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**Variation in severe maternal morbidity according to socioeconomic position: a UK national cohort study.**

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## ARTICLE SUMMARY

### Article Focus

- To explore the relationship between socioeconomic position and severe maternal morbidity using data obtained from a series of UKOSS studies
- To estimate the risk of severe morbidity by socioeconomic group, independent of ethnicity, maternal age, smoking, pre-existing medical condition, BMI, multiple pregnancy and past pregnancy complications

### Key Messages

- The results suggested that women from the lowest socioeconomic group are 1.22 times (95%CI: 0.92 – 1.61) more likely than women from the highest group to experience severe maternal morbidity
- The results also showed that other risk factors for severe maternal morbidity include non-white ethnicity, older maternal age ( $\geq 35$  years), BMI  $\geq 25$  kg/m<sup>2</sup>, pre-existing medical condition/s, multiple pregnancy or past pregnancy complications

### Strengths and Limitations

- Study strengths include the robust case ascertainment, minimisation of sampling bias through collection of data from all consultant-led obstetric units in the UK and the clear definition of severe maternal morbidity
- Limitations relate to the data available to define socioeconomic position and the application of the findings to other non-UK setting

**ABSTRACT**

**Objectives:** This study aimed to explore the independent association between socioeconomic position, defined by occupation, and severe maternal morbidity amongst women in the UK.

**Design:** Cohort study. **Setting:** The analysis was conducted as a case-control analysis, using data from a series of studies of direct causes of severe maternal morbidity undertaken through the UK Obstetric Surveillance System (UKOSS), with data collected throughout all consultant-led obstetric units in the UK. **Participants:** The analysis included 1,144 cases and 2,256 comparison women (controls). UKOSS studies from which data on case women were obtained included amniotic fluid embolism, acute fatty liver of pregnancy, eclampsia, peripartum hysterectomy, therapies for peripartum haemorrhage and uterine rupture. **Primary outcome measure:** Odds of severe maternal morbidity by socioeconomic group, independent of ethnicity, maternal age, smoking, pre-existing medical condition, BMI, multiple pregnancy and past pregnancy complications. Occupation was used to classify different socioeconomic groups.

**Secondary outcome measure:** Odds of morbidity related to ethnic group, maternal age, smoking, pre-existing medical condition, BMI, multiple pregnancy and past pregnancy complications. **Results:** Across the socioeconomic groups, adjusted odds ratios were 1.17 (95%CI: 0.94 – 1.45) for the 'intermediate group', 1.16 (95%CI: 0.93 – 1.45) for 'routine/manual', 1.22 (95%CI: 0.92 – 1.61) for 'unemployed' women, and 1.51 (95%CI: 1.18 – 1.94) for women with missing socioeconomic information. Women of non-white ethnicity, older maternal age ( $\geq 35$  years), BMI  $\geq 25$  kg/m<sup>2</sup> and those with pre-existing medical condition/s, multiple pregnancy or past pregnancy complications were shown to have a significantly increased odds of severe maternal morbidity. **Conclusion:** This study suggests that socioeconomic position may be an independently associated with an increased risk of severe maternal morbidity. Further research is warranted to investigate why this association exists in a country where healthcare is universal and free at the point of access.

## INTRODUCTION

Maternal morbidity is of major public health concern with some estimates suggesting that as many as ten million women worldwide suffer from pregnancy-related complications each year<sup>1</sup>. In countries where maternal deaths are infrequent, the information that mortality audits provide cannot be generalised readily across the general population of pregnant women, nor can it necessarily accurately predict which women are at increased risk of adverse outcomes<sup>2</sup>. In these circumstances, additional focus on maternal morbidity is now widely recognised as a robust approach to improving maternal health and the quality of healthcare provision<sup>3-6</sup>.

Identifying high-risk women is imperative to prevent adverse outcomes, since it allows more intensive clinical management to be directed towards these women and is the fundamental purpose of antenatal care. Extensive research has demonstrated convincing links between severe maternal morbidity and advanced maternal age, pre-existing medical conditions and obesity<sup>1 7 8</sup>. However, these factors are unable to explain entirely the differences in maternal morbidity found between different populations of women both within and between countries. Other evidence suggests that ethnicity and social disadvantage may have a significant role to play, with UK data suggesting that women from ethnic minority groups are more likely to experience severe maternal morbidity than white women, and also to be at higher risk of dying from specific maternal morbidities<sup>9-11</sup>. However, as highlighted by several studies exploring ethnic differences in maternal health, it is unclear whether ethnicity itself is directly related to poor outcomes or, more likely, whether it is a marker for factors such as low socioeconomic position and lower levels of education<sup>12</sup>. Because minority ethnic groups are often disproportionately represented in lower socioeconomic groups, ethnicity and socioeconomic position are often confounded and results attributed to ethnic difference, may be due, at least in part, to socioeconomic differences<sup>13</sup>. The aim of the analysis reported here was to explore the independent risk of severe maternal morbidity associated with socioeconomic position in the UK.

## METHODS

### Data Collection and Definitions

A case-control analysis was performed using data derived from a series of UK-wide studies of direct causes of maternal morbidity undertaken through the UK Obstetric Surveillance System (UKOSS)<sup>14</sup>. These studies were of amniotic fluid embolism (AFE), acute fatty liver of pregnancy (AFLP), eclampsia, peripartum hysterectomy, therapies for peripartum haemorrhage and uterine rupture<sup>15-21</sup>. For the purpose of this analysis, these cases are collectively referred to as 'severe maternal morbidity' (Box 1).

Box 1. Criteria used to define cases of severe maternal morbidity used in the UKOSS studies included in the analysis

Condition	Definition
<b>Amniotic fluid embolism</b>	EITHER a clinical diagnosis of AFE (acute hypotension or cardiac arrest, acute hypoxia or coagulopathy in the absence of any other potential explanation for the symptoms and signs observed) OR a pathological diagnosis (presence of fetal squames or hair in the lungs)
<b>Acute fatty liver of pregnancy</b>	EITHER AFLP confirmed by biopsy or post-mortem examination OR a clinician has made a diagnosis of AFLP with signs and symptoms consistent with AFLP
<b>Eclampsia</b>	The occurrence of convulsions during pregnancy or in the first ten days postpartum, together with at least two of the following features within 24 hours after the convulsions: 1) Hypertension (a booking diastolic pressure of <90mmHg, a maximum diastolic of ≥90mmHg, and a diastolic increment of ≥25mmHg) 2) Proteinuria (at least + protein in a random urine sample or ≥0.3g in a 24hr collection) 3) Thrombocytopenia (platelet count of <100X10 <sup>9</sup> /l) 4) An increased plasma alanine aminotransferase (ALT) concentration (≥42 iu/l) or an increased plasma aspartate transaminase aminotransferase (AST) concentration (≥42iu/l)

<b>Peripartum hysterectomy</b>	Any woman giving birth to an infant and having a hysterectomy during the same clinical episode
<b>Therapies for major peripartum haemorrhage</b>	All women in the UK treated therapeutically or prophylactically for major peripartum haemorrhage with: EITHER Activated factor VIIa OR B-lynch suture or other brace suture OR Arterial ligation or embolisation
<b>Uterine rupture</b>	Any woman in the UK identified as having a uterine rupture using the following definition: A complete separation of the wall of the pregnant uterus, with or without expulsion of the fetus, involving rupture of the membranes at the site of the uterine rupture or extension into the uterine muscle separate from any previous scar, and endangering the life of the mother or fetus Excluded: any asymptomatic palpable or visualised defect (for example dehiscence noted incidentally at caesarean delivery)

The UKOSS methodology used to carry out a following programme of studies has been described in detail elsewhere<sup>14</sup>. In brief, monthly mailing of notification cards to all obstetric-led maternity units in the UK enables identification of all cases under study. Data collection forms despatched in response to a case notification are used by the clinician responsible for care to report de-identified information from the woman's medical notes. Follow-up of non-responders maximises the completeness of data collection and confirms denominator numbers for incidence estimates.

For many of the UKOSS studies, data about a representative sample of women without the specific condition under study were obtained by collecting information on the two women delivering immediately prior to each case in the same obstetric unit; these were defined as controls. The same information was collected about the controls as for case women. Data were thus available for analysis on 1,144 cases of severe maternal morbidity and 2,256 controls.

Maternal occupation was used to classify cases and controls by socioeconomic group, where this was not available her husband's/partner's occupation was used. Using the National

Statistics Socio-Economic Classification (NS-SEC)<sup>22</sup>, occupation was coded into three categories plus an additional group for unemployed women.

### Statistical Analysis

All analyses were carried out using STATA version 11.0. Following univariable analysis multivariable unconditional logistic regression was used to examine the independent relationship between socioeconomic position and severe maternal morbidity and to generate adjusted odds ratios (aORs) with their 95% confidence intervals (95% CIs). Due to the high proportion and distribution of missing data for some variables, particularly for socio-economic position (Table 1), 'missing' was included as a separate group for each categorical variable.

Table 1. Distribution of missing data for cases of severe maternal morbidity and controls

Variable	Number of cases with missing information (% of total cases)	Number of controls with missing information (% of total controls)
Socioeconomic Group	193 (16.9)	303 (13.4)
Ethnicity	18 (1.6)	78 (3.5)
Age	3 (0.3)	18 (0.8)
Smoking	28 (2.5)	61 (2.7)
Pre-existing medical condition(s)	6 (0.5)	27 (1.2)
Body mass index	128 (11.2)	219 (9.7)
Parity	3 (0.3)	22 (1.0)
Multiple pregnancy	2 (0.2)	18 (0.8)
Past pregnancy complication(s)	9 (0.8)	31 (1.4)

Variables likely to confound the relationship between socioeconomic position and severe maternal morbidity were identified *a priori* based on evidence from the published literature and the results of univariable analysis. These included ethnicity, maternal age, smoking, pre-existing medical conditions and body mass index (BMI)<sup>17 10 23-35</sup>. Other variables included in the analysis as possible confounders were parity, multiple pregnancy and past pregnancy complications.



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3 Model building using unconditional logistic regression proceeded by the inclusion of severe  
4 maternal morbidity (outcome) and socio-economic position, with each of the potential  
5 confounders added in turn one-by-one. The individual effect of each variable on the fit of the  
6 data was assessed using the likelihood ratio test. The variables with *a priori* evidence of effect  
7 were included in the final model. Of the additional three variables examined (parity, multiple  
8 pregnancy and past pregnancy conditions) only those demonstrating a statistically significant  
9 effect ( $p < 0.05$ ) on the fit of the data were retained; parity was thus excluded.  
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19 Age, included as a continuous variable, was tested for departure from linearity; the results are  
20 for ease of presentation, shown as categorical. BMI was included as a categorical variable due  
21 to the high proportion of missing data and the concern that by including it as continuous, women  
22 with missing data would be excluded and the effect size underestimated. There was no  
23 evidence of a significant interaction between socioeconomic position and any of the other  
24 variables included in the final model ( $p < 0.01$ ).  
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33 Sensitivity analyses were conducted to explore the effect of missing data for socioeconomic  
34 position and BMI. The final multivariable regression model was re-run after redistributing  
35 missing values into the lowest and highest categories for each of these variables.  
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### 42 **Study Power**

43 The analysis included a fixed number of cases and controls derived from specified UKOSS  
44 studies. Based on 1,144 cases and 2,256 controls, and with a prevalence of 13% for the lowest  
45 (unemployed) socioeconomic group, the analysis had 80% power to detect as statistically  
46 significant ( $p < 0.05$ ) an odds of 1.32 or greater and 90% power to detect as statistically  
47 significant ( $p < 0.05$ ) an odds of 1.37 or greater, for comparisons with the highest  
48 (managerial/professional) socioeconomic group.  
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## RESULTS

The analysis included a total of 1,144 cases of severe maternal morbidity and 2,256 controls. Compared with the controls cases were more likely to be older, multiparous, of non-white ethnicity, to have missing socio-economic position and BMI information, to be non-smokers, to have a pre-existing medical condition, a history of past pregnancy complications and a multiple pregnancy in the index pregnancy (Table 2).

