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Complete List of Authors:	Giorda, Carlo; ASL TORINO 5, Diabetes and metabolism Unit Picariello, Roberta; ASL TO3, Regione Piemonte, Epidemiology Unit, Tartaglino, Barbara; Chaira Medica Association, Marafetti, Lisa; ASL torino 5, Diabetes and Metabolism Unit Di Noi, Fabiana; ASL torino 5, Diabetes and Metabolism Unit Alessiato, Annalisa; ASL TORINO 5, Diabetes and metabolism Unit Costa, Giuseppe; ASL TO3, Regione Piemonte, Epidemiology Unit, Gnavi, Roberto; ASL TO3, Regione Piemonte, Epidemiology Unit,
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Hospitalisation for heart failure and mortality associated with dipeptidyl peptidase 4 (DPP-4) inhibitor use in an unselected population of subjects with type 2 diabetes CARLO B. GIORDA¹, MD, ROBERTA PICARIELLO², BSc, BARBARA TARTAGLINO³ PhD, LISA MARAFETTI ¹ MD, FABIANA DI NOI¹ MD, ANNALISA ALESSIATO¹, PhD, GIUSEPPE COSTA^{2,4}, MD and ROBERTO GNAVI², MD

From:

¹ Metabolism and Diabetes Unit, ASL TO5, Regione Piemonte, Chieri, Italy; ² Epidemiology Unit, ASL TO3, Regione Piemonte, Grugliasco, Italy; ³ Chaira Medica Association, Chieri, Italy; ⁴ Department of Clinical and Biological Sciences, University of Torino, Torino, Italy.

Short Title: Heart failure and mortality in DPP-4i users

.nissio. Correspondence, phone calls and e-mails about the submission should be directed to me at the following address and phone number

Carlo Giorda, M.D.

Diabetologia

ASL TO5

Via De Maria, 1

10023 Chieri (To)

Italy

Tel. + 39 011 94293312

e-mail: giorda.carlobruno@aslto5.piemonte.it

giordaca@tin.it

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ABSTRACT

Objective. The SAVOR TIMI–53 study reported a significant increase in the risk of hospitalization for heart failure (HF) in patients treated with a DPP-4 inhibitor (DPP-4i) in comparison with placebo. A recent case-control study in part confirmed this risk signal. Our aim was to compare the occurrence of HF in relation to DPP-4i use versus any antidiabetic treatment.

Design. Population-based matched case-control study conducted using administrative data.

Setting. The Italian Region of Piedmont (4.4 million inhabitants).

Participants From a database of 282,000 patients treated with antidiabetic drugs, we identified 14,613 hospitalizations for HF, 7212 incident cases, and 1727 hospital re-admissions between 2008 and 2012; each case was matched for gender, age and antidiabetic therapy with ten controls; cases and controls were compared for exposure to DPP-4i.

Outcome measures Odds ratios (OR) and 95% confidence intervals were calculated by fitting a conditional logistic model. All analyses were adjusted for available risk factors for HF.

Results. We found no increased risk of hospitalisation for HF associated with the use of DPP-4i (OR for admission for HF 1.00 [0.94-1.07], incident HF1.01 [0.92-1.11], recurrent HF 1.02 [0.84-1.22]). All-cause mortality was 6% lower in DPP-4i users, whereas insulin users showed an excess of risk for any type of hospital admission (19%) and death (20%).

Conclusions. Our findings suggest that, in an unselected population of diabetic patients, the use of DPP-4i is not associated with an increased risk of HF. The favourable impact on all-cause mortality should be viewed with caution and also other explanations investigated.

Keywords: DPP-4 inhibitors, heart failure hospitalization, all-cause mortality, safety in diabetes treatment, case-control study.

Short title: A nested case-control study in a European setting

Strengths and limitations of this study

This study suggest that, in an unselected population of diabetic patients, the use of DPP-4i is not associated with an increased risk of HF

The study population was representative of the real type 2 diabetes population seen in Europe, without any selection based on insurance claims or age cut-off which distinguish it from other published studies.

Hospitalization for HF was evaluated as admission, incidence or recurrence and all five drugs currently available on the market were included

The main limitation is that weak associations between DPP-4i use and heart failure cannot be ruled out; but if they do exist, they are not so large.

While looking forward to the results of ongoing trials, practitioners can be reassured that the unexpected association reported in the SAVOR TIMI study has not been confirmed in the real world.

Introduction

DPP-4 inhibitors (DPP-4i), or gliptins, are oral agents that delay the catabolism of native GLP-1 by inhibiting the endogenous enzyme dipeptidyl peptidase 4 (DPP-4), thus extending the life of native GLP-1. They have attracted growing interest as first line therapies for type 2 diabetes largely because they are effective in controlling HbA1c while reducing the risk of hypoglicaemia and weight gain. Currently available data suggest that they also exert a protective effect on cardiovascular risk (1). In the first published trial with cardiovascular endpoints (2), saxagliptin did not increase the risk of a composite of nonfatal myocardial infarction (MI), nonfatal stroke or cardiovascular death, thereby meeting the primary safety objective. However, as compared with placebo, saxagliptin was unexpectedly associated with a 27% excess risk of hospitalization for heart failure (HF) (overall hazard ratio [HR] 1.27 [95% confidence interval [CI] 1.07-1.51]). Detailed sub-analyses (3,4) revealed that the absolute difference between groups was mainly seen during the first 6 months of therapy and that there was no increased risk of death due to heart failure. No clinically relevant factors predictive of increased relative risk with saxagliptin treatment could be definitively identified (5). A subsequent meta-analysis of all registrative trials with DPP-4i failed to rule out an association between the use of this class of drugs and an increased risk of heart failure (6).

Evidence for an association between situaliptin use and hospitalization for heart failure in some specific conditions comes from a recent observational study by Weir et al. (7) in which the authors

reported that in a case-control study, based on an administrative database of U.S. middle-aged adults with type 2 diabetes, treatment with sitagliptin was associated with an increased probability of hospitalization for heart failure only among patients with pre-existing heart failure.

This safety issue is relevant for diabetes care because the rates of heart failure and hospitalization are higher in patients with type 2 diabetes than in the non-diabetic population, regardless of treatment ((8,9,10)). Further complicating the question is that several drugs commonly used in the treatment of type 2 diabetes have been suspected to increase the risk of heart failure (11).

In light of the concerns over a possible HF risk associated with DPP-4i, we thought it useful to perform a matched case-control study to explore whether the same increased risk as seen in the SAVOR-TIMI study could be detectable in the real world and in an unselected population. To the best of our knowledge, no such analysis has been performed on data from European administrative databases which, by virtue of the universalistic care of European national health systems, encompass the whole population and include all types of available DPP-4i.

