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Developing and validating a risk score for prediction of preterm birth at Felege Hiwot Comprehensive Specialized Hospital, Northwest, Ethiopia: Retrospective follow up study, 2021

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Developing and validating a risk score for prediction of preterm birth at Felege Hiwot Comprehensive Specialized Hospital, Northwest Ethiopia: Retrospective follow- up study Sefineh F.Feleke¹, Zelalem A.Anteneh², Gizachew T.Wassie², Anteneh K.Yalew³, Anteneh M.Dessie⁴

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Objective: To develop and validate a risk score for the prediction of preterm birth using maternal characteristics.

Method: A retrospective follow-up study was conducted on March (1- 30) 2021 at Felege Hiwot comprehensive specialized hospital. The sample size was determined by assuming 10 events per predictor, based on this assumption total sample size was 1308. Data were collected using a structured checklist through chart review. Data were coded and entered into Epidata, version 3.02, and was analyzed by using R statistical programming language version 4.0.4 for further processing and analysis. Bivariable logistic regression was done to identify the relationship between each predictor and preterm birth. Variables with ($p \le 0.25$) from the bivariable analysis were entered into a backward stepwise multivariable logistic regression model, and significant variables (p < 0.05) were retained in the multivariable model. Model accuracy and goodness of fit were assessed by computing the area under the ROC curve (discrimination) and calibration plot (calibration) respectively

Results: Residence, gravidity, hemoglobin < 11 mg/dl, early rupture of membranes, antepartum hemorrhage, and pregnancy-induced hypertension remained in the final multivariable prediction model. The AUC of the model was 0.816 (95% confidence interval: 0.779 - 0.856).

Conclusion: These results suggest the possibility of predicting preterm birth using a simple prediction model constructed from maternal characteristics.

Strength and Limitations of the study

- ✓ Adequate number of participants with the outcome, which helped us to construct the model using a sufficient number of predictor variables.
- Prediction model is constructed from easily obtainable maternal characteristics that make it applicable in primary care settings.
- ✓ A single-site study, it is confined to a single area, which needs external validation before using it in another context.
- ✓ Furthermore, data were collected from each mother's card; due to this, some important variables were missed, such as previously highlighted factors with preterm birth in different studies

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Background

Preterm birth is described as babies that are born alive before the end of 37 weeks of pregnancy[1]. Preterm birth can be accidental (due to spontaneous preterm labor and/or preterm membrane rupture) or induced by the provider (by cesarean or labor induction)[2]. Most preterm births happen spontaneously[3].

An estimated 15 million babies worldwide are born too early per year. That's more than 1 in 10 infants. About 1 million newborns die per year because of preterm birth complications[4].

Across 184 countries, the rate of preterm birth ranges from 5% to 18% of babies born [5]. However, there are stark disparities in survival rates around the world. Half of the babies born at or below 32 weeks die in low-income settings due to a lack of practical, cost-effective, and critical care, such as comfort, breastfeeding assistance, basic infection care, and trouble Breathing[6].

In Ethiopia, every year, 320,000 babies are born too early and because of direct preterm complications, 24,400 children under five die [7]. According to the 2019 Mini Ethiopia Demographic and Health Survey, the neonatal mortality rate was 30 deaths per 1,000 live births and prematurity was the major cause of death[8]

Furthermore, the effect of preterm birth is also prolonged beyond the neonatal phase and throughout life[9]. Hence, the largest risk of severe health issues, including cerebral palsy, intellectual disability, chronic lung disease, and vision and hearing loss, is faced by babies born before maturity. This introduces a lifelong disability dimension. At some point in their lives, most people will face the struggles and potential disasters of preterm birth either directly in their families or indirectly through events for the nations[9, 10].

To alleviate this burden in the past few decades, numerous methods have been attempted internationally, including in Ethiopia, to prevent and enhance the treatment of preterm births [11-13]. Although several efforts were undertaken to prevent and reduce preterm birth, its rate appears to have increased over time [10, 14]. As part of the strategy, it is essential to diagnose or predict preterm birth earlier in pregnancy to take appropriate measures for high-risk groups.

There are clinical prediction models that attempt to predict the probability of preterm birth, however, all include laboratory tests that are generally not accessible in low-resource settings, like fetal fibronectin, insulin-like growth factor binding protein-1 (IGFBP-1), interleukin-6, and placental alpha-macroglobulin-1 to predict preterm birth[15-20].

3 | Page

Hence, because of limited resources, the use of easily accessible data to forecast preterm birth seems to be appealing in low- and middle-income areas. Although there are prediction models for preterm birth, variation in the occurrence of preterm birth globally is relevant, indicating variations in exposure to psychosocial, sociodemographic, and medical risk factors and genetic differences [21-23].

There is no prediction model for preterm birth in Ethiopia. Therefore, developing and validating a risk score for prediction of preterm birth using maternal(clinical and non-clinical) characteristics based on the available measurement is paramount to allow early preterm birth intervention such as utero transfer to tertiary care centers, appropriate corticosteroid administration while preventing excessive use, neuroprotective magnesium sulfate therapy, and antibiotic treatment in the event of infection[15, 24]

Methods and Materials Study setting

This retrospective study was conducted among 1260 pregnant women who did prenatal care and finally delivered at Felege Hiwot Comprehensive Specialized Hospital, Bahir Dar city, Northwest Ethiopia from January 30, 2019, to January 30, 2021. Bahir Dar is the capital city of Amhara national regional state and is found 575 km northwest of Addis Ababa.

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Felege Hiwot comprehensive specialized hospital was established with the German State government during the regime of Emperor H/ Selassie I in April 1963 G.C and is one of the oldest public hospitals in the Northwestern part of the country and located at the northern end of the city near Lake Tana and aspires to see a healthy, productive and prosperous society and become a center of medical service Excellency by 2029. During its establishment, it was planned to serve 25,000 people. The hospital has currently a total of 1431 manpower (5 obstetricians and gynecologists and 63 midwives among others) in different disciplines. It has a total of 500 formal beds, 11 wards (emergency ward and Inpatient wards such as Gynecological &Obstetric, Surgical, orthopedics, Medical, Pediatric, L&D, Eye unit, NICU, psychiatric, oncology, and 22 OPDS), 39 clinical and non-clinical departments /service units / providing laboratory, Diagnostic, curative & Rehabilitation service at outpatient & inpatient bases as well as disease prevention & health promotion services.

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The sample size required for model development was determined based on the minimum standard of 10 events per candidate predictor considered, according to the formula $N = (n \times 10)/I$ where N is the sample size, n is the number of candidate predictor variables and I is the estimated event rate in the population[25]. Since there were 17 candidate predictors considered and 10 events per candidate predictor, the estimated number of events for the study was 170. Based on a study done on the prevalence of preterm birth in Debre Tabor hospital was 13%[26], so taking into account this the required sample size was calculated as follows, n= 170*100/13= 1308.

Patient and public involvement

There was no direct interaction with patients in this study and no direct patient involvement in the design or conduct of this study.

Study Design and Participants

The theoretical design of the present study was; the incidence of preterm birth as a function of multiple predictors during pregnancy. The source population of the study was all pregnant mothers who gave birth at FHCSH. To be included in this study, mothers must meet all of the following eligibility criteria; All medical records of mothers who gave birth and had at least one ANC follow-up in FHCSH from January 30/2019 to January 30/2021.

Sampling method and procedures

A simple random sampling technique was employed to select participants using the medical registration number of a delivered mother from the delivery registration book. First, all mother delivered at FHCSH from January 30/2019 to January 30/20201 was identified from the delivery registration book. After that records of mothers who meet the inclusion criteria were included in the study. Subsequently, a sampling frame was prepared. Finally, the study unit was selected by using a computer-generated random number.

Data Collection

Outcome assessment: The outcome variable was attributed to women whose medical records indicated a physician or midwife diagnosis of preterm birth and delivery between 28 and 36 completed weeks of gestation. The gestational age (GA) was measured using either LNMP, which is found to be a more reliable measure of GA in a low-resource setting[27, 28], or an early ultrasound result.

Predictor assessment: Data was collected using a structured checklist through chart review. Checklists were developed after reviewing various relevant literature [29-33]. It consists of socio-demographic (Maternal age, Residence), Maternal obstetric characteristics : (History of preterm birth, History of abortion, history of stillbirth gravidity, Parity, Multiple pregnancy, APH, PROM, Gestational DM, and PIH), Maternal medical condition : (HGB level, Diabetic Mellitus, Chronic Hypertension, UTI and HIV).

Quality Assurance Mechanisms

To maintain the quality of data, the data collectors and supervisors were trained for a day on the objective of the study, the content of the checklists, how to fill the checklists. Afterward, reviewing 15 charts on medical records of mothers who gave birth at Felege Hiwot Comprehensive Specialized Hospital which is found in Northwest Ethiopia were done. After that, some adjustments were done accordingly. The checklist was developed in English.

Data Processing and Analysis

Data were entered into a software application (EPI DATA, version 3.02) and was analyzed by using R statistical programming language version 4.0.4 for further processing and analysis. There were 13(1%), 2(0.2%), 11(0.9%), 15(2.5%), 21(1.7%), 29(2.3%), 20(1.6%) and 20(1.6%)missing values for premature rupture of membranes, residence, chronic hypertension, multiple pregnancy gestational diabetes Mellitus, pregnancy-induced hypertension antepartum hemorrhage and hemoglobin respectively. We assumed data were missing at random, and we, therefore, performed a multivariate imputation by chained equations for all variables evaluated in the prediction model [34]. Sensitivity analysis was performed to assess whether the assumption of missing at random (MAR) is valid or not, and the results were reasonably comparable (Table1). Descriptive statistics including median, inter-quartile range (IQR), and percentages, were carried out.

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Table 1. Sensitivity analysis of the model to predict preterm birth: Comparison of the regression coefficients, standard errors (SE), and p-values for complete case analysis (CCA) and multiple imputed data (MI).

Predicator variables	Complete case analysis			Multiple imputations		
	В	SE	P value	В	SE	P value
Chronic hypertension	0.7313	0.6297	0.24	0.581	0.6285	0.92

6 | Page

(yes)						
Residence (rural)	0.815	0.1946	< 0.001	1.154	0.1958	< 0.001
GDM(yes)	0.709	0.4028	0.07	0.472	0.4236	0.26
HGB(<11g/dl)	0.497	0.2185	0.02	0.642	0.2153	0.001
PROM (yes)	1.898	0.2080	< 0.001	2.097	0.2129	< 0.001
APH (yes)	1.194	0.2858	< 0.001	1.298	0.2874	< 0.001
PIH (yes)	1.353	0.2600	< 0.001	1.368	0.2523	< 0.001
Multiple pregnancy (yes)	0.539	0.3173	0.08	0.446	0.3257	0.17
Gravidity(primigravida)	0.426	0.1944	0.02	0.711	0.1976	< 0.001

Model Development and Validation

For model development, bivariable logistic regression was done to obtain insight into the association of each potential predictor and preterm birth. Variables with (p < 0.25) from the bivariable analysis were entered into a backward stepwise multivariable logistic regression model, and significant variables (p < 0.05) were retained in the multivariable model. The results of significant predictors were reported as coefficients with 95% confidence intervals (CI). To check for the model accuracy and goodness of fit, we computed the area under the ROC curve (discrimination) and calibration plot (calibration) using "classifierplots" and "givitiR" packages of R respectively. The AUC ranged from 0.5 (no predictive ability) to 1 (perfect discrimination)[35]. The regression coefficients and their 95% confidence levels, and the AUC were adjusted for overfitting or over-optimism using bootstrapping technique. To make internal validation, we computed 1000 random bootstrap [36]samples with the replacement on all predictors in the data. The model's predictive performance after bootstrapping is considered as the performance that can be expected when the model is applied to future similar populations. To evaluate the clinical and public health impact of the model, we performed a decision curve analysis (DCA) [37] of standardized net benefit across a range of threshold probabilities (0 to 1). In the DCA, the model was compared against two extreme scenarios; "intervention for all" and "no intervention". In our case, the intervention considered is the referral of high-risk pregnant women to facilities where appropriate corticosteroid administration, antibiotic treatment.

Risk Score Development

To construct an easily applicable preterm birth prediction score, we transformed each coefficient from the model to a rounded number by dividing it by the lowest coefficient. The number of points was subsequently rounded to the nearest integer. We determined the total score for each individual by assigning the points for each variable present and adding them up. The score was transformed to a dichotomous, allowing each pregnant woman to be classified as having a high or low risk of preterm birth. The receiver operating characteristic curve (ROC) was plotted and the area under the curve (AUC) was calculated to measure the discriminatory power of the scoring system.

Ethical Approval

Ethical clearance was obtained from the Institutional Review Boards (IRB) of Bahir Dar University, College of Medicine and Health Sciences with Protocol number 083/ 2021) on February 26, 2021. Confidentiality was maintained by omitting the personal identifier of the participant during the data collection procedure and information was used only for research purposes. Data were collected from the register, which was kept in a secure place and all data were fully anonymized before we access them. After the collection of data, all the patient records and patients' cards were placed back into a secure place. Data were entered into a password-protected computer.

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Result

Demographic, Obstetric, and Clinical Characteristics of mothers who gave birth at Felege Hiwot Comprehensive Specialized Hospital.

A total of 1260 study cards were reviewed from a sample of 1308, about 48 cards were not reviewed due to the outcome of intrauterine fetal death, abortion. *Table (2)* shows the demographic, obstetric, and clinical characteristics of mothers who gave birth included in the analysis. The median age of the study participants was 26 years with IQR (24-30years); the majority of the participants 1086 (86.2%) were in the age group of 20-34 years.

More than three fourth of the participants 926 (73.49%) were urban residents. From the total of mothers who delivered at FHCSH, more than two-third 841 (66.7%) were multigravida. About parity, above half of them713 (56.6%) were multipara. Concerning past obstetric history, 55 (6.5%) of them had a history of previous preterm birth, 76 (9%) of them had a previous history of stillbirth and 162 (19.3%) of them had a previous history of abortion.

