Impact of obesity, overweight and underweight on life expectancy and lifetime medical expenditures: the Ohsaki Cohort Study

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ABSTRACT
Objectives: People who are obese have higher demands for medical care than those of the normal weight people. However, in view of their shorter life expectancy, it is unclear whether obese people have higher lifetime medical expenditure. We examined the association between body mass index, life expectancy and lifetime medical expenditure.

Design: Prospective cohort study using individual data from the Ohsaki Cohort Study.

Setting: Miyagi Prefecture, northeastern Japan.

Participants: The 41,965 participants aged 40–79 years.

Primary and secondary outcome measures: The life expectancy and lifetime medical expenditure aged from 40 years.

Results: In spite of their shorter life expectancy, obese participants might require higher medical expenditure than normal weight participants. In men aged 40 years, multiadjusted life expectancy for those who were obese participants was 41.4 years (95% CI 38.28 to 44.70), which was 1.7 years non-significantly shorter than that for normal weight participants (p=0.3184). Multiadjusted lifetime medical expenditure for obese participants was £112,858.9 (94,954.1–131,840.9), being 14.7% non-significantly higher than that for normal weight participants (p=0.1141). In women aged 40 years, multiadjusted life expectancy for those who were obese participants was 49.2 years (46.14–52.59), which was 3.1 years non-significantly shorter than for normal weight participants (p=0.0724), and multiadjusted lifetime medical expenditure was £137,765.9 (123,672.9–152,970.2), being 21.6% significantly higher (p=0.0005).

Conclusions: According to the point estimate, lifetime medical expenditure might appear to be higher for obese participants, despite their short life expectancy. With weight control, more people would enjoy their longevity with lower demands for medical care.

INTRODUCTION
Obesity is closely associated with an increased risk of cardiovascular disease, cancer, hypertension, diabetes mellitus and other medical problems. Previous studies have reported that obese and overweight people have higher needs and demands for medical care than normal weight people.1–5 However, it is unclear whether obese people have higher lifetime medical expenditure than those of the normal weight people because the former have a comparatively shorter life expectancy and lifetime medical expenditure. The Ohsaki Cohort Study included a large number of participants aged 40 years, who were followed up for 40 years.2 We estimated life expectancy and lifetime medical expenditure adjusted for gender and age, and examined the relationship between body mass index (BMI) and them. We also compared multiadjusted lifetime medical expenditure between normal and obese participants in men aged 40 years. This study is the first study to have investigated the association between body mass index, life expectancy and lifetime medical expenditure calculated from individual medical expenditure and mortality data over a long period in a general population. There was a limit to the accurate estimation of lifetime medical expenditure, because the Japanese population has a low prevalence of body mass index $\geq 30.0 \text{ kg/m}^2$. However, it is unclear whether obese people have higher lifetime medical expenditure than those of the normal weight people because the former have a comparatively shorter life expectancy.
expectancy. Additionally, underweight people have a higher risk of mortality and thus also tend to have higher medical expenditure per month or per person, based on a 10-year follow-up.

Although four previous studies have examined the association between obesity and lifetime medical expenditure, the results were inconsistent. One study showed that obese people had lower lifetime medical expenditure than those of the normal weight people, whereas the others indicated that obese people had higher lifetime medical expenditure. In addition, two of the four studies calculated estimated lifetime medical expenditure from excess risk of cause-specific mortality and mean medical expenditure for the index disease. Only the other two studies calculated lifetime medical expenditure on the basis of individual medical expenditure and mortality. However, one of those studies followed up the participants for only 2 years and the other calculated lifetime medical expenditure for elderly participants aged 70 years or over. Therefore, the association between body mass index (BMI) and lifetime medical expenditure remains to be fully clarified.

We therefore conducted a 13-year prospective observation of 41,965 Japanese adults aged 40–79 years living in the community, which accrued 392,860 person-years. We examined the association between BMI and lifetime medical expenditure, based on individual medical expenditure and life table analysis. We collected data for survival and all medical care utilisation and costs, excluding home care services provided home health aides, nursing home care and preventive health services in participants of this cohort study.

