Association between fasting blood glucose levels and stroke events: a large-scale community-based cohort study from China

Ya Zhang 1,2, Shujun Gu 3, Cuicui Wang 1,2, Dong Liu 1,2, Qiuyi Zhang 3, Man Yang 3, Zhengyuan Zhou 3, Hui Zuo 1,2

ABSTRACT

Objectives Diabetes mellitus has been associated with stroke. However, the association between fasting blood glucose (FBG) and stroke risk in a general population remains not clear. The purpose of our study was to examine the FBG levels on subsequent stroke risk in a community-based cohort in China.

Design Prospective cohort study, employing Cox proportional hazard model to analyse the association of FBG levels with stroke risk.

Setting A community-based cohort study included adults participating in a baseline survey conducted in 2013 in Changshu, eastern China.

Participants 16113 participants were recruited with a multistage sampling method, excluding participants with severe disability, severe cancer, severe psychiatric disturbance or previous stroke before enrolment.

Primary outcome measures Stroke events.

Results During a median follow-up of 5.5 years, 417 incident cases of stroke were identified. The adjusted HR for total and ischaemic stroke for participants in the fourth quartile of FBG compared with the first quartile was 1.44 (95% CI 1.07 to 1.94) and 1.57 (95% CI 1.11 to 2.21), respectively. FBG levels of ≥7.0 mmol/L were associated with an increased risk of stroke based on two clinical classifications (American Diabetes Association: 1.68 (1.24 to 2.27); WHO: 1.62 (1.21, 2.13)). In stratified analyses, risk associations existed in women (HR: 1.92, 95% CI 1.22 to 3.01) and postmenopausal women (HR: 1.68, 95% CI 1.06 to 2.68) for the fourth quartile versus the first. More importantly, the meta-analysis observed a positive association between FBG levels and stroke risk (pooled HR: 1.70, 95% CI 1.27 to 2.29; n=7).

Conclusions Higher FBG level was independently associated with an increased risk of stroke in Chinese adults, especially significant in women.

INTRODUCTION

Stroke is one of the most important causes of human diseases and death.1 According to the statistics of global disease burden in 2013, about 25.7 million patients with stroke and 6.5 million stroke deaths worldwide.2 There are 113 million disability-adjusted life years caused by stroke each year.3 It poses a heavy psychological and economic burden to patients’ families and society. Therefore, the primary prevention of stroke has now become a top priority for global public health. In China, it has become the leading cause of death4 and one of the important reasons for disability in adults.5

Studies have shown that diabetes mellitus can increase the risk of stroke.6,7 A cohort study involving 510,000 people observed that diabetes significantly increased the risk of ischaemic stroke and haematencephalon.7 A prospective study showed that insulin resistance or diagnosed diabetes can predict the first stroke.8 However, these studies focus on diabetic status instead of continuous glucose levels as the study exposure, which limits the generalisability of the research findings. Since fasting blood glucose (FBG) is the least interfered by diet,9 the FBG levels are considered as a more reliable tool to measure blood glucose levels than random blood glucose levels.10 FBG levels have a stronger predictive power for functional prognosis than random blood glucose levels. Therefore, rather than focusing on random blood glucose, it is important to explore the association between FBG levels and the risk of stroke.

In this study, a community population from China was followed up for 5.5 years to study the relationship between FBG levels and the
incidence of stroke, aiming at providing suggestions and a basis for primary prevention of stroke among people with different FBG levels.

METHODS
Study population
The baseline survey of the current cohort study was conducted in 2013. The source population was 20,943 participants aged 35–74 years in rural communities in Changshu, eastern China who participated in an earlier study on metabolic syndrome in 2008, in addition to a small proportion of new participants. A multistage sampling method was used to recruit participants. Individuals with severe disability, severe cancer or severe psychiatric disturbance were excluded. Of these, 16,457 subjects in total were included in the baseline survey. We excluded 137 patients with a previous stroke before enrolment. Participants with missing data on baseline health information (n=207) were also excluded, leaving 16,113 participants in total for final analyses (online supplemental appendix 1).

