

BMJ Open Preschool children's vision screening in New Zealand: a retrospective evaluation of referral accuracy

Miriam A Langeslag-Smith,¹ Alain C Vandal,^{2,3} Vincent Briane,² Benjamin Thompson,^{4,5} Nicola S Anstice⁵

To cite: Langeslag-Smith MA, Vandal AC, Briane V, *et al.* Preschool children's vision screening in New Zealand: a retrospective evaluation of referral accuracy. *BMJ Open* 2015;5:e009207. doi:10.1136/bmjopen-2015-009207

► Prepublication history for this paper is available online. To view these files please visit the journal online (<http://dx.doi.org/10.1136/bmjopen-2015-009207>).

Received 25 June 2015
Revised 17 August 2015
Accepted 21 September 2015

ABSTRACT

Objectives: To assess the accuracy of preschool vision screening in a large, ethnically diverse, urban population in South Auckland, New Zealand.

Design: Retrospective longitudinal study.

Methods: *B4 School Check* vision screening records (n=5572) were compared with hospital eye department data for children referred from screening due to impaired acuity in one or both eyes who attended a referral appointment (n=556). False positive screens were identified by comparing screening data from the eyes that failed screening with hospital data. Estimation of false negative screening rates relied on data from eyes that passed screening. Data were analysed using logistic regression modelling accounting for the high correlation between results for the two eyes of each child.

Primary outcome measure: Positive predictive value of the preschool vision screening programme.

Results: Screening produced high numbers of false positive referrals, resulting in poor positive predictive value (PPV=31%, 95% CI 26% to 38%). High estimated negative predictive value (NPV=92%, 95% CI 88% to 95%) suggested most children with a vision disorder were identified at screening. Relaxing the referral criteria for acuity from worse than 6/9 to worse than 6/12 improved PPV without adversely affecting NPV.

Conclusions: The *B4 School Check* generated numerous false positive referrals and consequently had a low PPV. There is scope for reducing costs by altering the visual acuity criterion for referral.

INTRODUCTION

Recently published best practice guidelines recommend two vision screening protocols in 3–6 year-old children:^{1–3} measurement of acuity with a single letter or picture surrounded by crowding bars at 1.5 m²; or auto-refraction.³ However, these guidelines have not been adopted in New Zealand (NZ) where vision screening is conducted on >90% of preschool children by lay-screeners as part of the *B4 School Check*, a national well-

Strengths and limitations of this study

- First assessment of the efficacy of the *B4 School Check* vision screening that was implemented in New Zealand in 2008.
- Retrospective longitudinal evaluation of vision screening results from 1 year of the *B4 School Check* programme with records available for over 5500 children.
- As full ophthalmic examination results were only available for children who were referred from vision screening, negative predictive value, sensitivity and specificity were estimated using data from both eyes of children and logistic regression modelling.
- Inter-eye correlation was taken into account by introducing a random effect assuming common positive predictive value and negative predictive value. However, using data from both eyes of an individual remains controversial.

child programme. The current NZ protocol uses the Parr Vision chart at 4 m to screen for amblyopia, strabismus and significant refractive errors. Preschool children who fail vision screening are referred to their local hospital eye department where they are further evaluated by an optometrist, orthoptist or ophthalmologist. The Parr Chart, which is the same as the Sheridan Gardiner test⁴ with the exception that it is used at 4 m rather than 3 m, contains a single letter surrounded by confusion bars on each page. However, there is no published validation of the Parr Chart against other commonly used paediatric tests. Neither the Parr Chart, nor the Sheridan Gardiner test, meet the International Visual Acuity Chart Guidelines⁵ due to the non-standardised progression of letter sizes and the use of letters which do not have equal legibility. Although the Sheridan Gardiner and Parr tests are quick and easy to administer,⁴ the uncrowded versions of these tests have been found to produce a high number of false positive



CrossMark

For numbered affiliations see end of article.

Correspondence to

Dr Nicola S Anstice;
n.anstice@auckland.ac.nz

results when used in a screening context.⁶ Screening with the crowded version of the Parr Chart was implemented with the introduction of the *B4 School Check* in 2008/2009 because crowding elements may improve the sensitivity of vision tests for the detection of amblyopia.^{7,8} However, a formal evaluation is required to examine the accuracy of vision screening for NZ children using the crowded Parr Chart.

The primary aim of this study was to evaluate the positive predictive value (PPV) of the *B4 School Check* pre-school vision screening programme and the precision of the acuity test used in a large cohort of ethnically diverse children living in an urban environment.