Characteristics	Category	Frequency	Percent
Gravidity	Primigravida	419	33.3
	Multigravida	841	66.7
Residence	Urban	926	73.5
	Rural	334	26.5
GDM	Yes	44	3.5
	No	1216	96.5
АРН	Yes	84	6.7
	No	1176	93.3
PIH	Yes	110	8.73
	No	1150	91.27
HGB level	<11d/d1	236	18.7
	>=11g/dl	1024	81.3
Chronic hypertension	Yes	21	1.7
	No	1239	98.3

Table 2. Demographic, obstetric, and clinical characteristics of mothers who gave birth at FHCSH from January 30/2019 to January 30/2021, in Northwest Ethiopia.

9 | Page

PROM	Yes	195	15.5	
	No	1065	84.5	
Multiple pregnancies	Yes	90	7.2	
	No	1170	92.8	

PROM: Premature rupture of membrane, HGB: hemoglobin, PIH: pregnancy-induced hypertension, APH: antepartum hemorrhage, GDM: gestational diabetes mellitus

Development of prediction model for preterm birth

Out of 1260 delivered neonates, 169 (13.4%) (95%, CI (11.6%, 15.4%) was preterm infants.

The bivariable logistic regression analysis found several factors were eligible to be included in the prediction model. Variables with $P \le 0.25$ in the bivariable logistic regression analysis were hemoglobin level, Gravidity, residence, gestational diabetes mellitus, APH, PIH, chronic hypertension, PROM, and multiple pregnancies. Using the results, a prediction model was developed an equation for the prediction model was obtained. *(Table 3)*

Table 3: Coefficients and risk-scores of each predictor included in the model to predictpreterm birth (n = 1260)

Predictors		Multiva	riable analysis			
Variables*	Original β	0.	Р-	Risk	Risk	
	(95 % CI)	Bootstrap β	value	score		
Residence		0				
(rural)	1.161 (0.780, 1.545)	1.148	< 0.001	2		
Gravidity	0.675 (0.291, 1.061)	0.666	0.01	1		
(primigravida)						
PROM (yes)	2.081 (1.669, 2.50)	2.051	< 0.001	3		
APH (yes)	1.364 (0.806, 1.915)	1.348	< 0.001	2		
PIH (yes)	1.387 (0.887, 1.879)	1.368	< 0.001	2		
HGB <11g/dl	0.676 (0.255, 1.09)	0.677	< 0.001	1		

*Variables retained in the reduced model are; residence, APH, hemoglobin, PIH, gravidity, and PROM. Both backward and forward selection showed the same results. β after internal validation with bootstrapping bootstrapping is shown. Simplified risk score: we divided the coefficient of predictors

10 | Page

included in the reduced model by the smallest (0.666). The probability or risk of preterm birth = 1/(1 + exp - (-3.517 + 1.148 * Residence (rural) + 0.666 * gravidity (primigravida) + 2.051*PROM (yes) + 1.348 * APH (yes) + 1.387*PIH + 0.677*HGB (<11g/dl)..

The AUC of the final reduced model was 0.816 (95% confidence interval: 0.779 - 0.856) (Figure 1a). The calibration test had a p-value of 0.6228, indicating that the model does not misrepresent the data or calibration of the model was visually accurate since observed and predicted probabilities were similar (Figure 1b).

Validation of the model with the bootstrap technique showed hardly any indication of undue influence by particular observations, with an optimism coefficient of 0.085, resulting AUC of 0.789 (corrected 95% CI: 0.748–0.83).

Using the coefficients (β) the predicted risk cutoff point was a probability of (SpEqualSe > 0.1320), the model has a sensitivity of 75.74%, specificity of 72.87%, a positive predictive value of 30.2%, and a negative predictive value of 95.1%.

When applying DCA, we first evaluate whether our model understudy has a higher net benefit than the default strategies (referring all and none). This model outperforms the default strategies across the relevant threshold range. The model has the highest net benefit across the entire range of threshold probabilities, which indicates that the model has the highest clinical and public health value. Hence, referral decision made using the model has a higher net benefit than not referring all or referring all regardless of their risk thresholds as shown in *figure (2)*

Risk Classification Using a Simplified Risk Score

We created a simplified risk score from the model for practical use. The reduced model's prediction score was simplified by rounding all regression coefficients. The simplified score had a considerably comparable prediction accuracy with the original β coefficients, with an AUC of 0.786 (95%CI: 0.729–0.827) (figure 3). The possible minimum and maximum scores a mother can have are 0 and 11, respectively.

When dichotomized to low risk (<3) and high risk (\geq 3) based on the risk score, 278 (14.36%) were categorized as high risk and 982 (77.9%) as low risk for preterm birth. Using "SpEqualSe", the suggested threshold score to predict preterm birth using risk scores is \geq 3with a sensitivity of 75.14 % and specificity of 67.46% (**table 4**).

Score*(risk	Prediction Model Based on Maternal Characteristics				
category)	Number of mothers	Incidence of preterm bir			
<3 (Low)	982 (77.9%)	72 (7.9%)			
>=3 (High)	278 (14.36%)	97 (53.59%)			
Total	1260 (100%)	169 (13.4%)			
* Score = (2*PIH) +	(3*PROM) + (hemoglobin < 11 mg/d)	l) + 2*residence + (2*APH) +			
gravidity.					

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Discussion

In this study, the incidence of preterm birth was found to be 13.4%. Maternal characteristics were identified in this retrospective study to build a preterm birth prediction risk score. The optimal combination of maternal factors to predict preterm birth include residency, gravidity, hemoglobin < 11 mg/dl, early rupture of membranes, antepartum hemorrhage, and pregnancy-induced hypertension, according to the prediction model. The model has an AUC of 0.816 (95%CI: 0.776 - 0.856). Predicting the probability of preterm birth in pregnant women is essential to take appropriate measures accordingly. Identifying women at risk of preterm birth is an important task for clinical care providers. However, in low and middle-income countries, there are only a few methods available for reliably predicting actual preterm labor in women. Previously, the focus of the research was to explain the maternal and fetal determinants of preterm birth. In recent years, the focus shifted to predicting preterm birth optimally using a combined set of characteristics.

Without any advanced laboratory or imaging testing, this study measured the predicted performance of a model based on maternal features during pregnancy. Furthermore, we discovered that utilizing SpEqualSe as an optimal cut point, the sensitivity and specificity of this prediction model achieved 75.14 percent and 67.46 percent, respectively, at the score threshold of 3.

In our study, a combination (residency, gravidity, hemoglobin < 11 mg/dl, early rupture of membranes, antepartum hemorrhage, and pregnancy-induced hypertension) of maternal characteristics results in an AUC of 0.816 (95%CI: 0.776 - 0.856), has an excellent accuracy according to diagnostic accuracy classification[38].

A study conducted in China showed that a model developed using advanced maternal age, lower maternal height, history of preterm delivery, amount of vaginal bleeding during pregnancy, and lack of folic acid intake before pregnancy for the prediction of overall preterm birth with AUC of (0.6)[39].

This difference may be due to some of the predictors they used such as lower maternal height, lack of folic acid intake before pregnancy, and advanced maternal age. However predictors they used such as lack of folic acid intake before pregnancy not easily obtainable information in routine clinical practice, which makes their model less practical in our setting. This prediction model constitutes variables that are easily obtainable and have reasonable accuracy to be used by

13 | Page

both mid-and lower-level health professionals in the primary care settings. Among the maternal characteristics included in our model, five can be easily found from history taking and one by test for hemoglobin.

The model's accuracy is consistent with a retrospective study done in China that established a preterm birth prediction model based on maternal characteristics, including demographics and clinical characteristics, and a model with predictors (gravidity, educational status, residency, previous history of preterm birth, twin pregnancy, pre-gestational diabetes mellitus (type I or II), chronic hypertension, and place of birth) with AUC of 0.749 (95%CI: 0.732–0.767) [40].

On the other hand, a model incorporating four predictors (cervical length at admission, gestational age, amniotic fluid glucose, and IL-6) has an area under the curve (AUROC) of 0.86[41] and similarly, the combination of biophysical, biochemical, immunological, microbiological, fetal cell, exosomal, or cell-free RNA at different gestational ages, integrated as part of a multivariable predictor model may be necessary to advance our attempts to predict sPTL and preterm birth. In the prediction of spontaneous preterm birth within 48 hours, a prognostic model including qfFN and clinical risk factors showed excellent results[42, 43]. Both models have higher discriminatory performance. The reason for the lower discriminatory performance in our study as compared to the studies described above could be because we used secondary data available from the register and as this dataset is limited and some variables that require advanced laboratory tests were not included in the model.

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Hence, predictors necessitate laboratory testing, which is often unavailable in low-resource settings. As a result, such predictors are difficult to come by in ordinary clinical and public health practice, making the model less useful.

In our prediction score, using 3 as a cutoff point has an acceptable level of specificity, sensitivity, PPV, and NPV to predict preterm birth. It is also possible to shift the cutoff point to increase either of the accuracy measures depending on the program aim and availability of resources.

Conclusion and recommendation

This study shows the possibility of predicting preterm birth using a simple prediction model constructed from maternal characteristics. Thus, the optimal combination of maternal characteristics such as residence, gravidity, hemoglobin < 11 mg/dl, premature rupture of membrane, antepartum hemorrhage, and pregnancy-induced hypertension shows the possibility of predicting preterm birth using a simple prediction model constructed from maternal characteristics. In addition, risk score calculations based on a combination of predictors were effective and had comparable accuracy with the model-based approach of original β coefficients. This score may assist in clinical decision-making. In addition, incorporating this convenient and easily applicable score in the health care system to be used by clinicians to inform pregnant mothers about the future course of their outcome after external validation. Doing further research is needed to validate the prediction tool using prospective follow-up studies in another context before introducing it to the clinical and public health practices.

Data Sharing Statement

The data will be available upon request from the corresponding author.

Author Contributions: S.F.F. conceived the study and wrote the manuscript. Z.A.A, S.F.F, G.T.W, A.K.Y, and A.M.D, all contribute to data analysis, study design, and supervision of data collection. All authors participated in manuscript revision for intellectual content and approval of the final version. All authors have read and agreed to the published version of the manuscript.

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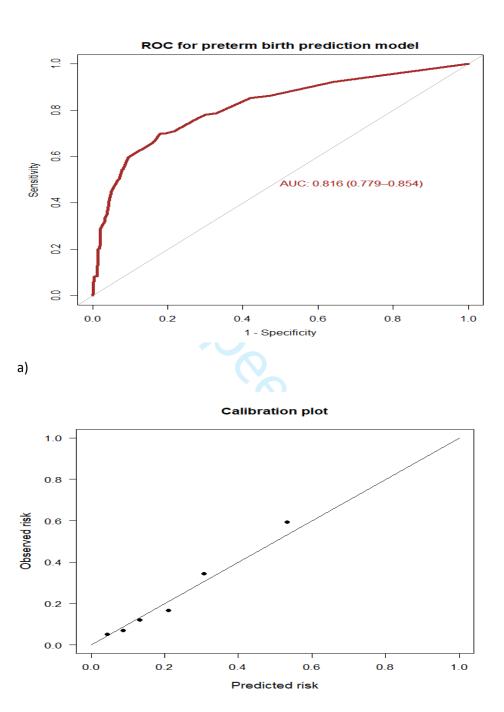
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b)

Figure 1: (a) Area under the ROC curve for the prediction model, and (b) Predicted versus observed preterm birth probability in the sample. This analysis includes mothers who gave birth at FHCSH from January 30/2019 to January 30/2021(n = 1260). Calibration plot created using "plot Calibration" in R programming.

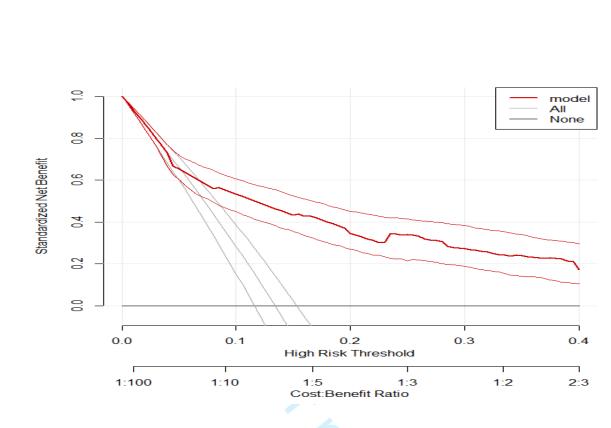


Figure 2: A decision curve plotting showing the net benefit of the model against threshold probability.

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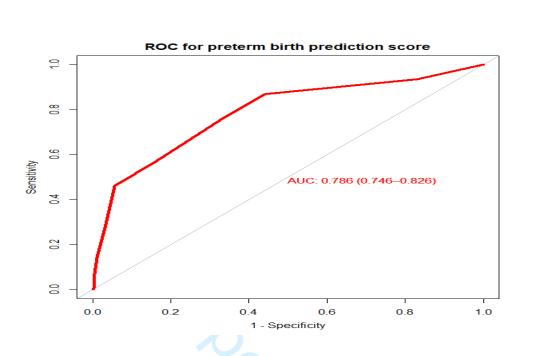


Figure 3: Area under the ROC curve for the simplified risk score to predict the risk of preterm birth among mothers who gave birth at FHCSH from January 30/2019 to January 30/2021.