Methods

Study design and participants

We conducted a population-based nested case-control study using regional administrative data from Piedmont (population about 4.4 million). The population is covered by an automated system of databases containing the records of all drugs dispensed from all regional pharmacies and hospital discharges reimbursed by the Italian National Health System. These archives can be linked together by a unique anonymous identifier that is encrypted to protect the patient's privacy. Because this automated system is anonymous, ethical committee approval and informed consent for this study were not required.

Procedures

We extracted information from the regional drug prescription database for individuals aged 56 years or older who were dispensed at least one dose of any drug to treat diabetes between January 1, 2009 and December 31, 2013 (DPP-4i were not available in Italy before 2008). Only Piedmont residents were included. To minimise the chance of inclusion of patients with type 1 diabetes, we linked the database to the regional hospital discharge database, which contains the records of all hospital admissions between 1995 and 2013. Excluded were individuals with an International Classification of Diseases, 9th Edition, Clinical Modification (ICD-9-CM) code for type 1 diabetes mellitus (250.x1 or 250.x3). Furthermore, as glitazones (TZD) increase the risk of heart failure, all patients who had received a prescription for TZD during the study period were excluded. *Selection of cases*

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We used four different definitions for cases. First, we identified all patients who had at least one discharge for heart failure, defined as ICD-9-CM codes 402.01, 402.11, 402.91, 425.4, 425.5, 425.9, 428 or 518.4 as the primary discharge diagnosis at any time after the first exposure to antidiabetic drugs (i.e., date of dispensation). For patients with more than one discharge for heart failure, we only included the first episode (i.e., the hospital admission closest to January 1 2009). Second, we identified "incident" cases of heart failure, defined as patients discharged with a diagnosis of heart failure (defined as above) during the study period, without a previous hospitalization for heart failure in the discharge diagnosis (either main or secondary) during the previous 60 months. Third, similarly to the study by Weir et al. (7) we followed up incident cases (defined as above) to identify "first re-hospitalisations" of those patients who had been admitted to hospital for a diagnosis of heart failure. Finally, we considered as cases all deaths (of any cause) that occurred in the population during the study period. Also included were Piedmont residents discharged from any hospital located outside Piedmont because information on exposure to dispensed drugs is available for all patients residing in the region. Similarly, we included the deaths of Piedmont residents wherever they occurred in Italy.

Selection of controls

To identify controls, we randomly selected ten controls from the same population source for each case, matched for year of birth (within a 5-year age band), sex, and year of first exposure to antidiabetic drugs. Controls were selected one subject at a time with replacement. The process was repeated for each outcome. Matching was done by the study statistician (RP) with the use of an automated computer program.

Exposure to DPP-4 inhibitors

We used the regional drug database to identify cases and controls who had been prescribed DPP-4i at any time in the 6 months before the hospital admission date. We used the hospital admission date of cases to calculate the exposure windows for controls. DPP-4i were selected according to the Anatomical Therapeutic Chemical (ATC) Classification System; ATC codes A10BH01 + A10BD07 (Sitagliptin), A10BH02 + A10BD08 (Vildagliptin), A10BH03 + A10BD10 (Saxagliptin), and A10BH05 + A10BD11 (Linagliptin) were considered.

Ascertainment of potential confounders

We defined potential confounders from the regional hospital discharge database as hospital admissions that occurred up to 5 years before the index date for ischemic heart diseases (ICD-9-CM 410-414). Likewise, we also included individuals who had been treated with glimepiride (ATC code A10BB12) or glibenclamide (ATC code A10BB01) in the 6 months prior to the date of hospital

admission. Treatment with insulin (ATC A10A) in the 6 months prior to the date of hospital admission was regarded as proxy of severity of disease.

Statistical analysis

 We calculated the proportions of categorical variables in cases and controls and assessed the differences in baseline characteristics with the chi square test. We estimated the risk of the four different outcomes associated with dispensation of any DPP-4i by fitting conditional logistic regression models, expressed as odds ratios (ORs), and corresponding 95% confidence intervals (CI). We adjusted the statistical models for the aforementioned confounders. Confounders included in the final models were past history of ischemic heart disease, insulin use, and glimepiride or glibenclamide (considered together) use. In the sensitivity analyses, we assessed the use of DPP-4i at any time before hospital admission or death. All analyses were done with the SAS PHREG procedure version 9.2.

Results

During the study period, 14,613 cases of hospital admission for heart failure, 7212 incident cases of heart failure, 1727 cases of re-admission, and 38,248 deaths occurred within this population of patients with type 2 diabetes. Compared with controls, the use, as well as the type, of DPP-4i did not differ between cases and controls for heart failure outcomes (between 1.8 and 2.0% of both cases and controls were on gliptins), whereas the use of insulin, but not of glibenclamide, was more frequent among the cases. There was a higher prevalence of ischemic heart disease among the cases. When mortality was considered as an outcome, exposure to DPP-4i was lower among the cases than the controls (0.8% vs. 1.8%) (Table 1).

After adjustment for available confounders, the use of gliptins up to 6 months before any of the outcomes considered was not associated with the risk of hospitalization for heart failure (OR 1.00, 95% CI 0.94-1.07; p = 0.9832), incident heart failure (OR 1.01, 95% CI 0.92-1.11; p = 0.7808), hospital re-admission for heart failure (OR 1.02, 95% CI 0.84-1.22; p = 0.8745), or death of any cause (OR 0.94, 95% CI 0.90-0.98; p = 0.0021). A history of ischemic heart disease or insulin use was associated with an increased risk of all the outcomes considered, except for hospital re-admission for heart failure (Table 2). In the sensitivity analysis, in which the time window was extended to include DPP-4i use at any time before the outcomes under study, the ORs were unchanged for all outcomes: hospital admission (OR 0.99, 95% CI 0.94-1.05; p = 0.8048), incident heart failure (OR 1.01, 95% CI 0.93-1.09; p = 0.8775), hospital re-admission (OR 0.97, 95% CI 0.85-1.16; p = 0.9558), or death (OR 0.94, 95% CI 0.91-0.98; p = 0.0018).

Discussion

The question whether therapy with DPP-4i may increase the risk of heart failure has raised concern over the safety of these drugs among practitioners (4). The key message of our analysis is that, in a large unselected population of treated individuals with type 2 diabetes, no association was found, regardless of casuality and as borne out by the neutral OR independent of the subgroup considered (Any admission, Incident or Re-admission for heart failure).

