised controlled trials being conducted.

CONCLUSION

Existing evidence suggests that, in patients undergoing joint replacement, the effect of prehabilitation (exercise/physiotherapy programmes in the months prior to surgery) on pain and function are too small to be considered clinically-important and were not robust over time. Prehabilitation did not result in clinically important (or statistically significant) differences in most measures of patient recovery, quality of life, length of stay and costs. Future research of sufficient power to measure clinically-relevant outcomes is required to identify which, if any, form of prehabilitation achieves better outcomes than in these trials. Jurisdictions considering implementation of prehabilitation services should consider whether resources could be better spent elsewhere on interventions of proven clinical benefit.

Author affiliations

1Centre for Medical Evidence, Decision Integrity and Clinical Impact (MEDICI), University of Western Ontario, London, Ontario, Canada
2Chinese Cochrane Centre, West China Hospital, Sichuan University, Chengdu, China
3Michael G DeGroote Institute for Pain Research and Care, McMaster University, Hamilton, Ontario, Canada
4Department of Anesthesiology and Pain Medicine, Konkuk University School of Medicine, Chungju, South Korea
5Department of Anesthesiology, Fuwai Hospital, Chinese Academy of Medical Sciences and Peking Union Medical College, Beijing, China
6Department of Anesthesiology & Perioperative Medicine, University of Western Ontario, London, Ontario, Canada
7Department of Epidemiology & Biostatistics, University of Western Ontario, London, Ontario, Canada

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Contributors

LW contributed to the study conception and design, literature screening, acquisition of data, analysis and interpretation of data, drafting of the manuscript and revision based on the comments of the co-authors. ML and ZZ participated in the literature screening, data acquisition and critical revision of the manuscript. JM performed the literature search, article retrieval and data acquisition. DC contributed to study conception and design, and critical revision of the manuscript. JM guided the methodology, and contributed to the study conception and design, data checking and interpretation, drafting and critical revision of the manuscript. All the authors approved the version submitted for publication and agreed to act as guarantors of the work.

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Competing interests

None declared.

Provenance and peer review

Not commissioned; externally peer reviewed.

Data sharing statement

No additional data are available.

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