**MATERIALS AND METHODS**

**Study cohort**

We used data from the Ohsaki National Health Insurance (NHI) Cohort Study. In brief, we sent a self-administered questionnaire on various lifestyle habits between October and December 1994 to all NHI beneficiaries living in the catchment area of Ohsaki Public Health Center, Miyagi Prefecture, northeastern Japan. A survey was conducted of NHI beneficiaries aged 40–79 years. Among 54,996 eligible individuals, 52,029 (95%) responded.

We excluded 776 participants who had withdrawn from the NHI before 1 January 1995, when we started the prospective collection of NHI claim files. Thus, 51,253 participants formed the study cohort. The study protocol was approved by the Ethics Committee of Tohoku University School of Medicine. The participants who had returned the self-administered questionnaires and had signed them were considered to have consented to participate in this study.

For the current analysis, we also excluded participants who did not provide information about body weight and height (n=3543), were at both extremes of the BMI range: lower than the 0.05th percentile for BMI (below 14.41 for men; below 13.67 for women) or higher than the 99.95th percentile for BMI (above 58.46 for men; above 62.00 for women; n=48), those who died within the first year (n=454) or those who had a history of cancer (n=1533), myocardial infarction (n=1233), stroke (n=831) or kidney disease (n=1646). Thus, a total of 41,965 participants (20,066 men and 21,899 women) participated.

**Body mass index**

The self-administered questionnaire included questions on weight and height, and BMI was calculated as weight divided by the square of height (kilograms per square metre). We divided the participants into groups according to the following BMI categories: <18.5 (underweight), 18.5–24.9 (normal weight), 25.0–29.9 (overweight) and ≥30.0 kg/m² (obesity). These BMI categories correspond to the cut-off points proposed by the WHO: normal BMI range (18.5–24.9 kg/m²), grade 1 overweight (25.0–29.9 kg/m²), grade 2 overweight (30.0–39.9 kg/m²) and grade 3 overweight (≥40.0 kg/m²). The validity of self-reported body weight and height has been reported earlier. Briefly, the weight and height of 14,883 participants, who were a subsample of the cohort, were measured during basic health examinations provided by local governments in 1995. The Pearson correlation coefficient (r) and weighted κ (k) between the self-reported values and measured values were r=0.96 (p<0.01) for weight, r=0.95 (p<0.01) for height and r=0.88 (p<0.01) and κ=0.72 for BMI categories.

**Health insurance system in Japan**

The details of the NHI system have been described previously. Briefly, everyone living in Japan is required to enrol in one health insurance system. The NHI covers 35% of the Japanese population for almost all medical treatment, including diagnostic tests, medication, surgery, supplies and materials, physicians and other personnel costs and most dental treatment. It also covers home care services provided by physicians and nurses but not those by other professionals such as home health aides. The NHI covers inpatient care but not nursing home care. Also, it does not cover preventive health services such as mass screening and health education. Payment to medical providers is made on a fee-for-service basis, where the price of each service is determined by a uniform national fee schedule.

If a participant withdrew from the NHI system because of death, emigration or employment, the withdrawal date and the reason for withdrawal were coded in the NHI withdrawal history files. We recorded any mortality or migration by reviewing the NHI withdrawal history files and collected data on the death of participants by reviewing the death certificates filed at Ohsaki Public Health Center. We then followed up the participants and prospectively collected data on medical care utilisation and its costs for all participants in the cohort from 1 January 1995 through 31 December 2007.
Statistical analysis
We conducted the same analysis as the previous study about the association between walking, life expectancy and lifetime medical expenditure. Briefly, we divided the age groups (x) from 40 years according to the following categories: 40–44, 45–49, 50–54, 55–59, 60, 64, 65–69, 70–74, 75–79, 80–84 and ≥85 years. Based on person-years and the number of deaths from 1996 until 2007, the multiajusted mortality rates for each age category were estimated from a Poisson regression model. The dependent variable was mortality, and independent variables were age groups, categories of BMI and the following covariates: smoking status (current and past smoker or never smoker), alcohol consumption (current drinker consuming 1–499 g/week, current drinker consuming ≥500 g/week or never and past drinker), sports and physical exercise (≥3 h/week or <3 h/week), time spent walking (≥1 h/week or <1 h/week) and education (junior high school, high school or college/university or higher). We did not adjust for hypertension and diabetes mellitus in the multivariate models because these variables are considered to occupy an intermediate position in the etiologic pathway between BMI and mortality.