Table 2. Unadjusted and adjusted odds of severe maternal morbidity by maternal characteristics

Characteristic	Cases n= (%) N= 1,144	Controls n= (%) N=2,256	Unadjusted OR (95% CI)	Adjusted* OR (95%CI)
<b>Socioeconomic group</b>				
Managerial/ Professional	292 (25.5)	567 (25.1)	1.0	1.0
Intermediate	244 (21.3)	482 (21.4)	0.98 (0.80 – 1.21)	1.17 (0.94 – 1.45)
Routine/Manual	273 (28.7)	595 (26.4)	0.89 (0.73 – 1.09)	1.16 (0.93 – 1.45)
Unemployed	142 (14.9)	309 (13.7)	0.89 (0.70 – 1.14)	1.22 (0.92 – 1.61)
Missing	193 (16.9)	303 (13.4)	1.24 (0.98 – 1.56)	1.51 (1.18 – 1.94)
<b>Ethnic group</b>				
White	827 (72.3)	1796 (79.6)	1.0	1.0
Asian	139 (12.2)	197 (8.7)	1.53 (1.22 – 1.93)	1.57 (1.23 – 2.00)
Black	108 (9.4)	116 (5.1)	2.02 (1.54 – 2.66)	1.77 (1.32 – 2.36)
Other	52 (4.6)	69 (3.1)	1.64 (1.13 – 2.37)	1.50 (1.02 – 2.19)
Missing	18 (1.6)	78 (3.5)	0.50 (0.30 – 0.84)	0.51 (0.28 – 0.91)
<b>Age</b>				
Age <20	56 (4.9)	139 (6.2)	0.98 (0.70 – 1.39)	1.08 (0.75 – 1.56)
Age 20 – 24	127 (11.1)	449 (19.9)	0.69 (0.54 – 0.89)	0.71 (0.55 – 0.92)
Age 25 – 29	225 (19.7)	549 (24.3)	1.0	1.0
Age 30 – 34	341 (29.8)	622 (27.6)	1.34 (1.09 – 1.64)	1.33 (1.08 – 1.64)
Age ≥35	392 (34.3)	479 (21.2)	2.00 (1.63 – 2.45)	1.98 (1.60 – 2.45)
Missing	3 (0.3)	18 (0.8)	0.41 (0.12 – 1.39)	2.66 (0.22 – 26.00)
<b>Smoking status</b>				
Non-Smokers	899 (78.6)	1,642 (72.8)	1.0	1.0
Smokers	217 (19.0)	553 (24.5)	0.72 (0.60 – 0.86)	0.89 (0.73 – 1.08)
Missing	28 (2.5)	61 (2.7)	0.84 (0.53 – 1.32)	0.95 (0.54 – 1.67)
<b>BMI</b>				
<b>Continuous (per kg/m<sup>2</sup> increase in BMI)</b>			1.01 (1.00 – 1.03)	
< 25 kg/m <sup>2</sup>	513 (44.8)	1,121 (49.7)	1.0	1.0
25-29.9 kg/m <sup>2</sup>	297 (26.0)	545 (24.2)	1.19 (1.00 – 1.42)	1.10 (0.92 – 1.32)
≥ 30 kg/m <sup>2</sup>	206 (18.0)	371 (16.5)	1.21 (0.99 – 1.48)	1.09 (0.88 – 1.34)
Missing	128 (11.2)	219 (9.7)	1.28 (1.00 – 1.63)	1.24 (0.95 – 1.61)
<b>Co-existing medical conditions**</b>				
No	1,045 (91.4)	2,126 (94.2)	1.0	1.0
Yes	93 (8.1)	103 (4.6)	1.84 (1.38 – 2.46)	1.60 (1.18 – 2.17)
Missing	6 (0.5)	27 (1.2)	0.45 (1.18 – 1.10)	0.76 (0.17 – 3.41)
<b>Parity</b>				
Nulliparous	403 (35.2)	984 (43.6)	1.0	No significant effect on fit of model (excluded)
Multiparous	738 (64.5)	1,250 (55.4)	1.44 (1.24 – 1.67)	
Missing	3 (0.3)	22 (1.0)	0.33 (0.10 – 1.12)	

<b>Multiple pregnancy</b>				
No	1,083 (94.7)	2,211 (98.0)	1.0	1.0
Yes	59 (5.2)	27 (1.2)	4.46 (2.81 – 7.08)	3.88 (2.42 – 6.22)
Missing	2 (0.2)	18 (0.8)	0.23 (0.53 – 0.98)	0.20 (0.02 – 2.54)
<b>Past pregnancy conditions***</b>				
No	1,041 (91.0)	2,121 (94.0)	1.0	1.0
Yes	94 (8.2)	104 (4.6)	1.81 (1.38 – 2.46)	1.58 (1.16 – 2.14)
Missing	9 (0.8)	31 (1.4)	0.59 (0.28 – 1.25)	1.12 (0.36 – 3.54)

\*All variables adjusted for all other variables in the table

\*\*Cardiac disease, diabetes mellitus, epilepsy, essential hypertension, haematological disorders, past thrombotic event, recent/current anticoagulation, IV drug use and/or alcohol abuse

\*\*\*Amniotic fluid embolism, gestational diabetes, pregnancy-induced hypertension, pre-eclampsia, eclampsia, thrombotic event, acute fatty liver of pregnancy, manual removal of placenta, placenta accreta, past classical caesarean section

Following adjustment, compared with the controls, cases were 1.17 (95%CI 0.94-1.45) times more likely to be in the 'intermediate' socioeconomic group than the 'managerial/professional' group, 1.16 (95%CI: 0.93 – 1.45) times more likely to be in the 'routine/manual' group, 1.22 (95%CI: 0.92 – 1.61) times more likely to be unemployed and 1.51 (95%CI: 1.18 – 1.94) times more likely to be have missing socioeconomic information. Compared with controls cases were 1.77 (95%CI 1.32-2.36) times more likely to be black than white, 1.57 (1.23-2.00) times more likely to be Asian and 1.50 (95%CI 1.02-2.19) times more likely to be from another minority ethnic group. The relationship between the odds severe maternal morbidity and maternal age was J-shaped with cases having significantly raised adjusted odds of being 30-34yrs (aOR 1.33; 95%CI 1.08-1.64) or 35+yrs (aOR 1.98; 95%CI 1.60-2.45). After adjustment, cases had a non-significant 10% increase in the odds of having a BMI of 25 or over and a 24% increase in the odds of having missing BMI data which was similarly non-significant. Cases were significantly more likely than controls to have a history of pre-existing medical conditions (aOR 1.60; 95%CI 1.18-2.17), past pregnancy complications (aOR 1.58; 95%CI 1.16-2.14) and to have a multiple pregnancy (aOR 3.54; 95%CI 2.18-5.76).

## Sensitivity analysis

Redistributing all the missing socioeconomic observations into the managerial/professional group reduced the aOR associated with unemployment to 1.02 (95%CI 0.79-1.32) whereas, redistribution of all the missing socioeconomic observations into the unemployed group produced an aOR of 1.38 (95%CI 1.11–1.72) associated with unemployment (Table 3).

Table 3. Sensitivity analysis for missing observations

Missing values	Sub-category	Cases (%)	Controls (%)	Adjusted OR (95% CI)
<b>Socioeconomic group</b>				
Missing SEP values recoded to 'Managerial/Professional' (best case)	Managerial	485 (42.4)	870 (38.6)	1.0
	Intermediate	244 (21.3)	482 (21.4)	1.01 (0.83 – 1.23)
	Routine/Manual	273 (23.9)	595 (26.4)	0.99 (0.82 – 1.21)
	Unemployed	142 (12.4)	309 (13.7)	1.02 (0.79 – 1.32)
Missing SEP values recoded to 'unemployed' (worst case)	Managerial	292 (25.5)	567 (25.1)	1.0
	Intermediate	244 (21.3)	482 (21.37)	1.17 (0.94 – 1.46)
	Routine/Manual	273 (23.9)	595 (26.4)	1.17 (0.94 – 1.46)
	Unemployed	335 (29.3)	612 (27.1)	1.38 (1.11 – 1.72)

Re-running the multivariable model including only those cases and controls with complete information for every variable generated an aOR of 1.31 (95%CI 0.97–1.77) associated with unemployment, an aOR of 1.13 (95%CI 0.89-1.43) for the routine/manual group and 1.19 (95%CI 0.94-1.49) for the intermediate group.

## DISCUSSION

### Findings

The results of this study suggest that there is an independent association between socioeconomic position and severe maternal morbidity. Although not statistically significant, this analysis showed that compared to women from the highest managerial/professional socioeconomic group, unemployed status was associated with a 22% increase in the odds of severe morbidity. Women with missing socioeconomic information had a more than 50% increase in the odds of severe morbidity and this finding was unlikely to be due to chance. Sensitivity analysis suggests that if all the women with missing socioeconomic data were

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3 unemployed, the estimate of the odds of severe morbidity increased to 38%; with a result  
4 unlikely to be due to chance.  
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9 Despite universal health care which is free at the point of delivery, and significant investment in  
10 researching social inequalities in health in the UK, socially-marginalised groups continue to fare  
11 poorly as compared with more affluent sections of society, across a range of health outcomes  
12 including, as suggested by our analysis, maternal health. The reasons for the social gradient in  
13 maternal health in the UK have not yet been thoroughly explored. It is therefore unclear whether  
14 the observed association is as a result of differences in the underlying risk of disease,  
15 differences in health-seeking behaviour, or differences in access to, or provision and quality of  
16 services during pregnancy, delivery and the post-partum period.  
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27 In addition to finding an association between severe morbidity and factors such as pre-existing  
28 medical conditions, multiple pregnancy and past pregnancy complications, this study also  
29 reflected the findings of past studies in France and the Netherlands<sup>6 26 32</sup> which have  
30 demonstrated a significantly increased risk of severe morbidity amongst women from minority  
31 ethnic groups. The association with socioeconomic position was, however, independent of  
32 these associations.  
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#### 42 **Implications of the Study Findings**

43 Health research in various settings worldwide has illuminated the ongoing existence of the  
44 'inverse care law', first described in 1971 by Tudor Hart as the tendency for the "availability of  
45 good medical care to vary inversely with the need for it in the population served"<sup>36</sup>.  
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51 Since the global financial crisis in 2008, the World Health Organisation has stressed that "as  
52 standards of living decrease in many countries, and government revenues are tightened, [we  
53 would argue that] it is even more urgent that the distributional effects on all policies are taken  
54 into account in policy decision making"<sup>37</sup>. In a paper that argues that social welfare spending is  
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3 as important as protecting health through safeguarding the healthcare budget, Stuckler et al  
4 have proposed that the economic crisis presents the opportunity to reorganise the provision of  
5 services to those most in need<sup>38</sup>. Our findings clearly identify one such group in need.  
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11 'Maternity Matters' guidance for service commissioning in England released in 2007  
12 emphasised the Government's commitment to providing choice guarantees for women  
13 regarding type of antenatal care and place of birth and also highlighted the need for future  
14 maternity services to address disproportionately higher rates of maternal morbidity amongst  
15 disadvantaged women<sup>39</sup>. Although it described a comprehensive approach to delivering  
16 improved maternity care the guidance did not provide any specific recommendations as to how  
17 to improve services for women from socially disadvantaged backgrounds. Our study  
18 emphasises the need for developing such recommendations within the currently changing  
19 structure of health service commissioning in England.  
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### 29 30 31 **Strengths and Limitations of the Study** 32