METHODS

A retrospective review of children who participated in the *B4 School Check* Vision Screening Programme from 1 March 2010 to 28 February 2011 was conducted. South Auckland was chosen as the study site to provide a wide demographic; it has the largest Māori and Pacifica populations in NZ and a large number of low socioeconomic status (SES) residents.⁹ The study was approved by the Health and Disability North X Regional Ethics Committee and all data was de-identified to preserve patients' anonymity.

Participants

Children who were referred from vision screening were identified through the Ministry of Health Database and data were collected on: age at screening, ethnicity, deprivation class (NZDep index)¹⁰ and unaided visual acuity (UVA) for the left and right eye. Basic demographic information was also gathered for children who declined vision screening. For children referred from the *B4 School Check* and seen at the hospital eye department in South Auckland, NZ, data were collected on UVA for the right and left eyes at the hospital eye assessment, ocular diagnosis and best corrected visual acuity (BCVA). Where possible, a reason was sought for non-attendance of children referred from the vision screening but not seen at Counties Manukau District Health Board (CMDHB).

B4 school check vision screening

The *B4 School Check* measured monocular UVA using the Parr Chart and either an optotype matching or letter naming task depending on the ability of the child. Uncrowded optotypes were used only when crowded measures could not be obtained. Referral for further assessment occurred if the UVA was worse than 6/9 in either eye or there was more than one line difference in UVA between the eyes. Unaided acuity measurements from vision screening were converted to logarithmic minimum angle of resolution (logMAR) results for comparison purposes.

Hospital eye examination

Initial assessment included a comprehensive ocular and medical history, monocular distance UVA with either the Crowded Keeler logMAR or Crowded Kay Pictures charts,^{11,12} assessment of ocular alignment and a red reflex check for media opacities. Additional tests were conducted at the discretion of the clinician. Children were discharged after initial consultation if UVA measured at the hospital eye examination was 6/9.5⁺¹ (0.18 logMAR units) or better and no other abnormalities were detected.¹³ Refraction was not performed on children if UVA measured at the hospital eye appointment was better than 6/9.5⁺¹ (0.18 logMAR), unless other ocular abnormalities were detected at the initial consultation. Children with UVA between 6/9.5⁻² (0.25 logMAR units) and 6/9.5 (0.20 logMAR units) were re-examined within 2 months and scheduled for a comprehensive assessment if no improvement was found. All other children received a cycloplegic refraction and ocular health examination. Where required, spectacles were prescribed based on standard guidelines¹⁴ adopted by the hospital eye department. Children prescribed spectacles and/or amblyopia therapy were examined every 6–8 weeks until BCVA improved to age-appropriate levels or it was deemed that BCVA had stabilised and was unlikely to improve further.

Data analysis

As the study cohort only included children with a positive screening test who were referred for further care, analysis of PPV, negative predictive value (NPV), sensitivity and specificity was only possible if each eye of a child was considered separately as a statistical unit.^{15,16} A mixed logistic regression model was constructed, taking into account the correlation between the two eye measurements by introducing a random effect and assuming common PPV and NPV (as well as sensitivity and specificity) regardless of eye. PPV and NPV were estimated by setting the results of the hospital eye examination as the binary dependent variable (diagnosed vision disorder or no disorder) and regressing it on the screening UVA pass/fail criterion. Sensitivity and specificity were estimated similarly, but setting the results of the UVA pass/fail criterion as a binary random variable and regressing on the hospital eye examination result. The effect of varying the pass/fail criterion for UVA measured at the vision screening on PPV and NPV was examined by calculating odds ratios (ORs) for a range of acuity referral criteria, which enabled the drawing of a receiver operating characteristic curve. Mixed logistic regression modelling to determine PPV, NPV, sensitivity and specificity data was performed using Statistical Analysis System (SAS Institute) V.9.4 and R (The R Foundation) V.3.1. The Bland Altman plot was constructed in SPSS (IBM) V.22.

Definition of terms

A vision disorder was classified as reduced UVA (worse than 0.20 logMAR) in one or both eyes that was due to

uncorrected refractive error or amblyopia, the presence of strabismus with or without reduced UVA, or ocular pathology. Children were classified as false positive referrals if they had normal UVA (better than 6/9.5 at the hospital eye department) or reduced UVA that normalised over time without intervention. The further classification of refractive error into either myopia or hyperopia depended on the spherical component of the cycloplegic retinoscopy. Myopia ≤ -1.00 DS, hyperopia $\geq +2.00$ DS, astigmatism ≥ 1.00 DC, and anisometropia ≥ 1.00 D difference between the eyes in any meridian were deemed significant refractive errors.¹⁷ PPV was defined as the proportion of eyes referred from vision screening that had a vision defect diagnosed and treated at the hospital eye department. NPV was the proportion of eyes that had normal vision at screening and were found to be visually normal at the hospital eye examination. Sensitivity provided an estimate of the true positive results and specificity an estimate of the true negative results of the acuity test used at screening.

