

BMJ Open Older doctors and progression through specialty training in the UK: a cohort analysis of General Medical Council data

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ABSTRACT

Objective: To determine whether older age at graduation is associated with any difference in outcomes from the annual specialty training progression assessment.

Design: An open cohort of 38 308 doctors who graduated from a UK medical school with annual assessments of progression in their specialty training programme with data centrally collected by the General Medical Council between 5 August 2009 to 31 July 2012.

Results: Mature junior doctors (≥ 29 years at graduation) were more likely to have problems with progression on their annual review of competence progression record of in training assessment (ARCP/RITA) than their younger colleagues (OR 1.34, 95% CI 1.22 to 1.49, $p < 0.001$). This association was, if anything, even stronger (OR 1.57, 95% CI 1.41 to 1.74, $p < 0.001$) after adjustment for gender, ethnicity, type of University and specialty. The same was true when only looking at the most extreme ARCP outcome (4) which is being asked to leave their specialist programme (OR 1.81, 95% CI 1.34 to 2.44, $p < 0.001$).

Conclusions: Mature doctors are a growing part of the medical workforce and they are likely to broaden the spectrum of doctors by bring different life experience to the profession. These results suggest that they are more likely to have problems with progressing through their specialist training programme. More research is required to determine the reasons behind these associations and how mature doctors can be supported both in choosing the best training programme and in coping with the complex demands of higher training at a later stage in their lives.

INTRODUCTION

Over the past decade, more mature students have been welcomed onto the medical training programme. While they only make up around 4% of medical students in the UK,¹ they are a more substantial proportion of graduates from the USA and Canada (16.7% and 14.2% were 30 or older at graduation, respectively^{2 3}). These students are often

Strengths and limitations of this study

- First study to look at how age at graduation affects a doctor's chance of succeeding in their annual revalidation.
- Large sample size with little missing data and minimal sources of bias for exposure and outcome variables.
- Results are counter to prevailing beliefs that mature medical students cope better with medical training as demonstrates greater problems with progression through the annual review of competence progression (ARCP) process.
- Highlights the importance of other demographic and clinical factors that determine progression in training.
- ARCP data is a simple measure of adequate progression and does not capture excellence so could hide a bimodal distribution whereby mature junior doctors are also more likely to excel as well as have problems of progression.

different in their outlook and abilities to a typical school leaver and may be better suited as both a student and future doctor. For example, the former director of the graduate entry programme at St George's Hospital Medical School has stated that 'mature students... are sooner and better able to handle the responsibilities of being a doctor' and are 'much more self-directed, challenging, demanding, questioning, and mature' than their younger counterparts.⁴ These subjective views have some limited support from both qualitative and quantitative research during the medical school years, for example, older students appear to do better at year 3 Objective Structured Clinical Examination exams.⁵ Two studies have suggested that mature students cope better with the transition to clinical responsibilities feeling less confused, daunted, anxious or intimidated and more likely to describe a positive transition.^{6 7} This may not merely reflect greater

academic experience; greater age at programme entry, as opposed to the presence of a previous degree, was a better predictor for positive attributes and attitudes related to being a doctor.⁸ This may reflect stronger motivational factors that lead them to positively choose medicine as a subsequent career.

Remarkably little is known about what happens to these mature graduates after they qualify. These positive attitudes could result in very focused and determined graduates who try to reach their choice of specialist career as quickly and efficiently as possible thereby progressing through their training rapidly. On the other hand, mature graduates are more likely to have established geographical roots and family commitments that may make handling the double burdens of career and family problematic even earlier in their training as compared to younger graduates. Anecdotal evidence from the Severn Deanery has suggested that some mature students required greater support with getting through their annual assessment (previously known as RITA—record of in training assessment) and now referred to as ARCP (annual review of competence progression). We objectively test the null hypothesis that the proportion of doctors who either require additional training time or who are asked to leave the programme is the same for both older and younger graduate doctors.

METHODS

Datasource and variable definitions

The General Medical Council (GMC), who collate the national data on ARCP/RITA, kindly provided us with an anonymised extract of data for all UK medical graduates who had a review between 5 August 2009 to 31 July 2012. In the UK, prior to 2013, the ARCP/RITA process begins at the start of specialty training (such as surgery or primary care) and continues until completion of training (obtaining a certificate of completion of training—CCT) that enables doctors to apply for a consultant post.