Laboratory measurements
Blood samples were collected in the morning after overnight fasting and were evaluated biochemical and clinical parameters on the same day. FBG was measured by an oxidase enzymatic method. The levels of high-density lipoprotein cholesterol (HDL-C), total cholesterol (TC) and triglyceride (TG) were assessed enzymatically using commercial reagents with an automatic biochemistry analyser (Hitachi, Tokyo, Japan). The levels of low-density lipoprotein cholesterol (LDL-C) were calculated by the Friedewald formula. The estimated glomerular filtration rate (eGFR) was calculated on the basis of the chronic kidney disease-epidemiology creatinine equation. The body mass index (BMI) was calculated as weight in kilograms divided by the square of height in metres, and weight and height were measured by standard methods.

Follow-up and ascertainment of stroke
Follow-up began at the baseline and ended at the date of stroke diagnosis, death or, 31 December 2018, whichever came first. We classified total stroke into an ischaemic stroke and non-ischaemic stroke (including haemorrhagic stroke, subarachnoid haemorrhage and unexplained types of stroke). The primary outcome was hospitalisation or death due to stroke collected by ICD-10 codes (total stroke: I60–I61, I63–I64; ischaemic stroke: I63 except I63.9). If a participant had more than one stroke event during the follow-up, only the first stroke was considered.

Assessments of covariates
Demographic characteristics, lifestyle factors, personal medical history (cardiovascular disease (CVD), hyperlipidaemia, hypertension, diabetes, chronic obstructive pulmonary disease (COPD), nephropathy and cancer) and family history data for all participants were obtained using standard questionnaires administered by trained staff. After the subjects rested for 5 min, the trained observers used an electronic sphygmomanometer (Omron HBP1300, OMRON Corporation, China) to measure the blood pressure in the sitting position three times every 30 s. The criteria for defining hypertension were an average measurement of systolic blood pressure ≥140 mmHg or diastolic blood pressure ≥90 mmHg, diagnosis of hypertension or antihypertensive medication. Impaired fasting glucose (IFG) is defined as an FBG of 6.1–6.9 mmol/L by WHO and an FBG of 5.6–6.9 mmol/L by the American Diabetes Association (ADA). The body mass index (BMI) was calculated as weight in kilograms divided by the square of height in metres, and weight and height were measured by standard methods.

Meta-analysis
We conducted a meta-analysis combining the current study findings with previous observations on FBG and stroke risk. We searched PubMed, Web of Knowledge, medRxiv and bioRxiv databases for related research articles published between 1 January 2000 and 10 September 2020 using the keywords ‘stroke’, ‘cardiovascular disease’ and ‘fasting blood glucose’, ‘blood glucose’, ‘fasting glucose’, ‘fasting plasma glucose’, ‘glucose’, in addition to ‘prospective study’, ‘cohort study’ or ‘follow-up’. The study subjects were restricted to apparently healthy adults in the community. Two authors extracted the data independently, and if they had disagreements, discussed it with the third author. This study was conducted following the Preferred Reporting Items for Systematic Reviews and Meta-Analyses guideline. Finally, seven studies were included. We extracted information about authors, sample size, the country where the study was implemented, time of follow-up, adjusted covariates, outcome (stroke events) and the effect size (HR). The measurement and definition of FBG levels were directly extracted from each literature. We used a random-effect model to calculate the pooled risk estimation based on the assumption of study heterogeneity. The Begg and Egger tests were performed to detect potential publication bias.

Statistical analysis
Baseline demographic characteristics for continuous variables were expressed as median (IQRs), and categorical variables were presented as percentages. Log-transformation was used to all laboratory measures to normalise distributions.

Cox proportional hazard analyses were conducted to evaluate FBG levels and risk of stroke and results presented as HR and 95% CIs. FBG levels were included in the models as a categorical measure (quartiles and three categories by clinical criteria). Multivariable models included conventional risk factors at baseline: age (continuous), sex, BMI (continuous), current smoking (yes/no), family history of stroke (yes/no), hypertension (yes/no), eGFR (mL/min/1.73 m²) LDL-C (mmol/L), HDL-C (mmol/L), TC (mmol/L), TG (mmol/L) and use of statins. We also used clinical classification (normal
FBG, IFG, hyperglycaemic) to classify FBG levels. In addition, we conducted stratified analyses by sex to check potential effect modification. Sensitivity analyses were conducted to determine the robustness of the primary analysis results. We limited the risk association analyses to participants who did not report their own medical history, including hypertension, hyperlipidaemia, cancer, nephropathy, CVD or COPD at baseline.

Statistical analyses were conducted with R (V.4.0.2, www.r-project.org). All tests were two sided, and a p value <0.05 was considered statistically significant.

**Patient and public involvement**

Patients and the public were not involved in the design, conduct or reporting of this study.