We separately calculated medical expenditure for participants who survived through the index year and for those who died because previous study showed that medical expenditure increased before death. The multiajusted medical expenditure per year was estimated using a linear regression model adjusted for the above covariates in survivors and decedents.

The estimates of multiajusted mortality and medical expenditure were used for estimating life expectancy and lifetime medical expenditure from 40 years of age. To estimate life expectancy and lifetime medical expenditure, we constructed life tables per 100 000 persons using Chiang's analytical method on the basis of the latest published complete life tables of Japan for the year 2000. Then, life expectancy (\(e_x\)) and lifetime medical expenditure (\(M_x\)) for each age groups (x) were estimated using the numbers of survivors (\(L_x\)), deaths (\(d_x\)), static population (\(L_x\)), multiajusted medical expenditure for survivors (\(a_x\)) and multiajusted medical expenditure for the deceased (\(b_x\)) as follows:

\[
\sum_i \text{is sum of } y = x
\]

\[
e_x = \frac{\sum L_x}{L_x}
\]

\[
M_x = \frac{\sum (L_x \cdot a_x + d_x \cdot b_x)}{L_x}
\]

The 95% CIs were estimated using a Monte Carlo simulation based on a Poisson regression model and linear regression model. We repeated 100 000 times, and all analysis were used the SAS V.9.1 statistical software package (SAS Institute Inc., 2004). All p values <0.05 were accepted as statistically significant.

We used a purchasing power parity rate of £1 = JPN £140.16

RESULTS
After 13 years of follow-up, we observed 5159 deaths (3356 men and 1803 women) among the 41 965 participants (20 066 men and 21 899 women).

The mean medical expenditure per year for survivors in men was £2393 in underweight, £2055 in normal weight, £2231 in overweight and £2334 in obesity, respectively. In women, it was £2375 in underweight, £1972 in normal weight, £2317 in overweight and £2733 in obesity, respectively. These differences of mean medical expenditure per year for survivors are statistically significant in men and women (ANOVA; p<0.0001). Also, the mean medical expenditure in the year of death for participants in men was £15445 in underweight, £16973 in normal weight, £17811 in overweight and £17878 in obesity, respectively. In women, it was £12833 in underweight, £15584 in normal weight, £17059 in overweight and £19635 in obesity, respectively. These differences of mean medical expenditure in the year of death for participants are statistically significant in only women (men, p=0.2241; women, p=0.0059).

Baseline characteristics by BMI category
The baseline characteristics of the study participants according to the BMI categories are shown for men and women (table 1), among whom 3.3% and 3.9% were underweight, 23.6% and 28.4% were overweight and 73.1% and 68.7% were normal weight, respectively. These characteristics showed statistically significant differences.

Mortality in terms of categories for BMI
Figure 1A for men and figure 1B for women show the mortality (per 1000 person-years) in each of the age groups according to the categories of BMI. In underweight participants, there was a tendency that the mortality was the highest in each age group.
Overweight participants showed similar mortality with normal weight participants, especially women. Overweight men showed slightly lower mortality than normal weight men. In obese participants, the mortality curve was not described smoothly because of small number of participants.

Table 2 shows the mortality ratio with 95% CIs according to the categories of BMI. In underweight participants, the multiadjusted mortality ratio was significantly higher than that in the normal weight participants (men, 1.62, 95% CI 1.41 to 1.86, p < 0.0001; women, 1.46, 1.22 to 1.76, p < 0.0001). In overweight participants, the multiadjusted mortality ratio was significantly lower in men and non-significantly lower in women than that in normal weight participants (men, 0.91, 0.83 to 0.99, p = 0.0260; women, 0.98, 0.88 to 1.10, p = 0.7841). In obese participants, the multiadjusted mortality ratio was non-significantly higher than that in normal weight participants (men, 1.14, 0.88 to 1.49, p = 0.3177; women, 1.23, 0.98 to 1.55, p = 0.0717).

Life expectancy and lifetime medical expenditure by BMI category

Table 3 shows life expectancy and lifetime medical expenditure with 95% CIs according to the BMI categories. By multiadjusted analysis, obese men and women had approximately 1.7 and 3.1 years non-significantly shorter life expectancy from the age of 40 years in comparison with men and women of normal weight, respectively (men, p = 0.3184; women, p = 0.0724). Meanwhile, obese men and women had approximately 14.7% non-significantly higher and 21.6% significantly higher lifetime medical expenditure in comparison with normal weight participants, respectively (men, p = 0.1141; women, p = 0.0005).