Since the coding of the outcomes for ARCP and RITA do not map directly onto each other, we had to use slightly different definitions for our outcome measure of poor progression. For ARCP we used codes 3 (requires additional training time), 4 (released from the programme) and 7.3 (inadequate progress) as a composite measure of poor progression. For the RITA we used codes D (targeted training) and E (intensified or repeat training) as our poor outcome measure (see online supplementary appendix 1 for the full coding scheme).^{9 10} We choose to exclude participants with a code for insufficient evidence (as this often reflects inadequate documentation rather than poor progress per se) and those trainees on an out-of-programme secondment.

Our exposure measure was based on an arbitrary age cut-off (coded as an integer value). There is no accepted standard definition of a 'mature' student so we chose to define this as a graduate who was 29 years or over at the

year of first registration (ie, year of graduation). By choosing this cut point we hoped to not include graduates who had simply taken a gap year, intercalated BSc or a prior degree before going straight into medicine (as this should mean they are not older than 27 years) but those who would have had some years of 'work' experience outside of medicine. This is similar to a previous study that defined the 'older mature' as 'students who have worked in other occupations for a number of years prior to making a decision to apply to medical school'¹¹ We further subdivided this 'mature' group into those aged between 29 and 31 years and those who were 32 years or older on date of first registration to examine for any dose–response effects with older age at registration. Finally for a sensitivity analysis we examined a more detailed classification of the younger baseline group into the following categories (≤ 23 , 24, 25, 26, 27, 28 years). We defined, a priori, a number of potential confounders or intermediaries that could be associated with being an older graduate and a greater probability of poor progression. These were gender, specialty, ethnicity and whether the graduate had qualified from a 'mature friendly' medical school that may be better able to help the older graduate cope with the future stresses of being a doctor. This last variable was operationalised as follows: We calculated the percentage of mature students graduating from the medical school and then created a binary variable if the percentage was greater than 10%—approximately the top quartile and these were mainly the new medical schools (eg, Exeter, Brighton & Sussex etc). We could not disaggregate all the London-based medical schools as they were all coded as University of London.

Statistical methods

The original data set had multiple records for a doctor for each assessment (long format) but this could be linked by an anonymous unique identifier. We reshaped the data into wide format (one row per doctor) so each doctor is only represented once in the data set. If the doctor had poor progression more than once, we only coded the first event. We compared simple proportions using χ^2 tests and linear regression for continuous variables. We then calculated the crude OR (95% CI, p values) for older age at graduation and poor progression and multivariable OR adjusting for gender, ethnicity (binary variable defined as non-ethnic if ticked any of the White ethnicity codes from census or ethnic minority, which included any other code), specialty (dummy variable) and mature friendly medical school (binary variable). For specialty we used hospital medicine as the baseline group as it had the largest number of doctors. We undertook a sensitivity analysis using the most extreme outcome—leaving the training programme. As this is only explicitly coded in the ARCP outcomes, we could not use participants with RITA assessments for this secondary analysis. We examined for potential interactions between age at registration with gender and ethnicity and either failure to progress or being asked to leave the specialty.

RESULTS

We received a total of 110 571 records (multiple assessments per doctor). We dropped 307 records (0.3%) without a specialty code and there were 5173 records with a missing outcome (4.7%) and 361 records (0.3%) with an ambiguous code that we could not use (99% of the missing outcome data came from 2012, when the GMC asked Deaneries to return forms even for doctors who were not having ARCPs as they were out of programme, on maternity leave or long-term sick so these are not really missing outcomes—Andy Knapton, GMC personal communication). In addition, there were 7072 records (6.4%) for out-of-programme secondments and 7737 records (7%) coded as insufficient evidence leaving us with 89 921 records. After removing incomplete data for ethnicity, year of birth, year of registration and graduating university, we were left with 83 702 records from 38 308 doctors (see figure 1) similar to the stated number of registered doctors (in approved practice settings) as listed by the GMC.¹² There were 2610 (6.8%) mature graduates (1414 between 29 and 31 years, and 1196 ≥ 32 years); 83.7% of assessments were ARCP and 16.3% were from the RITA. In total, 6045 doctors (15.8%) failed at least one ARCP or RITA during the 3 years of recorded data and of those, 491 (1.3%) were asked to leave the specialty programme (ARCP outcome 4).