33 UKOSS, the first system of its kind in obstetrics worldwide, has provided an invaluable wealth of  
34 information on uncommon but serious disorders of pregnancy. The active, negative surveillance  
35 approach has ensured robust case ascertainment, and many of the UKOSS studies are the  
36 largest-scale studies of rare disorders of pregnancy worldwide.  
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43 By studying disorders of pregnancy and childbirth in separate studies, UKOSS provided the  
44 means by which to explore 'severe maternal morbidity' due to direct pregnancy causes as an  
45 aggregate of multiple conditions, thereby avoiding the challenges encountered by previous  
46 studies which have used combined clinical and intervention criteria. It also avoids the perennial  
47 difficulty that there is no universally agreed definition of what constitutes 'severe maternal  
48 morbidity'. One disadvantage of this approach is that 'indirect causes' of maternal morbidity  
49 were excluded from our analysis, including several major contributors to maternal mortality in  
50 the UK such as cardiac disease, thromboembolic disease and puerperal psychosis/suicide<sup>40</sup>.  
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3 The design of UKOSS studies ensures that sampling bias is minimised by recruiting cases from  
4 all obstetric units across the UK, thus including women from all different ethnic, age and  
5 socioeconomic groups. Selection of controls is also aimed at attaining a representative sample  
6 of the general population of women giving birth by recruiting women from the same obstetric  
7 units as cases throughout the UK during the same time period. Comparison between UKOSS  
8 controls and national maternity data suggests that the controls recruited for UKOSS studies are  
9 closely representative of the general, child-bearing population<sup>19</sup>.  
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19 A universal challenge for studies that seek to explore the impact of social disadvantage in  
20 society is how to best capture the concept of 'social disadvantage' and define socioeconomic  
21 position. There are multiple dimensions to social disadvantage and there are many different  
22 approaches to defining a person's position in the social structure of society. The retrospective  
23 collection of data for UKOSS from patient records limited the definition of socioeconomic  
24 position to maternal occupation (or paternal where maternal was not available). Using the  
25 national classification of socioeconomic position (NS-SEC) ensured the study population was  
26 categorised using a system that is used nationally. However, the collection of other  
27 socioeconomic indicators such as level of education, years of schooling, residential location or  
28 income was not possible although this would have allowed further analysis of social  
29 disadvantage by using composite measures of disadvantage such as the Index of Multiple  
30 Deprivation. There is thus a place for further research using other indicators of socioeconomic  
31 position to see whether the results of this study are replicated.  
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48 The results of this study clearly apply only to high resource countries with well-developed  
49 healthcare systems. However, other high resource countries have different models of health  
50 care, such that pregnancy and delivery care may not be free at the point of access, as is the  
51 case in the UK. This study used only UK data and the results may therefore not be  
52 generalisable to women giving birth in other countries. Further research using information from  
53 other settings is therefore important to investigate whether socioeconomic position is associated  
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3 with severe maternal morbidity in countries with differing health systems as this may give  
4 insights into the causes and where on the causal pathway modifiable risks might lie.  
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## 8 9 **Conclusions**

10 This study suggests that socioeconomic position is associated with an increased risk of severe  
11 maternal morbidity, independent of risks associated with maternal age, ethnic group or known  
12 medical and pregnancy complications. However, the reasons for this association remain  
13 unclear, suggesting a need for further research to investigate whether there are differences in  
14 disease risks, health-seeking behaviour, access to care or the quality of care received by  
15 women from different socioeconomic groups. Further information about these differences would  
16 allow development of specific recommendations for the care of women from socially  
17 disadvantaged backgrounds in order to minimise future risks.  
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**STROBE 2007 (v4) checklist of items to be included in reports of observational studies in epidemiology\***  
**Checklist for cohort, case-control, and cross-sectional studies (combined)**

Section/Topic	Item #	Recommendation	Reported on page #
Title and abstract	1	(a) Indicate the study’s design with a commonly used term in the title or the abstract	1
		(b) Provide in the abstract an informative and balanced summary of what was done and what was found	3
<b>Introduction</b>			<b>4</b>
Background/rationale	2	Explain the scientific background and rationale for the investigation being reported	4
Objectives	3	State specific objectives, including any pre-specified hypotheses	4
<b>Methods</b>			
Study design	4	Present key elements of study design early in the paper	3
Setting	5	Describe the setting, locations, and relevant dates, including periods of recruitment, exposure, follow-up, and data collection	6
Participants	6	(a) <i>Cohort study</i> —Give the eligibility criteria, and the sources and methods of selection of participants. Describe methods of follow-up <i>Case-control study</i> —Give the eligibility criteria, and the sources and methods of case ascertainment and control selection. Give the rationale for the choice of cases and controls <i>Cross-sectional study</i> —Give the eligibility criteria, and the sources and methods of selection of participants	5-6
		(b) <i>Cohort study</i> —For matched studies, give matching criteria and number of exposed and unexposed <i>Case-control study</i> —For matched studies, give matching criteria and the number of controls per case	
Variables	7	Clearly define all outcomes, exposures, predictors, potential confounders, and effect modifiers. Give diagnostic criteria, if applicable	6-7
Data sources/ measurement	8*	For each variable of interest, give sources of data and details of methods of assessment (measurement). Describe comparability of assessment methods if there is more than one group	6
Bias	9	Describe any efforts to address potential sources of bias	7
Study size	10	Explain how the study size was arrived at	6 & 8
Quantitative variables	11	Explain how quantitative variables were handled in the analyses. If applicable, describe which groupings were chosen and why	7-8
Statistical methods	12	(a) Describe all statistical methods, including those used to control for confounding	7-8
		(b) Describe any methods used to examine subgroups and interactions	8
		(c) Explain how missing data were addressed	7
		(d) <i>Cohort study</i> —If applicable, explain how loss to follow-up was addressed <i>Case-control study</i> —If applicable, explain how matching of cases and controls was addressed	

		<i>Cross-sectional study</i> —If applicable, describe analytical methods taking account of sampling strategy	
		(e) Describe any sensitivity analyses	11
<b>Results</b>			
Participants	13*	(a) Report numbers of individuals at each stage of study—eg numbers potentially eligible, examined for eligibility, confirmed eligible, included in the study, completing follow-up, and analysed	9
		(b) Give reasons for non-participation at each stage	5
		(c) Consider use of a flow diagram	
Descriptive data	14*	(a) Give characteristics of study participants (eg demographic, clinical, social) and information on exposures and potential confounders	9-10
		(b) Indicate number of participants with missing data for each variable of interest	9-10
		(c) <i>Cohort study</i> —Summarise follow-up time (eg, average and total amount)	
Outcome data	15*	<i>Cohort study</i> —Report numbers of outcome events or summary measures over time	
		<i>Case-control study</i> —Report numbers in each exposure category, or summary measures of exposure	9-10
		<i>Cross-sectional study</i> —Report numbers of outcome events or summary measures	
Main results	16	(a) Give unadjusted estimates and, if applicable, confounder-adjusted estimates and their precision (eg, 95% confidence interval). Make clear which confounders were adjusted for and why they were included	9-10
		(b) Report category boundaries when continuous variables were categorized	9-10
		(c) If relevant, consider translating estimates of relative risk into absolute risk for a meaningful time period	
Other analyses	17	Report other analyses done—eg analyses of subgroups and interactions, and sensitivity analyses	11
<b>Discussion</b>			
Key results	18	Summarise key results with reference to study objectives	11-12
Limitations	19	Discuss limitations of the study, taking into account sources of potential bias or imprecision. Discuss both direction and magnitude of any potential bias	13-14
Interpretation	20	Give a cautious overall interpretation of results considering objectives, limitations, multiplicity of analyses, results from similar studies, and other relevant evidence	12-13
Generalisability	21	Discuss the generalisability (external validity) of the study results	14-15
<b>Other information</b>			
Funding	22	Give the source of funding and the role of the funders for the present study and, if applicable, for the original study on which the present article is based	15

\*Give information separately for cases and controls in case-control studies and, if applicable, for exposed and unexposed groups in cohort and cross-sectional studies.

**Note:** An Explanation and Elaboration article discusses each checklist item and gives methodological background and published examples of transparent reporting. The STROBE checklist is best used in conjunction with this article (freely available on the Web sites of PLoS Medicine at <http://www.plosmedicine.org/>, Annals of Internal Medicine at <http://www.annals.org/>, and Epidemiology at <http://www.epidem.com/>). Information on the STROBE Initiative is available at [www.strobe-statement.org](http://www.strobe-statement.org).



**Variation in severe maternal morbidity according to socioeconomic position: a UK national case-control study.**

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Keywords:	EPIDEMIOLOGY, OBSTETRICS, Maternal medicine < OBSTETRICS, PUBLIC HEALTH

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**Variation in severe maternal morbidity according to socioeconomic position: a UK national case-control study.**

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Correspondence to: Professor Marian Knight, NPEU, University of Oxford UK

Key words: maternal morbidity, socioeconomic position, UKOSS, NS-SEC

Word count: 2,606

Tables:

Table 1. Distribution of missing data

Table 2. Observed associations with severe maternal morbidity

Table 3. Sensitivity analysis of missing socioeconomic data

Figures:

Box 1. Criteria used to define cases in different UKOSS studies

## ARTICLE SUMMARY

### Article Focus

- To explore the relationship between socioeconomic position, defined by occupation, and severe maternal morbidity using data obtained from a series of UKOSS studies
- To estimate the risk of severe morbidity by socioeconomic group, independent of ethnicity, maternal age, smoking, pre-existing medical condition, BMI, multiple pregnancy and past pregnancy complications

### Key Messages

- The results suggested that women from the lowest socioeconomic group are 1.22 times (95%CI: 0.92 – 1.61) more likely than women from the highest group to experience severe maternal morbidity
- The results also showed that other risk factors for severe maternal morbidity include non-white ethnicity, older maternal age ( $\geq 35$  years), BMI  $\geq 25$  kg/m<sup>2</sup>, pre-existing medical condition/s, multiple pregnancy or previous pregnancy complications

### Strengths and Limitations

- Study strengths include the robust case ascertainment, minimisation of sampling bias through collection of data from all consultant-led obstetric units in the UK and the clear definition of severe maternal morbidity
- Limitations relate to the data available to define socioeconomic position and the application of the findings to other non-UK setting



**ABSTRACT**

**Objectives:** This study aimed to explore the independent association between socioeconomic position, defined by occupation, and severe maternal morbidity amongst women in the UK.

**Design:** Case-control study.

**Setting:** The analysis was conducted as a case-control analysis, using data from a series of studies of direct causes of severe maternal morbidity undertaken through the UK Obstetric Surveillance System (UKOSS), with data collected throughout all consultant-led obstetric units in the UK.

**Participants:** The analysis included 1,144 cases and 2,256 comparison women (controls). UKOSS studies from which data on case women were obtained included amniotic fluid embolism, acute fatty liver of pregnancy, eclampsia, peripartum hysterectomy, therapies for peripartum haemorrhage and uterine rupture.

**Primary outcome measure:** Odds of severe maternal morbidity by socioeconomic group, independent of ethnicity, maternal age, smoking, pre-existing medical condition, BMI, multiple pregnancy and past pregnancy complications. Occupation was used to classify different socioeconomic groups.

**Secondary outcome measure:** Odds of morbidity related to ethnic group, maternal age, smoking, pre-existing medical condition, BMI, multiple pregnancy and past pregnancy complications.

**Results:** Across the socioeconomic groups, compared to the 'managerial/professional' group, adjusted odds ratios were 1.17 (95%CI: 0.94 – 1.45) for the 'intermediate group', 1.16 (95%CI: 0.93 – 1.45) for 'routine/manual', 1.22 (95%CI: 0.92 – 1.61) for 'unemployed' women, and 1.51 (95%CI: 1.18 – 1.94) for women with missing socioeconomic information. Women of non-white ethnicity, older maternal age ( $\geq 35$  years), BMI  $\geq 25$  kg/m<sup>2</sup> and those with pre-existing medical condition/s, multiple pregnancy or past pregnancy complications were shown to have a significantly increased odds of severe maternal morbidity.

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3 **Conclusion:** This study suggests that socioeconomic position may be independently associated  
4 with an increased risk of severe maternal morbidity, although the observed association was not  
5 statistically significant. Further research is warranted to confirm this and investigate why this  
6 association might exist in a country where healthcare is universal and free at the point of  
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## INTRODUCTION

Maternal morbidity is of major public health concern with some estimates suggesting that as many as ten million women worldwide suffer from pregnancy-related complications each year<sup>1</sup>. In countries where maternal deaths are infrequent, the information that mortality audits provide cannot be generalised readily across the general population of pregnant women, nor can it necessarily accurately predict which women are at increased risk of adverse outcomes<sup>2</sup>. In these circumstances, additional focus on maternal morbidity is now widely recognised as a robust approach to improving maternal health and the quality of healthcare provision<sup>3-6</sup>.