In men, multiadjusted life expectancy was greatest for overweight, that is, 44.34 years (95% CI 43.11 to 45.54, p = 0.0264), followed by normal weight (43.03 years, 42.22 to 43.73) and obesity (41.36 years, 38.28 to 44.70, p = 0.3184) and was shortest for underweight (37.40 years, 35.80 to 38.87, p < 0.0001). The multiadjusted lifetime medical expenditure for overweight was the highest, that is, £114 766.9 (95% CI 107 754.1 to 121 966.6, p < 0.0001), followed by obesity (£112 858.9, 94 954.1 to 131 840.9, p = 0.1141) and normal weight (£98 355.0, 93 615.3 to 103 010.2) and was the lowest for underweight (£93 208.7, 81 704.9 to 104 706.4, p = 0.3916).

In women, multiadjusted life expectancy was greatest for overweight, that is, 52.56 years (50.67 to 54.46, p = 0.7797), followed by normal weight (52.31 years,
The present results indicate that (1) obese men and women have 14.7% non-significantly higher and 21.6% significantly higher multiadjusted lifetime medical expenditure than those of the normal weight participants (men, \( p=0.1141 \); women, \( p=0.0005 \)), even though their life expectancy is non-significantly shorter by 1.7 and 3.1 years than those of the normal weight participants, respectively (men, \( p=0.3184 \); women, \( p=0.0724 \)); (2) underweight men and women have 5.2% and 3.4% non-significantly lower lifetime medical expenditure than those of the normal weight participants (men, \( p=0.5174 \); women, \( p=0.3916 \)) because men and women live 5.6 and 5.3 years significantly less than those of the normal weight participants, respectively (men, \( p<0.0001 \); women, \( p<0.0001 \)).

Comparison with other studies

Obese participants had shorter life expectancy than normal weight participants, as has been observed in previous studies.\(^6\)\(^-\)\(^10\) Overweight participants had longer life expectancy than normal weight participants. Two of the four previous studies have reported that overweight participants had longer life expectancy than normal weight participants.\(^7\)\(^-\)\(^9\) These results support our finding of an association between being overweight and life expectancy. Additionally, an association between BMI and all-cause mortality in the Japanese population has been reported by other data sets.\(^23\)\(^-\)\(^29\) All seven previous studies showed that among the BMI categories, the lowest one had the highest mortality risk. These results are consistent with the fact that underweight participants have significantly the shortest life expectancy, as was observed in our study.

Thus, the association between BMI and life expectancy showed same trend with the pooled analyses of the association between BMI and all-cause mortality in Asia and Japan.\(^30\)\(^-\)\(^31\)

Our present results support three of the four previous studies of lifetime medical expenditure for obese

### Table 2: Mortality ratio for BMI categories in 41,965 participants

<table>
<thead>
<tr>
<th>BMI (kg/m(^2))</th>
<th>Univariate</th>
<th>Multiadjusted*</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Mortality ratio (95% CI)</td>
<td>p Value</td>
</tr>
<tr>
<td>Men</td>
<td></td>
<td></td>
</tr>
<tr>
<td>&lt;18.5</td>
<td>1.69 (1.47 to 1.93)</td>
<td>&lt;0.0001</td>
</tr>
<tr>
<td>18.5–24.9</td>
<td>1.00 (Reference)</td>
<td></td>
</tr>
<tr>
<td>25.0–29.9</td>
<td>0.90 (0.82 to 0.98)</td>
<td>0.0163</td>
</tr>
<tr>
<td>≥30.0</td>
<td>1.13 (0.87 to 1.47)</td>
<td>0.3712</td>
</tr>
<tr>
<td>Women</td>
<td></td>
<td></td>
</tr>
<tr>
<td>&lt;18.5</td>
<td>1.50 (1.25 to 1.81)</td>
<td>&lt;0.0001</td>
</tr>
<tr>
<td>18.5–24.9</td>
<td>1.00 (Reference)</td>
<td></td>
</tr>
<tr>
<td>25.0–29.9</td>
<td>1.00 (0.89 to 1.11)</td>
<td>0.9613</td>
</tr>
<tr>
<td>≥30.0</td>
<td>1.29 (1.03 to 1.62)</td>
<td>0.0273</td>
</tr>
</tbody>
</table>

*Adjusted for age groups, smoking status, alcohol drinking, sports and physical exercise, time spent walking and education.
BMI, body mass index.