Older doctors were more likely to be male, non-ethnic minority, and train in primary care or public health ($p < 0.001$) compared to younger doctors (see web table 1). Older doctors were more likely to have problems with progression (OR 1.34, 95% CI 1.22 to 1.49, p value < 0.001 ; table 1). After adjusting for gender, ethnicity, type of medical school, and choice of specialty, the OR was further increased (OR 1.57, 95% CI 1.41 to 1.74, $p < 0.001$). When we broke down the older age group into three categories (non-mature, 29 to 31 years, ≥ 32 years), the trend was even more marked both with and without adjustment for other covariates (OR 1.0, 1.43, 1.74, respectively, p value for trend < 0.001 after multivariable adjustment). Our more detailed breakdown of the younger age group suggested that increased problems with progression are evident at a younger age, 26 years and above, though the oldest group (≥ 32 years) appear to have additional problems (see web table 2).

Our secondary analysis using just the extreme outcome of leaving the training programme (ARCP-4) found an even greater OR of failing to progress for mature students compared to non-matures (OR 1.81, 95% CI 1.34 to 2.44, $p < 0.001$). When we examined this by our three level age group, we observed a non-linear trend (OR for non-mature, 29 to 31 years, ≥ 32 years: 1.0, 1.29, 2.48, respectively, p value for trend < 0.001) whereby the excess risk seemed mainly limited to the oldest group (≥ 32 years; see web table 3). There was no evidence of any interactions between maturity and either gender or ethnicity on failure to progress or being asked to leave the specialty. The results were essentially unchanged when we replaced the type of university with a dummy variable for all universities.

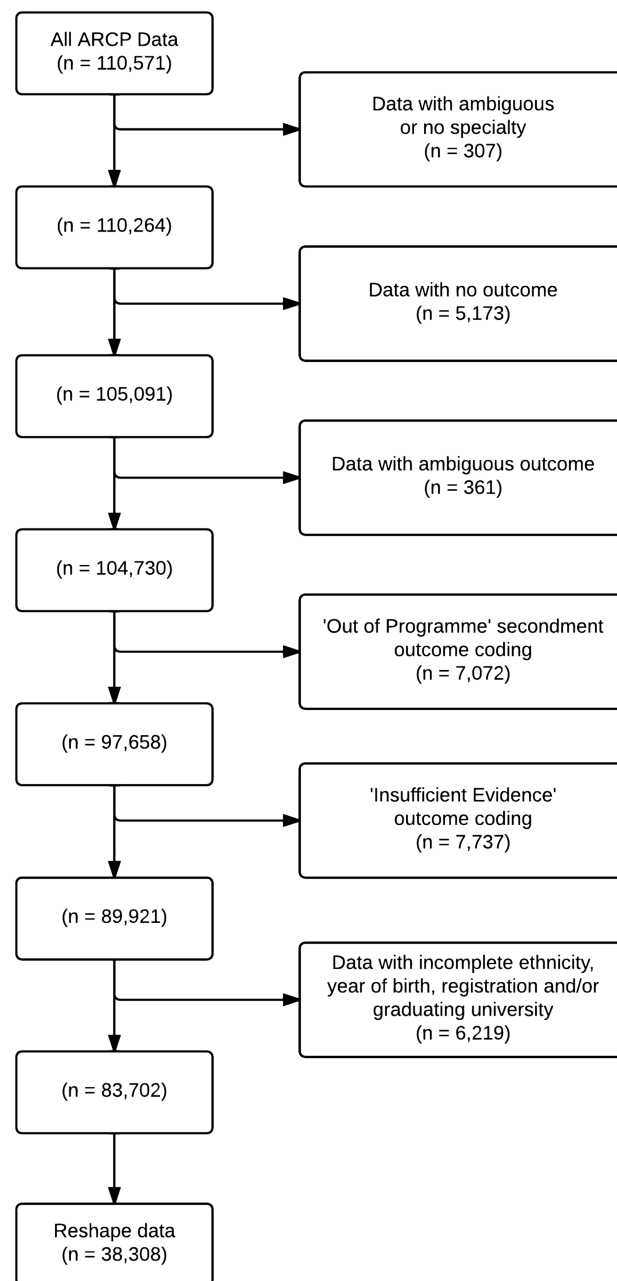


Figure 1 A flow chart showing losses of data due to incomplete or inadequate data to reach the final study sample.