Developing and validating risk prediction model for preterm birth at Felege Hiwot comprehensive specialized hospital, Northwest Ethiopia: A retrospective follow-up study

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Keywords:	OBSTETRICS, PUBLIC HEALTH, PERINATOLOGY





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7 8	3	Hiwot comprehensive specialized hospital, Northwest Ethiopia: A retrospective
9 10	4	follow-up study
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Abstract

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31 32	using maternal characteristics.
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	Design: A retrospective follow-up study was conducted. Data were coded and entered into
33	Epidata, version 3.02, and were analyzed by using R statistical programming language version
34	4.0.4 for further processing and analysis. Bivariable logistic regression was done to identify the
35	relationship between each predictor and preterm birth. Variables with (p \leq 0.25) from the
36	bivariable analysis were entered into a backward stepwise multivariable logistic regression
37	model, and significant variables ($p < 0.05$) were retained in the multivariable model. Model
38	accuracy and goodness of fit were assessed by computing the area under the ROC curve
39	(discrimination) and calibration plot (calibration) respectively.
40	Setting and participants: This retrospective study was conducted among 1260 pregnant women
41	who did prenatal care and finally delivered at Felege Hiwot Comprehensive Specialized
42	Hospital, Bahir Dar city, Northwest Ethiopia from January 30, 2019, to January 30, 2021.
43	Results : Residence, gravidity, haemoglobin $< 11 \text{ mg/dl}$, early rupture of membranes, antepartum
44	haemorrhage, and pregnancy-induced hypertension remained in the final multivariable prediction
45	model. The AUC of the model was 0.816 (95% confidence interval: $0.779 - 0.856$).
46	Conclusion: This study showed the possibility of predicting preterm birth using maternal
47	characteristics during pregnancy. Thus, using this model could help to identify pregnant women
48	at a higher risk of having a preterm birth to be linked to a center
49	Keywords: Prediction Model, Preterm birth, Risk score, Ethiopia
50	Strength and Limitations of the study
51	\checkmark An adequate number of participants with the outcome helped us to construct the
52	model using a sufficient number of predictor variables and inclusion of sensitivity
53	analyses.
54	\checkmark Multiple imputations was used to address missing data, which has been shown to be a
55	valid technique for dealing with missing data within logistic regression models, resulting
56	in less bias than excluding all women with missing data[1].
57	\checkmark The prediction model is constructed from easily obtainable maternal characteristics that
58	make it applicable in primary care settings.
	 33 34 35 36 37 38 39 40 41 42 43 44 45 46 47 48 49 50 51 52 53 54 55 56 57

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✓ A single-site study, it is confined to a single area, which needs external validation before
 using it in another context.

✓ Furthermore, data were collected from each mother's card; due to this, some important variables were missed, such as previously highlighted factors with preterm birth in different studies.

65 Introduction

 Preterm birth is described as babies that are born alive before the end of 37 weeks of
pregnancy[2]. Preterm birth can be accidental (due to spontaneous preterm labor and/or preterm
membrane rupture) or induced by the provider (by cesarean or labor induction)[3]. Most preterm
births happen spontaneously[4].

An estimated 15 million babies worldwide are born too early per year. That's more than 1 in 10
 infants. About 1 million newborns die per year because of preterm birth complications[5].

Across 184 countries, the rate of preterm birth ranges from 5% to 18% of babies born [6]. However, there are stark disparities in survival rates around the world. Half of the babies born at or below 32 weeks die in low-income settings due to a lack of practical, cost-effective, and critical care, such as comfort, breastfeeding assistance, basic infection care, and trouble Breathing[7].

Furthermore, the effect of preterm birth is also prolonged beyond the neonatal phase and throughout life[8]. Hence, the largest risk of severe health issues, including cerebral palsy, intellectual disability, chronic lung disease, and vision and hearing loss, is faced by babies born before maturity. This introduces a lifelong disability dimension. At some point in their lives, most people will face the struggles and potential disasters of preterm birth either directly in their families or indirectly through events for the nations[8, 9].

To alleviate this burden in the past few decades, numerous methods have been attempted internationally, including in Ethiopia, to prevent and enhance the treatment of preterm births [10-12]. As part of the strategy, it is essential to diagnose or predict preterm birth earlier in pregnancy to take appropriate measures for high-risk groups.

87 However, in most nations, predicting preterm birth is still largely based on subjective clinical
88 experience. This approach may increase unnecessary hospital admissions and unnecessary but

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potentially harmful treatments, such as the use of steroids for the maturation of the fetal lung and tocolysis[13, 14] There were clinical prediction models that aim to estimate the likelihood of preterm birth that include laboratory tests that are typically inaccessible in low-resource settings, such as fetal fibronectin, insulin-like growth factor binding protein-1 (IGFBP-1), interleukin-6, and placental alpha-macroglobulin-1[15-20]. Although there were prediction models for preterm birth, variation in the occurrence of preterm birth globally is relevant, indicating variations in exposure to psychosocial, sociodemographic, and medical risk factors and genetic differences [21-23]. Hence, because of limited resources, the use of easily accessible data to forecast preterm birth seems to be appealing in low- and middle-income areas. Therefore, developing and validating a risk prediction model for prediction of preterm birth using maternal(clinical and non-clinical) characteristics based on the available measurement is paramount to allow early preterm birth intervention such as utero transfer to tertiary care centers, appropriate corticosteroid administration while preventing excessive use, neuroprotective magnesium sulfate therapy, and antibiotic treatment in the event of infection[15, 24] **Methods and Materials Study setting** This retrospective study was conducted among 1260 pregnant women who did prenatal care and finally delivered at Felege Hiwot Comprehensive Specialized Hospital, Bahir Dar city, Northwest Ethiopia from January 30, 2019, to January 30, 2021. Bahir Dar is the capital city of Amhara national regional state and is found 575 km northwest of Addis Ababa.

The hospital has currently a total of 1431 manpower (5 Obstetricians and Gynaecologists and 63 midwives among others) in different disciplines. It has a total of 500 formal beds, 11 wards (emergency ward and Inpatient wards such as Gynecological &Obstetric, Surgical, Orthopaedics, Medical, Pediatric, L&D, Eye unit, NICU, psychiatric, oncology, and 22 OPDS), 39 clinical and non-clinical departments /service units / providing laboratory, Diagnostic, curative & Rehabilitation service at outpatient & inpatient bases as well as disease prevention & health promotion services.

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Sample size determination

The sample size required for model development was determined based on the minimum standard of 10 events per candidate predictor considered, according to the formula N = $(n \times 10)/I$ where N is the sample size, n is the number of candidate predictor variables and I is the estimated event rate in the population [25]. Since there were 17 candidate predictors considered and 10 events per candidate predictor, the estimated number of events for the study was 170. Based on a study done on the prevalence of preterm birth in Debre Tabor hospital was 13%[26], so taking into account this the required sample size was calculated as follows, n = 170*100/13 =1308.

5 130 Study Design and Participants

The theoretical design of the present study was; the incidence of preterm birth as a function of multiple predictors during pregnancy. The source population of the study was all pregnant mothers who gave birth at FHCSH. To be included in this study, mothers must meet all of the following eligibility criteria; All medical records of mothers who gave live birth and had at least one ANC follow-up in FHCSH from January 30/2019 to January 30/2021.

29 136 Sampling method and procedures

A simple random sampling technique was employed to select participants using the medical registration number of a delivered mother from the delivery registration book. First, all mother delivered at FHCSH in the last two years was identified from the delivery registration book. After that records of mothers who meet the inclusion criteria were included in the study. Subsequently, a sampling frame was prepared. Finally, the study unit was selected by using a computer-generated random number.

41 143 Data Collection

Outcome assessment: The outcome variable was attributed to women whose medical records indicated a physician or midwife diagnosis of preterm birth and delivery between 28 and 36 completed weeks of gestation. The gestational age (GA) was measured using either LNMP, which is found to be a more reliable measure of GA in a low-resource setting [27, 28], or an early ultrasound result(12 weeks).

Predictor assessment: Data were collected using a structured checklist through chart review.
 Checklists were developed after reviewing various relevant literatures [29-33]. It consists of
 socio-demographic (Maternal age, Residence), Maternal obstetric characteristics : (History of

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preterm birth, History of abortion, history of stillbirth gravidity, Parity, Multiple pregnancy, APH, PROM, Gestational DM, and PIH), Maternal medical condition : (HGB level, Diabetic Mellitus, Chronic Hypertension, UTI and HIV). **Quality Assurance Mechanisms** To maintain the quality of data, the data collectors and supervisors were trained for a day on the objective of the study, the content of the checklists, how to fill the checklists. Afterward, reviewing 15 charts on medical records of mothers who gave birth at Felege Hiwot Comprehensive Specialized Hospital which is found in Northwest Ethiopia were done. After that, some adjustments (removing variables that were not available in medical record of mothers) were done accordingly. The checklist was developed in English. **Data Processing and Analysis** Data were entered into a software application (EPI DATA, version 3.02) and was analyzed by using R statistical programming language version 4.0.4 for further processing and analysis. There were 13(1%), 2(0.2%), 11(0.9%), 15(2.5%), 21(1.7%), 29(2.3%), 20(1.6%) and 20(1.6%) missing values for premature rupture of membranes, residence, chronic hypertension, multiple pregnancy gestational diabetes Mellitus, pregnancy-induced hypertension ,antepartum hemorrhage and hemoglobin respectively. We assumed data were missing at random, and we, therefore, performed a multivariate imputation by chained equations for all variables evaluated in the prediction model [34]. Sensitivity analysis was performed to assess whether the assumption of missing at random (MAR) is valid or not, and the results were reasonably comparable table (1). Descriptive statistics including median, inter-quartile range (IQR), and percentages, were carried out. Table 1. Sensitivity analysis of the model to predict preterm birth: Comparison of the regression coefficients, standard errors (SE), and p-values for complete case analysis

177 (CCA) and multiple imputed data (MI).

Predicator variables	Comple	ete case ar	e analysis Multiple imputation		S	
	В	SE	P value	В	SE	P value
Chronic hypertension	0.7313	0.6297	0.24	0.581	0.6285	0.92
(yes)						

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Residence (rural)	0.815	0.1946	< 0.001	1.154	0.1958	< 0.001
GDM(yes)	0.709	0.4028	0.07	0.472	0.4236	0.26
HGB(<11g/dl)	0.497	0.2185	0.02	0.642	0.2153	0.001
PROM (yes)	1.898	0.2080	< 0.001	2.097	0.2129	< 0.001
APH (yes)	1.194	0.2858	< 0.001	1.298	0.2874	< 0.001
PIH (yes)	1.353	0.2600	< 0.001	1.368	0.2523	< 0.001
Multiple pregnancy (yes)	0.539	0.3173	0.08	0.446	0.3257	0.17
Gravidity(primigravida)	0.426	0.1944	0.02	0.711	0.1976	< 0.001

179 Model Development and Validation

For model development, bivariable logistic regression was done to obtain insight into the association between each potential predictor and preterm birth. Variables with (p < 0.25) from the bivariable analysis were entered into a backward stepwise multivariable logistic regression model, and significant variables (p < 0.05) were retained in the multivariable model. The results of significant predictors were reported as coefficients with 95% confidence intervals (CI). To check for the model accuracy and goodness of fit, we computed the area under the ROC curve (discrimination) and calibration plot (calibration) using "classifierplots" and "givitiR" packages of R respectively. The AUC ranged from 0.5 (no predictive ability) to 1 (perfect discrimination)[35]. The regression coefficients and their 95% confidence levels, and the AUC were adjusted for overfitting or over-optimism using bootstrapping technique. To make internal validation, we computed 1000 random bootstrap [36]samples with the replacement on all predictors in the data. The model's predictive performance after bootstrapping is considered as the performance that can be expected when the model is applied to future similar populations. To evaluate the clinical and public health impact of the model, we performed a decision curve analysis (DCA) [37] of standardized net benefit across a range of threshold probabilities (0 to 1). In the DCA, the model was compared against two extreme scenarios; "intervention for all" and "no intervention". In our case, the intervention considered is the referral of high-risk pregnant women to facilities where appropriate corticosteroid administration, antibiotic treatment.

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1 2		
3 4	201	Risk Score Development
5	202	To construct an easily applicable preterm birth prediction score, we transformed each coefficient
6 7	203	from the model into a rounded number by dividing it by the lowest coefficient. The number of
8 9	204	points was subsequently rounded to the nearest integer. We determined the total score for each
10 11	205	individual by assigning the points for each variable present and adding them up. The score was
12	206	transformed to a dichotomous, allowing each pregnant woman to be classified as having a high
13 14	207	or low risk of preterm birth. The receiver operating characteristic curve (ROC) was plotted and
15 16	208	the area under the curve (AUC) was calculated to measure the discriminatory power of the
17	209	scoring system.
18 19	210	Patient and public involvement
20 21	211	There was no direct interaction with patients in this study and no direct patient involvement in
22 23	212	the design or conduct of this study.
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Result

Demographic, Obstetric, and Clinical Characteristics of mothers

A total of 1260 study cards were reviewed from a sample of 1308, about 48 cards were not reviewed due to the outcome of intrauterine fetal death, and abortion. Table (2) shows the demographic, obstetric, and clinical characteristics of mothers who gave birth included in the analysis. The median age of the study participants was 26 years with IQR (24-30years); the majority of the participants 1086 (86.2%) were in the age group of 20-34 years. More than three fourth of the participants 926 (73.49%) were urban residents. Of the total of mothers who delivered at FHCSH, more than two-thirds of 841 (66.7%) were multigravida. About parity, above half of them713 (56.6%) were multipara. Concerning past obstetric history, 55 (6.5%) of them had a history of previous preterm birth, 76 (9%) of them had a previous history of stillbirth and 162 (19.3%) of them had a previous history of abortion.

Table 2. Demographic, obstetric, and clinical characteristics of mothers who gave birth at FHCSH, Northwest Ethiopia, 2021.