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**DISCUSSION**

Figure 1 Multiadjusted mortality by BMI categories in each age group in men (A) and women (B).

50.79 to 53.75) and obesity (49.23 years, 46.14 to 52.59, \( p=0.0724 \)) and was shortest for underweight (46.98 years, 44.63 to 49.29, \( p<0.0001 \)). The lifetime medical expenditure for obesity was the highest (\£137,765.9, 123,672.9 to 152,970.2, \( p=0.0005 \)), followed by overweight (\£129,964.6, 121,845.4 to 138,577.2, \( p<0.0001 \)) and normal weight (\£113,282.9, 106,668.0 to 120,054.6) and was lowest for underweight (\£109,382.2, 97,996.6 to 121,008.6, \( p=0.5174 \)).

participants.\textsuperscript{10,12,13} In comparison to previous studies, we calculated lifetime medical expenditure from individual medical expenditure and survival data covering longest follow-up period to date. Meanwhile, one study has shown that obese participants have lower lifetime medical expenditure than normal weight participants.\textsuperscript{11} However, that study limited the participants to medical expenditure than normal weight participants.\textsuperscript{10 12 13} In comparison to previous studies, overweight participants were found to have higher lifetime medical expenditure than normal weight participants, as had been reported previously.\textsuperscript{10,12,13} We consider that this was attributable to the higher medical expenditure per month or per person from the 10-year or 9-year follow-up than for normal weight participants.\textsuperscript{1} \textsuperscript{3} \textsuperscript{4} On the other hand, with regard to underweight participants, our present findings were inconsistent with those of a previous study that examined the association between being underweight and lifetime medical expenditure.\textsuperscript{13} However, that study calculated lifetime medical expenditure for elderly participants aged over 70 years. Elderly underweight participants have high mortality,\textsuperscript{32} and medical expenditure increases in the 1 year prior to death.\textsuperscript{20} Thus, lifetime medical expenditure from 70 years for underweight participants becomes higher than for participants of normal weight. Our study results are thus inconsistent with those reported previously.

We previously calculated life expectancy and lifetime medical expenditure for smokers and non-smokers from age 40 years by using the same data set as that for the present study.\textsuperscript{17} The results indicated that lifetime medical expenditure was non-significantly lower in smokers than in non-smokers, reflecting the 3.5 years shorter life expectancy of smokers. On the other hand, the present study indicated that lifetime medical expenditure was higher for obese participants in spite of their shorter life expectancy. This difference would result from the difference in which obesity and smoking affect one’s health and longevity. Previous studies of healthy and disability free life expectancy have agreed that smoking shortens life expectancy without affecting the years of life spent with ill-health or disability, while obesity shortens life expectancy and extends the years of life spent with ill-health or disability.\textsuperscript{33} On the basis of these differences, Reuser et al summarised the situation as ‘smoking kills and obesity disables’.\textsuperscript{7} Extended years with ill-health and/or disability must result in increased lifetime medical expenditure. All these findings suggest that weight control would bring about longer life expectancy and long-term enhancement of the quality of life and a cost saving.

**Strengths and limitations**

A major strength of our present study is that it is the first in the world to have clarified the association between BMI and lifetime medical expenditure calculated from individual medical expenditure and mortality data over

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**Table 3** Life expectancy and lifetime medical expenditure at age 40 years for BMI categories in 41,965 participants