**RESULTS**

**Characteristics of the study population**

After a median follow-up of 5.5 years (range: 0.1–5.7 years), a total of 417 (192 men and 225 women) stroke events were identified, of which 323 (77.5%) were ischaemic stroke. Table 1 shows the baseline characteristics of participants across the quartiles of FBG. Compared with the participants in the study with lower FBG, those with higher FBG levels were more likely to suffer from obesity, hypertension and hyperlipidaemia and had lower levels of eGFR and preferred to have a family history of stroke (p<0.001, table 1).

**FBG levels and risk of stroke**

As shown in table 2, we observed that higher FBG levels were significantly correlated with an increased risk of total and ischaemic stroke in both the original models and models adjusted for age, sex, BMI, HDL-C, LDL-C, TC, TG, eGFR, hypertension, smoking status, use of statins and family history of stroke. Compared with the lowest FBG quartile, the multivariable-adjusted HRs for the highest quartile were 1.44 (95% CI 1.07 to 1.94, p=0.021) for total and 1.57 (95% CI 1.11 to 2.21, p=0.016) for ischaemic stroke. A positive association between FBG levels and non-ischaemic stroke was observed in the crude model, but the risk association was rendered statistically non-significant in adjusted models. Furthermore, we examined the association between clinical classifications of FBG levels and stroke risk, and the results are presented in table 3. The multivariable-adjusted HR for total stroke was 1.68 (95% CI 1.24 to 2.27, p=0.001) and 1.62 (95% CI 1.21 to 2.13, p=0.002) for participants with hyperglycaemic compared with those with normal blood glucose, following the ADA and WHO recommendations, respectively.

Results of sex-stratified analyses are shown in table 4. In original models, higher levels of FBG were significantly associated with an increased risk of total stroke only in women. After adjustment covariate for age, BMI and other variates, the risk association remained significant (HR: 1.92, 95% CI 1.22 to 3.01, p=0.004) for the fourth level in female groups. Similar risk estimates were found in postmenopausal or oophorectomy women. But no significant association with stroke was observed for FBG levels in premenopausal women (online supplemental appendix 2).

In sensitivity analyses, the study population was restricted to 15404 subjects without a self-reported history of hyperlipidaemia, CVD, cancer, nephropathy and COPD yielded essentially no change in risk of total stroke (HR: 1.38, 95% CI 1.01 to 1.92) and ischaemic stroke (HR: 1.55, 95% CI 1.06 to 2.26). After further elimination of participants with hypertension, the association in women was slightly raised for total stroke (HR: 2.50, 95% CI 1.11 to 5.60) and ischaemic stroke (HR: 2.59, 95% CI 1.03 to 6.51) (data not shown).

**Meta-analysis**

Figure 1 shows the results of meta-analysis including six previously published prospective studies together with our current study. In a word, higher FBG levels were associated with an increased risk of stroke (pooled HR: 1.70, 95% CI 1.27 to 2.29). The Begg and Egger tests did not suggest the existence of potential publication bias (p>0.1 for both tests).

**DISCUSSION**

**Principal findings**

In our prospective community-based cohort, higher FBG levels were associated with an increased risk of total and ischaemic stroke. The risk association was prominent in women. After excluding the baseline of hypertension, hyperlipidaemia, CVD, cancer, nephropathy and COPD, these associations were still persisted in analyses of women. Our findings extend the previous knowledge on diabetes and stroke risk, suggesting an important role of routine glucose measurement in stroke prevention.

**Associations of FBG with stroke risk**

When analysing stroke subtypes, there is no evidence that FBG is associated with non-ischaemic stroke, but the association with ischaemic stroke persists. This is consistent with the results of a prospective cohort study among Korean men.14 It suggests that the correlation between FBG and stroke is mainly contributed by an ischaemic stroke. A large-scale prospective study in China (including 96110 participants) indicated that both hyperglycaemic and hypoglycaemic are associated with an increased risk of incident intracerebral haemorrhage, after adjustment for potential confounders.24

Results based on the ADA and WHO criteria demonstrated that FBG levels of ≥7.0 mmol/L were a significant risk factor for stroke (HR (95% CI) 1.68 (1.24 to 2.27) and 1.62 (1.21 to 2.13) following the ADA and WHO criteria, respectively), whereas the risk of stroke in patients with IFG had an increasing trend, but this trend was not statistically significant. A large number of studies showed that people with diabetes are at increased...
Table 1  Baseline characteristics of study population by FBG levels and stroke risk: Cohort Changshu

