BMJ Open  Opioid substitution therapy as a strategy to reduce deaths in prison: retrospective cohort study

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ABSTRACT

Objectives: To describe deaths in prison among opioid-dependent people, and examine associations between receipt of opioid substitution therapy (OST) and risk of death in prison.

Design: Retrospective cohort study.

Setting: Adult prisons in New South Wales (NSW), Australia.

Participants: 16 715 opioid-dependent people who were received to prison between 2000 and 2012.

Interventions: Opioid substitution therapy.

Primary outcome measures: Natural and unnatural (suicide, drug-induced, violent and other injury) deaths in prison.

Results: Cohort members were in prison for 30 998 person-years (PY), during which time there were 51 deaths. The all-cause crude mortality rate (CMR) in prison was 1.6/1000 PY (95% CI 1.2 to 2.2/1000 PY), and the unnatural death CMR was 1.1/1000 PY (95% CI 0.8 to 1.6/1000 PY). Compared to time out of OST, the hazard of all-cause death was 74% lower while in OST (adjusted HR (AHR): 0.26; 95% CI 0.13 to 0.50), and the hazard of unnatural death was 87% lower while in OST (AHR: 0.13; 95% CI 0.05 to 0.35). The all-cause and unnatural death CMRs during the first 4 weeks of incarceration were 6.6/1000 PY (95% CI 3.8 to 10.6/1000 PY) and 5.5/1000 PY (95% CI 2.9 to 9.4/1000 PY), respectively. Compared to periods not in OST, the hazard of all-cause death during the first 4 weeks of incarceration was 94% lower while in OST (AHR: 0.06; 95% CI 0.01 to 0.48), and the hazard of unnatural death was 93% lower while in OST (AHR: 0.07; 95% CI 0.01 to 0.53).

Conclusions: Mortality of opioid-dependent prisoners was significantly lower while in receipt of OST.

INTRODUCTION

Deaths in prison are a serious public health issue, raising questions as to the quality of care and supervision provided by correctional authorities. Reflecting such concerns, deaths in prison usually result in extensive postmortem inquiries and may also lead to litigation against correctional authorities and healthcare providers.1

Unnatural deaths in prison are of particular concern due to their preventable nature. Unnatural deaths include suicides, violent or drug-induced deaths and other injury-related deaths. Studies in developed countries have found that unnatural deaths comprise 48–59% of all deaths in prison, with suicide being the most common cause of unnatural death.2–5 These figures suggest considerable opportunities to reduce unnatural deaths in prison.

Opioid-dependent people commonly experience imprisonment,6 and there are several reasons to believe that opioid-dependent prisoners may be at particular risk of unnatural death in prison. Drug withdrawal has been implicated as a possible trigger for suicide in the first days of incarceration.7,8 Additionally, use of illicitly obtained opioids while in prison9 carries with it the risk of overdose. To the best of our knowledge, deaths in prison specifically among opioid-dependent people have not been described previously. Furthermore, no studies have considered whether treatment for opioid dependence during incarceration reduces mortality risk, as it does among opioid-dependent people residing in the community.10 In this study, we describe deaths in prison in a cohort of opioid-dependent people and examine whether receiving opioid substitution therapy (OST) reduces the risk of death in prison.
METHODS
Setting
New South Wales (NSW) is the most populous state in Australia and houses around one-third of the country’s prisoners, or 9645 people at 31 June 2012.11 Health services in NSW prisons are provided under the Ministry of Health, and there is a well-established prison OST programme that operates as part of the statewide opioid treatment programme.12 13 People who enter prison may continue treatment while incarcerated, and OST can be started during incarceration if clinically indicated.12 Clinical indications for OST are the same as in community settings.13 Most prisoners receiving OST are prescribed methadone due to the longer dosing procedure associated with buprenorphine and concerns about diversion of buprenorphine. In 2012, 16% of NSW prisoners were prescribed OST.11 12 Coverage of OST (the proportion of opioid-dependent prisoners receiving treatment) has been estimated at 43%.14

Data sources and linkage
Data for this retrospective cohort study were extracted from administrative datasets used to record OST programmes, offending and incarceration, and mortality. The Pharmaceutical Drugs of Addiction System (PHDAS) is a database of all methadone and buprenorphine recipients in NSW since 1985. The PHDAS records each patient’s full name, date of birth, gender and clinical variables including dates of OST entry and exit. Identifying variables in this dataset are considered to be of high accuracy as proof of identity must be shown to the prescribing doctor before a prescription can be issued. Data used for linkage were for the calendar years 1985–2010.