Our findings are in line with a very recent paper by Yu et al. who reported no association between the use of these incretin-based therapies, studied as a whole DPPi and GLP1 receptor agonists, and incident heart failure in patients with type 2 diabetes (12). The same concordance exists with the general message from animal studies and pathophysiological investigations which found no detrimental effect of DPP-4i on cardiac function (13,14). However, it's a puzzling link because two other recent physiopathological studies in patients produced unanticipated findings that treatment with DPP-4i could exacerbate heart failure (15,16). Moreover, since no plausible explanation for the SAVOR TIMI 53 finding has been forthcoming, it is difficult to speculate on the reasons for the differences between their and our findings in the drug-associated risk of heart failure.

Weir et al. found a significantly increased risk of hospitalization for heart failure associated with the use of sitagliptin among adults with type 2 diabetes, but only in patients with pre existing heart failure, with a relevant 84% excess of risk (7). One possible, though still speculative, explanation for this discrepancy could be an indication bias in that, before the SAVOR TIMI results were published, DPP-4i were well regarded and considered particularly safe in patients with left ventricular dysfunction or renal insufficiency. On closer analysis, the two study populations differ in average age (78 versus 54 years) and background antidiabetes therapies. An additional difference is in the study design: Weir's case-control study addressed only the effects of one gliptin (sitagliptin) and in a rather selected population of young insured individuals.

The favourable impact of DPP-4i use on all-cause mortality is welcome but still warrants caution. It could confirm recent observational data that suggested a reduction of mortality in the Danish diabetic population on incretins (17). It should also be remembered, however, that DPP-4i use can represent a marker of better specialty care (18).

Our analysis, by contrast, beyond the expected role of previous ischemic heart disease as a factor underlying impaired myocardial performance, revealed an important link between insulin use and risk of heart failure for both first and recurrent hospitalisation. This finding is neither new nor surprising. In the last years, numerous observational studies have found insulin to be a marker of poor cardiovascular outcomes, including heart failure (19). Perplexing is why the outcomes of

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insulin use in observational studies clash with the findings from randomised controlled trials such as the UKPDS (20) and ORIGIN (21) and seem to dismiss the role insulin can play in preventing complications. For instance, the DAI study, the largest cohort study on complications of diabetes in Italy, found a correlation between insulin therapy and the occurrence of coronary disease and stroke (22). Similar conclusions were reported for heart failure (23). A simple explanation would be that there is a typical indication bias. A more complex one would point to the problem of clinical inertia: insulin in the real world is given late, after chronic exposure to high glucose levels, thus marking subjects with poor irreversible legacy. In other words, the outcome after insulin therapy in observational studies may reflect the fact that it is initiated too late rather than its appropriate use.

Our study has several strengths. It was conducted using data retrieved from a population-based database representative of the real type 2 diabetes population seen in Europe, without any selection based on insurance claims, and, in particular, with no age cut-off, two factors which distinguish it from other published studies using administrative data. Our approach thus eliminated recall bias and minimised selection bias. All five drugs currently available on the market were included. Furthermore, as DPP-4i are dispensed and reimbursed only by prescription, we are confident to have included all dispensations.

Our study also has potential limitations, however. Though the sample size is rather large, this shouldn't be a concern: with the use of a one-sided test at the alpha = 5% level, the smallest risk, greater than one, that could be detected with 80% power is OR = 1.16 for hospital admissions, 1.24 for incident heart failure, and 1.49 for hospital re-admissions. We are aware that a weaker association between DPP-4i use and heart failure cannot be ruled out; but if it does exist, it is not so large. Another weakness is the missing data on metabolic control and other clinical variables such as NT-pro BNP levels, hypertension, heart valve defects, and renal failure, all of which could have impacted on hospitalisation rates for heart failure. Nonetheless, there is no reason why this could have favoured the control group and thus masked the association. As a proxy of diabetes severity, we adjusted for cardiovascular disease and level of therapy. We did not consider medication dose or adherence to therapy, and, given the low prevalence of exposed individuals, we were unable to factor in the effect of the different compounds separately. In addition, the use of a database of dispensed drugs rather than usage data might have overestimated the use of DPP-4i; however, it is unlikely that this would have affected cases and controls differently. Finally, only severe cases of heart failure hospitalisation were considered, leaving open the question whether milder episodes of cardiac insufficiency, not resulting in hospital admission, could have been increased in DPP-4i users.

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Contributors' Statement

CBG literature search, study design, data collection, data interpretation, writing.; RP data collection, data analysis; BT, LM and AA literature search, writing assistance; GC data interpretation, writing; RG literature search, study design, data collection, data interpretation, writing. CBG, RP, GC and RG had access to the raw data. The corresponding author had full access to all the data in the study and the final responsibility for the decision to submit for publication.

All authors approved the final version.

Competing interest statement

We have now read and understood BMJ policy on declaration of interests and declare that we have no competing interest.

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Table 1. Baseline Characteristics

1

2 3 4

	Any	y admission for HF		Incident HF			Re-admission for HF			All-cause mortality		
	Cases	Controls		Cases	Controls		Cases	Controls		Cases	Controls	
	(n =14,613)	(n =14,6130)	p value	(n =7212)	(n =72,120)	p value	(n =1727)	(n =17,222)	p value	(n =38,248)	(n =382,313)	p va
ge at recruitment, years	78.0	77.9		78.3	78.2		77.9	77.9		80.2	79.9	
ge at recruitment, years	(8.3)	(8.4)		(8.4)	(8.4)		(8.4)	(8.5)		(9.1)	(9.0)	
	7000 (52.0)	76,900		3577	35,770		914	9102		19,215	40,4002 (50,2)	
ex male	7690 (52.6)	(52.6)		(49.6)	(49.6)		(52.9)	(52.9)		(50.2)	19,1983 (50.2)	
PPP-4i use	256	2881		135	1285	0.5820	37	338		306	6717	<.0002
5 months)	(1.8)	(2.0)	0.0672	(1.9)	(1.8)		(2.1)	(2.0)	0.6090	(0.8)	(1.8)	
	328	3636		171	1657		47	470		477	8491	<.0001
PP-4i use (any)	(2.2)	(2.5)	0.0702	(2.4)	(2.3)	0.6917	(2.7)	(2.7)	0.9853	(1.3)	(2.2)	
revious disorders or reatments												
schemic heart disease		10,237		879	5492		281	2361		4270	23,938	
n the past 5 years)	3371 (23.1)	(7.0)	<.0001	(12.2)	(7.6)	<.0001	(16.3)	(13.7)	0.0034	(11.2)	(6.3)	<.0002
ilimepiride or glibenclamide	1531	16,756		916	8144		172	1691		3428	46,327	
n the past 6 months)	(10.5)	(11.5)	0.0003	(12.7)	(11.3)	0.0003	(10.0)	(9.8)	0.8516	(9.0)	(12.1)	<.0001
nsulin	5363	24,108		2177	12,042		730	4459		14,159	64,477	<.0002
n the past 6 months)	(36.7)	(16.5)	<.0001	(30.2)	(16.7)	<.0001	(42.3)	(25.9)	<.0001	(37.0)	(16.9)	
HF denotes heart f	ailure.											
Data are mean (SD) or no. (%)											
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Table 2. Matched Odd Ratios (OR) of different outcomes associated with exposure to DPP-4i in the 6 months before index date.