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>Total n=16113</th>
<th>Quartile of FBG levels (mmol/L)</th>
<th></th>
<th></th>
<th></th>
<th>P value</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td>Q1 (&lt;4.87) n=4097</td>
<td>Q2 (4.87–5.24) n=4027</td>
<td>Q3 (5.24–5.77) n=3979</td>
<td>Q4 (&gt;5.77) n=4010</td>
<td></td>
</tr>
<tr>
<td>Age, years</td>
<td>55.7 (46.3, 64.4)</td>
<td>48.6 (41.4, 59.1)</td>
<td>53.8 (44.8, 62.84)</td>
<td>57.8 (48.6, 65.9)</td>
<td>60.0 (51.4, 67.2)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Men (%)</td>
<td>40.9</td>
<td>42.1</td>
<td>39.4</td>
<td>40.2</td>
<td>42.1</td>
<td>0.002</td>
</tr>
<tr>
<td>Premenopausal in women (%)</td>
<td>37.5</td>
<td>55.3</td>
<td>42.3</td>
<td>29.7</td>
<td>22.3</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Follow-up time, years</td>
<td>5.53 (5.49, 5.56)</td>
<td>5.52 (5.49, 5.54)</td>
<td>5.52 (5.49, 5.55)</td>
<td>5.53 (5.50, 5.57)</td>
<td>5.55 (5.50, 5.58)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>BMI (kg/m²)</td>
<td>23.3 (21.2, 25.7)</td>
<td>22.6 (20.6, 24.9)</td>
<td>23.0 (21.0, 25.3)</td>
<td>23.6 (21.5, 25.9)</td>
<td>24.2 (22.0, 26.5)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Current smoking (%)</td>
<td>23.6</td>
<td>26.9</td>
<td>23.5</td>
<td>21.9</td>
<td>21.9</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Family history of stroke (%)</td>
<td>5.8</td>
<td>3.8</td>
<td>5.4</td>
<td>6.5</td>
<td>7.6</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Hypertension (%)</td>
<td>22.8</td>
<td>13.9</td>
<td>18.6</td>
<td>24.9</td>
<td>34.0</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Fasting blood glucose, mmol/L</td>
<td>5.24 (4.87, 5.77)</td>
<td>4.66 (4.47, 4.77)</td>
<td>5.06 (4.97, 5.15)</td>
<td>5.47 (5.34, 5.60)</td>
<td>6.37 (6.01, 7.20)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>impaired fasting glucose (%)</td>
<td>8.6</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Diabetes (%)</td>
<td>8.8</td>
<td>0.6</td>
<td>0.7</td>
<td>—</td>
<td>—</td>
<td>33.3</td>
</tr>
<tr>
<td>TC, mmol/L</td>
<td>4.70 (4.13, 5.36)</td>
<td>4.48 (3.95, 5.04)</td>
<td>4.62 (4.08, 5.23)</td>
<td>4.79 (4.22, 5.44)</td>
<td>4.99 (4.36, 5.70)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>TG, mmol/L</td>
<td>1.32 (0.96, 1.90)</td>
<td>1.20 (0.88, 1.68)</td>
<td>1.26 (0.93, 1.78)</td>
<td>1.33 (0.97, 1.91)</td>
<td>1.52 (1.08, 2.30)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>HDL-C, mmol/L</td>
<td>1.38 (1.14, 1.66)</td>
<td>1.36 (1.14, 1.62)</td>
<td>1.37 (1.14, 1.64)</td>
<td>1.40 (1.15, 2.68)</td>
<td>1.38 (1.14, 1.69)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>LDL-C, mmol/L</td>
<td>2.57 (1.92, 3.27)</td>
<td>2.64 (2.03, 3.34)</td>
<td>2.62 (1.97, 3.30)</td>
<td>2.55 (1.89, 3.26)</td>
<td>2.47 (1.79, 3.17)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>eGFR (mL/min/1.73 m²)</td>
<td>95.59 (84.39, 105.42)</td>
<td>100.88 (89.80, 110.06)</td>
<td>96.744 (86.38, 106.50)</td>
<td>93.60 (83.12, 102.95)</td>
<td>91.65 (79.13, 101.17)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Use of statins (%)</td>
<td>0.4</td>
<td>0.1</td>
<td>0.3</td>
<td>0.3</td>
<td>0.7</td>
<td>&lt;0.001</td>
</tr>
</tbody>
</table>

*P values were determined by χ² tests for categorical and Kruskal-Wallis tests for continuous variables.