Offending and incarceration data were sourced from the NSW Bureau of Crime Statistics and Research (BOCSAR) Re-offending Database (ROD). This database includes dates and types of offences in criminal cases heard by local, district and supreme courts of NSW, and dates of receptions to and discharges from the custody of Corrective Services NSW. All available data were provided for linkage; this included offending data for the period 1 December 1993 to 31 December 2011 and incarceration data for the period 1 January 2000 to 31 March 2012.

The National Death Index (NDI) is a database held by the Australian Institute of Health and Welfare (AIHW), containing mortality data collected from all the Registry of Births, Deaths and Marriages in each Australian State and Territory. Causes of death are recorded in the NDI by expert clinical coders on the basis of information contained in death certificates and, where available, coronial files, using the International Statistical Classification of Diseases and Related Health Problems 10th Edition (ICD-10). Dates of death were available for the entire period for which incarceration data were available (1 January 2000 to 31 March 2012); however, for administrative reasons, causes of death were available only for deaths occurring up to 31 December 2010.

Linkage between the OST and offending/incarceration data was performed by BOCSAR staff, and linkage between the OST and mortality data was undertaken by AIHW staff. Linkage was completed using probabilistic linkage software. Variables used for matching purposes included full name, date of birth, gender, and date and state of last known contact where available. These linked datasets were forwarded to the investigators with identifiers removed.

Cohort and data definitions
From the linked datasets, we identified a cohort of opioid-dependent people (people who were recorded in the PHDAS as having received OST at some point during 1986–2010) who had been incarcerated at least once (as recorded in the ROD). In defining receipt of OST, we excluded temporary treatment programmes (usually interstate visitors) and treatment episodes that were part of a buprenorphine clinical trial (during which the individual may have been allocated to placebo). Periods in the custody of Corrective Services NSW were determined using reception and discharge dates in the ROD. We used the first recorded OST episode as a proxy for the onset of opioid dependence and excluded prison episodes that occurred prior to the first OST episode; all subsequent prison episodes for an individual were included. If an individual’s first OST episode started while incarcerated, that prison episode was included. Given the chronic and relapsing nature of opioid dependence,15 16 we assumed that all cohort members remained opioid-dependent from their first episode of OST until death or the end of follow-up.

Deaths in prison were defined as deaths that occurred while the individual was recorded as being in the custody of Corrective Services NSW. This definition does not stipulate that a deceased individual was physically in prison at the time of death. Some decedents may have been receiving care in hospital at the time of death, while still in the custody of Corrective Services NSW. Deaths with known causes (ie, all deaths occurring in the calendar years 2000–2010) were categorised as natural (ie, disease-related) or unnatural deaths. We used the Corrective Services NSW definition of ‘unnatural deaths’, which includes suicides, drug-induced deaths, violent deaths and deaths due to accidental injuries or injuries in which no intent (ie, suicidal or violent) could be determined.17 ICD-10 codes for unnatural deaths are provided in the online supplementary materials.

Periods of OST while in prison, and treatment status (in OST/out of OST) at the time of death, were determined using treatment entry and exit dates from the PHDAS. New treatment episodes were defined when an individual started treatment 7 or more days after discharge from a previous treatment episode. A change in medicine (methadone to buprenorphine, or vice versa)
was considered a continuous episode if there were less than 7 days between ceasing one medicine and starting the other.