	Any admiss	sion for HF	Incide	nt HF	Re-admission for HF		All-cause mortality	
	Unadjusted OR	Adjusted OR	Unadjusted OR	Adjusted OR	Unadjusted OR	Adjusted OR	Unadjusted OR	Adjusted OF
	(95% CI);	(95% CI);						
	p value	p value						
	0.99	1.00	1.01	1.01	1.01	1.02	0.93	0.94
DPP-4i use	(0.92-1.05);	(0.94-1.07);	(0.92-1.11);	(0.92-1.11);	(0.84-1.22);	(0.84-1.22);	(0.89-0.97);	(0.90-0.98)
	p=0.6383	p=0.9832	p=0.8867	p=0.7808	p=0.8944	p=0.8745	p=0.0005	p=0.0021
Previous disorders or treatments								
Ischemic heart disease	1.36	1.34	1.09	1.08	1.03	1.02	1.11	1.09
(in the past 5 years)	(1.31-1.40); p=<.0001	(1.29-1.38); p=<.0001	(1.04-1.14); p=0.0003	(1.03-1.13); p=0.0014	(0.96-1.11); p=0.4455	(0.95-1.10); p=0.6067	(1.09-1.14); p=<.0001	(1.07-1.12) p=<.0001
Glimepiride or glibenclamide	0.99	1.01	1.02	1.03	1.00	1.02	0.96	0.98
(in the past 6 months)	(0.96-1.02); p=0.3555	(0.98-1.04); p=0.4844	(0.98-1.06); p=0.3520	(0.99-1.08); p=0.0998	(0.92-1.09); p=0.9614	(0.93-1.11); p=0.6751	(0.95-0.98); p=<.0001	(0.97-1.00) p=0.0540
Insulin	1.21	1.19	1.13	1.13	1.12	1.12	1.20	1.20
(in the past 6 months)	(1.18-1.24); p=<.0001	(1.17-1.22); p=<.0001	(1.10-1.17); p=<.0001	(1.10-1.17); p=<.0001	(1.06-1.19); p=0.0001	(1.06-1.19); p=0.0002	(1.19-1.22); p=<.0001	(1.18-1.21) p=<.0001

HF denotes heart failure

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Hospitalisation for heart failure and mortality associated with dipeptidyl peptidase 4 (DPP-4) inhibitor use in an unselected population of subjects with type 2 diabetes. A nested case-controll study.

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Hospitalisation for heart failure and mortality associated with dipeptidyl peptidase 4 (DPP-4) inhibitor use in an unselected population of subjects with type 2 diabetes. A nested casecontrolled study.

CARLO B. GIORDA¹, MD, ROBERTA PICARIELLO², BSc, BARBARA TARTAGLINO³ PhD, LISA MARAFETTI¹ MD, FABIANA DI NOI¹ MD, ANNALISA ALESSIATO¹, PhD, GIUSEPPE COSTA^{2,4}, MD and ROBERTO GNAVI², MD

From:

¹ Metabolism and Diabetes Unit, ASL TO5, Regione Piemonte, Chieri, Italy; ² Epidemiology Unit, ASL TO3, Regione Piemonte, Grugliasco, Italy; ³ Chaira Medica Association, Chieri, Italy; ⁴ Department of Clinical and Biological Sciences, University of Torino, Torino, Italy.

Short Title: Heart failure and mortality in DPP-4i users

unission s. Correspondence, phone calls and e-mails about the submission should be directed to me at the following address and phone number

Carlo Giorda, M.D.

Diabetologia

ASL TO5

Via De Maria, 1

10023 Chieri (To)

Italy

Tel. + 39 011 94293312

e-mail: giorda.carlobruno@aslto5.piemonte.it

giordaca@tin.it

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ABSTRACT

Objective. The SAVOR TIMI–53 study reported a significant increase in the risk of hospitalization for heart failure (HF) in patients treated with a DPP-4 inhibitor (DPP-4i) in comparison with placebo. A recent case-control study in part confirmed this risk signal. Our aim was to compare the occurrence of HF in relation to DPP-4i use versus any antidiabetic treatment.

Design. Population-based matched case-control study conducted using administrative data.

Setting. The Italian Region of Piedmont (4.4 million inhabitants).

Participants From a database of 282,000 patients treated with antidiabetic drugs, we identified 14,613 hospitalizations for HF, 7212 incident cases, and 1727 hospital re-admissions between 2008 and 2012; each case was matched for gender, age and antidiabetic therapy with ten controls; cases and controls were compared for exposure to DPP-4i.

Outcome measures Odds ratios (OR) and 95% confidence intervals were calculated by fitting a conditional logistic model. All analyses were adjusted for available risk factors for HF.

Results. We found no increased risk of hospitalisation for HF associated with the use of DPP-4i (OR for admission for HF 1.00 [0.94-1.07], incident HF1.01 [0.92-1.11], recurrent HF 1.02 [0.84-1.22]). All-cause mortality was 6% lower in DPP-4i users (p<0.001), whereas insulin users showed an excess of risk for any type of hospital admission (19%) and death (20%) (p<0.001).

Conclusions. Our findings suggest that, in an unselected population of diabetic patients, the use of DPP-4i is not associated with an increased risk of HF. The favourable impact on all-cause mortality should be viewed with caution and also other explanations investigated.

Keywords: DPP-4 inhibitors, heart failure hospitalization, all-cause mortality, safety in diabetes treatment, case-control study.

Short title: A nested case-control study in a European setting

 Strengths and limitations of this study

This study suggest that, in an unselected population of diabetic patients, the use of DPP-4i is not associated with an increased risk of HF

The study population was representative of the real type 2 diabetes population seen in Europe, without any selection based on insurance claims or age cut-off which distinguish it from other published studies.

Hospitalization for HF was evaluated as admission, incidence or recurrence and all five drugs currently available on the market were included

The main limitation is that weak associations between DPP-4i use and heart failure cannot be ruled out; but if they do exist, they are not so large.

While looking forward to the results of ongoing trials, practitioners can be reassured that the unexpected association reported in the SAVOR TIMI study has not been confirmed in the real world.