Identifying high-risk women is imperative to prevent adverse outcomes, since it allows more intensive clinical management to be directed towards these women and is the fundamental purpose of antenatal care. Extensive research has demonstrated convincing links between severe maternal morbidity and advanced maternal age, pre-existing medical conditions and obesity<sup>1 7 8</sup>. However, these factors are unable to explain entirely the differences in maternal morbidity found between different populations of women both within and between countries. Other evidence suggests that ethnicity and social disadvantage may have a significant role to play, with UK data suggesting that women from ethnic minority groups and from socially disadvantaged groups are more likely to experience severe maternal morbidity than white women, and also to be at higher risk of dying from specific maternal morbidities<sup>9-11</sup>. However, as highlighted by several studies exploring ethnic differences in maternal health, it is unclear whether ethnicity itself is directly related to poor outcomes or, more likely, whether it is a marker for factors such as low socioeconomic position and lower levels of education<sup>12</sup>. Because minority ethnic groups are often disproportionately represented in lower socioeconomic groups, results attributed to ethnic differences are likely to be confounded by socioeconomic differences<sup>13</sup>. The aim of the analysis reported here was to explore whether there is an independent risk of severe maternal morbidity associated with socioeconomic position in the UK.

## METHODS

### Data Collection and Definitions

A case-control analysis was performed using data derived from a series of UK-wide studies of direct causes of maternal morbidity undertaken through the UK Obstetric Surveillance System (UKOSS)<sup>14</sup>. These studies were of amniotic fluid embolism (AFE), acute fatty liver of pregnancy (AFLP), eclampsia, peripartum hysterectomy, therapies for peripartum haemorrhage and uterine rupture<sup>15-21</sup>. For the purpose of this analysis, these cases are collectively referred to as 'severe maternal morbidity' (Box 1).

Box 1. Criteria used to define cases of severe maternal morbidity used in the UKOSS studies included in the analysis

Condition	Definition	UKOSS Study Timeframe
<b>Amniotic fluid embolism</b>	EITHER a clinical diagnosis of AFE (acute hypotension or cardiac arrest, acute hypoxia or coagulopathy in the absence of any other potential explanation for the symptoms and signs observed) OR a pathological diagnosis (presence of fetal squames or hair in the lungs)	01/02/2005 – 31/01/2010
<b>Acute fatty liver of pregnancy</b>	EITHER AFLP confirmed by biopsy or post-mortem examination OR a clinician has made a diagnosis of AFLP with signs and symptoms consistent with AFLP	01/02/2005 – 30/04/2006
<b>Eclampsia</b>	The occurrence of convulsions during pregnancy or in the first ten days postpartum, together with at least two of the following features within 24 hours after the convulsions: 1) Hypertension (a booking diastolic pressure of <90mmHg, a maximum diastolic of ≥90mmHg, and a diastolic increment of ≥25mmHg) 2) Proteinuria (at least + protein in a random urine sample or ≥0.3g in a 24hr collection) 3) Thrombocytopenia (platelet count of <100X10 <sup>9</sup> /l)	01/02/2005 – 29/02/2006

	4) An increased plasma alanine aminotransferase (ALT) concentration ( $\geq 42$ iu/l) or an increased plasma aspartate transaminase aminotransferase (AST) concentration ( $\geq 42$ iu/l)	
<b>Peripartum hysterectomy</b>	Any woman giving birth to an infant and having a hysterectomy during the same clinical episode	01/02/2005 – 29/02/2006
<b>Therapies for major peripartum haemorrhage</b>	All women in the UK treated therapeutically or prophylactically for major peripartum haemorrhage with: EITHER Activated factor VIIa OR B-lynch suture or other brace suture OR Arterial ligation or embolisation	01/09/2007 – 30/09/2008
<b>Uterine rupture</b>	Any woman in the UK identified as having a uterine rupture using the following definition: A complete separation of the wall of the pregnant uterus, with or without expulsion of the fetus, involving rupture of the membranes at the site of the uterine rupture or extension into the uterine muscle separate from any previous scar, and endangering the life of the mother or fetus Excluded: any asymptomatic palpable or visualised defect (for example dehiscence noted incidentally at caesarean delivery)	01/04/2009 – 30/04/2010

The UKOSS methodology used to carry out a rolling programme of studies has been described in detail elsewhere<sup>14</sup>. In brief, monthly mailing of notification cards to all obstetric-led maternity units in the UK enables identification of all cases under study. Data collection forms despatched in response to a case notification are used by the clinician responsible for care to report de-identified information from the woman's medical notes. Follow-up of non-responders maximises the completeness of data collection and confirms denominator numbers for incidence estimates.

For many of the UKOSS studies, data about a representative sample of women without the specific condition under study were obtained by collecting information on the two women

delivering immediately prior to each case in the same obstetric unit; these were defined as controls. The same information was collected about the controls as for case women. Data were thus available for analysis on 1,144 cases of severe maternal morbidity and 2,256 controls.

Maternal occupation was used to classify cases and controls by socioeconomic group, where this was not available her husband's/partner's occupation was used. Using the National Statistics Socio-Economic Classification (NS-SEC)<sup>22</sup>, occupation was coded into three categories plus an additional group for unemployed women.

### Statistical Analysis

All analyses were carried out using STATA version 11.0. Following univariable analysis multivariable unconditional logistic regression was used to examine the independent relationship between socioeconomic position and severe maternal morbidity and to generate adjusted odds ratios (aORs) with their 95% confidence intervals (95% CIs). Due to the high proportion and distribution of missing data for some variables, particularly for socio-economic position (Table 1), 'missing' was included as a separate group for each categorical variable.

Table 1. Distribution of missing data for cases of severe maternal morbidity and controls

Variable	Number of cases with missing information (% of total cases)	Number of controls with missing information (% of total controls)
Socioeconomic Group	193 (16.9)	303 (13.4)
Ethnicity	18 (1.6)	78 (3.5)
Age	3 (0.3)	18 (0.8)
Smoking	28 (2.5)	61 (2.7)
Pre-existing medical condition(s)	6 (0.5)	27 (1.2)
Body mass index	128 (11.2)	219 (9.7)
Parity	3 (0.3)	22 (1.0)
Multiple pregnancy	2 (0.2)	18 (0.8)
Past pregnancy complication(s)	9 (0.8)	31 (1.4)

Variables likely to confound the relationship between socioeconomic position and severe maternal morbidity were identified *a priori* based on evidence from the published literature and

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3 the results of univariable analysis. These included ethnicity, maternal age, smoking, pre-existing  
4 medical conditions and body mass index (BMI)<sup>17 10 23-36</sup>. Other variables included in the analysis  
5 as possible confounders were parity, multiple pregnancy and past pregnancy complications.  
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11 Model building using unconditional logistic regression proceeded by the inclusion of severe  
12 maternal morbidity (outcome) and socio-economic position, with each of the potential  
13 confounders added in turn one-by-one. The individual effect of each variable on the fit of the  
14 data was assessed using the likelihood ratio test. The variables with *a priori* evidence of effect  
15 were included in the final model. Of the additional three variables examined (parity, multiple  
16 pregnancy and past pregnancy conditions) only those demonstrating a statistically significant  
17 effect ( $p < 0.05$ ) on the fit of the data were retained; parity was thus excluded.  
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27 Age, included as a continuous variable, was tested for departure from linearity; the results are  
28 for ease of presentation, shown as categorical. BMI was included as a categorical variable due  
29 to the high proportion of missing data and the concern that by including it as continuous, women  
30 with missing data would be excluded. There was no evidence of a significant interaction  
31 between socioeconomic position and any of the other variables included in the final model  
32 ( $p < 0.01$ ).  
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42 Sensitivity analyses were conducted to explore the effect of missing data for socioeconomic  
43 position and BMI. The final multivariable regression model was re-run after redistributing  
44 missing values into the lowest and highest categories for each of these variables.  
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54 The findings demonstrated that the distribution and impact of missing data was  
55 substantial and that the pattern of missing data was unlikely to be 'missing at random'.  
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3 The results of multivariable analysis were therefore likely to be biased by the high  
4 proportion and non-random pattern of missing data if the missing observations were not  
5 accounted for. Due to the evidence against a 'missing at random' assumption, Multiple  
6 Imputation was not an appropriate solution to deal with the effects of missing data. We  
7 therefore included the missing data as a separate category for the relevant variables,  
8 treating them as 'proxy indicators' in the subsequent model-building processes.  
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### 18 **Study Power**

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20 The analysis included a fixed number of cases and controls derived from specified UKOSS  
21 studies. Based on 1,144 cases and 2,256 controls, and with a prevalence of 13% for the lowest  
22 (unemployed) socioeconomic group, the analysis had 80% power to detect as statistically  
23 significant ( $p < 0.05$ ) an odds of 1.32 or greater and 90% power to detect as statistically  
24 significant ( $p < 0.05$ ) an odds of 1.37 or greater, for comparisons with the highest  
25 (managerial/professional) socioeconomic group.  
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## 35 **RESULTS**

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37 The analysis included a total of 1,144 cases of severe maternal morbidity and 2,256 controls.  
38 Compared with the controls cases were more likely to be older, multiparous, of non-white  
39 ethnicity, to have missing socio-economic position and BMI information, to be non-smokers, to  
40 have a pre-existing medical condition, a history of past pregnancy complications and a multiple  
41 pregnancy in the index pregnancy (Table 2).  
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57 Table 2. Maternal characteristics associated with severe maternal morbidity  
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Characteristic	Cases n= (%) N= 1,144	Controls n= (%) N=2,256	Unadjusted OR (95% CI)	Adjusted* OR (95%CI)
<b>Socioeconomic group</b>				
Managerial/ Professional	292 (25.5)	567 (25.1)	1.0	1.0
Intermediate	244 (21.3)	482 (21.4)	0.98 (0.80 – 1.21)	1.17 (0.94 – 1.45)
Routine/Manual	273 (28.7)	595 (26.4)	0.89 (0.73 – 1.09)	1.16 (0.93 – 1.45)
Unemployed	142 (14.9)	309 (13.7)	0.89 (0.70 – 1.14)	1.22 (0.92 – 1.61)
Missing	193 (16.9)	303 (13.4)	1.24 (0.98 – 1.56)	1.51 (1.18 – 1.94)
<b>Ethnic group</b>				
White	827 (72.3)	1796 (79.6)	1.0	1.0
Asian	139 (12.2)	197 (8.7)	1.53 (1.22 – 1.93)	1.57 (1.23 – 2.00)
Black	108 (9.4)	116 (5.1)	2.02 (1.54 – 2.66)	1.77 (1.32 – 2.36)
Other	52 (4.6)	69 (3.1)	1.64 (1.13 – 2.37)	1.50 (1.02 – 2.19)
Missing	18 (1.6)	78 (3.5)	0.50 (0.30 – 0.84)	0.51 (0.28 – 0.91)
<b>Age</b>				
Age <20	56 (4.9)	139 (6.2)	0.98 (0.70 – 1.39)	1.08 (0.75 – 1.56)
Age 20 – 24	127 (11.1)	449 (19.9)	0.69 (0.54 – 0.89)	0.71 (0.55 – 0.92)
Age 25 – 29	225 (19.7)	549 (24.3)	1.0	1.0
Age 30 – 34	341 (29.8)	622 (27.6)	1.34 (1.09 – 1.64)	1.33 (1.08 – 1.64)
Age ≥35	392 (34.3)	479 (21.2)	2.00 (1.63 – 2.45)	1.98 (1.60 – 2.45)
Missing	3 (0.3)	18 (0.8)	0.41 (0.12 – 1.39)	2.66 (0.22 – 26.00)
<b>Smoking status</b>				
Non-Smokers	899 (78.6)	1,642 (72.8)	1.0	1.0
Smokers	217 (19.0)	553 (24.5)	0.72 (0.60 – 0.86)	0.89 (0.73 – 1.08)
Missing	28 (2.5)	61 (2.7)	0.84 (0.53 – 1.32)	0.95 (0.54 – 1.67)
<b>BMI</b>				
<b>Continuous (per kg/m<sup>2</sup> increase in BMI)</b>			1.01 (1.00 – 1.03)	
< 25 kg/m <sup>2</sup>	513 (44.8)	1,121 (49.7)	1.0	1.0
25-29.9 kg/m <sup>2</sup>	297 (26.0)	545 (24.2)	1.19 (1.00 – 1.42)	1.10 (0.92 – 1.32)
≥ 30 kg/m <sup>2</sup>	206 (18.0)	371 (16.5)	1.21 (0.99 – 1.48)	1.09 (0.88 – 1.34)
Missing	128 (11.2)	219 (9.7)	1.28 (1.00 – 1.63)	1.24 (0.95 – 1.61)
<b>Co-existing medical conditions**</b>				
No	1,045 (91.4)	2,126 (94.2)	1.0	1.0
Yes	93 (8.1)	103 (4.6)	1.84 (1.38 – 2.46)	1.60 (1.18 – 2.17)
Missing	6 (0.5)	27 (1.2)	0.45 (1.18 – 1.10)	0.76 (0.17 – 3.41)
<b>Parity</b>				
Nulliparous	403 (35.2)	984 (43.6)	1.0	No significant effect on fit of model
Multiparous	738 (64.5)	1,250 (55.4)	1.44 (1.24 – 1.67)	(excluded)
Missing	3 (0.3)	22 (1.0)	0.33 (0.10 – 1.12)	
<b>Multiple pregnancy</b>				
No	1,083 (94.7)	2,211 (98.0)	1.0	1.0
Yes	59 (5.2)	27 (1.2)	4.46 (2.81 – 7.08)	3.88 (2.42 – 6.22)
Missing	2 (0.2)	18 (0.8)	0.23 (0.53 – 0.98)	0.20 (0.02 – 2.54)
<b>Past pregnancy conditions***</b>				
No	1,041 (91.0)	2,121 (94.0)	1.0	1.0
Yes	94 (8.2)	104 (4.6)	1.81 (1.38 – 2.46)	1.58 (1.16 – 2.14)
Missing	9 (0.8)	31 (1.4)	0.59 (0.28 – 1.25)	1.12 (0.36 – 3.54)