Data analysis
Statistical analyses were undertaken in SAS V.9.3. We used descriptive statistics to examine demographic characteristics and incarceration. Participants with any record of receiving OST in prison were compared with those who did not receive OST in prison using χ² tests for categorical variables and the Wilcoxon rank-sum test for continuous variables. We constructed a frequency histogram to illustrate time from entry to prison until death. For dates for which comparable data were available (financial years 2000–2001 to 2009–2010), we compared the total number of unnatural deaths among prisoners in NSW with the number of unnatural deaths in the opioid-dependent cohort.

We calculated all-cause and cause-specific crude mortality rates (CMR) with Poisson 95% CI for all time in prison and for the first 4 weeks in prison. Rates were also calculated by treatment status (in/out of OST). Person-years (PY) used in the calculation of CMRs accrued whenever participants were in the custody of Corrective Services NSW. Prison episodes where the participant was received and released on the same day were counted as 1 day. PY ceased to accrue at death on 31 March 2012 for all-cause mortality rates, and on 31 December 2010 for cause-specific mortality rates. CMR ratios with 95% CI were calculated to assess unadjusted risk of death in treatment compared with that out of treatment. The rate ratio (RR) was not calculated if there were zero deaths in either the in-treatment or out-of-treatment condition.

We used Cox regression to examine the adjusted association between OST exposure and two in-prison mortality outcomes: deaths from all causes and unnatural deaths. Two models were developed for each outcome: the first considering mortality risk across all time in prison, and the second examining mortality risk during the first 4 weeks of incarceration. All prison episodes were included, with days in prison counted as discontinuous intervals at risk. Although participants could have multiple observations within the dataset, the internal computations of a Cox model are such that when there is only one event of interest per person (in this case, death), there is no need for adjustments for multiple observations.17 OST exposure while in prison was considered a time-dependent variable, with periods out of treatment coded as 0 and periods in treatment coded as 1. Demographic variables included in the models were gender, indigenous (ie, Aboriginal or Torres Strait Islander) status and age at entry to prison; this latter variable increased as necessary with subsequent incarcerations of an individual. We also included variables that described prior experiences of incarceration during our observation periods, namely, the number and duration (days) of prior incarcerations. These variables increased with each incarceration per person. We also included dichotomous variables that recorded if a participant had been charged with a drug, property or violent offence prior to incarceration. These were included because of possible associations between death in prison (particularly suicide) and prior offending (particularly violent offending).18 19 The offending history variables could change from no to yes with subsequent incarcerations. Observations with missing covariate data were excluded from the analyses. We tested each variable for its bivariate association with mortality and entered all variables into a multivariate model. We tested the proportional hazards assumption for each static predictor variable in the multivariate model by including an interaction between it and log (time) in the model. All variables in both models were found to satisfy the proportional hazards assumption.

RESULTS
The cohort comprised 16 715 opioid-dependent people who had been incarcerated at least once between January 2000 and March 2012 (table 1). The majority (79%; n=13 199) of the cohort was male, and 30% (n=5011) were identified as indigenous. The median age at first entry to prison during our observation period was 30 years (ranging from 16 to 64 years).

Cohort members were received in prison a median of two times (ranging between 1 and 34). Duration of incarceration ranged from 1 day to the entire length of follow-up (12 years and 3 months), with a median duration of 71 days. Most participants (76.9%; n=12 852) received OST at some point while incarcerated. Compared to participants who received OST in prison, participants who did not receive OST in prison were significantly older, less likely to be indigenous and have fewer incarcerations of shorter duration (table 1).

Deaths in prison
Fifty-one cohort members died in prison (table 2). Most were men (90%; n=46) and non-indigenous (73%; n=37). The median age at death was 34 years (ranging between 20 and 54). Only 22% (n=11) of decedents were receiving OST at the time of death. Cause of death was available for 43 deaths in prison. Of these, nearly three-quarters (74%; n=32) were unnatural deaths: 19 suicides, 6 drug-related deaths, 2 violent deaths and 5 deaths due to other injuries. Between the financial years 2000–2001 and 2009–2010, unnatural deaths in this cohort accounted for 35% of all unnatural deaths in NSW prisons (see online supplementary materials for further details).