Introduction

DPP-4 inhibitors (DPP-4i), or gliptins, are oral agents that delay the catabolism of native GLP-1 by inhibiting the endogenous enzyme dipeptidyl peptidase 4 (DPP-4), thus extending the life of native GLP-1. They have attracted growing interest as first line therapies for type 2 diabetes largely because they are effective in controlling HbA1c while reducing the risk of hypoglicaemia and weight gain. Data from trials with glycemic endpoints suggest that they also exert a protective effect on cardiovascular risk (1). In the first published trial with cardiovascular endpoints (2), saxagliptin neither increase, nor reduce, the risk of a composite of nonfatal myocardial infarction (MI), nonfatal stroke or cardiovascular death, thereby meeting the primary safety objective. However, as compared with placebo, saxagliptin was unexpectedly associated with a 27% excess risk of hospitalization for heart failure (HF) (overall hazard ratio [HR] 1.27 [95% confidence interval [CI] 1.07-1.51]). Detailed sub-analyses (3,4) revealed that the absolute difference between groups was mainly seen during the first 6 months of therapy and that there was no increased risk of death due to heart failure. No clinically relevant factors predictive of increased relative risk with saxagliptin treatment could be definitively identified (5). A subsequent meta-analysis of all registrative trials with DPP-4i showed an association between the use of this class if drugs and increased risk of HF which disappears when excluding cardiovascular outcome trials. (6).

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Conflicting evidence for an association between sitagliptin use and hospitalization for heart failure in some specific conditions comes from recent observational studies. Wang et al (7) found an increased risk of hospitalization for HF in with sitagliptin use in Taiwanese insured individuals whereas Chen et al. (8) reported no increased risk in patient with history of chronic kidney disease and myocardial infarction. Weir et al. (9) in particular reported that in a case-control study, based on an administrative database of U.S. middle-aged adults with type 2 diabetes, treatment with sitagliptin was associated with an increased probability of hospitalization for heart failure only among patients with pre-existing heart failure.

This safety issue is relevant for diabetes care because the rates of heart failure and hospitalization are higher in patients with type 2 diabetes than in the non-diabetic population, regardless of treatment (10,11,12). Further complicating the question is that several drugs commonly used in the treatment of type 2 diabetes have been suspected to increase the risk of heart failure (13).

In light of the concerns over a possible HF risk associated with DPP-4i, we thought it useful to perform a matched case-control study to explore whether the same increased risk as seen in the SAVOR-TIMI study could be detectable in the real world and in an unselected population. To the best of our knowledge, no such analysis has been performed on data from European administrative databases which, by virtue of the universalistic care of European national health systems, encompass the whole population and include all types of available DPP-4i.

Methods

Study design and participants

We conducted a population-based nested case-control study using regional administrative data from Piedmont (population about 4.4 million). The population is covered by an automated system of databases containing the records of all drugs dispensed from all regional pharmacies and hospital discharges reimbursed by the Italian National Health System. These archives can be linked together by a unique anonymous identifier that is encrypted to protect the patient's privacy. Because this automated system is anonymous, ethical committee approval and informed consent for this study were not required.

Procedures

We extracted information from the regional drug prescription database for individuals aged 56 years or older who were dispensed at least one dose of any drug to treat diabetes between January 1, 2009 and December 31, 2013 (DPP-4i were not available in Italy before 2008). Only Piedmont residents were included. To minimise the chance of inclusion of patients with type 1 diabetes, we linked the database to the regional hospital discharge database, which contains the records of all hospital admissions between 1995 and 2013. Excluded were individuals with an International Classification

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of Diseases, 9th Edition, Clinical Modification (ICD-9-CM) code for type 1 diabetes mellitus (250.x1 or 250.x3). Furthermore, as glitazones (TZD) increase the risk of heart failure, all patients who had received a prescription for TZD during the study period were excluded.

Selection of cases

We used four different definitions for cases. First, we identified all patients who had at least one discharge for heart failure, defined as ICD-9-CM codes 402.01, 402.11, 402.91, 425.4, 425.5, 425.9, 428 or 518.4 as the primary discharge diagnosis at any time after the first exposure to antidiabetic drugs (i.e., date of dispensation). For patients with more than one discharge for heart failure, we only included the first episode (i.e., the hospital admission closest to January 1 2009). Second, we identified "incident" cases of heart failure, defined as patients discharged with a diagnosis of heart failure (defined as above) during the study period, without a previous hospitalization for heart failure in the discharge diagnosis (either main or secondary) during the previous 60 months. Third, similarly to the study by Weir et al. (9) we followed up incident cases (defined as above) to identify "first re-hospitalisations" of those patients who had been admitted to hospital for a diagnosis of heart failure. Finally, we considered as cases all deaths (of any cause) that occurred in the population during the study period. Also included were Piedmont residents discharged from any hospital located outside Piedmont because information on exposure to dispensed drugs is available for all patients residing in the region. Similarly, we included the deaths of Piedmont residents wherever they occurred in Italy.

Selection of controls

To identify controls, we randomly selected ten controls from the same population source for each case, matched for year of birth (within a 5-year age band), sex, and year of first exposure to antidiabetic drugs. Controls were selected one subject at a time with replacement. The process was repeated for each outcome. Matching was done by the study statistician (RP) with the use of an automated computer program.

Exposure to DPP-4 inhibitors

We used the regional drug database to identify cases and controls who had been prescribed DPP-4i at any time in the 6 months before the hospital admission date. We used the hospital admission date of cases to calculate the exposure windows for controls. DPP-4i were selected according to the Anatomical Therapeutic Chemical (ATC) Classification System; ATC codes A10BH01 + A10BD07 (Sitagliptin), A10BH02 + A10BD08 (Vildagliptin), A10BH03 + A10BD10 (Saxagliptin), and A10BH05 + A10BD11 (Linagliptin) were considered.

Ascertainment of potential confounders

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We defined potential confounders from the regional hospital discharge database as hospital admissions that occurred up to 5 years before the index date for ischemic heart diseases (ICD-9-CM 410-414). Likewise, we also included individuals who had been treated with glimepiride (ATC code A10BB12) or glibenclamide (ATC code A10BB01) in the 6 months prior to the date of hospital admission. Treatment with insulin (ATC A10A) in the 6 months prior to the date of hospital admission was regarded as proxy of severity of disease.

Statistical analysis

We calculated the proportions of categorical variables in cases and controls and assessed the differences in baseline characteristics with the chi square test. We estimated the risk of the four different outcomes associated with dispensation of any DPP-4i by fitting conditional logistic regression models, expressed as odds ratios (ORs), and corresponding 95% confidence intervals (CI). We adjusted the statistical models for the aforementioned confounders. Confounders included in the final models were past history of ischemic heart disease, insulin use, and glimepiride or glibenclamide (considered together) use. In the sensitivity analyses, we assessed the use of DPP-4i at any time before hospital admission or death. All analyses were done with the SAS PHREG procedure version 9.2.