BMI, body mass index; eGFR, estimated glomerular filtration rate; FBG, fasting blood glucose; HDL-C, high-density lipoprotein cholesterol; LDL-C, low-density lipoprotein cholesterol; Q, quartile; TC, total cholesterol; TG, triglyceride.
A meta-analysis including 698,782 participants reported that the risk of ischaemic stroke in diabetic patients was 2.27 times higher than normal people. A cohort study in Sweden also showed that poor blood glucose control in patients with type 2 diabetes was associated with an increased risk of stroke and death. Although diabetes is associated with stroke, the influence of pre-diabetes on future stroke risk has not been clear yet. Some large sample meta-analyses reported pre-diabetes is associated with the risk of CVD and heart failure. But in another meta-analysis, including 15 cohort studies, more than half of the studies showed that after adjusting for cardiovascular risk factors, there was no significant association between pre-diabetes and stroke.

### Table 2

<table>
<thead>
<tr>
<th>Quartile of FBG levels</th>
<th>Q1 (HR (95% CI))</th>
<th>Q2 (HR (95% CI))</th>
<th>Q3 (HR (95% CI))</th>
<th>Q4 (HR (95% CI))</th>
<th>P trend</th>
</tr>
</thead>
<tbody>
<tr>
<td>Total stroke</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Cases/N</td>
<td>57/4097</td>
<td>89/4027</td>
<td>110/3979</td>
<td>161/4010</td>
<td></td>
</tr>
<tr>
<td>Model 1</td>
<td>1 (Ref.)</td>
<td>1.60 (1.15 to 2.22)</td>
<td>2.00 (1.45 to 2.75)</td>
<td>2.92 (2.16 to 3.94)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Model 2</td>
<td>1 (Ref.)</td>
<td>1.21 (0.87 to 1.69)</td>
<td>1.17 (0.85 to 1.62)</td>
<td>1.46 (1.07 to 1.98)</td>
<td>0.018</td>
</tr>
<tr>
<td>Model 3</td>
<td>1 (Ref.)</td>
<td>1.19 (0.86 to 1.63)</td>
<td>1.16 (0.86 to 1.62)</td>
<td>1.44 (1.07 to 1.94)</td>
<td>0.021</td>
</tr>
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<td>Ischaemic stroke</td>
<td></td>
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<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Cases/N</td>
<td>42/4097</td>
<td>69/4027</td>
<td>83/3979</td>
<td>129/4010</td>
<td></td>
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<tr>
<td>Model 1</td>
<td>1 (Ref.)</td>
<td>1.68 (1.15 to 2.47)</td>
<td>2.05 (1.41 to 2.96)</td>
<td>3.17 (2.24 to 4.49)</td>
<td>&lt;0.001</td>
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<tr>
<td>Model 2</td>
<td>1 (Ref.)</td>
<td>1.27 (0.87 to 1.87)</td>
<td>1.19 (0.82 to 1.73)</td>
<td>1.58 (1.11 to 2.25)</td>
<td>0.013</td>
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<tr>
<td>Model 3</td>
<td>1 (Ref.)</td>
<td>1.25 (0.91 to 1.86)</td>
<td>1.17 (0.81 to 1.70)</td>
<td>1.57 (1.11 to 2.21)</td>
<td>0.016</td>
</tr>
<tr>
<td>Non-ischaemic stroke</td>
<td></td>
<td></td>
<td></td>
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<td></td>
</tr>
<tr>
<td>Cases/N</td>
<td>15/4097</td>
<td>20/4027</td>
<td>27/3979</td>
<td>32/4010</td>
<td></td>
</tr>
<tr>
<td>Model 1</td>
<td>1 (Ref.)</td>
<td>1.36 (0.70 to 2.66)</td>
<td>1.86 (0.99 to 3.50)</td>
<td>2.20 (1.19 to 4.07)</td>
<td>0.006</td>
</tr>
<tr>
<td>Model 2</td>
<td>1 (Ref.)</td>
<td>1.04 (0.53 to 2.04)</td>
<td>1.13 (0.60 to 2.14)</td>
<td>1.11 (0.59 to 2.07)</td>
<td>0.713</td>
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<tr>
<td>Model 3</td>
<td>1 (Ref.)</td>
<td>1.10 (0.57 to 2.14)</td>
<td>1.12 (0.64 to 2.24)</td>
<td>1.07 (0.57 to 2.19)</td>
<td>0.844</td>
</tr>
</tbody>
</table>

Model 1: unadjusted model; model 2: adjusted for age, sex and BMI; model 3: further adjusted for HDL-C, TC, TG, LDL-C, eGFR, hypertension, smoking status, use of statins and family history of stroke.