Characteristics	Category	Frequency	Percent
Gravidity	Primigravida 🗸	419	33.3
	Multigravida	841	66.7
Residence	Urban	926	73.5
	Rural	334	26.5
DM	Yes	44	3.5
	No	1216	96.5
V PH	Yes	84	6.7
	No	1176	93.3
IH	Yes	110	8.73
	No	1150	91.27
GB level	<11d/d1	236	18.7
	>=11g/dl	1024	81.3
hronic hypertension	Yes	21	1.7
	No	1239	98.3
PROM	Yes	195	15.5

9 | Page

Page 11 of 27

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		No	1065	84.5	
	Multiple pregnat	ncies Yes	90	7.2	
		No	1170	92.8	
243	PROM: Pren	nature rupture of membro	nne, HGB: hemoglobi	in, PIH: pregna	ncy-induce
244	hypertension,	APH: antepartum hemorr	hage, GDM: gestation	nal diabetes mell	litus
245					
246	Development of pr	ediction model for preter	rm birth		
247		ed neonates, 169 (13.4%)		5.4%) was pret	erm infants
248	The bivariable logist	ic regression analysis fou	nd several factors we	re eligible to be	included i
249	the prediction model	. These variables were ha	emoglobin level, Grav	vidity, residence	, gestationa
250	diabetes mellitus, AF	PH, PIH, chronic hyperten	sion, PROM, and mul	ltiple pregnancie	s. Using th
251	results, a prediction	model was developed an	equation for the pre	diction model w	vas obtaine
252	table (3).				
253		s and risk-scores of each	predictor included i	n the model to p	oredict
253 254			predictor included i	n the model to p	predict
	Table 3: Coefficient		0	n the model to p able analysis	predict
	Table 3: Coefficientpreterm birth (n = 1		0		oredict Risk
	Table 3: Coefficient preterm birth (<i>n</i> = 1 Predictors	(260)	0	able analysis	
	Table 3: Coefficient preterm birth (<i>n</i> = 1 Predictors Variables*	2 60) Original β	Multivari	able analysis P-	Risk
	Table 3: Coefficient preterm birth (<i>n</i> = 1 Predictors	2 60) Original β	Multivari Bootstrap β	able analysis P-	Risk
	Table 3: Coefficient preterm birth (<i>n</i> = 1 Predictors Variables* Residence	1260) Οriginal β (95 % CI)	Multivari Bootstrap β	able analysis P- value	Risk score
	Table 3: Coefficient preterm birth (<i>n</i> = 1 Predictors Variables* Residence (rural)	1 260) Οriginal β (95 % CI) 1.161 (0.780, 1.545	Multivari Bootstrap β	able analysis P- value <0.001	Risk score 2
	Table 3: Coefficient preterm birth (n = 1) Predictors Variables* Residence (rural) Gravidity	1 260) Οriginal β (95 % CI) 1.161 (0.780, 1.545	Multivari Bootstrap β 5) 1.148 1) 0.666	able analysis P- value <0.001	Risk score 2
	Table 3: Coefficient preterm birth (n = 1 Predictors Variables* Residence (rural) Gravidity (primigravida)	Criginal β (95 % CI) 1.161 (0.780, 1.545 0.675 (0.291, 1.06	Multivari Bootstrap β 5) 1.148 1) 0.666) 2.051	able analysis P- value <0.001 0.01	Risk score 2 1
	Table 3: Coefficient preterm birth (n = 1 Predictors Variables* Residence (rural) Gravidity (primigravida) PROM (yes)	Original β (95 % CI) 1.161 (0.780, 1.545 0.675 (0.291, 1.06 2.081 (1.669, 2.50	Multivari Bootstrap β 5) 1.148 1) 0.666) 2.051 5) 1.348	able analysis P- value <0.001	Risk score 2 1 3
	Table 3: Coefficientpreterm birth (n = 1PredictorsVariables*Residence(rural)Gravidity(primigravida)PROM (yes)APH (yes)	Original β (95 % CI) 1.161 (0.780, 1.545 0.675 (0.291, 1.06 2.081 (1.669, 2.50 1.364 (0.806, 1.915	Multivari Bootstrap β 5) 1.148 1) 0.666) 2.051 5) 1.348 9) 1.368	able analysis P- value <0.001	Risk score 2 1 3 2
	Table 3: Coefficientpreterm birth (n = 1PredictorsVariables*Residence(rural)Gravidity(primigravida)PROM (yes)APH (yes)PIH (yes)HGB <11g/dl	Original β (95 % CI) 1.161 (0.780, 1.545 0.675 (0.291, 1.06 2.081 (1.669, 2.50) 1.364 (0.806, 1.915) 1.387 (0.887, 1.87)	Multivari Bootstrap β 5) 1.148 1) 0.666) 2.051 5) 1.348 9) 1.368) 0.677	able analysis P- value <0.001	Risk score 2 1 3 2 2 2 1

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1		
2 3 4	258	reduced model by the smallest (0.666). The probability or risk of preterm birth = $1/(1 + exp - (-$
5	259	3.517+ 1.148 * Residence (rural) + 0.666 *gravidity (primigravida) + 2.051*PROM (yes) + 1.348
6 7	260	* APH (yes) + 1.387*PIH +0.677*HGB (<11g/dl)
8 9	261	The AUC of the final reduced model was 0.816 (95% confidence interval: 0.779 - 0.856)
10	262	(Figure 1a). The calibration test had a p-value of 0.492, indicating that the model does not
11 12	263	misrepresent the data or calibration of the model was visually accurate since observed and
13 14	264	predicted probabilities were similar (Figure 1b).
15	265	Validation of the model with the bootstrap technique showed hardly any indication of undue
16 17	266	influence by particular observations, with an optimism coefficient of 0.085, resulting AUC of
18 19	267	0.789 (corrected 95% CI: 0.748–0.83).
20 21	268	Using the coefficients (β) the predicted risk cutoff point was a probability of (SpEqualSe >
22	269	0.1320), the model has a sensitivity of 75.74%, specificity of 72.87%, a positive predictive value
23 24	270	of 30.2%, and a negative predictive value of 95.1%.
25 26	271	When applying DCA, we first evaluate whether our model understudy has a higher net benefit
27	272	than the default strategies (referring all and none). This model outperforms the default strategies
28 29	273	across the relevant threshold range. The model has the highest net benefit across the entire range
30 31	274	of threshold probabilities, which indicates that the model has the highest clinical and public
32 33	275	health value. Hence, referral decision made using the model has a higher net benefit than not
34	276	referring at all or referring all regardless of their risk thresholds as shown in <i>figure (2)</i> .
35 36	277	Risk Classification Using a Simplified Risk Score
37 38	278	We created a simplified risk score from the model for practical use. The reduced model's
39	279	prediction score was simplified by rounding all regression coefficients. The simplified score had

prediction score was simplified by rounding all regression coefficients. The simplified score had a considerably comparable prediction accuracy with the original β coefficients, with an AUC of 0.786 (95%CI: 0.729–0.827) (figure 3). The possible minimum and maximum scores a mother can have are 0 and 11, respectively.

- Using "SpEqualSe", the suggested threshold score to predict preterm birth using risk scores is 284 \geq 3with a sensitivity of 75.14 % and specificity of 67.46% table (4).
- When dichotomized to low risk (<3) and high risk (≥3) based on the risk score, 278 (14.36%) were categorized as high risk and 982 (77.9%) as low risk for preterm birth.

287Table 4: Risk classification of preterm birth using simplified prediction score (n = 1260)Score*(riskPrediction Model Based on Maternal Characteristics

11 | Page

60

	category)	Number of mothers	Incidence of preterm birth					
	<3 (Low)	982 (77.9%)	72 (7.9%)					
	>=3 (High)	278 (14.36%)	97 (53.59%)					
288	Total * Score = (2*PIH) + (1260 (100%) (3*PROM) + (hemoglobin < 11 mg/di	$\frac{169 (13.4\%)}{10 + 2*residence + (2*APH) + 10}$					
289	gravidity.							
290	Discussion							
291	In this study, the incidence of preterm birth was found to be 13.4%. Maternal characteristics							
292	were identified in this retrospective study to build a preterm birth prediction risk score. The							
293	optimal combination of maternal factors to predict preterm birth include residency, gravidity,							
294	hemoglobin < 11 mg/dl, early rupture of membranes, antepartum hemorrhage, and pregnancy-							
295	induced hypertension, accord	rding to the prediction model.	The model has an AUC of 0.8					
296	(95%CI: 0.776 - 0.856). H	Predicting the probability of pret	term birth in pregnant women					
297	essential to take appropriate	measures accordingly. Identifying	g women at risk of preterm birth					
298	an important task for clinic	cal care providers. However, in l	low and middle-income countrie					
299	there are only a few metho	ds available for reliably predictin	ng actual preterm labor in wome					
300	Previously, the focus of th	e research was to explain the m	naternal and fetal determinants					
301	preterm birth. In recent year	ars, the focus shifted to predictin	g preterm birth optimally using					
302	combined set of characterist	ics.						
303	Without any advanced lat	poratory or imaging testing, this	is study measured the predict					
304	performance of a model	based on maternal features dur	ring pregnancy. Furthermore, w					
305	discovered that utilizing S	pEqualSe as an optimal cut poin	t, the sensitivity and specificity					
306	this prediction model achieved	eved 75.14 percent and 67.46 p	percent, respectively, at the sco					
307	threshold of 3.							
308	In our study, a combinatio	n (residency, gravidity, hemoglo	bin < 11 mg/dl, early rupture					
309	membranes, antepartum h	emorrhage, and pregnancy-indu	uced hypertension) of matern					
310	characteristics results in an	AUC of 0.816 (95%CI: 0.776 -	0.856), has an excellent accura					
311	according to diagnostic accu	racy classification[38].						
312	A study conducted in China	showed that a model developed u	sing advanced maternal age, low					
313	maternal height, history of preterm delivery, amount of vaginal bleeding during pregnancy, and							
314	lack of folic acid intake befo	pre pregnancy for the prediction of	overall preterm birth with AUC					
315	(0.6)[39].							

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This difference may be due to some of the predictors they used such as lower maternal height, lack of folic acid intake before pregnancy, and advanced maternal age. However predictors they used such as lack of folic acid intake before pregnancy not easily obtainable information in routine clinical practice, which makes their model less practical in our setting. This prediction model constitutes variables that are easily obtainable and have reasonable accuracy to be used by both mid-and lower-level health professionals in the primary care settings. Among the maternal characteristics included in our model, five can be easily found from history taking and one by test for hemoglobin.

The model's accuracy is consistent with a retrospective study done in China that established a preterm birth prediction model based on maternal characteristics, including demographics and clinical characteristics, and a model with predictors (gravidity, educational status, residency, previous history of preterm birth, twin pregnancy, pre-gestational diabetes mellitus (type I or II), chronic hypertension, and place of birth) with AUC of 0.749 (95%CI: 0.732–0.767) [40].

On the other hand, a model incorporating four predictors (cervical length at admission, gestational age, amniotic fluid glucose, and IL-6) has an area under the curve (AUROC) of 0.86[41] and similarly, the combination of biophysical, biochemical, immunological, microbiological, fetal cell, exosomal, or cell-free RNA at different gestational ages, integrated as part of a multivariable predictor model may be necessary to advance our attempts to predict sPTL and preterm birth. In the prediction of spontaneous preterm birth within 48 hours, a prognostic model including qfFN and clinical risk factors showed excellent results[42, 43]. Both models have higher discriminatory performance. The reason for the lower discriminatory performance in our study as compared to the studies described above could be because we used secondary data available from the register and as this dataset is limited and some variables that require advanced laboratory tests were not included in the model.

Hence, predictors necessitate laboratory testing, which is often unavailable in low-resource
settings. As a result, such predictors are difficult to come by in ordinary clinical and public
health practice, making the model less useful.

A study conducted in the UK found that data on maternal characteristics and obstetric history at 11–13 weeks of gestation were predictive of spontaneous early preterm deliveries; this model had an AUC of 0.67[44] which had lower discriminatory performance than the present study. This difference may be difference in study population.

13 | Page

Page 15 of 27

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A model that predict a risk of preterm delivery in women with a multiple pregnancy incorporates previous preterm delivery, monochorionicity, smoking, educational level, and triplet pregnancy for preterm and very preterm delivery had a c-index of 0.68 (95% CI 0.63 to 0.72) and 0.68 (95% CI 0.62 to 0.75) respectively[45]. It had lower discriminatory performance than the present study. This might be due to study population difference. In the present study the study populations were both women who had multiple pregnancies and singleton pregnancy.

In our prediction score, using 3 as a cutoff point has an acceptable level of specificity, sensitivity, PPV, and NPV to predict preterm birth. It is also possible to shift the cutoff point to increase either of the accuracy measures depending on the program aim and availability of resources.

¹ 357 Conclusion and recommendation

This study shows the possibility of predicting preterm birth using a simple prediction model constructed from maternal characteristics. Thus, the optimal combination of maternal characteristics such as residence, gravidity, haemoglobin < 11 mg/dl, premature rupture of membrane, antepartum haemorrhage, and pregnancy-induced hypertension shows the possibility of predicting preterm birth using a simple prediction model constructed from maternal characteristics. In addition, risk score calculations based on a combination of predictors were effective and had comparable accuracy with the model-based approach of original β coefficients. This score may assist in clinical decision-making. In addition, incorporating this convenient and easily applicable score in the health care system to be used by clinicians to inform pregnant mothers about the future course of their outcome after external validation. Doing further research is needed to validate the prediction tool using prospective follow-up studies in another context before introducing it to the clinical and public health practices.

43 370 Data Sharing Statement

The data will be available upon request from the corresponding author.

Author Contributions: S.F.F. conceived the study and wrote the manuscript. Z.A.A, S.F.F,
 G.T.W, A.K.Y, and A.M.D, all contribute to data analysis, study design, and supervision of data
 collection. All authors participated in manuscript revision for intellectual content and approval of
 the final version. All authors have read and agreed to the published version of the manuscript.

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1 2			
2 3 4	378	Comp	peting interest's statement: The author reports no conflicts of interest in this work.
5	379	Ethic	cal approval
6 7	380	Ethica	al clearance was obtained from the Institutional Review Boards (IRB) of Bahir Dar
8 9	381	Unive	ersity, College of Medicine and Health Sciences with Protocol number 083/ 2021) on
10	382		ary 26, 2021. Confidentiality was maintained by omitting the personal identifier of the
11 12	383		ipant during the data collection procedure and information was used only for research
13 14	384	1	ses. Data were collected from the register, which was kept in a secure place and all data
15	385		fully anonymized before we access them. After the collection of data, all the patient records
16 17			
18	386	1	batient cards were placed back in a secure place. Data were entered into a password-
19 20	387	protec	cted computer.
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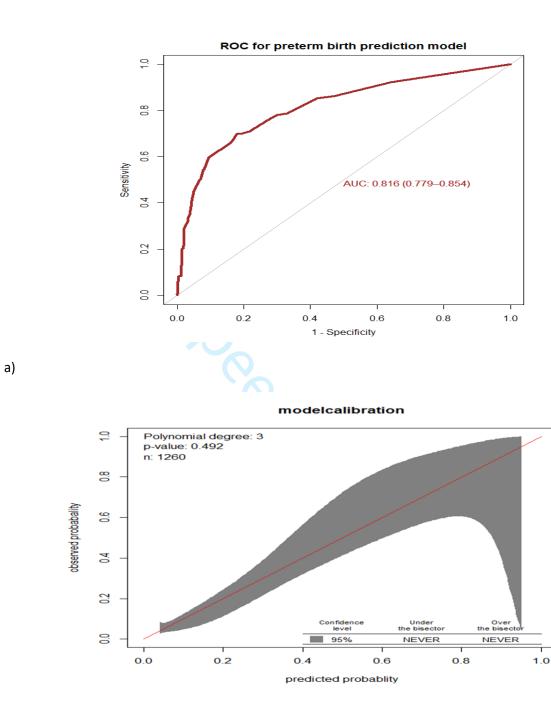
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b)

Figure 1: (a) Area under the ROC curve for the prediction model, and (b) Predicted versus observed preterm birth probability in the sample. This analysis includes mothers who gave birth at FHCSH, 2021(n = 1260). Calibration plot created using "givitiCalibrationBelt" in R programming.