RESULTS

Cohort

A breakdown of the records that were available for retrospective review is shown in figure 1. Six hundred and two children declined vision screening and over half were of Māori (23.1%) or Pacifica ethnicity (37.4%), which is higher than the proportion of these children who received vision screening (15.7% and 25%, respectively). SES appeared to be a factor affecting whether vision screening was accepted or declined, as 54% of the

families who declined screening lived in the most deprived areas, whereas only 12% of the families who declined lived in the most affluent areas.

The average age at vision screening was 52 ± 4 months (range 37–70 months). More than 14% of children were referred from vision screening ($n=698$) and the majority of these (80%) were seen at the hospital eye department, CMDHB. For the children not evaluated at the ophthalmology department ($n=140$) the most common reasons were: unable to contact parents/caregivers (26%), failure to attend scheduled appointments (22%), seen by another eye care provider (9%), moved outside the area (9%) or ineligible for subsidised health-care (8%). No information on non-attendance was available for 35 children.

Estimates of PPV, NPV, sensitivity and specificity

PPV of the *B4 School Check* vision screening was low (31%, 95% CIs 26% to 38%) while estimates of NPV were high (92%, 95% CI 88% to 95%). OR analysis found that increasing the referral cut-off to UVA worse than 6/12 would improve the PPV significantly (53%, 95% CI 42% to 64%) but would not lower the NPV markedly (86%, 95% CI 82% to 89%). Sensitivity for the screening was good (0.89, 95% CI 0.84 to 0.93), however, specificity was low (0.47, 95% CI 0.41 to 0.53). A large number of children ($n=112$) had reduced UVA at the initial eye examination, but UVA measurement normalised over time without intervention. These children were classified as false positive referrals, as no treatment was provided and UVA improved with time.

Figure 1 Retrospective records review of *B4 School Check* vision screening results in the South Auckland region of New Zealand from 1 March 2010 to 28 February 2011.

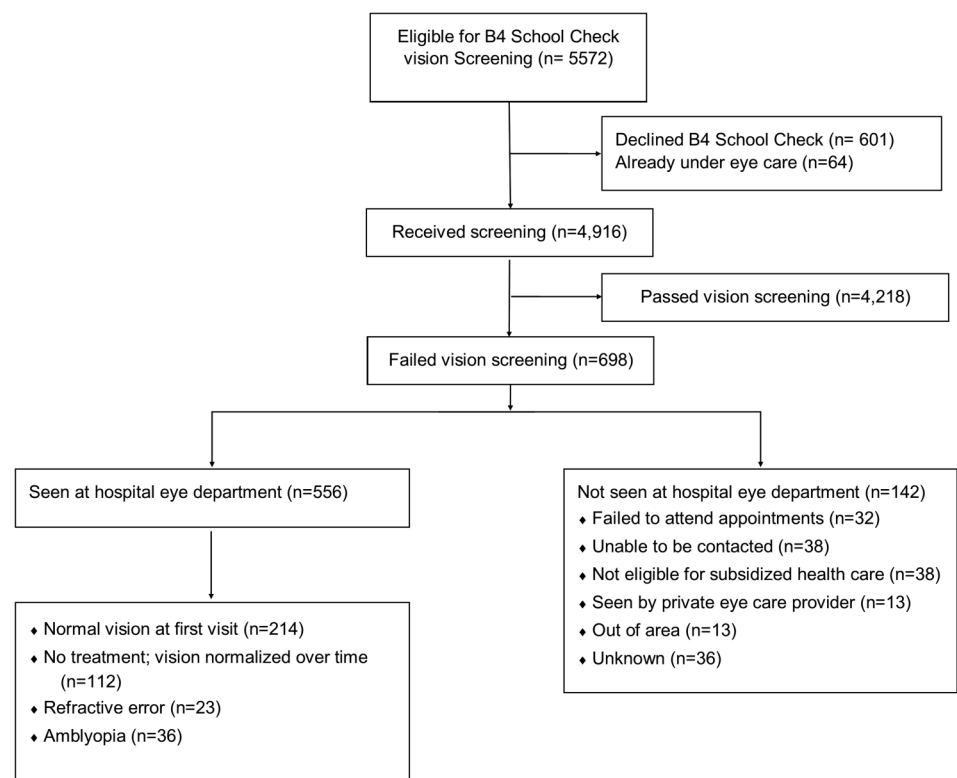
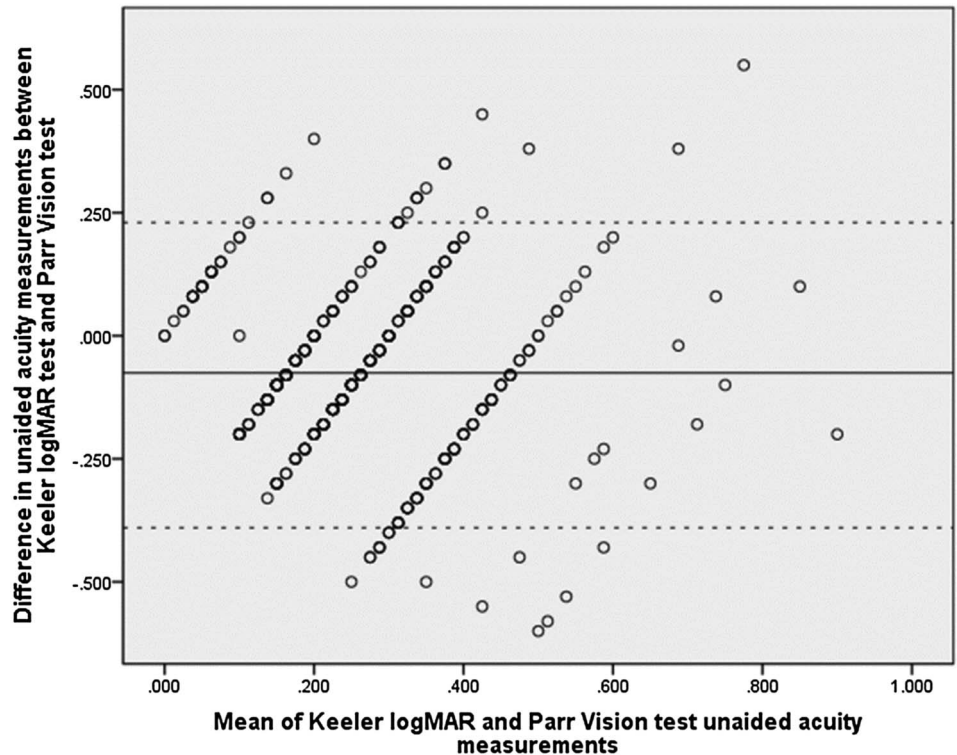