### Table 3

<table>
<thead>
<tr>
<th>ADA definition</th>
<th>&lt;5.6 mmol/L</th>
<th>Cases/N</th>
<th>Model 1</th>
<th>Model 2</th>
<th>Model 3</th>
</tr>
</thead>
<tbody>
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<td>HR (95% CI)</td>
<td>220/11079</td>
<td>1 (Ref.)</td>
<td>1 (Ref.)</td>
<td>1 (Ref.)</td>
<td></td>
</tr>
<tr>
<td>WHO definition</td>
<td>&lt;6.1 mmol/L</td>
<td>Cases/N</td>
<td>301/13422</td>
<td>1 (Ref.)</td>
<td>1 (Ref.)</td>
</tr>
<tr>
<td>HR (95% CI)</td>
<td>1 (Ref.)</td>
<td>1 (Ref.)</td>
<td>1 (Ref.)</td>
<td>1 (Ref.)</td>
<td></td>
</tr>
</tbody>
</table>

Model 1: unadjusted model; model 2: adjusted for age, sex and BMI; model 3: further adjusted for HDL-C, TC, TG, LDL-C, eGFR, hypertension, smoking status, use of statins and family history of stroke.

ADA, American Diabetes Association; eGFR, estimated glomerular filtration rate; FBG, fasting blood glucose; HDL-C, high-density lipoprotein cholesterol; LDL-C, low-density lipoprotein cholesterol; TC, total cholesterol; TG, triglyceride.
A cohort study investigated the sex-specific associations of pre-diabetes with major clinical outcomes reached the same conclusion.31

Stratified analysis showed that the correlation between FBG levels and stroke risk only existed in women. According to the restratification of menopause, only postmenopausal women with the highest FBG level had a higher risk of stroke (HR (95% CI) 1.68 (1.06 to 2.68)), no associations were found in premenopausal women. It is probably due to the low number of stroke events in premenopausal women. Few published studies reported sex-stratified results. A sex-specific cohort study in the Middle East showed that only men with higher FBG levels showed an increased risk of stroke (HR: 2.15 (1.26 to 4.67) in FBG level of 6.1–6.9 mmol/L, HR: 2.38 (1.08 to 5.25) in FBG level of >7.0 mmol/L). 32 Korean Heart Study showed that IFG was associated with an increased risk of ischaemic stroke in men but not in women.33 We suspect that the disparities in sex between studies could be due to different definitions of outcome (all stroke vs ischaemic stroke) and the variety in covariate adjustments. A meta-analysis including 64 cohort studies provides the clearest evidence yet for the sex difference of diabetes and stroke risk. The excess risk of diabetes-related stroke in women is significantly higher than that in men and independent of sex difference in other cardiovascular risk factors.34 Our results support the findings of this meta-analysis, indicating that the association between FBG levels and stroke cases in women, and the mechanism may be related to the protective effect of oestrogen.

**Possible mechanisms**

Accumulating evidence suggests that FBG levels can be used to predict the risk of CVD.20 25 35 As an adverse consequence of CVD, stroke is closely related to FBG levels.36 37 Its mechanism may be related to the dysfunction of cerebral microvasculature caused by hyperglycaemic, insulin resistance, obesity and hypertension.38 The microvasculature is involved in the regulation of many brain processes. Once damaged, it is easy to cause cerebrovascular accidents.39 The dysfunction is also obvious in the pre-diabetes population, which may suggest that IFG is the risk factor for stroke.40 Besides, diabetic retinopathy was proved to be a sign of lacunar ischaemic stroke, which had the same effects as cerebral microvascular dysfunction.41 Compared with men, women with diabetes are at higher risk of stroke.34 Previously, it was thought that the sex specificity of the severity of diabetes-related CVDs was caused by the differences between men and women in treatment and chronic disease management. However, as the gradual equality of men and women, the risk of serious CVD in postmenopausal women is still higher.34 Sex hormones may explain the sex differences observed.