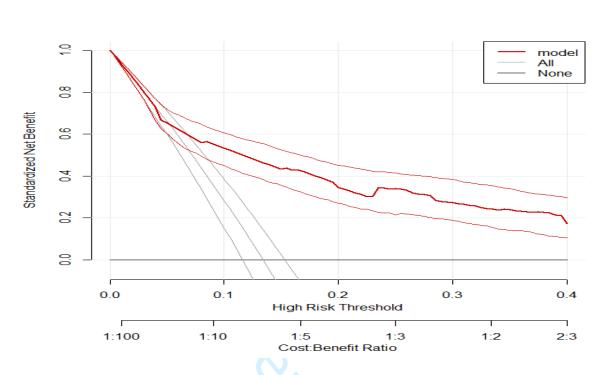


Figure 2: A decision curve plotting the net benefit of the model against threshold probability.

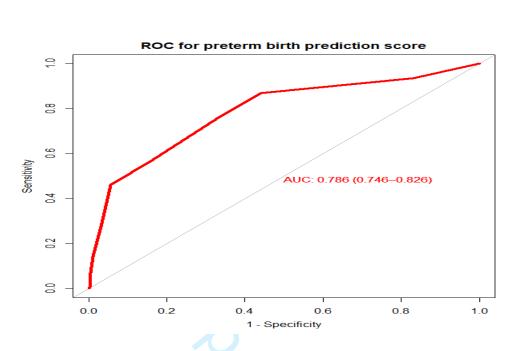


Figure 3: Area under the ROC curve for the simplified risk score to predict the risk of preterm birth among mothers who gave birth at FHCSH, 2021.

Page 2	23 of 27
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3 of 27			BMJ Open			
The RECORD routinely colle		– checklist of items, extended fron data.	n the STROBE statem	ent that should be reported	n observatio	nal studies using
	Item No.	STROBE items	Location in manuscript where items are reported	RECORD items	2	Location in manuscript where items are reported
Title and abst	ract			epr	2	
	1	(a) Indicate the study's design with a commonly used term in the title or the abstract (b) Provide in the abstract an informative and balanced summary of what was done and what was found	Line 1-59	RECORD 1.1: The type of a should be specified in the tit abstract. When possible, the the databases used should be RECORD 1.2: If applicable geographic region and time which the study took place s reported in the title or abstra RECORD 1.3: If linkage be databases was conducted for this should be clearly stated or abstract.	the same of the same within hould be t. veen the study,	Line 1-59
Introduction			1		2.	
Background rationale	2	Explain the scientific background and rationale for the investigation being reported	Line 60-100			
Objectives	3	State specific objectives, including any prespecified hypotheses	Line 96-98		÷	
Methods						
Study Design	4	Present key elements of study design early in the paper	Line 103	guest.		
Setting	5	Describe the setting, locations, and relevant dates, including periods of recruitment, exposure, follow-up, and data collection	Line 103-113 Line 123-128	Protected by a		
Participants	6	(a) Cohort study - Give the eligibility criteria, and the	Line 123-128	RECORD 6.1: The methods population selection (such as		Line 129-135

			BMJ Open	5/bmjc	Page 24 of
		 sources and methods of selection of participants. Describe methods of follow-up <i>Case-control study</i> - Give the eligibility criteria, and the sources and methods of case ascertainment and control selection. Give the rationale for the choice of cases and controls <i>Cross-sectional study</i> - Give the eligibility criteria, and the sources and methods of selection of participants <i>(b) Cohort study</i> - For matched studies, give matching criteria and number of exposed and unexposed <i>Case-control study</i> - For matched studies, give matching criteria and the number of controls per case 	er cuio	algorithms used to identify subjects) should be listed in detail. If this is not possible, an explanation should be provided. RECORD 6.2: Any validation studies of the codes or algorithms used to select the population should be referenced. If validation was conducted for this study and not published elsewhere detailed methods and results should be provided. RECORD 6.3: If the study involved linkage of databases, consider use of a flow diagram or other graphical display to demonstrate the data linkage process, including the number of individuals with linked data at each stage.	
Variables	7	Clearly define all outcomes, exposures, predictors, potential confounders, and effect modifiers. Give diagnostic criteria, if applicable.	Line 137-147	RECORD 7.1: A complete fist of codes and algorithms used to classify exposures, outcomes, confognders, and effect modifiers should be provided. If these cannot be reported, an explanation should be provided.	Line 137-147
Data sources/ measurement	8	 For each variable of interest, give sources of data and details of methods of assessment (measurement). Describe comparability of assessment methods if there is more than one group 	Line 142-147	2024 by guest. Protected	
Bias	9	Describe any efforts to address potential sources of bias	Line 123-141	by copyright	
Study size	10	Explain how the study size was	Line 114-122	yrig	

e 25 of 27			BMJ Open	6/bmjo	
		arrived at			
Quantitative variables	11	Explain how quantitative variables were handled in the analyses. If applicable, describe which groupings were chosen, and why	Line 155-198	2022-061061 on 2	
Statistical methods	12	 (a) Describe all statistical methods, including those used to control for confounding (b) Describe any methods used to examine subgroups and interactions (c) Explain how missing data were addressed (d) <i>Cohort study</i> - If applicable, explain how loss to follow-up was addressed <i>Case-control study</i> - If applicable, explain how matching of cases and controls was addressed <i>Cross-sectional study</i> - If applicable, describe analytical methods taking account of sampling strategy (e) Describe any sensitivity analyses 	Line 155-198	com/ on April 1	
Data access and cleaning methods		Line 148-154		RECORD 12.1: Authors should describe the extent to which the investigators had access to the database population used to create the study population. RECORD 12.2: Authors should provide information on the data cleaning methods used in the study. ₹	Line 130-132
Linkage				RECORD 12.3: State whether the study included person-level, institutional-	Line 130-132

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				level, or other data linkage across two or more databases. The methods of linkage and methods of linkage quality evaluation should be provided.	
Results				Q	
Participants	13	 (a) Report the numbers of individuals at each stage of the study (<i>e.g.</i>, numbers potentially eligible, examined for eligibility, confirmed eligible, included in the study, completing follow-up, and analysed) (b) Give reasons for non- participation at each stage. (c) Consider use of a flow diagram 	Line 130-135	RECORD 13.1: Describe indetail the selection of the persons included in the study (<i>i.e.</i> , study population selection) including filtering based on glata quality, data availability and linkage. The selection of included persons can be described in the text and/or by means of the study flow diagram.	Line 130-135
Descriptive data	14	 (a) Give characteristics of study participants (<i>e.g.</i>, demographic, clinical, social) and information on exposures and potential confounders (b) Indicate the number of participants with missing data for each variable of interest (c) <i>Cohort study</i> - summarise follow-up time (<i>e.g.</i>, average and total amount) 	Line 232-241	om http://bmjopen.bmj.com/ on April 1	
Outcome data	15	Cohort study - Report numbers of outcome events or summary measures over time Case-control study - Report numbers in each exposure category, or summary measures of exposure Cross-sectional study - Report numbers of outcome events or summary measures	Line 247	9, 2024 by guest. Protected by copyright	
Main results	16	(a) Give unadjusted estimates	Line 246-286	Угі	

27 of 27			BMJ Open	/bmjo	
		and, if applicable, confounder- adjusted estimates and their precision (e.g., 95% confidence interval). Make clear which confounders were adjusted for and why they were included (b) Report category boundaries when continuous variables were categorized (c) If relevant, consider translating estimates of relative risk into absolute risk for a meaningful time period		pen-2022-061061 on 26 September 2022. Dov	
Other analyses	17	Report other analyses done—e.g., analyses of subgroups and interactions, and sensitivity analyses	Line 162-166	vnloaded from	
Discussion			·		
Key results	18	Summarise key results with reference to study objectives	Line 351-362	p://bmj	
Limitations	19	Discuss limitations of the study, taking into account sources of potential bias or imprecision. Discuss both direction and magnitude of any potential bias	Line 50-59	RECORD 19.1: Discuss the implications of using data that were not created or collected to answer the specific research question(s) Include discussion of misclassification bias, unmeasured confounding, massing data, and changing eligibility over time, as they pertain to the study being reported.	
Interpretation	20	Give a cautious overall interpretation of results considering objectives, limitations, multiplicity of analyses, results from similar studies, and other relevant evidence	Line 351-362	024 by guest. Protected	
Generalisability	21	Discuss the generalisability (external validity) of the study results	Line 351-362	by copyright	

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Other Information	n					
Funding	22	Give the source of funding and the role of the funders for the present study and, if applicable, for the original study on which the present article is based	Line 369	2022-061061 on 2		
Accessibility of protocol, raw data, and programming code		Line 364		RECORD 22.1: Authors shoul information on how to access a supplemental information such study protocol, raw data, or programming code.	any	Line 364

*Reference: Benchimol EI, Smeeth L, Guttmann A, Harron K, Moher D, Petersen I, Sørensen HT, von Elm E, Langen SM, the RECORD Working ricense. Committee. The REporting of studies Conducted using Observational Routinely-collected health Data (RECORD) Statement. PLoS Medicine 2015; in press. ded from http://bmjopen.bmj.com/ on April 19, 2024 by guest. Protected by copyright.

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Developing and validating risk prediction model for preterm birth at Felege Hiwot comprehensive specialized hospital, Northwest Ethiopia: A retrospective follow-up study

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2	Developing and validating risk prediction model for preterm birth at Felege				
3	Hiwot Comprehensive Specialized Hospital, Northwest Ethiopia: A				
4	retrospective follow-up study				
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3 4	29	Abstract
5 6	30	Objective: To develop and validate a risk prediction model for the prediction of preterm birth
7	31	using maternal characteristics.
8 9	32	Design: A retrospective follow-up study was conducted. Data were coded and entered into
10 11 12 13	33	Epidata, version 3.02, and were analyzed by using R statistical programming language version
	34	4.0.4 for further processing and analysis. Bivariable logistic regression was done to identify the
14	35	relationship between each predictor and preterm birth. Variables with (p \leq 0.25) from the
15 16 17 18	36	bivariable analysis were entered into a backward stepwise multivariable logistic regression
	37	model, and significant variables ($p < 0.05$) were retained in the multivariable model. Model
19 20	38	accuracy and goodness of fit were assessed by computing the area under the ROC curve
21	39	(discrimination) and calibration plot (calibration), respectively.
22 23	40	Setting and participants: This retrospective study was conducted among 1260 pregnant women
24 25 26	41	who did prenatal care and finally delivered at Felege Hiwot Comprehensive Specialized
	42	Hospital, Bahir Dar city, Northwest Ethiopia, from January 30, 2019, to January 30, 2021.
27 28	43	Results : Residence, gravidity, haemoglobin < 11 mg/dl, early rupture of membranes, antepartum
29 30	44	haemorrhage, and pregnancy-induced hypertension remained in the final multivariable prediction
31 32	45	model. The AUC of the model was 0.816 (95% confidence interval: $0.779 - 0.856$).
33	46	Conclusion: This study showed the possibility of predicting preterm birth using maternal
34 35	47	characteristics during pregnancy. Thus, using this model could help to identify pregnant women
36 37	48	at a higher risk of having a preterm birth to be linked to a center
38	49	Keywords: Prediction Model, Preterm birth, Risk score, Ethiopia
39 40	50	Strength and Limitations of the study

- \checkmark An adequate number of participants with the outcome helped us to construct the 51 model using a sufficient number of predictor variables and the inclusion of sensitivity 52 analyses. 53
- 54 \checkmark Multiple imputation were used to address missing data, which has been shown to be a valid technique for dealing with missing data within logistic regression models, resulting 55 in less bias than excluding all women with missing data. 56
 - \checkmark The prediction model is constructed from easily obtainable maternal characteristics that make it applicable in primary care settings.

 \checkmark A single-site study, it is confined to a single area, which needs external validation before

 \checkmark Furthermore, data were collected from each mother's card; due to this, some important

variables were missed, such as previously highlighted factors of preterm birth in

10 63 different studies.
11
12 64
13
14 65 Introduction
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using it in another context.

 Preterm birth is described as babies that are born alive before the end of 37 weeks of
pregnancy[1]. Preterm birth can be accidental (due to spontaneous preterm labor and/or preterm
membrane rupture) or induced by the provider (by cesarean or labor induction)[2]. Most preterm
births happen spontaneously[3].

An estimated 15 million babies worldwide are born too early per year. That is more than 1 in 10
 infants. About 1 million newborns die per year because of preterm birth complications[4].

Across 184 countries, the rate of preterm birth ranges from 5% to 18% of babies born [5]. However, there are stark disparities in survival rates around the world. Half of the babies born at or below 32 weeks die in low-income settings due to a lack of practical, cost-effective, and critical care, such as comfort, breastfeeding assistance, basic infection care, and trouble Breathing[6].