Results

During the study period, 14,613 cases of hospital admission for heart failure, 7212 incident cases of heart failure, 1727 cases of re-admission, and 38,248 deaths occurred within this population of patients with type 2 diabetes. Compared with controls, the use, as well as the type, of DPP-4i did not differ between cases and controls for heart failure outcomes (between 1.8 and 2.0% of both cases and controls were on gliptins), whereas the use of insulin, but not of glibenclamide, was more frequent among the cases. There was a higher prevalence of ischemic heart disease among the cases than the controls (0.8% vs. 1.8%) (Table 1).

After adjustment for available confounders, the use of gliptins up to 6 months before any of the outcomes considered was not associated with the risk of hospitalization for heart failure (OR 1.00, 95% CI 0.94-1.07; p = 0.9832), incident heart failure (OR 1.01, 95% CI 0.92-1.11; p = 0.7808), hospital re-admission for heart failure (OR 1.02, 95% CI 0.84-1.22; p = 0.8745), or death of any cause (OR 0.94, 95% CI 0.90-0.98; p = 0.0021). A history of ischemic heart disease or insulin use was associated with an increased risk of all the outcomes considered, except for hospital re-admission for heart failure (Table 2). In the sensitivity analysis, in which the time window was extended to include DPP-4i use at any time before the outcomes under study, the ORs were

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unchanged for all outcomes: hospital admission (OR 0.99, 95% CI 0.94-1.05; p = 0.8048), incident heart failure (OR 1.01, 95% CI 0.93-1.09; p = 0.8775), hospital re-admission (OR 0.97, 95% CI 0.85-1.16; p = 0.9558), or death (OR 0.94, 95% CI 0.91-0.98; p = 0.0018).

Discussion

The question whether therapy with DPP-4i may increase the risk of heart failure has raised concern over the safety of these drugs among practitioners (4). The key message of our analysis is that, in a large unselected population of treated individuals with type 2 diabetes, no association was found, regardless of casuality and as borne out by the neutral OR independent of the subgroup considered (Any admission, Incident or Re-admission for heart failure).

Our findings are in line with a very recent paper by Yu et al. who reported no association between the use of these incretin-based therapies, studied as a whole DPPi and GLP1 receptor agonists, and incident heart failure in patients with type 2 diabetes (14). The same concordance exists with the general message from animal studies and pathophysiological investigations which found no detrimental effect of DPP-4i on cardiac function (15,16). However, it's a puzzling link because two other recent physiopathological studies in patients produced unanticipated findings that treatment with DPP-4i could exacerbate heart failure (17,18). Moreover, since no plausible explanation for the SAVOR TIMI 53 finding has been forthcoming, it is difficult to speculate on the reasons for the differences between their and our findings in the drug-associated risk of heart failure.

Weir et al. found a significantly increased risk of hospitalization for heart failure associated with the use of sitagliptin among adults with type 2 diabetes, but only in patients with pre existing heart failure, with a relevant 84% excess of risk (9). One possible, though still speculative, explanation for this discrepancy could be an indication bias in that, before the SAVOR TIMI results were published, DPP-4i were well regarded and considered particularly safe in patients with left ventricular dysfunction or renal insufficiency. On closer analysis, the two study populations differ in average age (78 versus 54 years) and background antidiabetes therapies. An addiction difference lies in the selection of cases with the American population appearing as a rather selected population of young selected insure individuals.

The favourable impact of DPP-4i use on all-cause mortality is welcome but still warrants caution. It could confirm recent observational data that suggested a reduction of mortality in the Danish diabetic population on incretins (19). It should also be remembered, however, that DPP-4i use can represent a marker of better specialty care (20).

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Our analysis, by contrast, beyond the expected role of previous ischemic heart disease as a factor underlying impaired myocardial performance, revealed an important link between insulin use and risk of heart failure for both first and recurrent hospitalisation. This finding is neither new nor surprising. In the last years, numerous observational studies have found insulin to be a marker of poor cardiovascular outcomes, including heart failure (21) raising the question of the possible pathogenetic role of hypoglycaemic events Perplexing is why the outcomes of insulin use in observational studies clash with the findings from randomised controlled trials such as the UKPDS (22) and ORIGIN (23) and seem to dismiss the role insulin can play in preventing complications. For instance, the DAI study, the largest cohort study on complications of diabetes in Italy, found a correlation between insulin therapy and the occurrence of coronary disease and stroke (24). Similar conclusions were reported for heart failure (25). A simple explanation would be that there is a typical indication bias. A more complex one would point to the problem of clinical inertia: insulin in the real world is given late, after chronic exposure to high glucose levels, thus marking subjects with poor irreversible legacy. In other words, the outcome after insulin therapy in observational studies may reflect the fact that it is initiated too late rather than its appropriate use.

Our study has several strengths. It was conducted using data retrieved from a population-based database representative of the real type 2 diabetes population seen in Europe, without any selection based on insurance claims, and, in particular, with no age cut-off, two factors which distinguish it from other published studies using administrative data. Our approach thus eliminated recall bias and minimised selection bias. All five drugs currently available on the market were included. Furthermore, as DPP-4i are dispensed and reimbursed only by prescription, we are confident to have included all dispensations.

Our study also has potential limitations, however. Though the sample size is rather large, this shouldn't be a concern: with the use of a one-sided test at the alpha = 5% level, the smallest risk, greater than one, that could be detected with 80% power is OR = 1.16 for hospital admissions, 1.24 for incident heart failure, and 1.49 for hospital re-admissions. We are aware that a weaker association between DPP-4i use and heart failure cannot be ruled out; but if it does exist, it is not so large. Another weakness is the missing data on metabolic control and other clinical variables such as NT-pro BNP levels, hypertension, heart valve defects, and renal failure, all of which could have impacted on hospitalisation rates for heart failure. Nonetheless, there is no reason why this could have favoured the control group and thus masked the association. As a proxy of diabetes severity, we adjusted for cardiovascular disease and level of therapy. We did not consider medication dose or adherence to therapy, and, given the low prevalence of exposed individuals, we were unable to

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factor in the effect of the different compounds separately. In addition, the use of a database of dispensed drugs rather than usage data might have overestimated the use of DPP-4i; however, it is unlikely that this would have affected cases and controls differently. Finally, only severe cases of heart failure hospitalisation were considered, leaving open the question whether milder episodes of cardiac insufficiency, not resulting in hospital admission, could have been increased in DPP-4i users.

Ongoing trials including the Sitagliptin Cardiovascular Outcome Study (26) and the Functional Impact of GLP-1 for Heart Failure Treatment (27) may help to clarify the conflicting findings. Meanwhile, practitioners can be reassured that the unexpected association reported in the SAVOR TIMI study has not been confirmed in the real world.