Figure 2 Bland-Altman analysis of unaided visual acuity (UVA) of the right eye of children screened as part of the *B4 School Check* (Parr Vision Test) who attended the ophthalmology department where UVA was measured with the Keeler logMAR Test. The solid black line represents the average difference between these two measures (in logMAR), while the dashed black lines show the 95% confidence limits of agreement between the two measures.



Comparison of UVA between vision screening and hospital eye assessment

There was poor agreement between UVA measured at vision screening using the Parr Chart and UVA measured at the hospital eye department using the Crowded Keeler (n=495) or Crowded Kay Picture tests (n=56). **Figure 2** shows the level of agreement between UVA measurements made using the Parr Chart and the Crowded Keeler logMAR test (mean difference in UVA = -0.08 ± 0.16 logMAR; $t = -10.97$, $p < 0.0001$). Wide 95% limits of agreement (-0.39 to $+0.23$ logMAR) suggested the two measures did not produce equivalent results in the same cohort of children. As only a small proportion of children had UVA measured with the Crowded Kay Pictures test they were not included in the Bland-Altman analysis.

Prevalence of vision disorders in the cohort

The most common ocular diagnosis was significant refractive error (n=184, [table 1](#)), while strabismus (n=23), amblyopia (n=36) and ocular pathology (n=6) were less prevalent in this cohort of children.

Reduced UVA was present at the hospital eye assessment in all children with myopia; however, 12 children with other refractive errors had UVA better than 6/9 (0.18 logMAR). **Figure 3** demonstrates that the different ocular diagnoses, including unexplained reduced UVA, had overlapping acuity levels. Therefore, UVA alone could not be used to differentiate between different visual disorders. Even when acuity was measured with standardised charts by experienced eye care providers, there was no reliable difference in UVA in

children with refractive error and those with unexplained reduced UVA which normalised over time.

Overall, 230 children required ophthalmic treatment (spectacle prescription, patching or strabismus surgery) which improved acuity from 0.34 ± 0.20 logMAR at initial assessment to 0.12 ± 0.11 logMAR at discharge. When a success criterion of BCVA in the better eye of 6/12 (0.30 logMAR which is equal to the NZ driving standard) was applied, treatment was unsuccessful in only two children due to pathology. A further 10 children had BCVA of worse than 6/12 in the poorer eye, primarily due to poor compliance with amblyopia treatment (n=9).