Furthermore, the effect of preterm birth is also prolonged beyond the neonatal phase and throughout life[7]. Hence, the largest risk of severe health issues, including cerebral palsy, intellectual disability, chronic lung disease, and vision and hearing loss, is faced by babies born before maturity. This introduces a lifelong disability dimension. At some point in their lives, most people will face the struggles and potential disasters of preterm birth either directly in their families or indirectly through events for the nations[7, 8].

To alleviate this burden, in the past few decades, numerous methods have been attempted internationally, including in Ethiopia, to prevent and enhance the treatment of preterm births [9-11]. As part of the strategy, it is essential to diagnose or predict preterm birth earlier in pregnancy to take appropriate measures for high-risk groups. However, in most nations, predicting preterm birth is still largely based on subjective clinical experience. This approach may increase unnecessary hospital admissions and unnecessary but potentially harmful treatments, such as the use of steroids for the maturation of the fetal lung and tocolysis[12, 13].

3 | Page

Page 5 of 33

BMJ Open

There were clinical prediction models that aim to estimate the likelihood of preterm birth that include laboratory tests that are typically inaccessible in low-resource settings, such as fetal fibronectin, insulin-like growth factor binding protein-1 (IGFBP-1), interleukin-6, and placental alpha-macroglobulin-1[14-19]. Most current research on PTB prediction focuses on finding PTB risk factors using a hypothesis-testing methodology in highly controlled environments. PTB has been linked to a number of risk factors, including previous preterm labor, multiple gestation (carrying several children), and diabetes, problems with the cervix, uterus, or placenta, smoking, and infections [20-22]. However, women who have preterm delivery often have no known risk factors[23]. In addition, some of the predictors (such as prior PTB) do not apply for first-time mothers.

Predicting the risk of PTB in pregnant women has been the subject of numerous studies[24], but no model exists that is accurate enough to be used in clinical settings. Most research (e.g., cervical length or fetal fibronectin) have concentrated on predictors during the second or third trimester[25]. These predictors, however, can only forecast PTB at intermediate risk and have only been shown to be reliable in high-risk populations. Unfortunately, the majority of women who give birth early have no evident risk factors, and more than half of PTBs happen in low-risk pregnancies, indicating the limited usefulness of using fetal fibronectin or cervical length in the general population[26].

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Due to scarce resources, using readily available data to predict PTB seems appealing in low- and middle-income communities. But relatively few models have been made public. The considerable range in PTB occurrence across the globe, which suggests differences in exposure to psychosocial, sociodemographic, and medical risk factors as well as genetic variations, is also significant [27-29]. As a result, it is necessary to develop and evaluate PTB prediction models in various populations.

Therefore, developing and validating a risk prediction model for the prediction of preterm birth using maternal(clinical and nonclinical) characteristics based on the available measurements is paramount to allow early preterm birth interventions such as in utero transfer to tertiary care appropriate corticosteroid administration while preventing centers, excessive use, neuroprotective magnesium sulfate therapy, and antibiotic treatment in the event of infection[14, 30]

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Methods and Materials

Study setting

This retrospective study was conducted among 1260 pregnant women who did prenatal care and finally delivered at Felege Hiwot Comprehensive Specialized Hospital, Bahir Dar city, Northwest Ethiopia, from January 30, 2019, to January 30, 2021. Bahir Dar is the capital city of Amhara national regional state and is found 575 km northwest of Addis Ababa. The hospital has currently a total of 1431 manpower (5 Obstetricians and Gynaecologists and 63 midwives among others) in different disciplines. It has a total of 500 formal beds, 11 wards (emergency ward and Inpatient wards such as Gynecological &Obstetric, Surgical, Orthopaedics, Medical, Pediatric, L&D, Eye unit, NICU, psychiatric, oncology, and 22 OPDS), 39 clinical and non-clinical departments /service units / providing laboratory, Diagnostic, curative & Rehabilitation service at outpatient & inpatient bases as well as disease prevention & health promotion services.

Sample size determination

The sample size required for model development was determined based on the minimum standard of 10 events per candidate predictor considered, according to the formula N = (n \times /I, where N is the sample size, n is the number of candidate predictor variables and I is the estimated event rate in the population[31]. Since there were 17 candidate predictors considered and 10 events per candidate predictor, the estimated number of events for the study was 170. Based on a study done on the prevalence of preterm birth in Debre Tabor hospital was 13%[32], so taking into account this the required sample size was calculated as follows, n = 170*100/13 =1308.

Study Design and Participants

The theoretical design of the present study was; the incidence of preterm birth as a function of multiple predictors during pregnancy. The source population of the study was all pregnant mothers who gave birth at FHCSH. To be included in this study, mothers must meet all of the following eligibility criteria; all medical records of mothers who gave live birth and had at least one ANC follow-up in FHCSH from January 30/2019 to January 30/2021.

Sampling method and procedures

A simple random sampling technique was employed to select participants using the medical registration number of a delivered mother from the delivery registration book. First, all mothers

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delivered at FHCSH in the last two years was identified from the delivery registration book.
After that, records of mothers who met the inclusion criteria were included in the study.
Subsequently, a sampling frame was prepared. Finally, the study unit was selected by using a computer-generated random number.

10 156 Data Collection

Outcome assessment: The outcome variable was attributed to women whose medical records indicated a physician or midwife diagnosis of preterm birth and delivery between 28 and 36 completed weeks of gestation. The gestational age (GA) was measured using either LNMP, which is found to be a more reliable measure of GA in a low-resource setting[33, 34], or an early ultrasound result(12 weeks).

Predictor assessment: Data were collected using a structured checklist through chart review. Checklists were developed after reviewing various relevant literatures [35-39]. It consists of socio-demographic (Maternal age, Residence), Maternal obstetric characteristics : (History of preterm birth, History of abortion, history of stillbirth, gravidity, Parity, Multiple pregnancy, APH, PROM, Gestational DM, and PIH), Maternal medical condition : (HGB level, Diabetic Mellitus, Chronic Hypertension, UTI and HIV).

32 169 Quality Assurance Mechanisms

To maintain the quality of data, the data collectors and supervisors were trained for a day on the objective of the study, the content of the checklists, and how to fill the checklists. Afterward, reviewing 15 charts medical records of mothers who gave birth at Felege Hiwot Comprehensive Specialized Hospital which is found in Northwest Ethiopia were done. After that, some adjustments (removing variables that were not available in the medical records of mothers) were done accordingly. The checklist was developed in English.

44 176 **Data Processing and Analysis**

Data were entered into a software application (EPI DATA, version 3.02) and was analyzed by using R statistical programming language version 4.0.4 for further processing and analysis. There were 13(1%), 2(0.2%), 11(0.9%), 15(2.5%), 21(1.7%), 29(2.3%), 20(1.6%) and 20(1.6%) missing values for premature rupture of membranes, residence, chronic hypertension, multiple pregnancy gestational diabetes Mellitus, pregnancy-induced hypertension ,antepartum hemorrhage and hemoglobin respectively.

6 | Page

We assumed the data were missing at random, and we, therefore, performed a multivariate imputation by chained equations for all variables evaluated in the prediction model [40]. Sensitivity analysis was performed to assess whether the assumption of missing at random (MAR) is valid or not, and the results were reasonably comparable table (1). Descriptive statistics including median, interquartile range (IQR), and percentages, were carried out.

Table 1. Sensitivity analysis of the model to predict preterm birth: Comparison of the regression coefficients, standard errors (SE), and p-values for complete case analysis (CCA) and multiple imputed data (MI).

Predicator variables	Complete case analysis			Multiple imputations		
	В	SE	P value	В	SE	P value
Chronic hypertension	0.7313	0.6297	0.24	0.581	0.6285	0.92
(yes)						
Residence (rural)	0.815	0.1946	< 0.001	1.154	0.1958	< 0.001
GDM(yes)	0.709	0.4028	0.07	0.472	0.4236	0.26
HGB(<11g/dl)	0.497	0.2185	0.02	0.642	0.2153	0.001
PROM (yes)	1.898	0.2080	< 0.001	2.097	0.2129	< 0.001
APH (yes)	1.194	0.2858	< 0.001	1.298	0.2874	< 0.001
PIH (yes)	1.353	0.2600	< 0.001	1.368	0.2523	< 0.001
Multiple pregnancy (yes)	0.539	0.3173	0.08	0.446	0.3257	0.17
Gravidity(primigravida)	0.426	0.1944	0.02	0.711	0.1976	< 0.001

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192 Model Development and Validation

For model development, bivariable logistic regression was done to obtain insight into the association between each potential predictor and preterm birth. Variables with ($p \le 0.25$) from the bivariable analysis were entered into a backward stepwise multivariable logistic regression model, and significant variables (p < 0.05) were retained in the multivariable model. The results of significant predictors were reported as coefficients with 95% confidence intervals (CI). To check for the model accuracy and goodness of fit, we computed the area under the ROC curve (discrimination) and calibration plot (calibration) using "classifierplots" and "givitiR" packages of R respectively. The AUC ranged from 0.5 (no predictive ability) to 1 (perfect

7 | Page

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discrimination)[41]. The regression coefficients and their 95% confidence levels, and the AUC were adjusted for overfitting or over optimism using the bootstrapping technique. To make internal validation, we computed 1000 random bootstrap [42]samples with the replacement of all predictors in the data. The model's predictive performance after bootstrapping is considered as the performance that can be expected when the model is applied to future similar populations. To evaluate the clinical and public health impact of the model, we performed a decision curve analysis (DCA) [43] of standardized net benefits across a range of threshold probabilities (0 to 1). In the DCA, the model was compared with two extreme scenarios; "intervention for all" and "no intervention". In our case, the intervention considered is the referral of high-risk pregnant women to facilities where appropriate, corticosteroid administration, antibiotic treatment.

Risk Score Development

To construct an easily applicable preterm birth prediction score, we transformed each coefficient of the model into a rounded number by dividing it by the lowest coefficient. The number of points was subsequently rounded to the nearest integer. We determined the total score for each individual by assigning points for each variable present and adding them up. The score was transformed to dichotomous, allowing each pregnant woman to be classified as having a high or low risk of preterm birth. The receiver operating characteristic curve (ROC) was plotted and the area under the curve (AUC) was calculated to measure the discriminatory power of the scoring system.

5 220 Patient and public involvement

There was no direct interaction with patients in this study and no direct patient involvement in the design or conduct of this study.

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Result

Demographic, Obstetric, and Clinical Characteristics of mothers

A total of 1260 study cards were reviewed from a sample of 1308, about 48 cards were not reviewed due to the outcome of intrauterine fetal death, and abortion. Table (2) shows the demographic, obstetric, and clinical characteristics of mothers who gave birth included in the analysis. The median age of the study participants was 26 years with IQR (24-30years); the majority of the participants 1086 (86.2%) were in the age group of 20-34 years. More than three-fourth of the participants 926 (73.49%) were urban residents. Of the total of mothers who delivered at FHCSH, more than two-thirds of 841 (66.7%) were multigravida. About parity, above, half of them713 (56.6%) were multipara. Concerning past obstetric history, 55 (6.5%) of them had a history of previous preterm birth, 76 (9%) of them had a history of stillbirth and 162 (19.3%) of them had a history of abortion.

Table 2. Demographic, obstetric, and clinical characteristics of mothers who gave birth at FHCSH, Northwest Ethiopia, 2021.

Characteristics	Category	Frequency	Percent
Gravidity	Primigravida	419	33.3
	Multigravida	841	66.7
Residence	Urban	926	73.5
	Rural	334	26.5
GDM	Yes	44	3.5
	No	1216	96.5
APH	Yes	84	6.7
	No	1176	93.3
PIH	Yes	110	8.73
	No	1150	91.27
HGB level	<11d/d1	236	18.7
	>=11g/dl	1024	81.3
Chronic hypertension	Yes	21	1.7
	No	1239	98.3
PROM	Yes	195	15.5

9 | Page

	No	1065	84.5
Multiple pregnancies	Yes	90	7.2
	No	1170	92.8

hypertension, APH: antepartum hemorrhage, GDM: gestational diabetes mellitus

245 PROM: Premature rupture of membrane, HGB: hemoglobin, PIH: pregnancy-induced

247 Development of prediction model for preterm birth

Out of 1260 delivered neonates, 169 (13.4%) (95%, CI (11.6%, 15.4%) was preterm infants. The bivariable logistic regression analysis found several factors were eligible to be included in the prediction model. These variables were haemoglobin level, Gravidity, residence, gestational diabetes mellitus, APH, PIH, chronic hypertension, PROM, and multiple pregnancies. Using the results, a prediction model was developed, and equation for the prediction model was obtained table (3).

254 Table 3: Coefficients and risk scores of each predictor included in the model to predict

255 preterm birth (*n* = 1260)

Predictors	Multivariable analysis				
Variables [*]	Original β		Р-	Risk	
	(95 % CI)	Bootstrap β	value	score	
Residence					
(rural)	1.161 (0.780, 1.545)	1.148	< 0.001	2	
Gravidity	0.675 (0.291, 1.061)	0.666	0.01	1	
(primigravida)					
PROM (yes)	2.081 (1.669, 2.50)	2.051	< 0.001	3	
APH (yes)	1.364 (0.806, 1.915)	1.348	< 0.001	2	
PIH (yes)	1.387 (0.887, 1.879)	1.368	< 0.001	2	
HGB <11g/dl	0.676 (0.255, 1.09)	0.677	< 0.001	1	

*Variables retained in the reduced model are; residence, APH, hemoglobin, PIH, gravidity, and PROM. Both backward and forward selection showed the same results. β after internal validation with bootstrapping is shown. Simplified risk score: we divided the coefficient of predictors included in the reduced model by the smallest (0.666). The probability or risk of preterm birth = 1/(1 + exp - (-

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3 3.517+1.148 * Residence (rural) + 0.666 *gravidity (primigravida) + 2.051*PROM (yes) + 1.348
 261 * APH (yes) + 1.387*PIH +0.677*HGB (<11g/dl).

The AUC of the final reduced model was 0.816 (95% confidence interval: 0.779 - 0.856)
(Figure 1a). The calibration test had a p-value of 0.492, indicating that the model does not
misrepresent the data or the calibration of the model was visually accurate since the observed
and predicted probabilities were similar (Figure 1b).