<table>
<thead>
<tr>
<th>Study</th>
<th>Quartile of FBG levels</th>
<th>HR (95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Nielson, 2007</td>
<td>Q1</td>
<td>1.24 (1.05, 1.46)</td>
</tr>
<tr>
<td>Boden-Albala, 2006</td>
<td>Q2</td>
<td>2.70 (1.90, 3.80)</td>
</tr>
<tr>
<td>Sung, 2008</td>
<td>Q3</td>
<td>2.39 (2.13, 2.69)</td>
</tr>
<tr>
<td>Liu, 2007</td>
<td>Q4</td>
<td>1.51 (1.07, 2.13)</td>
</tr>
<tr>
<td>Lee, 2016</td>
<td></td>
<td>1.19 (1.02, 1.38)</td>
</tr>
<tr>
<td>Peng, 2019</td>
<td></td>
<td>1.89 (1.09, 3.29)</td>
</tr>
<tr>
<td>Our Study, 2020</td>
<td></td>
<td>1.44 (1.07, 1.94)</td>
</tr>
<tr>
<td>Combined</td>
<td></td>
<td>1.70 (1.27, 2.29)</td>
</tr>
</tbody>
</table>

**Figure 1** Meta-analysis of fasting blood glucose at the highest vs lowest levels and the risk of developing stroke. The horizontal lines indicate the lower and upper limits of the 95% CI, and the grey squares reflect the HR of each study.
in human epidemiological studies, indicating the oestrogen protection in women, preventing the progression of non-diabetic CVDs at least before menopause. Animal experiments show that the decrease in oestrogen levels in ovariecortexed mice can lead to impaired glycemia, reduced glucose tolerance, hyperinsulinemia and impaired insulin secretion, and then develop into hyperglycaemic and atherosclerosis. However, diabetes seemingly attenuates the protective effect of oestrogen on women in the development of CVD.

Strengths and limitations

The main strengths of this study include the large sample size, measurement of plasma FBG and the complete follow-up. We followed up the cases by recording linkages with hospital discharge diagnoses and the Cause of Death Registry. We used both quartiles and clinical classifications (normal FBG, IFG, hyperglycaemic) to classify FBG levels and obtained similar results. We considered different stroke subtypes as the study outcome. To our knowledge, this study provides the first meta-analysis investigating the association between FBG levels and stroke risk in the general population; therefore, results are likely generalizable to all races.

Our study has several limitations. First, we do not have HbA1c measurements in this study. Among indicators to assess glucose variability, HbA1c is the gold standard for clinical evaluation of long-term blood glucose control, while FBG may overlook the potential effects of postprandial blood glucose levels and other meaningful factors. Reports revealed that HbA1c played an important role in distinguishing people with high cardiovascular risk from those without diabetes. Second, the levels of estradiol and testosterone in the study subjects were not measured, although we speculated that the sex difference was probably due to the protective effect of sex hormones. Third, due to a low number of stroke cases, the real risk association of the premenopausal population may be biased.

CONCLUSIONS

Our study indicates that higher FBG levels are associated with an increased stroke in the general population and meta-analysis corroborated this finding. Furthermore, we observed that the association was more prominent in women than in men. We also observed strong risk associations among postmenopausal or oophorectomy women. FBG levels as a routine and low-cost measurement could be used to identify individuals with higher stroke risk in the future.

Contributors

Conceptualisation: YZ and HZ; methodology: YZ; software, YZ; validation: YZ; formal analysis: YZ; resources: SG, QZ, MY and ZZ; data curation: YZ; administration: HZ; funding acquisition: SG and HZ. All authors have agreed on the validation: YZ; formal analysis: YZ; resources: SG, QZ, MY and ZZ; data curation: YZ; administration: HZ; funding acquisition: SG and HZ. All authors have agreed on the

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Competing interests

None declared.

Patient consent for publication

Not required.

Ethics approval

The study protocol conforms to the ethical guidelines of the Declaration of Helsinki. The baseline survey and record linkages with study outcomes were approved by the Ethics Committee of Fuwai Cardiovascular Hospital, Beijing, China (number 2012-399). After a detailed explanation of our study, we obtained a written informed consent form from each participant.

Provenance and peer review

Not commissioned; externally peer reviewed.

Data availability statement

Data may be obtained from a third party and are not publicly available. The datasets generated and/or analyzed during the current study are not publicly available due to the restrictions of containing information that could compromise the privacy of research participants.

Supplemental material

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