DISCUSSION

This study provides strong evidence that doctors who are older at graduation were more likely to have problems with progression at their annual assessment and were more likely to leave their initial specialist training programme. These findings appeared to be independent of other factors, such as gender, ethnicity, type of medical school and specialty. The last showed wide variability with some specialties having higher (obstetrics and gynaecology) and others lower (general practice and public health) rates of problems with progression. This finding is consistent with the results of a recent analysis comparing doctors who obtained their medical degree

Table 1 Association between 'Mature status' and failure to progress at ARCP adjusted for a range of potential confounders

	Model 1* OR (95% CI)	p Value	Model 2* OR (95% CI)	p Value
Older (≥ 29 years) (2610)	1.34 (1.22 to 1.49)	<0.001	1.57 (1.41 to 1.74)	<0.001
Younger group (≤ 28 years) (35 698)	1.00		1.00	
Older group (29–31 years) (1414)	1.27 (1.11 to 1.46)	0.001	1.43 (1.24 to 1.65)	<0.001
Oldest group (≥ 32 years) (1196)	1.43 (1.24 to 1.65)	<0.001	1.74 (1.50 to 2.02)	<0.001
p Value for trend		<0.001		<0.001
Female gender (21 470)			0.82 (0.77 to 0.87)	<0.001
Ethnic minority (11 338)			1.59 (1.49 to 1.68)	<0.001
Mature friendly university (35 745)			1.18 (1.06 to 1.32)	0.003
First specialty				
Medicine (10 135)			1.00	
ACCS and related (5827)			1.00 (0.92 to 1.08)	0.93
Surgery (6077)			0.84 (0.77 to 0.91)	<0.001
GP and public health (9094)			0.26 (0.24 to 0.29)	<0.001
O and G (1528)			2.16 (1.91 to 2.43)	<0.001
Paediatrics (2791)			0.81 (0.72 to 0.90)	<0.001
Pathology (564)			0.84 (0.67 to 1.06)	0.14
Psychiatry (966)			0.51 (0.42 to 0.63)	<0.001
Radiology (1326)			0.88 (0.76 to 1.02)	0.10

*Model 1, simple OR; Model 2 for binary age group after adjustment for all covariates as shown in table except for the three level age group variable. This model was then rerun with the three level age group and other covariates to examine for a dose-response effect. ACCS, acute care common stem; GP, general practitioner; O and G, obstetrics and gynaecology.

either in or outside of the UK and testing whether the Performance and Linguistics Assessments Board examination system explained performance at ARCP.¹³ It is also consistent with the recent GMC Report on the state of medical education and practice in the UK¹⁴ which found (in Figure 46) that doctors who were over 30 when joining the register were more likely than their younger counterparts to receive a sanction or a warning. While the null hypothesis defined 'mature' graduates as those over 28 years at first registration, additional analysis has highlighted this effect is evident for doctors as young as 26 on registration, who make up over 20% of the doctor population in this sample.

As these results are unlikely to be due to chance, one must consider other possible explanations. Bias in either measurement of exposure or outcome is very unlikely as age at graduation is taken from year of registration and year of birth so should be well recorded and any coding errors are likely to be random in nature. Similarly any coding errors in the ARCP/RITA outcomes are unlikely to be differential according to age at graduation. A very small proportion of outcome data were missing and this is unlikely to have been systematically biased. Though we attempted to control for a variety of covariates that could influence the outcome, we did not have reliable data on whether trainees were in full or part time training. The latter may be more common in mature graduates and may influence progression in training. Similarly we could not explore if there was an interaction between mature graduate status and full-time or part-time training. In addition, ARCP is not intended to capture excellence in training but merely adequate progression. It is possible that the performance of mature graduates is bimodal so

that some mature doctors actually have better outcomes but this would not be evident in our analysis.