In addition, to verify whether any maternal characteristics were used as a specific predictor of preterm birth we performed an ROC analysis. The analysis indicated that, residence (AUC=0.604, 95% CI 0.564 to 0.643), gravidity (AUC=0.59, 95% CI 0.571 to 0.628), PROM (AUC=0.580, 95% CI 0.544 to 0.616), APH (AUC= 0.695, 95% CI 0.661 to 0.729), PIH (AUC= 0.721, 95% CI 0.685 to 0.757), and HGB (AUC=0.630, 95% CI 0.591 to 0.668) emerged as better predictors of preterm birth (Figure 2).

- 27 27 Validation of the model with the bootstrap technique showed hardly any indication of undue
 27 27 27 273 influence by particular observations, with an optimism coefficient of 0.085, resulting AUC of
 29 274 0.789 (corrected 95% CI: 0.748–0.83).
- ³⁰ ³¹ 275 Using the coefficient (β), the predicted risk cutoff point was a probability of (SpEqualSe > ³² 276 0.1320), the model has a sensitivity of 75.74%, specificity of 72.87%, a positive predictive value ³⁴ 277 of 30.2%, and a negative predictive value of 95.1%.
- When applying DCA, we first evaluate whether our model understudy has a higher net benefit than the default strategies (referring all and none). This model outperforms the default strategies across the relevant threshold range. The model has the highest net benefit across the entire range of threshold probabilities, which indicates that the model has the highest clinical and public health value. Hence, the referral decision made using the model has a higher net benefit than not referring at all or referring all regardless of their risk threshold as shown in *figure (3)*.
- ⁴⁶ 284 Risk Classification Using a Simplified Risk Score
 ⁴⁷
- We created a simplified risk score from the model for practical use. The reduced model's prediction score was simplified by rounding all regression coefficients. The simplified score had a considerably comparable prediction accuracy to the original β coefficients, with an AUC of 0.786 (95%CI: 0.729–0.827) (figure 4). The possible minimum and maximum scores a mother can have are 0 and 11, respectively.

1 2							
3 4	290	Using "SpEqualSe", the suggested threshold score to predict preterm birth using risk scores is					
4 5 6 7 8 9 10 11 12 13 14	291	\geq 3 with a sensitivity of 75.14 % and specificity of 67.46% table (4).					
	292	When dichotomized into low risk (<3) and high risk (\geq 3) based on the risk score, 278 (14.36%)					
	293	were categorized as high risk and 982 (77.9%) as low risk for preterm birth.					
	294	Table 4: Risk classification of preterm birth using simplified prediction score (n = 1260)					
		Score*(risk Prediction Model Based on Maternal Characteristics					
		category) Number of mothers Incidence of preterm birth					
15		<3 (Low) 982 (77.9%) 72 (7.9%)					
16 17		>=3 (High) 278 (14.36%) 97 (53.59%)					
18	•••	Total 1260 (100%) 169 (13.4%)					
19 20	295 296	* Score = $(2*PIH) + (3*PROM) + (hemoglobin < 11 mg/dl) + 2*residence + (2*APH) +$					
21 22 23 24 25 26 27 28 29 30 31 32 33 34 35	290	gravidity. Discussion					
	298	In this study, the incidence of preterm birth was found to be 13.4%. Maternal characteristics					
	299	were identified in this retrospective study to build a preterm birth prediction risk score. We					
	300	intended to employ maternal features that are easily accessible and pertinent to clinical practice					
	301	in countries with constrained resources, including Ethiopia. These nations may not have the					
	302	financial resources to pay for ultrasound exams and laboratory tests. The optimal combination of					
	303	maternal factors to predict preterm birth includes residency, gravidity, and hemoglobin < 11					
	304	mg/dl, early rupture of membranes, antepartum hemorrhage, and pregnancy-induced					
36	305	hypertension, according to the prediction model. The model has an AUC of 0.816 (95%CI:					
37 38	306	0.776 - 0.856). Predicting the probability of preterm birth in pregnant women is essential to take					
39 40	307	appropriate measures accordingly. Identifying women at risk of preterm birth is an important					
41	308	task for clinical care providers. However, in low and middle-income countries, there are only a					
42 43	309	few methods available for reliably predicting actual preterm labor in women. Previously, the					
44 45	310	focus of the research was to explain the maternal and fetal determinants of preterm birth. In					
46 47	311	recent years, the focus shifted to predicting preterm birth optimally using a combined set of					
47 48	312	characteristics.					
49 50	313	Without any advanced laboratory or imaging testing, this study measured the predicted					
51 52	314	performance of a model based on maternal features during pregnancy. Furthermore, we					
52 53 54 55 56	315	discovered that utilizing SpEqualSe as an optimal cut point, the sensitivity and specificity of					

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this prediction model achieved 75.14 percent and 67.46 percent, respectively, at the score threshold of 3.

In our study, a combination (residency, gravidity, hemoglobin < 11 mg/dl, early rupture of membranes, antepartum hemorrhage, and pregnancy-induced hypertension) of maternal characteristics resulted in an AUC of 0.816 (95%CI: 0.776 - 0.856), has an excellent accuracy according to diagnostic accuracy classification[44].

We found that early rupture of membrane is strong predictors of preterm birth. Similar evidence was found in different studies [36, 37, 45, 46]. The effect of a burst membrane on uterine contraction could explain this. Existing scientific evidence confirms that when a membrane ruptures, natural uterotonic chemicals are released, and these uterotonic chemicals drive uterine contraction, resulting in PTB. This finding suggested that due attention should be given to women with premature rupture of membrane.

In our study, pregnancy-induced hypertension is strong predictors of preterm birth. Similar studies have demonstrated that pregnancy-induced hypertension was predictive of subsequent preterm birth[47, 48]. This could be related to vascular injury to the placenta caused by pregnancy-induced hypertension issues or iatrogenesis caused by the severity of hypertension or its complications. As a result, the oxytocin receptors are activated, resulting in preterm labor and delivery. Or else this conclusion could be explained by current scientific evidence suggesting that PIH is linked to vascular and placental injury, which causes oxytocin receptors to be activated, resulting in PTB. Therefore, it is imperative to identify populations at risk pregnancy-induced hypertension and introduce risk lowering interventions.

Another strong predictor of preterm birth is the place of residence. Existed evidence shows that there is an association between preterm birth and rural residence [49-53]. This gap may be explained by the greater accessibility and availability of maternal health service in metropolitan regions. It has long been understood that social deprivation and the nuanced interactions between them affect prenatal outcomes, including premature birth[54]. Hence, accessing maternal health services targeted to rural women could improve prenatal outcomes including the risk of preterm birth.

Antepartum hemorrhage is the predictor of preterm birth which is supported by different studies[55]. Identification of groups at risk for antepartum hemorrhage and the introduction of risk-reducing measures are therefore essential. Other predictors of preterm birth are gravidity and

13 | Page

Page 15 of 33

BMJ Open

hemoglobin <11 g/dl (anemia) which is in line with different studies[32, 56]. The molecular factors that could explain how anemia, iron deficiency, or both, could result in preterm delivery. In reality, a number of plausible molecular processes have linked anemia to a higher risk of premature birth. Accordingly, maternal and fetal stress can be caused by anemia (by resulting in hypoxia) and iron deficiency (by increasing serum nor-epinephrine concentrations), which in turn induces the production of corticotrophin-releasing hormone (CRH). Additionally, iron deficiency may raise the risk of maternal infections, which can again boost the synthesis of CRH. High levels of CRH are known to be a risk factor for PTB since they increase the likelihood of PTB [57]. Thus, we can conclude that, in order to prevent PTB, routine ANC services need to place a greater emphasis on anemia prevention.

A study conducted in China showed that a model developed using advanced maternal age, lower maternal height, history of preterm delivery, amount of vaginal bleeding during pregnancy, and lack of folic acid intake before pregnancy for the prediction of overall preterm birth with AUC of (0.6)[58]. Which had lower discriminatory performance than the present study, this difference may be due to some of the predictors they used such as lower maternal height, lack of folic acid intake before pregnancy, and advanced maternal age. However, the predictors they used such as lack of folic acid intake before pregnancy are not easily obtainable information in routine clinical practice, which makes their model less practical in our setting. This prediction model constitutes variables that are easily obtainable and have reasonable accuracy to be used by both mid-and lower-level health professionals in primary care settings. Among the maternal characteristics included in our model, five can be easily found by history taking and one by test for hemoglobin. The model's accuracy is consistent with a retrospective study done in China that established a preterm birth prediction model based on maternal characteristics, including demographics and clinical characteristics, and a model with predictors (gravidity, educational status, residency, history of preterm birth, twin pregnancy, pre-gestational diabetes mellitus (type I or II), chronic hypertension, and place of birth) with AUC of 0.749 (95%CI: 0.732–0.767) [48].

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On the other hand, a model incorporating four predictors (cervical length at admission, gestational age, amniotic fluid, glucose, and IL-6) has an area under the curve (AUROC) of 0.86[59] and similarly, the combination of biophysical, biochemical, immunological, microbiological, fetal cell, exosomal, or cell-free RNA at different gestational ages, integrated as part of a multivariable predictor model may be necessary to advance our attempts to predict

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sPTL and preterm birth. In the prediction of spontaneous preterm birth within 48 hours, a prognostic model including qfFN and clinical risk factors showed excellent results[60, 61]. Both models have higher discriminatory performance. The reason for the lower discriminatory performance in our study compared to the studies described above could be because we used secondary data available from the register and as this dataset is limited and some variables that require advanced laboratory tests were not included in the model.

Hence, predictors necessitate laboratory testing, which is often unavailable in low-resource settings. As a result, such predictors are difficult to come by in ordinary clinical and public health practice, making the model less useful.

A study conducted in the UK found that data on maternal characteristics and obstetric history at 11–13 weeks of gestation were predictive of spontaneous early preterm delivery; this model had an AUC of 0.67[62] which had lower discriminatory performance than the present study. This difference may be the difference in the study population.

A model that predicts a risk of preterm delivery in women with multiple pregnancy incorporating previous preterm delivery, monochorionicity, smoking, educational level, and triplet pregnancy for preterm and very preterm delivery had a c-index of 0.68 (95% CI 0.63 to 0.72) and 0.68 (95% CI 0.62 to 0.75) respectively [63]. It had lower discriminatory performance than the present study. This might be due to the study population difference. In the present study, the study populations were both women who had multiple pregnancies and singleton pregnancy.

In our prediction score, using 3 as a cutoff point has an acceptable level of specificity, sensitivity, PPV, and NPV to predict preterm birth. It is also possible to shift the cutoff point to increase either of the accuracy measures depending on the program aim and availability of resources.

The strength of the study was using an adequate number of participants with the outcome, which helped us to construct the model using a sufficient number of predictor variables. In addition, our prediction model was constructed from easily obtainable maternal characteristics that make it applicable in primary care setting and multiple imputation were used to address missing data, which has been shown to be a valid technique for dealing with missing data within logistic regression models, resulting in less bias than excluding all women with missing data.

However, the findings from this study should be interpreted with the perspective of the following limitations. As a single-site study, it is confined to a single area, which needs external validation

15 | Page

BMJ Open

before using it in another context. Furthermore, data were collected from each mother's card; due
to this, some important variables were missed, such as previously highlighted factors with
preterm birth in different studies.

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- 10 413
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- 14 415 Conclusions and recommendations

This study shows the possibility of predicting preterm birth using a simple prediction model constructed from maternal characteristics. Thus, the optimal combination of maternal characteristics such as residence, gravidity, haemoglobin < 11 mg/dl, premature rupture of membrane, antepartum haemorrhage, and pregnancy-induced hypertension shows the possibility of predicting preterm birth using a simple prediction model constructed from maternal characteristics. In addition, risk score calculations based on a combination of predictors was effective and had comparable accuracy with the model-based approach of the original β coefficients. This score may assist in clinical decision-making. In addition, incorporating this convenient and easily applicable score in the health care system to be used by clinicians to inform pregnant mothers about the future course of their outcome after external validation. Doing further research is needed to validate the prediction tool using prospective follow-up studies in another context before introducing it to clinical and public health practices.

- 36 428 Data Sharing Statement
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- 429 Data will be available upon request from the corresponding author.

Author Contributions: S.F.F. conceived the study and wrote the manuscript. Z.A.A, S.F.F,
G.T.W, A.K.Y, and A.M.D, all contribute to data analysis, study design, and supervision of data
collection. All authors participated in manuscript revision for intellectual content and approval of
the final version. All authors have read and agreed to the published version of the manuscript.

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50 436 **Competing interest's statement:** The author reports no conflicts of interest in this work.

52 437 Ethical approval

Ethical clearance was obtained from the Institutional Review Board (IRB) of Bahir Dar
University, College of Medicine and Health Sciences with Protocol number 083/ 2021) on

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February 26, 2021. It is a retrospective study of medical records and personal identifiers were not used on the data collection checklist. So, the IRB waived the requirement for informed consent from each participant. Confidentiality was maintained by omitting the personal identifier of the participant during the data collection procedure and the information was used only for research purposes. Data were collected from the register, which was kept in a secure place and all data were fully anonymized before we accessed them. After the collection of data, all patient records and patient cards were placed back in a secure place. Data were entered into a password-protected computer.