<table>
<thead>
<tr>
<th>BMI (kg/m(^2))</th>
<th>Univariate</th>
<th>Multiadjusted*</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Estimate</td>
<td>95% CI</td>
</tr>
<tr>
<td>Men</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Life expectancy at age 40 years (years)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>&lt;18.5</td>
<td>36.72</td>
<td>35.10 to 38.17</td>
</tr>
<tr>
<td>18.5–24.9</td>
<td>42.70</td>
<td>41.91 to 43.37</td>
</tr>
<tr>
<td>25.0–29.9</td>
<td>44.09</td>
<td>42.89 to 45.25</td>
</tr>
<tr>
<td>≥30.0</td>
<td>41.23</td>
<td>38.16 to 44.54</td>
</tr>
<tr>
<td>Lifetime medical expenditure at age 40 years (£)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>&lt;18.5</td>
<td>94,877.5</td>
<td>83,411.4 to 106,275.7</td>
</tr>
<tr>
<td>18.5–24.9</td>
<td>97,244.1</td>
<td>92,662.5 to 101,774.0</td>
</tr>
<tr>
<td>25.0–29.9</td>
<td>114,398.2</td>
<td>107,490.1 to 121,505.3</td>
</tr>
<tr>
<td>≥30.0</td>
<td>115,362.6</td>
<td>97,361.8 to 134,555.0</td>
</tr>
<tr>
<td>Women</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Life expectancy at age 40 years (years)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>&lt;18.5</td>
<td>46.26</td>
<td>43.98 to 48.43</td>
</tr>
<tr>
<td>18.5–24.9</td>
<td>51.70</td>
<td>50.28 to 53.02</td>
</tr>
<tr>
<td>25.0–29.9</td>
<td>51.74</td>
<td>49.98 to 53.48</td>
</tr>
<tr>
<td>≥30.0</td>
<td>48.13</td>
<td>45.23 to 51.22</td>
</tr>
<tr>
<td>Lifetime medical expenditure at age 40 years (£)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>&lt;18.5</td>
<td>108,278.3</td>
<td>97,142.8 to 119,593.7</td>
</tr>
<tr>
<td>18.5–24.9</td>
<td>111,512.8</td>
<td>105,303.4 to 117,910.4</td>
</tr>
<tr>
<td>25.0–29.9</td>
<td>127,869.3</td>
<td>120,236.3 to 135,932.3</td>
</tr>
<tr>
<td>≥30.0</td>
<td>134,887.1</td>
<td>121,318.4 to 149,838.6</td>
</tr>
</tbody>
</table>

*Adjusted for age groups, smoking status, alcohol drinking, sports and physical exercise, time spent walking and education. BMI, body mass index.
a long period in a general population from the age of 40 years.1 14–18 The NHI covers almost all medical care utilisation.1 14 16 18 Additionally, in order to reduce bias, we adjusted confounders by including various covariates in our Poisson regression model and linear regression mode.16 On the other hand, several limitations of our study should also be considered. First, we used self-reported BMI which in a source of error.34 35 We consider this error to be a non-differential classification. This misclassification would lead to attenuation of the true association towards the null. To address this problem, van Dam et al16 studied the association between BMI and mortality using lower BMI cut-off points: 24.5 kg/m2 to reflect a measured BMI of 25.0 kg/m2, and 29.0 kg/m2 to reflect a measured BMI of 30.0 kg/m2. The association showed similar with original cut-off points. Second, the 95% CI was wide, and there was a limit to the accurate estimation of life expectancy and lifetime medical expenditure for obese participants. Additionally, we did not observe significant association in obese participants without lifetime medical expenditure in women. However, our results are consistent with those of the previous studies.6–8 10–12 13 In Japan, prevalence of obesity is only 3%.37 Thus, the reason for non-significant association might be β error because of the lack of statistical power due to small number of obese participants.

Conclusions and policy implication

In summary, even though we observed non-significant association between obesity, life expectancy and lifetime medical expenditure without lifetime medical expenditure in women, lifetime medical expenditure might appear to be higher for obese participants, despite their short life expectancy. With better weight control, more people would enjoy their longevity with lower needs and demands for medical care.

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Contributors All authors contributed to the design of the study. MN, SK, MK, KO-M, TS and IT participated in data collection. MN, SK, AH, MK and SH participated in data analysis. MN, MK, KO-M, TS, AH, MK and SH participated in the writing of the report. SK and IT participated in critical revision of the manuscript. All authors approved the final version of the report for submission.

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Competing interests None.

Ethics approval The study protocol was approved by the Ethics Committee of Tohoku University School of Medicine. Participants who had returned the self-administered questionnaires and signed them were considered to have consented to participate.

Provenance and peer review Not commissioned; externally peer reviewed.

Data sharing statement No additional data available.

REFERENCES