One must consider several possible explanations as to why older graduates have more problems progressing through higher training if we assume our observed associations are truly causal. (1) They may have more commitments outside of work (caring commitments for either children or parents or other personal relationship issues) that may make it harder to successfully complete all the assessments required for ARCP.¹¹ (2) They may find themselves committing to a specialty that may not have been their first choice in order to stay in a certain part of the country for their children or spouse. This could result in them doing less well in ARCP due a degree of ambivalence to this specialty. (3) They may have more problems passing postgraduate specialist exams or completing more technical skills competencies which result in either additional training time or in the worst case leaving the specialty. This may be one explanation why we observed the same difficulty with progression for doctors of ethnic minority background who are known to have a higher failure rate with the Membership of the Royal College of General Practitioners (MRCGP) exam.^{13 15} (4) Being older, these doctors may find it harder to engage with the informal social support groups among junior doctors (either due to personal commitments or the age gap) and thus have fewer resources to call on during challenging rotations or clinical situations. (5) The higher rate of leaving the programme in the oldest age group may reflect an inappropriate choice of specialty or that older graduates, having had a past career and already made one major change, have more confidence to switch specialties than younger graduates.

These results should not be interpreted as older graduates are therefore less competent doctors. The ARCP/RITA assessments are there to monitor training progression against specific competencies and milestones and are not a direct measure of the quality of doctors. Some excellent doctors simply take longer to complete their training and may have gained additional skills and life experiences on this journey, learning more from their mistakes than their successes.

These results, however, should not be a cause for complacency. Longer training programmes exert additional financial pressures on training budgets and any doctor who leaves medicine altogether at this stage has had a lot of time and money invested into their training. The problem is not unique to older graduates as we observed that men, ethnic minorities and some specialties showed the same pattern of results.

In conclusion, mature doctors are part of the makeup of the NHS workforce and they widen the variety of doctors as well as bringing insights from past careers that is to be welcomed. While they appear to do better than their younger counterparts at university, they are more likely to have problems with specialty training in the UK. We believe that the causes for this are multifactorial and probably not unique to the UK but generalisable to other high-income countries like the USA and Canada, though this requires empirical confirmation. These results should be an impetus for further qualitative research to provide greater insights into why older graduates are more like to have difficulties in progression and direct action from training programmes so that they can identify problems at an earlier stage and provide greater support for such trainees as appropriate.

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Contributors YP conceived of the idea for this piece of work as her student SSC and approached the GMC for data access. YB-S acted as her SSC supervisor and provided support and training for the statistical analysis. YP undertook the initial data cleaning and analysis. YB-S checked the analyses and undertook some additional analyses. YP drafted the first version of the paper that was then edited by YB-S. All authors approved the final version of the manuscript. YP acts as the guarantor.

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Competing interests YB-S regularly sits on ARCP committees for Public Health trainees.

Ethics approval The authors did not seek formal NHS ethical approval for this study as it was a secondary data analysis of existing data that had been anonymised to ensure data confidentiality.

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Data sharing statement No additional data are available.

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Web Table 1: Association between mature status and other covariates*

	Normal age (≤28 years)	Older group (29 to 31 years)	Oldest group (≥32 years)
Gender			
Male	43.4% (15,484)	50.7 % (717)	53.3% (637)
Female	56.6% (20,214)	49.3% (697)	46.7% (559)
Ethnic Minority			
Non-Ethnic Minority	69.7% (24,883)	77.8% (1,100)	82.5% (987)
Ethnic Minority	30.3% (10,815)	22.2% (314)	17.5% (209)
Graduating University			
Mature Friendly University	94.1% (33,580)	84.8% (1,199)	80.8% (966)
Non-Mature Friendly University	5.9% (2,118)	15.2% (215)	19.2% (230)
First Specialty			
Medicine	94.8% (9,604)	2.9 (296)	2.6% (235)
ACCS & related	94.2% (5,487)	3.3% (191)	2.6% (149)
Surgery	93.0% (5,648)	4.2% (252)	2.9% (177)
GP & Public Health	90.1% (8,190)	4.9% (446)	5% (458)
O&G	93.6% (1,430)	3.3% (50)	3.1% (48)
Paediatrics	95.9% (2,674)	2.6% (73)	1.6% (44)
Pathology	89.5% (505)	5.1% (29)	5.3% (30)
Psychiatry	93.5% (903)	3.9% (38)	2.6% (25)
Radiology	94.8% (1,257)	2.9% (39)	2.3% (30)
TOTAL			100% (35,698)

*All associations were unlikely to have occurred by chance (p<0.001)

Web Table 2: Multivariable association of age at graduation and failure to progress at ARCP using more detailed age-bands and adjusted for a range of potential confounders.