\*All variables adjusted for all other variables in the table

\*\*Cardiac disease, diabetes mellitus, epilepsy, essential hypertension, haematological disorders, past thrombotic event, recent/current anticoagulation, IV drug use and/or alcohol abuse

\*\*\*Amniotic fluid embolism, gestational diabetes, pregnancy-induced hypertension, pre-eclampsia, eclampsia, thrombotic event, acute fatty liver of pregnancy, manual removal of placenta, placenta accreta, past classical caesarean section

Following adjustment, compared with the controls, cases were 1.17 (95%CI 0.94-1.45) times more likely to be in the 'intermediate' socioeconomic group than the 'managerial/professional' group, 1.16 (95%CI: 0.93 – 1.45) times more likely to be in the 'routine/manual' group, 1.22 (95%CI: 0.92 – 1.61) times more likely to be unemployed, noting that these were not statistically significant increases. Cases were statistically significantly 1.51 (95%CI: 1.18 – 1.94) times more likely to have missing socioeconomic information than controls. Compared with controls cases were 1.77 (95%CI 1.32-2.36) times more likely to be black than white, 1.57 (95%CI 1.23-2.00) times more likely to be Asian and 1.50 (95%CI 1.02-2.19) times more likely to be from another minority ethnic group, all statistically significant associations. The relationship between the odds severe maternal morbidity and maternal age was J-shaped with cases having significantly raised adjusted odds of being 30-34yrs (aOR 1.33; 95%CI 1.08-1.64) or 35+yrs (aOR 1.98; 95%CI 1.60-2.45). After adjustment, cases had a non-significant 10% increase in the odds of having a BMI of 25 or over and a 24% increase in the odds of having missing BMI data which was similarly non-significant. Cases were significantly more likely than controls to have a history of pre-existing medical conditions (aOR 1.60; 95%CI 1.18-2.17), past pregnancy complications (aOR 1.58; 95%CI 1.16-2.14) and to have a multiple pregnancy (aOR 3.54; 95%CI 2.18-5.76).

### Sensitivity analysis

Redistributing all the missing socioeconomic observations into the managerial/professional group reduced the aOR associated with unemployment to 1.02 (95%CI 0.79-1.32), a non-significant association, whereas, redistribution of all the missing socioeconomic observations into the unemployed group produced an aOR of 1.38 (95%CI 1.11–1.72) associated with unemployment (Table 3), a statistically significant increase.

Table 3. Sensitivity analysis for missing observations

Missing values	Sub-category	Cases (%)	Controls (%)	Adjusted OR (95% CI)
<b>Socioeconomic group</b>				

Missing SEP values recoded to 'Managerial/Professional' (best case)	Managerial	485 (42.4)	870 (38.6)	1.0
	Intermediate	244 (21.3)	482 (21.4)	1.01 (0.83 – 1.23)
	Routine/Manual	273 (23.9)	595 (26.4)	0.99 (0.82 – 1.21)
	Unemployed	142 (12.4)	309 (13.7)	1.02 (0.79 – 1.32)
Missing SEP values recoded to 'unemployed' (worst case)	Managerial	292 (25.5)	567 (25.1)	1.0
	Intermediate	244 (21.3)	482 (21.37)	1.17 (0.94 – 1.46)
	Routine/Manual	273 (23.9)	595 (26.4)	1.17 (0.94 – 1.46)
	Unemployed	335 (29.3)	612 (27.1)	1.38 (1.11 – 1.72)

Re-running the multivariable model including only those cases and controls with complete information for every variable generated an aOR of 1.31 (95%CI 0.97–1.77) associated with unemployment, an aOR of 1.13 (95%CI 0.89-1.43) for the routine/manual group and 1.19 (95%CI 0.94-1.49) for the intermediate group. These results are very similar to the estimates from the analysis using proxy indicators.

## DISCUSSION

### Findings

The results of this study suggest that there may be an independent association between socioeconomic position and severe maternal morbidity. We found that, compared to women from the highest managerial/professional socioeconomic group, unemployed status was associated with a 22% increase in the odds of severe morbidity. However, this was not a statistically significant association and may therefore represent a chance finding. Women with missing socioeconomic information had a more than 50% increase in the odds of severe morbidity and this finding was unlikely to be due to chance. The women with missing socioeconomic information were clearly different in a range of characteristics than the women with complete information, and we hypothesised that these women were more likely to be unemployed. To explore what the impact of this might be in our analysis if this hypothesis were correct, we undertook a sensitivity analysis which suggested that if all the women with missing socioeconomic data were unemployed, the estimate of the odds of severe morbidity increased to 38%; with a result unlikely to be due to chance. It is important to note, however, that as

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3 information about the covariates was obtained by an obstetrician knowing the outcome status of  
4 the women, the missing values and/or potential misclassification might be differential according  
5 to outcome. We are unable to assess any further whether our hypothesis about those with  
6 missing information is correct; nevertheless, we believe this potential association is worthy of  
7 further investigation amongst different populations and using different methodologies to  
8 determine whether it can be reproduced.  
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17 Despite universal health care which is free at the point of delivery, and significant investment in  
18 researching social inequalities in health in the UK, socially-marginalised groups continue to fare  
19 poorly as compared with more affluent sections of society, across a range of health outcomes  
20 including, as suggested by our analysis, maternal health. The reasons for the social gradient in  
21 maternal health in the UK have not yet been thoroughly explored. It is therefore unclear whether  
22 the observed potential association is as a result of differences in the underlying risk of disease,  
23 including, but not limited to, lifestyle factors, differences in health-seeking behaviour, or  
24 differences in access to, or provision and quality of services during pregnancy, delivery and the  
25 post-partum period.  
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37 In addition to finding an association between severe morbidity and factors such as pre-existing  
38 medical conditions, multiple pregnancy and past pregnancy complications, this study also  
39 reflected the findings of past studies in France and the Netherlands<sup>6 26 32</sup> which have  
40 demonstrated a significantly increased risk of severe morbidity amongst women from minority  
41 ethnic groups. The suggested association with socioeconomic position was, however,  
42 independent of these associations.  
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### 52 **Implications of the Study Findings**

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54 Health research in various settings worldwide has illuminated the ongoing existence of the  
55 'inverse care law', first described in 1971 by Tudor Hart as the tendency for the "availability of  
56 good medical care to vary inversely with the need for it in the population served"<sup>37</sup>.  
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5 Since the global financial crisis in 2008, the World Health Organisation has stressed that “as  
6 standards of living decrease in many countries, and government revenues are tightened, [we  
7 would argue that] it is even more urgent that the distributional effects on all policies are taken  
8 into account in policy decision making”<sup>38</sup>. In a paper that argues that social welfare spending is  
9 as important as protecting health through safeguarding the healthcare budget, Stuckler et al  
10 have proposed that the economic crisis presents the opportunity to reorganise the provision of  
11 services to those most in need<sup>39</sup>. Our findings identify one such group who may be in need.  
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21 ‘Maternity Matters’ guidance for service commissioning in England released in 2007  
22 emphasised the Government’s commitment to providing choice guarantees for women  
23 regarding type of antenatal care and place of birth and also highlighted the need for future  
24 maternity services to address disproportionately higher rates of maternal morbidity amongst  
25 disadvantaged women<sup>40</sup>. Although it described a comprehensive approach to delivering  
26 improved maternity care the guidance did not provide any specific recommendations as to how  
27 to improve services for women from socially disadvantaged backgrounds. Our study  
28 emphasises the need for developing such recommendations within the currently changing  
29 structure of health service commissioning in England.  
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#### 41 **Strengths and Limitations of the Study**

42 UKOSS, the first system of its kind in obstetrics worldwide, has provided an invaluable wealth of  
43 information on uncommon but serious disorders of pregnancy. The active, negative surveillance  
44 approach has ensured robust case ascertainment, and many of the UKOSS studies are the  
45 largest-scale studies of rare disorders of pregnancy worldwide.  
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54 By studying disorders of pregnancy and childbirth in separate studies, UKOSS provided the  
55 means by which to explore ‘severe maternal morbidity’ due to direct pregnancy causes as an  
56 aggregate of multiple conditions, thereby avoiding the challenges encountered by previous  
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3 studies which have used combined clinical and intervention criteria. It also avoids the perennial  
4 difficulty that there is no universally agreed definition of what constitutes 'severe maternal  
5 morbidity'. One disadvantage of this approach is that 'indirect causes' of maternal morbidity  
6 were excluded from our analysis, including several major contributors to maternal mortality in  
7 the UK such as cardiac disease, thromboembolic disease and puerperal psychosis/suicide<sup>41</sup>.  
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12 The design of UKOSS studies ensures that sampling bias is minimised by recruiting cases from  
13 all obstetric units across the UK, thus including women from all different ethnic, age and  
14 socioeconomic groups. Selection of controls is also aimed at attaining a representative sample  
15 of the general population of women giving birth by recruiting women from the same obstetric  
16 units as cases throughout the UK during the same time period. Comparison between UKOSS  
17 controls and national maternity data suggests that the controls recruited for UKOSS studies are  
18 closely representative of the general, child-bearing population<sup>19</sup>.  
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31 A universal challenge for studies that seek to explore the impact of social disadvantage in  
32 society is how to best capture the concept of 'social disadvantage' and define socioeconomic  
33 position. There are multiple dimensions to social disadvantage and there are many different  
34 approaches to defining a person's position in the social structure of society. The retrospective  
35 collection of data for UKOSS from patient records limited the definition of socioeconomic  
36 position to maternal occupation (or paternal where maternal was not available). Using the  
37 national classification of socioeconomic position (NS-SEC) ensured the study population was  
38 categorised using a system that is used nationally. However, the collection of other  
39 socioeconomic indicators such as level of education, years of schooling, residential location or  
40 income was not possible although this would have allowed further analysis of social  
41 disadvantage by using composite measures of disadvantage such as the Index of Multiple  
42 Deprivation. There is thus a place for further research using other indicators of socioeconomic  
43 position to see whether the results of this study are replicated.  
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3 The results of this study clearly apply only to high resource countries with well-developed  
4 healthcare systems. However, other high resource countries have different models of health  
5 care, such that pregnancy and delivery care may not be free at the point of access, as is the  
6 case in the UK. This study used only UK data and the results may therefore not be  
7 generalisable to women giving birth in other countries. Further research using information from  
8 other settings is therefore important to investigate whether socioeconomic position is associated  
9 with severe maternal morbidity in countries with differing health systems as this may give  
10 insights into the causes and where modifiable risks might lie. The use of data collected through  
11 the International Network of Obstetric Surveillance Systems<sup>42</sup>, which use similar definitions and  
12 methodologies, may be a route to achieve this.  
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## 25 **Conclusions**