One-third of all deaths in prison (n=17) occurred during the first 4 weeks of incarceration (figure 1). Of these, three-quarters (13/17; 76%) were unnatural deaths: 10 suicides, 2 other injury-related deaths and 1 drug-related death.
Crude mortality rates

The all-cause CMR in prison was 1.6/1000 PY (95% CI 1.2 to 2.2/1000 PY) (table 3). The all-cause CMR was significantly lower while in OST (CMR: 0.7/1000 PY; 95% CI 0.3 to 1.2/1000 PY), compared to time out of OST (CMR: 2.7/1000 PY; 95% CI 2.0 to 3.7/1000 PY) (RR: 0.24; 95% CI 0.12 to 0.47).

Deaths from natural causes occurred at similar rates whether in (0.3/1000 PY; 95% CI 0.1 to 0.8/1000 PY) or out (0.5/1000 PY; 95% CI 0.2 to 1.0/1000 PY) of OST (RR: 0.68; 95% CI 0.21 to 2.21). There was, however, a marked effect of OST on rates of unnatural death. The in-treatment CMR for unnatural deaths was 0.3/1000 PY (95% CI 0.1 to 0.8/1000 PY), compared to 2.2/1000 PY (95% CI 1.4 to 3.2/1000 PY) while out of OST (RR: 0.15; 95% CI 0.06 to 0.39). Of the specific unnatural causes of death, suicides were significantly less frequent while in OST (CMR: 0.2/1000 PY; 95% CI 0.04 to 0.6/1000 PY) than while not in OST (CMR: 0.2/1000 PY; 95% CI 0.04 to 0.6/1000 PY) (RR: 0.15; 95% CI 0.04 to 0.52). Of the six observed drug-induced deaths and five other injury deaths, none occurred during OST, but both violent deaths occurred while in OST.

During the first 4 weeks of incarceration, the all-cause CMR was 6.6/1000 PY (95% CI 3.8 to 10.6/1000 PY). Of the 17 cohort members who died during the first 4 weeks of incarceration, only one was receiving OST at the time of death. The in-treatment all-cause CMR in the first 4 weeks of incarceration was 0.8 deaths/1000 PY (95% CI 0.02 to 4.3/1000 PY), significantly lower than the CMR of 12.7/1000 PY (95% CI 7.2 to 20.6/1000 PY) observed among out-of-treatment cohort members (RR: 0.06; 95% CI 0.01 to 0.45). Across all time in prison, the natural mortality rate did not differ by OST status, but the in-treatment CMR for unnatural deaths (11/1000 PY; 95% CI 5.7 to 19.2/1000 PY) (RR: 0.07; 95% CI 0.04 to 0.6/1000 PY) (RR: 0.15; 95% CI 0.04 to 0.52). Of the six observed drug-induced deaths and five other injury deaths, none occurred during OST, but both violent deaths occurred while in OST.

### Table 1
Demographic characteristics of opioid-dependent persons received to prison, New South Wales, Australia, 2000–2012 (n=16 715)

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>Total (n=16 715)</th>
<th>Ever received OST in prison (n=12 852)</th>
<th>No OST in prison (n=3863)</th>
<th>p Value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Male, n (%)</td>
<td>13 199 (79)</td>
<td>10 109 (78.7)</td>
<td>3090 (80)</td>
<td>0.07</td>
</tr>
<tr>
<td>Age (years)</td>
<td>30 (16–64)</td>
<td>29 (16–63)</td>
<td>32 (17–64)</td>
<td>&lt;0.0001</td>
</tr>
<tr>
<td>Indigenous,* n (%)</td>
<td>5011 (30)</td>
<td>4208 (32.8)</td>
<td>803 (20.8)</td>
<td>&lt;0.0001</td>
</tr>
<tr>
<td>Number of prison episodes, median (minimum–maximum)</td>
<td>2 (1–34)</td>
<td>3 (1–34)</td>
<td>1 (1–23)</td>
<td>&lt;.0001</td>
</tr>
<tr>
<td>Duration (days) of prison episodes, median (minimum–maximum)</td>
<td>71 (1–4473)</td>
<td>77 (1–4473)</td>
<td>46 (1–4473)</td>
<td>&lt;0.0001</td>
</tr>
<tr>
<td>Any receipt of OST in prison</td>
<td>12 852 (76.9)</td>
<td>–</td>
<td>–</td>
<td>–</td>
</tr>
</tbody>
</table>

*Indigenous status missing for 10 individuals. p Values refer to the comparison between participants who did or did not receive OST in prison (χ² test for categorical variables; Wilcoxon rank-sum test for continuous variables).