DISCUSSION

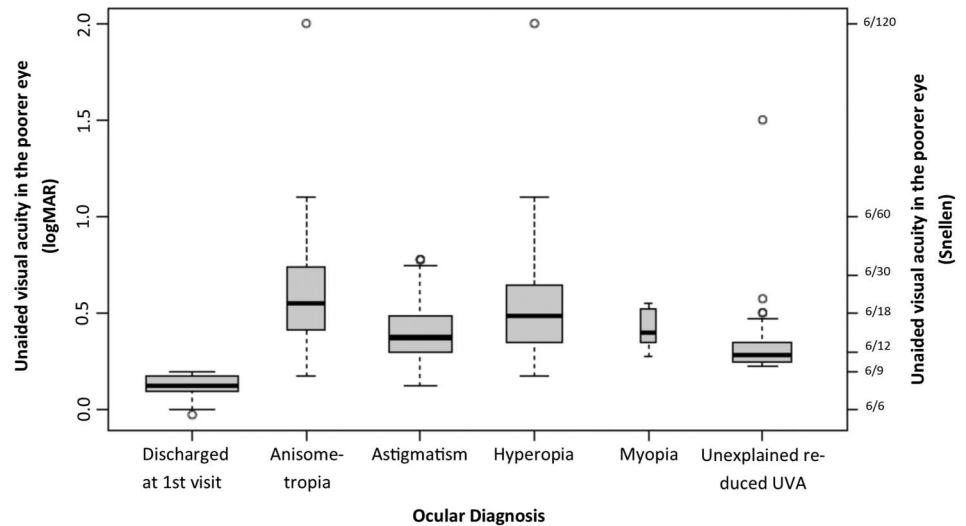
The principal finding of this study was the low PPV of the current NZ preschool vision screening; only 31% of children who were referred from vision screening had a

Table 1 Prevalence of significant refractive errors in children (n=184) referred from *B4 School Check* in Counties Manukau region who had visual acuity worse than 0.2 logMAR (6/9.5) in at least one eye

Refractive error	Definition	Number of children
Hyperopia	$\geq +2.00$ DS	95
Myopia	≤ -1.00 DS	7
Astigmatism	≥ 1.00 DS	121
Anisometropia	Interocular difference of ≥ 1.00 D	42

Some children had more than one type of refractive error.

Figure 3 Boxplot representing the level of unaided visual acuity (UVA) for children with refractive errors, unexplained reduced UVA and for those discharged after the initial consultation. The width of each boxplot corresponds to the number of observations, although some children had more than one refractive diagnosis.



diagnosed visual deficit. The poor PPV appeared to be primarily due to the poor specificity of the screening test and the acuity criteria used, as many visually normal children failed screening with the cut-off criterion of UVA worse than 6/9. Modelling of our data suggests UVA worse than 6/12, one line larger than the current criteria, may provide a more accurate referral cut-off. The low PPV was also due to the significant number of children (16%) who had unexplained reduced UVA at the initial hospital eye examination which normalised over time. It is unclear the reason for this reduced UVA and what processes occur that lead to acuity normalising over time. Retinal structure is fully developed by 4 years of age,^{18 19} although other aspects of the visual pathway are still developing and the effects of letter crowding may be more pronounced in children up until the teenage years.^{20 21} Other factors may also play a role in the poor specificity of the current preschool vision screening programme. For example, the Vision in Preschoolers (VIP) group has recommended 1.5 m testing distance when screening is conducted by nurses or lay screeners² which is much shorter than the 4 m distance employed in the *B4 School Check* vision screening. The VIP study group has also found that testing by nurses and lay screeners was less successful compared with licensed eye-care professionals,² although trained nurses were better at performing acuity measures than lay screeners. This suggests that additional training for vision screeners could improve outcomes from preschool vision screening in NZ. Finally our findings suggest, that while the *B4 School Check* reaches the majority of children in South Auckland approximately 10% of children declined vision screening and half of these children are of Māori/Pasifika ethnicity. This is consistent with Ministry of Health (NZ) data from 2011 which showed that Māori children have a risk ratio (RR) for declining vision screening of 2.7 (95% CI (2.3 to 3.1), $p < 0.0000001$) compared to non-Māori, non-Pasifika children. Pasifika children had an RR of 3.1 (95% CI (2.7 to

3.6), $p < 0.0000001$) compared to the same reference group, as well as a higher RR of declining screening than Māori (95% CI (1.04 to 1.33), $p = 0.011$). Māori and Pasifika face significant barriers to healthcare including socioeconomic disadvantage,²² lack of cultural awareness in health care²² and cost of care.^{23 24} Although the *B4 School Check* is free, the cost of lost earnings for parents, transport and treatment (spectacle prescriptions and patching), need to be considered when identifying barriers to service utilisation.²⁵ As such, socioeconomic factors also played a role in the uptake of preschool vision screening and the Ministry of Health 2011 *B4 School Check* data found the RR of declining between quintile 5 (lowest) and quintiles 1 (highest) to 3 aggregated was 2.8 (95% CI (2.4 to 3.2), $p < 0.0000001$).