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27 This study suggests that socioeconomic position may be associated with an increased risk of  
28 severe maternal morbidity, independent of risks associated with maternal age, ethnic group or  
29 known medical and pregnancy complications, although the association we observed was not  
30 statistically significant. However, the reasons for this possible association remain unclear,  
31 suggesting a need, not only to see if the observed association can be reproduced in other  
32 populations or settings, but also for further research to investigate whether there are differences  
33 in disease risks, lifestyle and health-seeking behaviour, access to care or the quality of care  
34 received by women from different socioeconomic groups. Further information about these  
35 differences would allow development of specific recommendations for the care of women from  
36 socially disadvantaged backgrounds in order to minimise future risks.  
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**Variation in severe maternal morbidity according to socioeconomic position: a UK national case-control study.**

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Key words: maternal morbidity, socioeconomic position, UKOSS, NS-SEC

Word count: 2,606

Tables:

Table 1. Distribution of missing data

Table 2. Observed associations with severe maternal morbidity

Table 3. Sensitivity analysis of missing socioeconomic data

Figures:

Box 1. Criteria used to define cases in different UKOSS studies

## ARTICLE SUMMARY

### Article Focus

- To explore the relationship between socioeconomic position, [defined by occupation](#), and severe maternal morbidity using data obtained from a series of UKOSS studies
- To estimate the risk of severe morbidity by socioeconomic group, independent of ethnicity, maternal age, smoking, pre-existing medical condition, BMI, multiple pregnancy and past pregnancy complications

### Key Messages

- The results suggested that women from the lowest socioeconomic group are 1.22 times (95%CI: 0.92 – 1.61) more likely than women from the highest group to experience severe maternal morbidity
- The results also showed that other risk factors for severe maternal morbidity include non-white ethnicity, older maternal age ( $\geq 35$  years), BMI  $\geq 25$  kg/m<sup>2</sup>, pre-existing medical condition/s, multiple pregnancy or [previous](#) pregnancy complications

### Strengths and Limitations

- Study strengths include the robust case ascertainment, minimisation of sampling bias through collection of data from all consultant-led obstetric units in the UK and the clear definition of severe maternal morbidity
- Limitations relate to the data available to define socioeconomic position and the application of the findings to other non-UK setting

## ABSTRACT

**Objectives:** This study aimed to explore the independent association between socioeconomic position, defined by occupation, and severe maternal morbidity amongst women in the UK.

**Design:** [Case-control](#) study. **Setting:** The analysis was conducted as a case-control analysis, using data from a series of studies of direct causes of severe maternal morbidity undertaken through the UK Obstetric Surveillance System (UKOSS), with data collected throughout all consultant-led obstetric units in the UK. **Participants:** The analysis included 1,144 cases and 2,256 comparison women (controls). UKOSS studies from which data on case women were obtained included amniotic fluid embolism, acute fatty liver of pregnancy, eclampsia, peripartum hysterectomy, therapies for peripartum haemorrhage and uterine rupture. **Primary outcome measure:** Odds of severe maternal morbidity by socioeconomic group, independent of ethnicity, maternal age, smoking, pre-existing medical condition, BMI, multiple pregnancy and past pregnancy complications. Occupation was used to classify different socioeconomic groups.

**Secondary outcome measure:** Odds of morbidity related to ethnic group, maternal age, smoking, pre-existing medical condition, BMI, multiple pregnancy and past pregnancy complications. **Results:** Across the socioeconomic groups, [compared to the 'managerial/professional' group](#), adjusted odds ratios were 1.17 (95%CI: 0.94 – 1.45) for the 'intermediate group', 1.16 (95%CI: 0.93 – 1.45) for 'routine/manual', 1.22 (95%CI: 0.92 – 1.61) for 'unemployed' women, and 1.51 (95%CI: 1.18 – 1.94) for women with missing socioeconomic information. Women of non-white ethnicity, older maternal age ( $\geq 35$  years), BMI  $\geq 25$  kg/m<sup>2</sup> and those with pre-existing medical condition/s, multiple pregnancy or past pregnancy complications were shown to have a significantly increased odds of severe maternal morbidity. **Conclusion:** This study suggests that socioeconomic position may be independently associated with an increased risk of severe maternal morbidity, [although the observed association was not statistically significant](#). Further research is warranted to [confirm this and investigate](#) why this

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association might exist in a country where healthcare is universal and free at the point of access.

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## INTRODUCTION

Maternal morbidity is of major public health concern with some estimates suggesting that as many as ten million women worldwide suffer from pregnancy-related complications each year<sup>1</sup>. In countries where maternal deaths are infrequent, the information that mortality audits provide cannot be generalised readily across the general population of pregnant women, nor can it necessarily accurately predict which women are at increased risk of adverse outcomes<sup>2</sup>. In these circumstances, additional focus on maternal morbidity is now widely recognised as a robust approach to improving maternal health and the quality of healthcare provision<sup>3-6</sup>.

Identifying high-risk women is imperative to prevent adverse outcomes, since it allows more intensive clinical management to be directed towards these women and is the fundamental purpose of antenatal care. Extensive research has demonstrated convincing links between severe maternal morbidity and advanced maternal age, pre-existing medical conditions and obesity<sup>1 7 8</sup>. However, these factors are unable to explain entirely the differences in maternal morbidity found between different populations of women both within and between countries. Other evidence suggests that ethnicity and social disadvantage may have a significant role to play, with UK data suggesting that women from ethnic minority groups and from socially disadvantaged groups are more likely to experience severe maternal morbidity than white women, and also to be at higher risk of dying from specific maternal morbidities<sup>9-11</sup>. However, as highlighted by several studies exploring ethnic differences in maternal health, it is unclear whether ethnicity itself is directly related to poor outcomes or, more likely, whether it is a marker for factors such as low socioeconomic position and lower levels of education<sup>12</sup>. Because minority ethnic groups are often disproportionately represented in lower socioeconomic groups, results attributed to ethnic differences are likely to be confounded by socioeconomic differences<sup>13</sup>. The aim of the analysis reported here was to explore whether there is an independent risk of severe maternal morbidity associated with socioeconomic position in the UK.



## METHODS

### Data Collection and Definitions

A case-control analysis was performed using data derived from a series of UK-wide studies of direct causes of maternal morbidity undertaken through the UK Obstetric Surveillance System (UKOSS)<sup>14</sup>. These studies were of amniotic fluid embolism (AFE), acute fatty liver of pregnancy (AFLP), eclampsia, peripartum hysterectomy, therapies for peripartum haemorrhage and uterine rupture<sup>15-21</sup>. For the purpose of this analysis, these cases are collectively referred to as 'severe maternal morbidity' (Box 1).

Box 1. Criteria used to define cases of severe maternal morbidity used in the UKOSS studies included in the analysis

Condition	Definition	UKOSS Study Timeframe
<b>Amniotic fluid embolism</b>	EITHER a clinical diagnosis of AFE (acute hypotension or cardiac arrest, acute hypoxia or coagulopathy in the absence of any other potential explanation for the symptoms and signs observed) OR a pathological diagnosis (presence of fetal squames or hair in the lungs)	<a href="#">01/02/2005 – 31/01/2010</a>
<b>Acute fatty liver of pregnancy</b>	EITHER AFLP confirmed by biopsy or post-mortem examination OR a clinician has made a diagnosis of AFLP with signs and symptoms consistent with AFLP	<a href="#">01/02/2005 – 30/04/2006</a>
<b>Eclampsia</b>	The occurrence of convulsions during pregnancy or in the first ten days postpartum, together with at least two of the following features within 24 hours after the convulsions: 1) Hypertension (a booking diastolic pressure of <90mmHg, a maximum diastolic of ≥90mmHg, and a diastolic increment of ≥25mmHg) 2) Proteinuria (at least + protein in a random urine sample or ≥0.3g in a 24hr collection) 3) Thrombocytopenia (platelet count of <100X10 <sup>9</sup> /l)	<a href="#">01/02/2005 – 29/02/2006</a>

	4) An increased plasma alanine aminotransferase (ALT) concentration ( $\geq 42$ iu/l) or an increased plasma aspartate transaminase aminotransferase (AST) concentration ( $\geq 42$ iu/l)	
<b><u>Peripartum hysterectomy</u></b>	<u>Any woman giving birth to an infant and having a hysterectomy during the same clinical episode</u>	<u>01/02/2005 – 29/02/2006</u>
<b><u>Therapies for major peripartum haemorrhage</u></b>	<u>All women in the UK treated therapeutically or prophylactically for major peripartum haemorrhage with: EITHER Activated factor VIIa OR B-lynch suture or other brace suture OR Arterial ligation or embolisation</u>	<u>01/09/2007 – 30/09/2008</u>
<b><u>Uterine rupture</u></b>	<u>Any woman in the UK identified as having a uterine rupture using the following definition: A complete separation of the wall of the pregnant uterus, with or without expulsion of the fetus, involving rupture of the membranes at the site of the uterine rupture or extension into the uterine muscle separate from any previous scar, and endangering the life of the mother or fetus Excluded: any asymptomatic palpable or visualised defect (for example dehiscence noted incidentally at caesarean delivery)</u>	<u>01/04/2009 – 30/04/2010</u>

The UKOSS methodology used to carry out a rolling programme of studies has been described in detail elsewhere<sup>14</sup>. In brief, monthly mailing of notification cards to all obstetric-led maternity units in the UK enables identification of all cases under study. Data collection forms despatched in response to a case notification are used by the clinician responsible for care to report de-identified information from the woman's medical notes. Follow-up of non-responders maximises the completeness of data collection and confirms denominator numbers for incidence estimates.

For many of the UKOSS studies, data about a representative sample of women without the specific condition under study were obtained by collecting information on the two women

delivering immediately prior to each case in the same obstetric unit; these were defined as controls. The same information was collected about the controls as for case women. Data were thus available for analysis on 1,144 cases of severe maternal morbidity and 2,256 controls.

Maternal occupation was used to classify cases and controls by socioeconomic group, where this was not available her husband's/partner's occupation was used. Using the National Statistics Socio-Economic Classification (NS-SEC)<sup>22</sup>, occupation was coded into three categories plus an additional group for unemployed women.

### Statistical Analysis

All analyses were carried out using STATA version 11.0. Following univariable analysis multivariable unconditional logistic regression was used to examine the independent relationship between socioeconomic position and severe maternal morbidity and to generate adjusted odds ratios (aORs) with their 95% confidence intervals (95% CIs). Due to the high proportion and distribution of missing data for some variables, particularly for socio-economic position (Table 1), 'missing' was included as a separate group for each categorical variable.