Contributorship Statement

CBG literature search, study design, data collection, data interpretation, writing.; RP data collection, data analysis; BT, LM and AA literature search, writing assistance; GC data interpretation, writing; RG literature search, study design, data collection, data interpretation, writing. CBG, RP, GC and RG had access to the raw data. The corresponding author had full access to all the data in the study and the final responsibility for the decision to submit for publication.

All authors approved the final version.

Competing interests

We have now read and understood BMJ policy on declaration of interests and declare that we have no competing interest.

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Data sharing statement

Technical appendix, statistical code, and dataset were not made available in the Dryad repository. However these data can be requested to corresponding author.

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Table 1. Baseline Characteristics

Controls 3) (n =14,6130) 77.9 (8.4) 76,900 6) (52.6) 2881 (2.0) 3636 (2.5)	p value	Cases (n =7212) 78.3 (8.4) 3577 (49.6) 135 (1.9) 171 (2.4)	Controls (n =72,120) 78.2 (8.4) 35,770 (49.6) 1285 (1.8) 1657 (2.3)	p value 0.5820 0.6917	Cases (n =1727) 77.9 (8.4) 914 (52.9) 37 (2.1) 47	Controls (n =17,222) 77.9 (8.5) 9102 (52.9) 338 (2.0) 470	p value 0.6090	Cases (n =38,248) 80.2 (9.1) 19,215 (50.2) 306 (0.8)	Controls (n =382,313) 79.9 (9.0) 19,1983 (50.2) 6717 (1.8)	p valu <.000
77.9 (8.4) 76,900 6) (52.6) 2881 (2.0) 3636	0.0672	78.3 (8.4) 3577 (49.6) 135 (1.9) 171	78.2 (8.4) 35,770 (49.6) 1285 (1.8) 1657	0.5820	77.9 (8.4) 914 (52.9) 37 (2.1)	77.9 (8.5) 9102 (52.9) 338 (2.0)		80.2 (9.1) 19,215 (50.2) 306	79.9 (9.0) 19,1983 (50.2) 6717	
(8.4) 76,900 6) (52.6) 2881 (2.0) 3636		(8.4) 3577 (49.6) 135 (1.9) 171	(8.4) 35,770 (49.6) 1285 (1.8) 1657		(8.4) 914 (52.9) 37 (2.1)	(8.5) 9102 (52.9) 338 (2.0)	0.6090	(9.1) 19,215 (50.2) 306	(9.0) 19,1983 (50.2) 6717	<.000
76,900 6) (52.6) 2881 (2.0) 3636		3577 (49.6) 135 (1.9) 171	35,770 (49.6) 1285 (1.8) 1657		914 (52.9) 37 (2.1)	9102 (52.9) 338 (2.0)	0.6090	19,215 (50.2) 306	19,1983 (50.2) 6717	<.000
6) (52.6) 2881 (2.0) 3636		(49.6) 135 (1.9) 171	(49.6) 1285 (1.8) 1657		(52.9) 37 (2.1)	(52.9) 338 (2.0)	0.6090	(50.2) 306	6717	<.00
(52.6) 2881 (2.0) 3636		135 (1.9) 171	1285 (1.8) 1657		37 (2.1)	338 (2.0)	0.6090	306	6717	<.00
(2.0) 3636		(1.9) 171	(1.8) 1657		(2.1)	(2.0)	0.6090			<.00
3636		171	1657				0.6090	(0.8)	(1.8)	<.00
	0.0702			0.6917	47	470				
(2.5)	0.0702	(2.4)	(2.3)	0.6917				477	8491	<.0001
					(2.7)	(2.7)	0.9853	(1.3)	(2.2)	
10,237		879	5492		281	2361		4270	23,938	
1) (7.0)	<.0001	(12.2)	(7.6)	<.0001	(16.3)	(13.7)	0.0034	(11.2)	(6.3)	<.0001
16,756		916	8144		172	1691		3428	46,327	
(11.5)	0.0003	(12.7)	(11.3)	0.0003	(10.0)	(9.8)	0.8516	(9.0)	(12.1)	<.00
24,108		2177	12,042		730	4459		14,159	64,477	_
(16.5)	<.0001	(30.2)	(16.7)	<.0001	(42.3)	(25.9)	<.0001	(37.0)	(16.9)	<.000
										13
										13
For	peer revie	w only - ht	tp://bmjoper	n.bmj.com	n/site/about/	guidelines.xl	html			
)	(11.5) 24,108 (16.5)	(7.0) 16,756 0.0003 (11.5) 24,108 (.0001 (16.5) 5) For peer revie	(7.0) (12.2) 16,756 916 0.0003 (11.5) (12.7) 24,108 2177 <.0001 (16.5) (30.2) 5) For peer review only - htt	(7.0) (12.2) (7.6) 16,756 916 8144 0.0003 (11.5) (12.7) (11.3) 24,108 2177 12,042 <.0001 (16.5) (30.2) (16.7) 5) For peer review only - http://bmjoper	(7.0) (12.2) (7.6) 16,756 916 8144 0.0003 (11.5) (12.7) (11.3) 24,108 2177 12,042 <.0001 (16.5) (30.2) (16.7) For peer review only - http://bmjopen.bmj.com	(7.0) (12.2) (7.6) (16.3) 16,756 916 8144 172 0.0003 0.0003 0.0003 0.0003 0 (11.5) (12.7) (11.3) (10.0) 24,108 2177 12,042 730 <.0001	(7.0) (12.2) (7.6) (16.3) (13.7) 16,756 916 8144 172 1691 0.0003 (11.5) (12.7) (11.3) (10.0) (9.8) 24,108 <.0001	(7.0) (12.2) (7.6) (16.3) (13.7) 16,756 916 8144 172 1691 0.8516 (11.5) (12.7) (11.3) (10.0) (9.8) 0.8516 24,108 2177 12,042 730 4459 .0001 (16.5) (30.2) (16.7) (42.3) (25.9) .0001	(7.0) (12.2) (7.6) (16.3) (13.7) (11.2) 16,756 916 8144 0.0003 172 1691 3428 (11.5) (12.7) (11.3) (10.0) (9.8) (9.0) 24,108 <0.001	$ \begin{array}{cccccccccccccccccccccccccccccccccccc$