OST, opioid substitution therapy.

### Table 2
Demographic characteristics and causes of death of opioid-dependent persons who died in prison, New South Wales, Australia, 2000–2012 (n=51)

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>N (%)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Male</td>
<td>46 (90)</td>
</tr>
<tr>
<td>Age (years)</td>
<td>34 (20–54)</td>
</tr>
<tr>
<td>Indigenous</td>
<td>14 (27)</td>
</tr>
<tr>
<td>In OST at time of death</td>
<td>11 (22)</td>
</tr>
<tr>
<td>Deaths from natural causes*</td>
<td>11 (26)</td>
</tr>
<tr>
<td>Unnatural deaths*</td>
<td>32 (74)</td>
</tr>
<tr>
<td>Suicide</td>
<td>19 (44)</td>
</tr>
<tr>
<td>Drug-related</td>
<td>6 (14)</td>
</tr>
<tr>
<td>Violent</td>
<td>2 (5)</td>
</tr>
<tr>
<td>Other injury-related</td>
<td>5 (12)</td>
</tr>
</tbody>
</table>

*Cause of death unknown for eight decedents. Percentage is based on n=43.

OST, opioid substitution therapy.

### Figure 1
Deaths from natural, unnatural and unknown causes, by time in prison prior to death among opioid-dependent people who died in prison, New South Wales, Australia, 2000–2012 (n=51).
Adjusted effect of OST on deaths in custody

In the first 4 weeks of a prison episode, the all-cause mortality hazard was 94% lower while in OST, compared to being out of OST (adjusted HR (AHR): 0.06; 95% CI 0.01 to 0.48), regardless of gender, indigenous status, age, incarceration history or offending history (table 4). A similar finding was observed in relation to unnatural deaths; during the first 4 weeks of a prison episode, while in OST, the hazard of unnatural death was 93% lower than while not in OST (AHR 0.07; 95% CI 0.01 to 0.59).

Across all time in prison, the all-cause mortality hazard during periods of OST was reduced by three-quarters compared with periods not in OST (AHR: 0.26; 95% CI 0.13 to 0.50), regardless of gender, indigenous status, age, incarceration history or offending history. Across all time in prison, compared with periods not in OST, the hazard of unnatural death was 87% lower while in OST (AHR: 0.13; 95% CI 0.05 to 0.35).

DISCUSSION

We have demonstrated a very strong association between receipt of OST and lowered mortality among opioid-dependent prisoners. After adjusting for demographic and criminal history factors, compared to time not in OST, being in OST was associated with a 74% lower hazard of dying in prison. The association between mortality risk and OST was driven by the reduction in unnatural deaths, with periods in OST associated with an 87% decrease in mortality hazard compared to time not in OST. One-third of deaths in prison in this cohort occurred during the first 4 weeks of incarceration; in these initial weeks in prison, compared to time not in OST, being in OST was associated with a 94% lower all-cause mortality hazard and a 93% lower hazard of unnatural death. Although the CIs around the HRs were wide, even the uppermost limits of the intervals suggested substantial reductions in mortality while in OST. For example, across all time in prison, there was at least a 65% reduction in unnatural mortality hazard while in OST compared to time not in OST.