Table 1. Distribution of missing data for cases of severe maternal morbidity and controls

Variable	Number of cases with missing information (% of total cases)	Number of controls with missing information (% of total controls)
Socioeconomic Group	193 (16.9)	303 (13.4)
Ethnicity	18 (1.6)	78 (3.5)
Age	3 (0.3)	18 (0.8)
Smoking	28 (2.5)	61 (2.7)
Pre-existing medical condition(s)	6 (0.5)	27 (1.2)
Body mass index	128 (11.2)	219 (9.7)
Parity	3 (0.3)	22 (1.0)
Multiple pregnancy	2 (0.2)	18 (0.8)
Past pregnancy complication(s)	9 (0.8)	31 (1.4)

Variables likely to confound the relationship between socioeconomic position and severe maternal morbidity were identified *a priori* based on evidence from the published literature and

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3 the results of univariable analysis. These included ethnicity, maternal age, smoking, pre-existing  
4 medical conditions and body mass index (BMI)<sup>17 10 23-36</sup>. Other variables included in the analysis  
5 as possible confounders were parity, multiple pregnancy and past pregnancy complications.  
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10 Model building using unconditional logistic regression proceeded by the inclusion of severe  
11 maternal morbidity (outcome) and socio-economic position, with each of the potential  
12 confounders added in turn one-by-one. The individual effect of each variable on the fit of the  
13 data was assessed using the likelihood ratio test. The variables with *a priori* evidence of effect  
14 were included in the final model. Of the additional three variables examined (parity, multiple  
15 pregnancy and past pregnancy conditions) only those demonstrating a statistically significant  
16 effect ( $p < 0.05$ ) on the fit of the data were retained; parity was thus excluded.  
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25 Age, included as a continuous variable, was tested for departure from linearity; the results are  
26 for ease of presentation, shown as categorical. BMI was included as a categorical variable due  
27 to the high proportion of missing data and the concern that by including it as continuous, women  
28 with missing data would be excluded. There was no evidence of a significant interaction  
29 between socioeconomic position and any of the other variables included in the final model  
30 ( $p < 0.01$ ).  
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41 Sensitivity analyses were conducted to explore the effect of missing data for socioeconomic  
42 position and BMI. The final multivariable regression model was re-run after redistributing  
43 missing values into the lowest and highest categories for each of these variables.  
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54 The findings demonstrated that the distribution and impact of missing data was  
55 substantial and that the pattern of missing data was unlikely to be 'missing at random'.  
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The results of multivariable analysis were therefore likely to be biased by the high proportion and non-random pattern of missing data if the missing observations were not accounted for. Due to the evidence against a 'missing at random' assumption, Multiple Imputation was not an appropriate solution to deal with the effects of missing data. We therefore included the missing data as a separate category for the relevant variables, treating them as 'proxy indicators' in the subsequent model-building processes.

### Study Power

The analysis included a fixed number of cases and controls derived from specified UKOSS studies. Based on 1,144 cases and 2,256 controls, and with a prevalence of 13% for the lowest (unemployed) socioeconomic group, the analysis had 80% power to detect as statistically significant ( $p < 0.05$ ) an odds of 1.32 or greater and 90% power to detect as statistically significant ( $p < 0.05$ ) an odds of 1.37 or greater, for comparisons with the highest (managerial/professional) socioeconomic group.

### RESULTS

The analysis included a total of 1,144 cases of severe maternal morbidity and 2,256 controls. Compared with the controls cases were more likely to be older, multiparous, of non-white ethnicity, to have missing socio-economic position and BMI information, to be non-smokers, to have a pre-existing medical condition, a history of past pregnancy complications and a multiple pregnancy in the index pregnancy (Table 2).

Table 2. Maternal characteristics associated with severe maternal morbidity

Characteristic	Cases n= (%) N= 1,144	Controls n= (%) N=2,256	Unadjusted OR (95% CI)	Adjusted* OR (95%CI)
<b>Socioeconomic group</b>				
Managerial/ Professional	292 (25.5)	567 (25.1)	1.0	1.0
Intermediate	244 (21.3)	482 (21.4)	0.98 (0.80 – 1.21)	1.17 (0.94 – 1.45)
Routine/Manual	273 (28.7)	595 (26.4)	0.89 (0.73 – 1.09)	1.16 (0.93 – 1.45)
Unemployed	142 (14.9)	309 (13.7)	0.89 (0.70 – 1.14)	1.22 (0.92 – 1.61)
Missing	193 (16.9)	303 (13.4)	1.24 (0.98 – 1.56)	1.51 (1.18 – 1.94)
<b>Ethnic group</b>				
White	827 (72.3)	1796 (79.6)	1.0	1.0
Asian	139 (12.2)	197 (8.7)	1.53 (1.22 – 1.93)	1.57 (1.23 – 2.00)
Black	108 (9.4)	116 (5.1)	2.02 (1.54 – 2.66)	1.77 (1.32 – 2.36)
Other	52 (4.6)	69 (3.1)	1.64 (1.13 – 2.37)	1.50 (1.02 – 2.19)
Missing	18 (1.6)	78 (3.5)	0.50 (0.30 – 0.84)	0.51 (0.28 – 0.91)
<b>Age</b>				
Age <20	56 (4.9)	139 (6.2)	0.98 (0.70 – 1.39)	1.08 (0.75 – 1.56)
Age 20 – 24	127 (11.1)	449 (19.9)	0.69 (0.54 – 0.89)	0.71 (0.55 – 0.92)
Age 25 – 29	225 (19.7)	549 (24.3)	1.0	1.0
Age 30 – 34	341 (29.8)	622 (27.6)	1.34 (1.09 – 1.64)	1.33 (1.08 – 1.64)
Age ≥35	392 (34.3)	479 (21.2)	2.00 (1.63 – 2.45)	1.98 (1.60 – 2.45)
Missing	3 (0.3)	18 (0.8)	0.41 (0.12 – 1.39)	2.66 (0.22 – 26.00)
<b>Smoking status</b>				
Non-Smokers	899 (78.6)	1,642 (72.8)	1.0	1.0
Smokers	217 (19.0)	553 (24.5)	0.72 (0.60 – 0.86)	0.89 (0.73 – 1.08)
Missing	28 (2.5)	61 (2.7)	0.84 (0.53 – 1.32)	0.95 (0.54 – 1.67)
<b>BMI</b>				
<b>Continuous (per kg/m<sup>2</sup> increase in BMI)</b>			1.01 (1.00 – 1.03)	
< 25 kg/m <sup>2</sup>	513 (44.8)	1,121 (49.7)	1.0	1.0
25-29.9 kg/m <sup>2</sup>	297 (26.0)	545 (24.2)	1.19 (1.00 – 1.42)	1.10 (0.92 – 1.32)
≥ 30 kg/m <sup>2</sup>	206 (18.0)	371 (16.5)	1.21 (0.99 – 1.48)	1.09 (0.88 – 1.34)
Missing	128 (11.2)	219 (9.7)	1.28 (1.00 – 1.63)	1.24 (0.95 – 1.61)
<b>Co-existing medical conditions**</b>				
No	1,045 (91.4)	2,126 (94.2)	1.0	1.0
Yes	93 (8.1)	103 (4.6)	1.84 (1.38 – 2.46)	1.60 (1.18 – 2.17)
Missing	6 (0.5)	27 (1.2)	0.45 (1.18 – 1.10)	0.76 (0.17 – 3.41)
<b>Parity</b>				
Nulliparous	403 (35.2)	984 (43.6)	1.0	No significant effect on fit of model
Multiparous	738 (64.5)	1,250 (55.4)	1.44 (1.24 – 1.67)	(excluded)
Missing	3 (0.3)	22 (1.0)	0.33 (0.10 – 1.12)	
<b>Multiple pregnancy</b>				
No	1,083 (94.7)	2,211 (98.0)	1.0	1.0
Yes	59 (5.2)	27 (1.2)	4.46 (2.81 – 7.08)	3.88 (2.42 – 6.22)
Missing	2 (0.2)	18 (0.8)	0.23 (0.53 – 0.98)	0.20 (0.02 – 2.54)
<b>Past pregnancy conditions***</b>				
No	1,041 (91.0)	2,121 (94.0)	1.0	1.0
Yes	94 (8.2)	104 (4.6)	1.81 (1.38 – 2.46)	1.58 (1.16 – 2.14)
Missing	9 (0.8)	31 (1.4)	0.59 (0.28 – 1.25)	1.12 (0.36 – 3.54)

\*All variables adjusted for all other variables in the table

\*\*Cardiac disease, diabetes mellitus, epilepsy, essential hypertension, haematological disorders, past thrombotic event, recent/current anticoagulation, IV drug use and/or alcohol abuse

\*\*\*Amniotic fluid embolism, gestational diabetes, pregnancy-induced hypertension, pre-eclampsia, eclampsia, thrombotic event, acute fatty liver of pregnancy, manual removal of placenta, placenta accreta, past classical caesarean section

Following adjustment, compared with the controls, cases were 1.17 (95%CI 0.94-1.45) times more likely to be in the 'intermediate' socioeconomic group than the 'managerial/professional' group, 1.16 (95%CI: 0.93 – 1.45) times more likely to be in the 'routine/manual' group, 1.22 (95%CI: 0.92 – 1.61) times more likely to be unemployed, noting that these were not statistically significant increases. Cases were statistically significantly 1.51 (95%CI: 1.18 – 1.94) times more likely to have missing socioeconomic information than controls. Compared with controls cases were 1.77 (95%CI 1.32-2.36) times more likely to be black than white, 1.57 (95%CI 1.23-2.00) times more likely to be Asian and 1.50 (95%CI 1.02-2.19) times more likely to be from another minority ethnic group, all statistically significant associations. The relationship between the odds severe maternal morbidity and maternal age was J-shaped with cases having significantly raised adjusted odds of being 30-34yrs (aOR 1.33; 95%CI 1.08-1.64) or 35+yrs (aOR 1.98; 95%CI 1.60-2.45). After adjustment, cases had a non-significant 10% increase in the odds of having a BMI of 25 or over and a 24% increase in the odds of having missing BMI data which was similarly non-significant. Cases were significantly more likely than controls to have a history of pre-existing medical conditions (aOR 1.60; 95%CI 1.18-2.17), past pregnancy complications (aOR 1.58; 95%CI 1.16-2.14) and to have a multiple pregnancy (aOR 3.54; 95%CI 2.18-5.76).

### Sensitivity analysis

Redistributing all the missing socioeconomic observations into the managerial/professional group reduced the aOR associated with unemployment to 1.02 (95%CI 0.79-1.32), a non-significant association, whereas, redistribution of all the missing socioeconomic observations into the unemployed group produced an aOR of 1.38 (95%CI 1.11–1.72) associated with unemployment (Table 3), a statistically significant increase.

Table 3. Sensitivity analysis for missing observations

Missing values	Sub-category	Cases (%)	Controls (%)	Adjusted OR (95% CI)
<b>Socioeconomic group</b>				

Missing SEP values recoded to 'Managerial/Professional' (best case)	Managerial	485 (42.4)	870 (38.6)	1.0
	Intermediate	244 (21.3)	482 (21.4)	1.01 (0.83 – 1.23)
	Routine/Manual	273 (23.9)	595 (26.4)	0.99 (0.82 – 1.21)
	Unemployed	142 (12.4)	309 (13.7)	1.02 (0.79 – 1.32)
Missing SEP values recoded to 'unemployed' (worst case)	Managerial	292 (25.5)	567 (25.1)	1.0
	Intermediate	244 (21.3)	482 (21.37)	1.17 (0.94 – 1.46)
	Routine/Manual	273 (23.9)	595 (26.4)	1.17 (0.94 – 1.46)
	Unemployed	335 (29.3)	612 (27.1)	1.38 (1.11 – 1.72)

Re-running the multivariable model including only those cases and controls with complete information for every variable generated an aOR of 1.31 (95%CI 0.97–1.77) associated with unemployment, an aOR of 1.13 (95%CI 0.89-1.43) for the routine/manual group and 1.19 (95%CI 0.94-1.49) for the intermediate group. These results are very similar to the estimates from the analysis using proxy indicators.

## DISCUSSION

### Findings

The results of this study suggest that there may be an independent association between socioeconomic position and severe maternal morbidity. We found that, compared to women from the highest managerial/professional socioeconomic group, unemployed status was associated with a 22% increase in the odds of severe morbidity. However, this was not a statistically significant association and may therefore represent a chance finding. Women with missing socioeconomic information had a more than 50% increase in the odds of severe morbidity and this finding was unlikely to be due to chance. The women with missing socioeconomic information were clearly different in a range of characteristics than the women with complete information, and we hypothesised that these women were more likely to be unemployed. To explore what the impact of this might be in our analysis if this hypothesis were correct, we undertook a sensitivity analysis which suggested that if all the women with missing socioeconomic data were unemployed, the estimate of the odds of severe morbidity increased to 38%; with a result unlikely to be due to chance. It is important to note, however, that as



information about the covariates was obtained by an obstetrician knowing the outcome status of the women, the missing values and/or potential misclassification might be differential according to outcome. We are unable to assess any further whether our hypothesis about those with missing information is correct; nevertheless, we believe this potential association is worthy of further investigation amongst different populations and using different methodologies to determine whether it can be reproduced.