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Page 21 of 33

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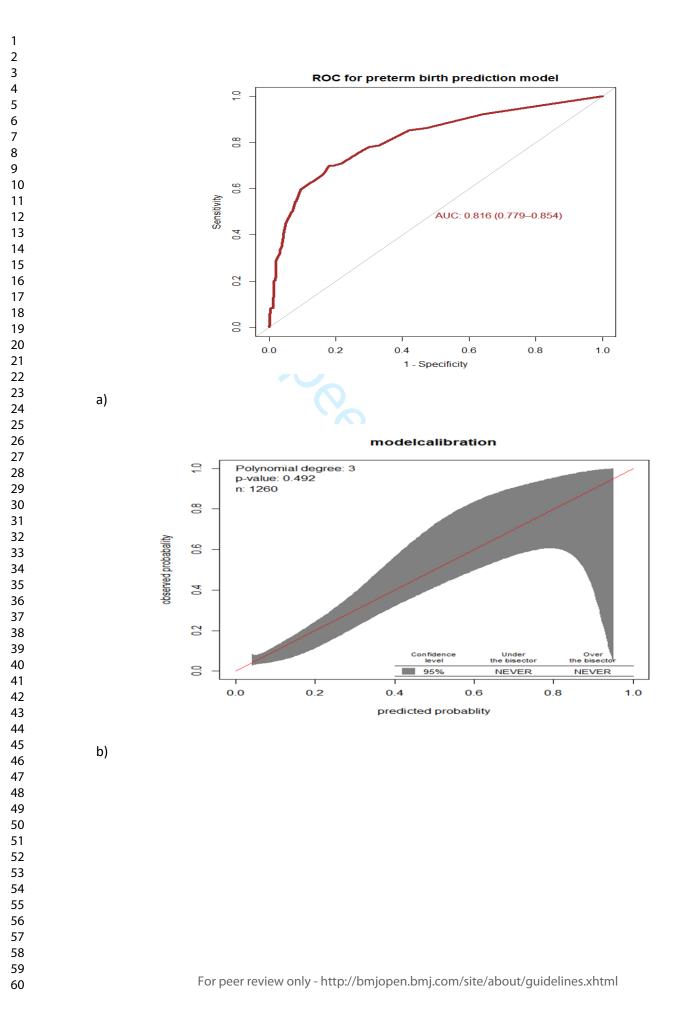
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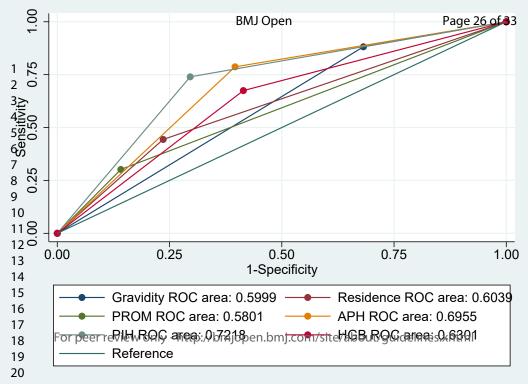
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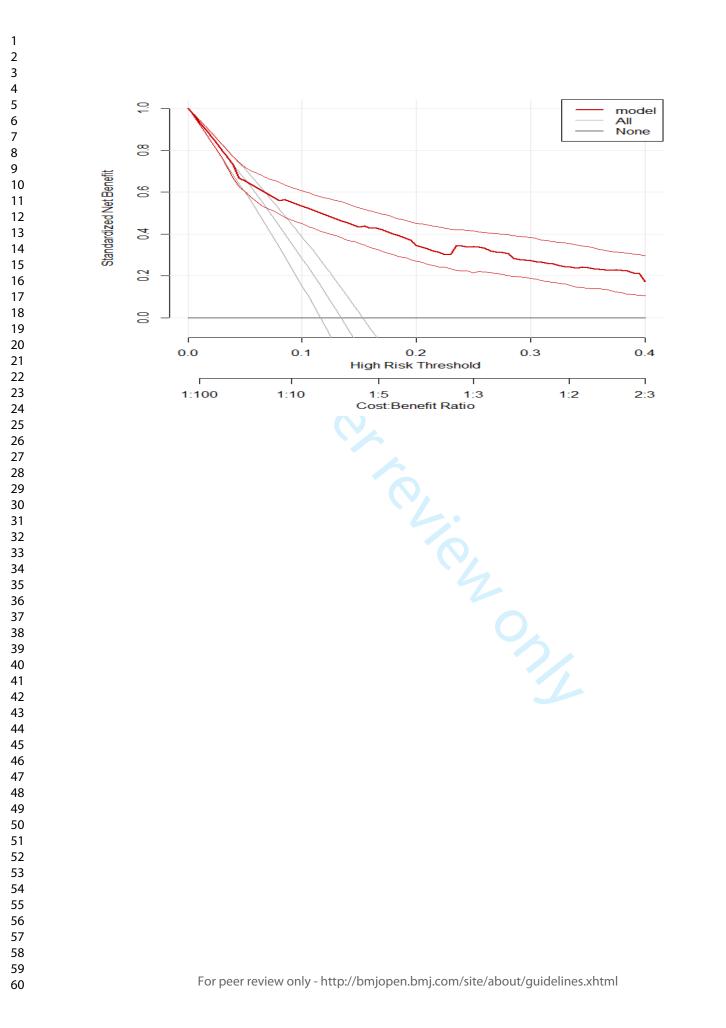
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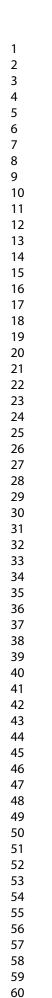
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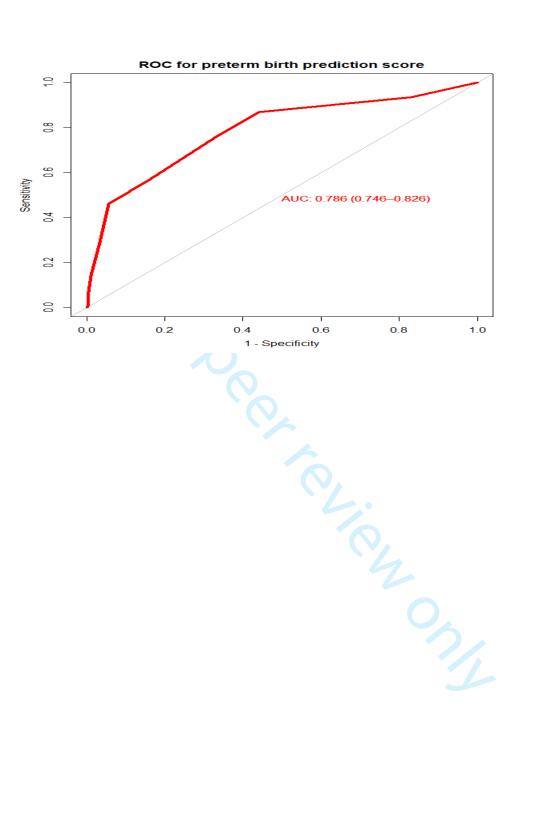
638	Figure 1: (a) Area under the ROC curve for the prediction model, and (b) Predicted versus
639	observed preterm birth probability in the sample. This analysis includes mothers who gave birth
640	at FHCSH, $2021(n = 1260)$. Calibration plot created using "givitiCalibrationBelt" in R
641	programming.
642	Figure 2: Receiver operating characteristic curve of maternal parameters for prediction of
643	postpartum glucose intolerance. Residence, PROM, APH, PIH, HGB and Gravidity.
644	Figure 3: A decision curve plotting the net benefit of the model against threshold probability.
645	Figure 4: Area under the ROC curve for the simplified risk score to predict the risk of preterm
646	birth among mothers who gave birth at FHCSH, 2021.
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Page 29 of 33

9 of 33 The RECORI routinely colle		– checklist of items, extended fron data.	BMJ Open n the STROBE statem	ent that should be reported	n observatio	nal studies using
	Item No.	STROBE items	Location in manuscript where items are reported	RECORD items		Location in manuscript where items are reported
Title and abst	ract			ept		-
	1	(a) Indicate the study's design with a commonly used term in the title or the abstract (b) Provide in the abstract an informative and balanced summary of what was done and what was found	Line 1-59	RECORD 1.1: The type of a should be specified in the tit abstract. When possible, the the databases used should be RECORD 1.2: If applicable geographic region and time which the study took place s reported in the title or abstra RECORD 1.3: If linkage be databases was conducted for this should be clearly stated or abstract.	e or hame of included. the ame within hould be et. ween the study,	Line 1-59
Introduction				<u>ج</u> 0	• •	
Background rationale	2	Explain the scientific background and rationale for the investigation being reported	Line 60-100		>	
Objectives	3	State specific objectives, including any prespecified hypotheses	Line 96-98			
Methods				by		-
Study Design	4	Present key elements of study design early in the paper	Line 103	guest.		
Setting	5	Describe the setting, locations, and relevant dates, including periods of recruitment, exposure, follow-up, and data collection	Line 103-113 Line 123-128	Protected by o		
Participants	6	(a) Cohort study - Give the eligibility criteria, and the	Line 123-128	RECORD 6.1: The methods		Line 129-135

			BMJ Open	5/bmjc	Page 30 of
		 sources and methods of selection of participants. Describe methods of follow-up <i>Case-control study</i> - Give the eligibility criteria, and the sources and methods of case ascertainment and control selection. Give the rationale for the choice of cases and controls <i>Cross-sectional study</i> - Give the eligibility criteria, and the sources and methods of selection of participants (b) Cohort study - For matched studies, give matching criteria and unexposed <i>Case-control study</i> - For matched studies, give matching criteria and the number of controls per case 	Provio	algorithms used to identify subjects) should be listed in detail. If this is not possible, an explanation should be provided. RECORD 6.2: Any validation studies of the codes or algorithms used to select the population should be referenced. If validation was conducted for this study and not published elsewhere detailed methods and results should be provided. RECORD 6.3: If the study involved linkage of databases, consider use of a flow diagram or other graphical display to demonstrate the data linkage process, including the number of individuals with linked data at each stage.	
Variables	7	Clearly define all outcomes, exposures, predictors, potential confounders, and effect modifiers. Give diagnostic criteria, if applicable.	Line 137-147	RECORD 7.1: A complete fist of codes and algorithms used to classify exposures, outcomes, confognders, and effect modifiers should be provided. If these cannot be reported, an explanation should be provided.	Line 137-147
Data sources/ measurement	8	 For each variable of interest, give sources of data and details of methods of assessment (measurement). Describe comparability of assessment methods if there is more than one group 	Line 142-147	2024 by guest. Protected	
Bias	9	Describe any efforts to address potential sources of bias	Line 123-141	by copyright.	
Study size	10	Explain how the study size was	Line 114-122	yrig	

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		arrived at		р р л.	
Quantitative variables	11	Explain how quantitative variables were handled in the analyses. If applicable, describe which groupings were chosen, and why	Line 155-198	2022-061061 on 2	
Statistical methods	12	 (a) Describe all statistical methods, including those used to control for confounding (b) Describe any methods used to examine subgroups and interactions (c) Explain how missing data were addressed (d) <i>Cohort study</i> - If applicable, explain how loss to follow-up was addressed <i>Case-control study</i> - If applicable, explain how matching of cases and controls was addressed <i>Cross-sectional study</i> - If applicable, describe analytical methods taking account of sampling strategy (e) Describe any sensitivity analyses 	Line 155-198	om/ on April 1	
Data access and cleaning methods		Line 148-154		RECORD 12.1: Authors should describe the extent to which the investigators had access to the database population used to create the study population.	Line 130-132
Linkage				RECORD 12.3: State whether the study included person-level, institutional-	Line 130-132

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				level, or other data linkage across two or more databases. The methods of linkage and methods of linkage quality evaluation should be provided.	
Results				Q	
Participants	13	 (a) Report the numbers of individuals at each stage of the study (<i>e.g.</i>, numbers potentially eligible, examined for eligibility, confirmed eligible, included in the study, completing follow-up, and analysed) (b) Give reasons for non- participation at each stage. (c) Consider use of a flow diagram 	Line 130-135	RECORD 13.1: Describe indetail the selection of the persons included in the study (<i>i.e.</i> , study population selection) including filtering based on glata quality, data availability and linkage. The selection of included persons can be described in the text and/or by means of the study flow diagram.	Line 130-135
Descriptive data	14	 (a) Give characteristics of study participants (<i>e.g.</i>, demographic, clinical, social) and information on exposures and potential confounders (b) Indicate the number of participants with missing data for each variable of interest (c) <i>Cohort study</i> - summarise follow-up time (<i>e.g.</i>, average and total amount) 	Line 232-241	om http://bmjopen.bmj.com/ on April 1	
Outcome data	15	Cohort study - Report numbers of outcome events or summary measures over time Case-control study - Report numbers in each exposure category, or summary measures of exposure Cross-sectional study - Report numbers of outcome events or summary measures	Line 247	9, 2024 by guest. Protected by copyright	
Main results	16	(a) Give unadjusted estimates	Line 246-286	yrig	

33 of 33			BMJ Open	/bmjo	
		and, if applicable, confounder- adjusted estimates and their precision (e.g., 95% confidence interval). Make clear which confounders were adjusted for and why they were included (b) Report category boundaries when continuous variables were categorized (c) If relevant, consider translating estimates of relative risk into absolute risk for a meaningful time period		pen-2022-061061 on 26 September 2022. Dov	
Other analyses	17	Report other analyses done—e.g., analyses of subgroups and interactions, and sensitivity analyses	Line 162-166	vnloaded from	
Discussion			·		
Key results	18	Summarise key results with reference to study objectives	Line 351-362	p://bmj	
Limitations	19	Discuss limitations of the study, taking into account sources of potential bias or imprecision. Discuss both direction and magnitude of any potential bias	Line 50-59	RECORD 19.1: Discuss the implications of using data that were not created or collected to answer the specific research question(s) Include discussion of misclassification bias, unmeasured confounding, massing data, and changing eligibility over time, as they pertain to the study being reported.	
Interpretation	20	Give a cautious overall interpretation of results considering objectives, limitations, multiplicity of analyses, results from similar studies, and other relevant evidence	Line 351-362	024 by guest. Protected	
Generalisability	21	Discuss the generalisability (external validity) of the study results	Line 351-362	by copyright	

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Other Information								
Funding	22	Give the source of funding and the role of the funders for the present study and, if applicable, for the original study on which the present article is based	Line 369					
Accessibility of protocol, raw data, and programming code		Line 364		RECORD 22.1: Authors sho information on how to access supplemental information su study protocol, raw data, or programming code.	s any the as the	Line 364		

*Reference: Benchimol EI, Smeeth L, Guttmann A, Harron K, Moher D, Petersen I, Sørensen HT, von Elm E, Langen SM, the RECORD Working ricense. Committee. The REporting of studies Conducted using Observational Routinely-collected health Data (RECORD) Statement. PLoS Medicine 2015; in press. ded from http://bmjopen.bmj.com/ on April 19, 2024 by guest. Protected by copyright.

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