	OR (95% CI)	p Value
≤ 23 years old (8,453)	1.05 (0.97 to 1.14)	0.23
24 years old (13,997)	1.0	
25 years old (7,951)	1.12 (1.04 to 1.21)	<0.001
26 years old (2,738)	1.50 (1.35 to 1.67)	<0.001
27 years old (1,614)	1.50 (1.31 to 1.72)	<0.001
28 years old (945)	1.55 (1.30 to 1.84)	<0.001
29 to 31 years old (Older group) (1,414)	1.6 (1.38 to 1.85)	<0.001
≥ 32 years old (Oldest group) (1,196)	1.95 (1.67 to 2.28)	<0.001
p-value for trend	<0.001	
Female gender	0.83 (0.78 to 0.88)	<0.001
Ethnic Minority	1.59 (1.50 to 1.69)	<0.001
Mature Friendly University	1.12 (1.00 to 1.26)	0.04
First Specialty		
ACCS & related	1.00 (0.92 to 1.08)	0.91
Medicine	1.0	
Surgery	0.84 (0.77 to 0.91)	<0.001
GP & Public Health	0.26 (0.23 to 0.29)	<0.001
O&G	2.13 (1.89 to 2.40)	<0.001
Paediatrics	0.81 (0.72 to 0.91)	<0.001
Pathology	0.83 (0.66 to 1.05)	0.12
Psychiatry	0.50 (0.41 to 0.62)	<0.001
Radiology	0.89 (0.77 to 1.03)	0.13

Web Table 3: Multivariable association of ‘Mature status’ and being asked to leave specialty at ARCP (code 4) adjusted for a range of covariates*.

	OR (95% CI)	p Value
Older (≥ 29 years)	1.81 (1.34 to 2.44)	<0.001
Younger group (≤ 28 years)		
Older group (29 to 31 years)	1.29 (0.82 to 2.03)	0.28
Oldest group (≥ 32 years)	2.48 (1.69 to 3.62)	<0.001
p-value for trend	<0.001	
Female gender	0.78 (0.65 to 0.94)	0.01
Ethnic Minority	1.52 (1.26 to 1.83)	<0.001
Mature Friendly University	1.24 (0.89 to 1.73)	0.21
First Specialty		
ACCS & related	0.99 (0.78 to 1.25)	0.95
Medicine	1.0	
Surgery	0.51 (0.39 to 0.68)	<0.001
GP & Public Health	0.22 (0.16 to 0.31)	<0.001
O&G	0.49 (0.25 to 0.79)	0.005
Paediatrics	0.74 (0.62 to 1.25)	0.48
Pathology	-	-
Psychiatry	0.47 (0.24 to 0.91)	0.03
Radiology	0.18 (0.07 to 0.43)	<0.0001

* Model run with binary age-group after adjustment for all covariates as shown in table except for the three level age group variable. This model was then rerun with the three level age-group and other covariates to examine for a dose-response effect

Appendix 1: ACRP/RITA Outcome categories

ACRP	
Outcome 1	Satisfactory Progress
Outcome 2	Unsatisfactory Progress - Development of specific competences required, additional training time not required
Outcome 3	Unsatisfactory Progress - Inadequate progress by the trainee, additional training time required
Outcome 4	Unsatisfactory Progress - Released from the training programme with or without specified competences; trainee will be required to give up their National Training Number.
Outcome 5	Incomplete evidence presented.
Outcome 6	Recommendation for completion of training.
Outcome 7	Fixed-term specialty outcome:
Outcome 7.1	- Satisfactory progress in or completion of the LAT / FTSTA placement.
Outcome 7.2	- Development of Specific Competences Required – additional training time not required
Outcome 7.3	- Inadequate progress by trainee
Outcome 7.4	- Incomplete evidence presented
Outcome 8	Out of programme for research, approved clinical training or a career break (OOPR/OOPT/OOPC).
Outcome 9	For doctors undertaking top-up training in a training post.
RITA	
C	Satisfactory progress
D	Recommendation for targeted training
E	Records a recommendation for intensified supervision/repeated experience.
F	Records out-of-programme experience (including maternity leave)
G	Provides a final record of satisfactory progress on completion of training.