This is the first study to evaluate the efficacy of the *B4 School Check* vision screening since the programme was implemented throughout NZ in 2008. It involved analysis of screening results from a large number of ethnically diverse children over a 1 year period and was primarily designed to investigate the positive predictive value of the current preschool vision screening protocol. Additionally, most studies find a proportion of children (2–30%^{26–28}) with unexplained reduced UVA that ranges although few follow these children to see if acuity improves with time. One of the strengths of this study was that these children received follow-up care and normalisation of UVA was able to be observed. PPV, NPV, sensitivity and specificity were all obtained by data modelling from children referred from the *B4 School Vision Check* and therefore are only estimates of the true values. Each eye was considered as an individual statistical unit, which made it possible to evaluate true positive, false positive, true negative and false negative results from the data available. Nevertheless, the use of data from both eyes of the same individual is controversial. Recent publications argue that using data from both eyes is valid, as long as appropriate inter-eye correlations

are incorporated into the analysis^{15 16 29} and in this study was taken into account by introducing a random effect into the logistic regression model. Based on our results, approximately 4.5% of children screened had a vision deficit, which is consistent with the prevalence of vision disorders found in other studies.^{17 30 31} However, some studies have shown a much higher prevalence of vision problems and a recent study from the USA found 20% of 3–5 year-old children had a vision disorder.³² If the true prevalence of refractive error is closer to one in five, then our estimates of NPV, sensitivity and specificity are too low. It is possible that some children with mild-moderate hyperopia and astigmatism passed the *B4 School Check* vision screening and data from these children were not available for analysis.³³ As data were only available from individual eyes that had ‘passed’ vision screening we were only able to estimate the false negative rate of the *B4 School Check* programme. Studies designed to specifically address false negative rates are required to evaluate the accuracy of our estimates.

The poor PPV and low specificity found in NZ preschool vision screening is also seen in many other screening programmes, such as breast screening³⁴ and newborn hearing screening.³⁵ Improvements in the specificity of newborn hearing screening occurred by rescreening all children who failed the initial test, with 80% passing a rescreening test.³⁶ Similarly, the Multi-Ethnic Pediatric Eye Disease Study found 25% of children had reduced UVA at initial assessment, which improved on retesting in 81% of participants.³⁷ A similar strategy for preschool vision screening has been proposed whereby those children with reduced vision at screening should undergo a full refractive examination before referral to hospital eye services.³⁸ Alternatively evaluating the use of other acuity charts for performing preschool vision screening in NZ may be warranted. There was poor agreement between UVA measurements with the screening test and the acuity chart used by hospital staff to assess vision. The difference between the results for the two charts is likely to be due to a combination of: (1) differences between the tests themselves for example, the absence of a 0.1 logMAR line on the Parr Chart, (2) the different training levels of the staff administering the tests and (3) different testing environments. While our study design does not allow us to identify which of these was the most important factor, the results indicate that the screening and clinical tests should be made as similar as possible to minimise false positive referrals. Other studies show higher specificity (61%³ to 68%³⁹) using alternative acuity charts when sensitivity was set to a similar level (90%) as that found in the current study, and this may reflect unequal progression of letter sizes and the absence of the 6/7.5 (0.1 logMAR) line on the NZ screening test.

The PPV of preschool vision screening in South Auckland, NZ was low, with many visually normal children referred due to the poor specificity of the current B4 School screening protocol. This has significant

implications for the healthcare system, by putting an unnecessary burden on hospital eye departments. Altering the referral criteria, to refer children with UVA worse than 6/12, improved the PPV without affecting the NPV and this simple solution should be further investigated. In school-aged children, with more advanced cognitive development and attention levels than the preschool children who were the focus of the present study, acuity of 6/9.5 (0.2 logMAR) is rarely associated with a visual problem that requires treatment.^{40 41} Internationally, a variety of acuity cut-offs are employed for vision screening; a recent vision survey in East Timor found that only 0.2% of students with a significant refractive error passed a screening using a 6/12 Tumbling E target.⁴² In a Native American population with a high prevalence of astigmatism, the highest sensitivity (97%) and specificity (79%) was found with a fail criteria of UVA worse than 6/19 combined with automated measures of corneal curvature.⁴³

Data modelling suggested that both the NPV and sensitivity of vision screening were high. However, as data were only available from children who were referred from vision screening, analysis of NPV, sensitivity and specificity relied on utilising data from both eyes of each child, taking into account the correlation between the two eyes. While this should not affect our estimates of PPV, a large scale study of vision disorders in NZ children and prospective evaluation of preschool screening protocols is needed to find the actual prevalence of eye conditions in NZ and provide accurate estimates of NPV, sensitivity and specificity of the *B4 School Check*.