Despite universal health care which is free at the point of delivery, and significant investment in researching social inequalities in health in the UK, socially-marginalised groups continue to fare poorly as compared with more affluent sections of society, across a range of health outcomes including, as suggested by our analysis, maternal health. The reasons for the social gradient in maternal health in the UK have not yet been thoroughly explored. It is therefore unclear whether the observed potential association is as a result of differences in the underlying risk of disease, including, but not limited to, lifestyle factors, differences in health-seeking behaviour, or differences in access to, or provision and quality of services during pregnancy, delivery and the post-partum period.

In addition to finding an association between severe morbidity and factors such as pre-existing medical conditions, multiple pregnancy and past pregnancy complications, this study also reflected the findings of past studies in France and the Netherlands<sup>6 26 32</sup> which have demonstrated a significantly increased risk of severe morbidity amongst women from minority ethnic groups. The suggested association with socioeconomic position was, however, independent of these associations.

### **Implications of the Study Findings**

Health research in various settings worldwide has illuminated the ongoing existence of the 'inverse care law', first described in 1971 by Tudor Hart as the tendency for the "availability of good medical care to vary inversely with the need for it in the population served"<sup>37</sup>.

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5 Since the global financial crisis in 2008, the World Health Organisation has stressed that “as  
6 standards of living decrease in many countries, and government revenues are tightened, [we  
7 would argue that] it is even more urgent that the distributional effects on all policies are taken  
8 into account in policy decision making”<sup>38</sup>. In a paper that argues that social welfare spending is  
9 as important as protecting health through safeguarding the healthcare budget, Stuckler et al  
10 have proposed that the economic crisis presents the opportunity to reorganise the provision of  
11 services to those most in need<sup>39</sup>. Our findings identify one such group who may be in need.  
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21 ‘Maternity Matters’ guidance for service commissioning in England released in 2007  
22 emphasised the Government’s commitment to providing choice guarantees for women  
23 regarding type of antenatal care and place of birth and also highlighted the need for future  
24 maternity services to address disproportionately higher rates of maternal morbidity amongst  
25 disadvantaged women<sup>40</sup>. Although it described a comprehensive approach to delivering  
26 improved maternity care the guidance did not provide any specific recommendations as to how  
27 to improve services for women from socially disadvantaged backgrounds. Our study  
28 emphasises the need for developing such recommendations within the currently changing  
29 structure of health service commissioning in England.  
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#### 41 **Strengths and Limitations of the Study**

42 UKOSS, the first system of its kind in obstetrics worldwide, has provided an invaluable wealth of  
43 information on uncommon but serious disorders of pregnancy. The active, negative surveillance  
44 approach has ensured robust case ascertainment, and many of the UKOSS studies are the  
45 largest-scale studies of rare disorders of pregnancy worldwide.  
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53 By studying disorders of pregnancy and childbirth in separate studies, UKOSS provided the  
54 means by which to explore ‘severe maternal morbidity’ due to direct pregnancy causes as an  
55 aggregate of multiple conditions, thereby avoiding the challenges encountered by previous  
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3 studies which have used combined clinical and intervention criteria. It also avoids the perennial  
4 difficulty that there is no universally agreed definition of what constitutes 'severe maternal  
5 morbidity'. One disadvantage of this approach is that 'indirect causes' of maternal morbidity  
6 were excluded from our analysis, including several major contributors to maternal mortality in  
7 the UK such as cardiac disease, thromboembolic disease and puerperal psychosis/suicide<sup>41</sup>.  
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15 The design of UKOSS studies ensures that sampling bias is minimised by recruiting cases from  
16 all obstetric units across the UK, thus including women from all different ethnic, age and  
17 socioeconomic groups. Selection of controls is also aimed at attaining a representative sample  
18 of the general population of women giving birth by recruiting women from the same obstetric  
19 units as cases throughout the UK during the same time period. Comparison between UKOSS  
20 controls and national maternity data suggests that the controls recruited for UKOSS studies are  
21 closely representative of the general, child-bearing population<sup>19</sup>.  
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31 A universal challenge for studies that seek to explore the impact of social disadvantage in  
32 society is how to best capture the concept of 'social disadvantage' and define socioeconomic  
33 position. There are multiple dimensions to social disadvantage and there are many different  
34 approaches to defining a person's position in the social structure of society. The retrospective  
35 collection of data for UKOSS from patient records limited the definition of socioeconomic  
36 position to maternal occupation (or paternal where maternal was not available). Using the  
37 national classification of socioeconomic position (NS-SEC) ensured the study population was  
38 categorised using a system that is used nationally. However, the collection of other  
39 socioeconomic indicators such as level of education, years of schooling, residential location or  
40 income was not possible although this would have allowed further analysis of social  
41 disadvantage by using composite measures of disadvantage such as the Index of Multiple  
42 Deprivation. There is thus a place for further research using other indicators of socioeconomic  
43 position to see whether the results of this study are replicated.  
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3 The results of this study clearly apply only to high resource countries with well-developed  
4 healthcare systems. However, other high resource countries have different models of health  
5 care, such that pregnancy and delivery care may not be free at the point of access, as is the  
6 case in the UK. This study used only UK data and the results may therefore not be  
7 generalisable to women giving birth in other countries. Further research using information from  
8 other settings is therefore important to investigate whether socioeconomic position is associated  
9 with severe maternal morbidity in countries with differing health systems as this may give  
10 insights into the causes and where modifiable risks might lie. The use of data collected through  
11 the International Network of Obstetric Surveillance Systems<sup>42</sup>, which use similar definitions and  
12 methodologies, may be a route to achieve this.

## 23 24 25 **Conclusions**

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27 This study suggests that socioeconomic position may be associated with an increased risk of  
28 severe maternal morbidity, independent of risks associated with maternal age, ethnic group or  
29 known medical and pregnancy complications. although the association we observed was not  
30 statistically significant. However, the reasons for this possible association remain unclear,  
31 suggesting a need, not only to see if the observed association can be reproduced in other  
32 populations or settings, but also for further research to investigate whether there are differences  
33 in disease risks, lifestyle and health-seeking behaviour, access to care or the quality of care  
34 received by women from different socioeconomic groups. Further information about these  
35 differences would allow development of specific recommendations for the care of women from  
36 socially disadvantaged backgrounds in order to minimise future risks.

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49 **Competing interests:** None declared

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56 maternity care through national studies of “near-miss” maternal morbidity” programme  
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For peer review only

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**STROBE 2007 (v4) checklist of items to be included in reports of observational studies in epidemiology\***  
**Checklist for cohort, case-control, and cross-sectional studies (combined)**

Section/Topic	Item #	Recommendation	Reported on page #
Title and abstract	1	(a) Indicate the study’s design with a commonly used term in the title or the abstract	1
		(b) Provide in the abstract an informative and balanced summary of what was done and what was found	3
<b>Introduction</b>			<b>4</b>
Background/rationale	2	Explain the scientific background and rationale for the investigation being reported	4
Objectives	3	State specific objectives, including any pre-specified hypotheses	4
<b>Methods</b>			
Study design	4	Present key elements of study design early in the paper	3
Setting	5	Describe the setting, locations, and relevant dates, including periods of recruitment, exposure, follow-up, and data collection	6
Participants	6	(a) <i>Cohort study</i> —Give the eligibility criteria, and the sources and methods of selection of participants. Describe methods of follow-up <i>Case-control study</i> —Give the eligibility criteria, and the sources and methods of case ascertainment and control selection. Give the rationale for the choice of cases and controls <i>Cross-sectional study</i> —Give the eligibility criteria, and the sources and methods of selection of participants	5-6
		(b) <i>Cohort study</i> —For matched studies, give matching criteria and number of exposed and unexposed <i>Case-control study</i> —For matched studies, give matching criteria and the number of controls per case	
Variables	7	Clearly define all outcomes, exposures, predictors, potential confounders, and effect modifiers. Give diagnostic criteria, if applicable	6-7
Data sources/ measurement	8*	For each variable of interest, give sources of data and details of methods of assessment (measurement). Describe comparability of assessment methods if there is more than one group	6
Bias	9	Describe any efforts to address potential sources of bias	7
Study size	10	Explain how the study size was arrived at	6 & 8
Quantitative variables	11	Explain how quantitative variables were handled in the analyses. If applicable, describe which groupings were chosen and why	7-8
Statistical methods	12	(a) Describe all statistical methods, including those used to control for confounding	7-8
		(b) Describe any methods used to examine subgroups and interactions	8
		(c) Explain how missing data were addressed	7
		(d) <i>Cohort study</i> —If applicable, explain how loss to follow-up was addressed <i>Case-control study</i> —If applicable, explain how matching of cases and controls was addressed	

		<i>Cross-sectional study</i> —If applicable, describe analytical methods taking account of sampling strategy	
		(e) Describe any sensitivity analyses	11
<b>Results</b>			
Participants	13*	(a) Report numbers of individuals at each stage of study—eg numbers potentially eligible, examined for eligibility, confirmed eligible, included in the study, completing follow-up, and analysed	9
		(b) Give reasons for non-participation at each stage	5
		(c) Consider use of a flow diagram	
Descriptive data	14*	(a) Give characteristics of study participants (eg demographic, clinical, social) and information on exposures and potential confounders	9-10
		(b) Indicate number of participants with missing data for each variable of interest	9-10
		(c) <i>Cohort study</i> —Summarise follow-up time (eg, average and total amount)	
Outcome data	15*	<i>Cohort study</i> —Report numbers of outcome events or summary measures over time	
		<i>Case-control study</i> —Report numbers in each exposure category, or summary measures of exposure	9-10
		<i>Cross-sectional study</i> —Report numbers of outcome events or summary measures	
Main results	16	(a) Give unadjusted estimates and, if applicable, confounder-adjusted estimates and their precision (eg, 95% confidence interval). Make clear which confounders were adjusted for and why they were included	9-10
		(b) Report category boundaries when continuous variables were categorized	9-10
		(c) If relevant, consider translating estimates of relative risk into absolute risk for a meaningful time period	
Other analyses	17	Report other analyses done—eg analyses of subgroups and interactions, and sensitivity analyses	11
<b>Discussion</b>			
Key results	18	Summarise key results with reference to study objectives	11-12
Limitations	19	Discuss limitations of the study, taking into account sources of potential bias or imprecision. Discuss both direction and magnitude of any potential bias	13-14
Interpretation	20	Give a cautious overall interpretation of results considering objectives, limitations, multiplicity of analyses, results from similar studies, and other relevant evidence	12-13
Generalisability	21	Discuss the generalisability (external validity) of the study results	14-15
<b>Other information</b>			
Funding	22	Give the source of funding and the role of the funders for the present study and, if applicable, for the original study on which the present article is based	15

\*Give information separately for cases and controls in case-control studies and, if applicable, for exposed and unexposed groups in cohort and cross-sectional studies.

**Note:** An Explanation and Elaboration article discusses each checklist item and gives methodological background and published examples of transparent reporting. The STROBE checklist is best used in conjunction with this article (freely available on the Web sites of PLoS Medicine at <http://www.plosmedicine.org/>, Annals of Internal Medicine at <http://www.annals.org/>, and Epidemiology at <http://www.epidem.com/>). Information on the STROBE Initiative is available at [www.strobe-statement.org](http://www.strobe-statement.org).

## Correction

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Lindquist A, Knight M, Kurinczuk JJ. Variation in severe maternal morbidity according to socioeconomic position: a UK national case-control study. *BMJ Open* 2013;**3**:e002742. The order of author names is incorrect. It should be Lindquist A, Kurinczuk JJ, Knight M.

*BMJ Open* 2013;**3**:e002742corr1. doi:10.1136/bmjopen-2013-002742corr1