The cohort included in this study was drawn from a large, statewide administrative dataset of people receiving treatment for opioid dependence. Although precise data are not available on the representativeness of this cohort, in sentinel surveillance studies of people who inject drugs in NSW (98% of whom have a history of illicit opioid use), almost 60% of participants are currently in OST, and more than 80% have a history of OST, suggesting that the cohort is highly representative of the NSW opioid-dependent population. The subset of this population that had been incarcerated was included...
Table 4 Cox proportional hazards models of predictors of death in prison among opioid-dependent prisoners in New South Wales, Australia, 2000–2012 (n=16,715)

<table>
<thead>
<tr>
<th>Predictor</th>
<th>All-cause deaths</th>
<th>Unnatural deaths</th>
<th>All-cause deaths</th>
<th>Unnatural deaths</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Total time in prison</td>
<td>First 4 weeks in prison</td>
<td>Total time in prison</td>
<td>First 4 weeks in prison</td>
</tr>
<tr>
<td></td>
<td>Adjusted HR (95% CI)</td>
<td>HR (95% CI)</td>
<td>Adjusted HR (95% CI)</td>
<td>HR (95% CI)</td>
</tr>
<tr>
<td>Receipt of OST*</td>
<td>0.25 (0.13 to 0.48)</td>
<td>0.26 (0.13 to 0.5)</td>
<td>0.15 (0.06 to 0.39)</td>
<td>0.26 (0.13 to 0.5)</td>
</tr>
<tr>
<td>Male</td>
<td>1.21 (0.48 to 3.04)</td>
<td>0.94 (0.36 to 2.41)</td>
<td>0.75 (0.29 to 1.96)</td>
<td>0.57 (0.22 to 1.54)</td>
</tr>
<tr>
<td>Indigenous</td>
<td>0.64 (0.34 to 1.18)</td>
<td>0.84 (0.44 to 1.59)</td>
<td>0.92 (0.44 to 1.9)</td>
<td>0.78 (0.37 to 1.68)</td>
</tr>
<tr>
<td>Age at prison reception</td>
<td>1.06 (1.03 to 1.1)</td>
<td>1.05 (1.01 to 1.09)</td>
<td>1.03 (0.98 to 1.07)</td>
<td>1.02 (0.97 to 1.07)</td>
</tr>
<tr>
<td>No of prior incarcerations</td>
<td>1.01 (0.91 to 1.12)</td>
<td>1.09 (0.99 to 1.21)</td>
<td>1.10 (0.97 to 1.24)</td>
<td>1.13 (0.99 to 1.28)</td>
</tr>
<tr>
<td>Duration of prior incarcerations‡</td>
<td>0.98 (0.96 to 1)</td>
<td>0.98 (0.96 to 1)</td>
<td>1.01 (0.98 to 1.04)</td>
<td>1.01 (0.98 to 1.04)</td>
</tr>
<tr>
<td>Any prior drug offences</td>
<td>0.95 (0.53 to 1.7)</td>
<td>0.97 (0.53 to 1.76)</td>
<td>0.96 (0.46 to 2.01)</td>
<td>0.92 (0.43 to 1.95)</td>
</tr>
<tr>
<td>Any prior property offences</td>
<td>0.47 (0.25 to 0.92)</td>
<td>0.76 (0.37 to 1.57)</td>
<td>0.71 (0.29 to 1.75)</td>
<td>0.81 (0.31 to 2.14)</td>
</tr>
<tr>
<td>Any prior violent offences</td>
<td>0.53 (0.3 to 0.94)</td>
<td>0.69 (0.37 to 1.26)</td>
<td>0.65 (0.31 to 1.34)</td>
<td>0.70 (0.32 to 1.5)</td>
</tr>
</tbody>
</table>

*Time-dependent variable.
‡HR refers to change in hazard associated with 28-day increase in duration of prior incarcerations.
OST, opioid substitution treatment.


Ethics approval Research ethics committees of the University of New South Wales (NSW) (HC12019), the NSW Ministry of Health (2011/11/360), the Australian Institute of Health and Welfare (EC2011/2/13), the Alfred Hospital (165/11), Corrective Services NSW (11/62607), Justice Health (NSW) (G219/11), the NSW Aboriginal Health and Medical Research Council (AHMRC) (793/11) and Justice Health (Victoria) (CF/13/3440C).

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