Author affiliations

¹Department of Ophthalmology, Counties Manukau District Health Board, Auckland, New Zealand

²Ko Awatea, Counties Manukau District Health Board, Auckland, New Zealand

³Faculty of Health and Environmental Sciences, Auckland University of Technology, Auckland, New Zealand

⁴Department of Optometry and Vision Science, The University of Waterloo, Waterloo, Ontario, Canada

⁵School of Optometry and Vision Science, The University of Auckland, Auckland, New Zealand

Acknowledgements The authors thank Associate Professor Robert Jacobs and Dr John Phillips for commenting on an earlier draft of this manuscript.

Contributors MAL-S, BT and NSA designed the study; MAL-S collected data; ACV and VB performed the analysis; MAL-S and NSA drafted the manuscript and all authors approved the final version to be published.

Funding This work was supported by the Arthur D Bronlund Trust, CCRep and University of Auckland Faculty Research Development Fund Grants (3704420).

Competing interests None declared.

Ethics approval This study received ethics approval from the Northern X Regional Ethics Committee (NZ).

Provenance and peer review Not commissioned; externally peer reviewed.

Data sharing statement Additional data can be accessed via the Dryad data repository at <http://datadryad.org/> with the doi:10.5061/dryad.ht357.

Open Access This is an Open Access article distributed in accordance with the Creative Commons Attribution Non Commercial (CC BY-NC 4.0) license,

which permits others to distribute, remix, adapt, build upon this work non-commercially, and license their derivative works on different terms, provided the original work is properly cited and the use is non-commercial. See: <http://creativecommons.org/licenses/by-nc/4.0/>