	Any admiss	ion for HF	Incide	Incident HF		Re-admission for HF		mortality
	Unadjusted OR	Adjusted OR						
	(95% CI);							
	p value							
	0.99	1.00	1.01	1.01	1.01	1.02	0.93	0.94
DPP-4i use	(0.92-1.05);	(0.94-1.07);	(0.92-1.11);	(0.92-1.11);	(0.84-1.22);	(0.84-1.22);	(0.89-0.97);	(0.90-0.98);
	p=0.6383	p=0.9832	p=0.8867	p=0.7808	p=0.8944	p=0.8745	p=0.0005	p=0.0021
Previous disorders or treatments								
Ischemic heart disease	1.36	1.34	1.09	1.08	1.03	1.02	1.11	1.09
(in the past 5 years)	(1.31-1.40); p=<.0001	(1.29-1.38); p=<.0001	(1.04-1.14); p=0.0003	(1.03-1.13); p=0.0014	(0.96-1.11); p=0.4455	(0.95-1.10); p=0.6067	(1.09-1.14); p=<.0001	(1.07-1.12); p=<.0001
Glimepiride or glibenclamide	0.99	1.01	1.02	1.03	1.00	1.02	0.96	0.98
(in the past 6 months)	(0.96-1.02); p=0.3555	(0.98-1.04); p=0.4844	(0.98-1.06); p=0.3520	(0.99-1.08); p=0.0998	(0.92-1.09); p=0.9614	(0.93-1.11); p=0.6751	(0.95-0.98); p=<.0001	(0.97-1.00); p=0.0540
Insulin	1.21	1.19	1.13	1.13	1.12	1.12	1.20	1.20
(in the past 6 months)	(1.18-1.24); p=<.0001	(1.17-1.22); p=<.0001	(1.10-1.17); p=<.0001	(1.10-1.17); p=<.0001	(1.06-1.19); p=0.0001	(1.06-1.19); p=0.0002	(1.19-1.22); p=<.0001	(1.18-1.21); p=<.0001

Table 2. Matched Odd Ratios (OR) of different outcomes associated with exposure to DPP-4i in the 6 months before index date.

HF denotes heart failure

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Section/Topic	Itom #	Recommendation	Bonortad on page #
Title and abstract	1 Item #	(a) Indicate the study's design with a commonly used term in the title or the abstract	Reported on page #
	1		1
		(b) Provide in the abstract an informative and balanced summary of what was done and what was found	2
ntroduction			
Background/rationale	2	Explain the scientific background and rationale for the investigation being reported	3 and 4
Objectives	3	State specific objectives, including any pre-specified hypotheses	4
Methods			
Study design	4	Present key elements of study design early in the paper	4
Setting	5	Describe the setting, locations, and relevant dates, including periods of recruitment, exposure, follow-up, and data collection	4 and 5
Participants	6	(a) Cohort study—Give the eligibility criteria, and the sources and methods of selection of participants. Describe methods of follow-up Case-control study—Give the eligibility criteria, and the sources and methods of case ascertainment and control selection. Give the rationale for the choice of cases and controls Cross-sectional study—Give the eligibility criteria, and the sources and methods of selection of participants	5
		(b) Cohort study—For matched studies, give matching criteria and number of exposed and unexposed Case-control study—For matched studies, give matching criteria and the number of controls per case	5
Variables	7	Clearly define all outcomes, exposures, predictors, potential confounders, and effect modifiers. Give diagnostic criteria, if applicable	6
Data sources/ measurement	8*	For each variable of interest, give sources of data and details of methods of assessment (measurement). Describe comparability of assessment methods if there is more than one group	5 and 6
Bias	9	Describe any efforts to address potential sources of bias	4 and 5 and 13 (table
			1) and 14 (table 2)
tudy size	10	Explain how the study size was arrived at	Discussed on page 8
Quantitative variables	11	Explain how quantitative variables were handled in the analyses. If applicable, describe which groupings were chosen and why	6
Statistical methods	12	(a) Describe all statistical methods, including those used to control for confounding	6
		(b) Describe any methods used to examine subgroups and interactions	No subgroups were considered and no interactions were

STROBE 2007 (v4) checklist of items to be included in reports of observational studies in epidemiology*

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			examined
		(c) Explain how missing data were addressed	No missing data
		(d) Cohort study—If applicable, explain how loss to follow-up was addressed	
		Case-control study—If applicable, explain how matching of cases and controls was addressed	5
		Cross-sectional study—If applicable, describe analytical methods taking account of sampling strategy	
		(e) Describe any sensitivity analyses	6
Results			
Participants	13*	(a) Report numbers of individuals at each stage of study—eg numbers potentially eligible, examined for eligibility, confirmed eligible, included in the study, completing follow-up, and analysed	6, and 13 (table 1)
		(b) Give reasons for non-participation at each stage	Doesn't apply: the
			study is based on
			administrative data
		(c) Consider use of a flow diagram	We didn't use a flow
			diagram.
Descriptive data	14*	(a) Give characteristics of study participants (eg demographic, clinical, social) and information on exposures and potential confounders	13 (table 1)
		(b) Indicate number of participants with missing data for each variable of interest	No missing data
		(c) Cohort study—Summarise follow-up time (eg, average and total amount)	
Outcome data	15*	Cohort study—Report numbers of outcome events or summary measures over time	
		Case-control study—Report numbers in each exposure category, or summary measures of exposure	6 and 13 (table 1)
		Cross-sectional study—Report numbers of outcome events or summary measures	
Main results	16	(<i>a</i>) Give unadjusted estimates and, if applicable, confounder-adjusted estimates and their precision (eg, 95% confidence interval). Make clear which confounders were adjusted for and why they were included	6, 7 and 14 (table 2)
		(b) Report category boundaries when continuous variables were categorized	Not applicable
		(c) If relevant, consider translating estimates of relative risk into absolute risk for a meaningful time period	Not done
Other analyses	17	Report other analyses done—eg analyses of subgroups and interactions, and sensitivity analyses	6 and 7
Discussion	•		
Key results	18	Summarise key results with reference to study objectives	7
Limitations	19	Discuss limitations of the study, taking into account sources of potential bias or imprecision. Discuss both direction and magnitude of any potential bias	8
Interpretation	20	Give a cautious overall interpretation of results considering objectives, limitations, multiplicity of analyses, results from similar studies, and other relevant evidence	8 and 9
Generalisability	21	Discuss the generalisability (external validity) of the study results	8

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Other information			
Funding	22	Give the source of funding and the role of the funders for the present study and, if applicable, for the original study on 9	
		which the present article is based	

*Give information separately for cases and controls in case-control studies and, if applicable, for exposed and unexposed groups in cohort and cross-sectional studies. Note: An Explanation and Elaboration article discusses each checklist item and gives methodological background and published examples of transparent reporting. The STROBE checklist is best used in conjunction with this article (freely available on the Web sites of PLoS Medicine at http://www.plosmedicine.org/, Annals of Internal Medicine at http://www.annals.org/, and Epidemiology at http://www.epidem.com/). Information on the STROBE Initiative is available at www.strobe-statement.org.

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