REFERENCES

- Cotter SA, Cyert LA, Miller JM, *et al.* Vision screening for children 36 to <72 months: recommended practices. *Optom Vis Sci* 2015;92:6–16.
- Vision in Preschoolers (VIP) Study Group. Preschool vision screening tests administered by nurse screeners compared with lay screeners in the vision in preschoolers study. *Invest Ophthalmol Vis Sci* 2005;46:2639–48.
- Schmidt PP, Maguire M, Dobson V, *et al.* Comparison of preschool vision screening tests as administered by licensed eye care professionals in the Vision in Preschoolers study. *Ophthalmology* 2004;111:637–50.
- Sheridan MD, Gardiner PA. Sheridan-Gardiner test for visual acuity. *BMJ* 1970;2:108–9.
- Universal CO. Visual Acuity Measurement Standard. Visual Functions Committee. 1984.
- Anstice N, Spink J, Abdul-Rahman A. Review of preschool vision screening referrals in South Auckland, New Zealand. *Clin Exp Optom* 2012;95:442–8.
- Simmers AJ, Gray LS, McGraw PV, *et al.* Contour interaction for high and low contrast optotypes in normal and amblyopic observers. *Ophthalmic Physiol Opt* 1999;19:253–60.
- Formankiewicz MA, Waugh SJ. The effects of blur and eccentric viewing on adult acuity for pediatric tests: implications for amblyopia detection. *Invest Ophthalmol Vis Sci* 2013;54:6934–43.
- Atkinson J, Salmond C, Crampton P. *NZDep2013 index of deprivation*. New Zealand: Ministry of Health, 2014.
- Salmond C, Crampton P, Atkinson J. *NZDep2006 index of deprivation*. Department of Public Health, University of Otago Wellington, 2007.
- McGraw PV, Winn B. Glasgow Acuity Cards: a new test for the measurement of letter acuity in children. *Ophthalmic Physiol Opt* 1993;13:400–4.
- Kay H. New method of assessing visual acuity with pictures. *Br J Ophthalmol* 1983;67:131–3.
- Shea SJ, Gaccon L. In the absence of strabismus what constitutes a visual deficit in children? *Br J Ophthalmol* 2006;90:40–3.
- Donahue SP. Prescribing spectacles in children: a pediatric ophthalmologist's approach. *Optom Vis Sci* 2007;84:110–4.
- Murdoch IE, Morris SS, Cousens SN. People and eyes: statistical approaches in ophthalmology. *Br J Ophthalmol* 1998;82:971–3.
- Glynn RJ, Rosner B. Regression methods when the eye is the unit of analysis. *Ophthalmic Epidemiol* 2012;19:159–65.
- Williams C, Northstone K, Howard M, *et al.* Prevalence and risk factors for common vision problems in children: data from the ALSPAC study. *Br J Ophthalmol* 2008;92:959–64.
- Cornish EE, Hendrickson AE, Provis JM. Distribution of short-wavelength-sensitive cones in human fetal and postnatal retina: early development of spatial order and density profiles. *Vision Res* 2004;44:2019–26.
- Diaz-Araya C, Provis JM. Evidence of photoreceptor migration during early foveal development: a quantitative analysis of human fetal retinae. *Vis Neurosci* 1992;8:505–14.
- Jeon ST, Hamid J, Maurer D, *et al.* Developmental changes during childhood in single-letter acuity and its crowding by surrounding contours. *J Exp Child Psychol* 2010;107:423–37.
- Norgett Y, Siderov J. Crowding in children's visual acuity tests—effect of test design and age. *Optom Vis Sci* 2011;88:920–7.
- Ellison-Loschmann L, Pearce N. Improving access to health care among New Zealand's Maori population. *Am J Public Health* 2006;96:612.
- Jatrana S, Crampton P. Primary health care in New Zealand: Who has access? *Health Policy* 2009;93:1–10.
- Jatrana S, Crampton P, Norris P. Ethnic differences in access to prescription medication because of cost in New Zealand. *J Epidemiol Community Health* 2011;65:454–60.
- Mitsuhiro MH, Berezovsky A, Belfort R, Jr., *et al.* Uptake, barriers and outcomes in the follow-up of patients referred for free-of-cost cataract surgery in the Sao Paulo eye study. *Ophthalmic Epidemiol* 2015;22:253–9.
- Maul E, Barroso S, Munoz SR, *et al.* Refractive error study in children: results from La Florida, Chile. *Am J Ophthalmol* 2000;129:445–54.
- Paudel P, Ramson P, Naduvilath T, *et al.* Prevalence of vision impairment and refractive error in school children in Ba Ria—Vung Tau province, Vietnam. *Clin Experiment Ophthalmol* 2014;42:217–26.
- Gao Z, Meng N, Muecke J, *et al.* Refractive error in school children in an urban and rural setting in Cambodia. *Ophthalmic Epidemiol* 2012;19:16–22.
- Kerr NM, Chew SS, Eady EK, *et al.* Diagnostic accuracy of confrontation visual field tests. *Neurology* 2010;74:1184–90.
- Robaei D, Rose K, Ojaimi E, *et al.* Visual acuity and the causes of visual loss in a population-based sample of 6-year-old Australian children. *Ophthalmology* 2005;112:1275–82.
- Robaei D, Rose KA, Ojaimi E, *et al.* Causes and associations of amblyopia in a population-based sample of 6-year-old Australian children. *Arch Ophthalmol* 2006;124:878–84.
- Ying GS, Maguire MG, Cyert LA, *et al.* Prevalence of vision disorders by racial and ethnic group among children participating in head start. *Ophthalmology* 2014;121:630–6.
- O'Donoghue L, Rudnicka AR, McClelland JF, *et al.* Visual acuity measures do not reliably detect childhood refractive error—an epidemiological study. *PLoS ONE* 2012;7:e34441.
- Elmore JG, Barton MB, Mocerri VM, *et al.* Ten-year risk of false positive screening mammograms and clinical breast examinations. *N Engl J Med* 1998;338:1089–96.
- Clemens CJ, Davis SA, Bailey AR. The false-positive in universal newborn hearing screening. *Pediatrics* 2000;106:e7.
- Clemens CJ, Davis SA. Minimizing false-positives in universal newborn hearing screening: a simple solution. *Pediatrics* 2001;107:e29.
- Tarczy-Hornoch K, Cotter SA, Borchert M, *et al.* Prevalence and causes of visual impairment in Asian and Non-Hispanic white preschool children: Multi-Ethnic Pediatric Eye Disease Study. *Ophthalmology* 2013;120:1220–6.
- Williamson TH, Andrews R, Dutton GN, *et al.* Assessment of an inner city visual screening programme for preschool children. *Br J Ophthalmol* 1995;79:1068–73.
- Ying GS, Kulp MT, Maguire M, *et al.* Sensitivity of screening tests for detecting vision in preschoolers-targeted vision disorders when specificity is 94%. *Optom Vis Sci* 2005;82:432–8.
- Saxena R, Vashist P, Tandon R, *et al.* Accuracy of visual assessment by school teachers in school eye screening program in Delhi. *Indian J Community Med* 2015;40:38–42.
- Murthy GV. Vision testing for refractive errors in schools: 'screening' programmes in schools. *Community Eye Health* 2000;13:3–5.
- Ramke J, du Toit R, Roberts B, *et al.* Vision screening of children attending primary school in rural Timor-Leste. *Clin Experiment Ophthalmol* 2011;39:377–8.
- Miller JM, Dobson V, Harvey EM, *et al.* Cost-efficient vision screening for astigmatism in Native American preschool children. *Invest Ophthalmol Vis Sci* 2003;44:3756–63.