Development and validation of a patient-reported measure of compassion in healthcare: the Sinclair Compassion Questionnaire (SCQ)

Shane Sinclair 1,2,3 Thomas F Hack 4,5 Cara C MacInnis 6,7 Priya Jaggi,1,2 Harrison Boss 6,8 Susan McClement,4 Aynharan Sinnarajah 1,3 Genevieve Thompson 4,9 The COMPASS Research Team

ABSTRACT

Objectives  Compassion is a key indicator of quality care that is reportedly eroding from patients’ care experience. While the need to assess compassion is recognised, valid and reliable measures are lacking. This study developed and validated a clinically informed, psychometrically rigorous, patient-reported compassion measure.

Design  Data were collected from participants living with life-limiting illnesses over two study phases across four care settings (acute care, hospice, long term care (LTC) and homecare). In phase 1, data were analysed through exploratory factor analysis (EFA), with the final items analysed via confirmatory factor analysis (CFA) in phase 2. The Schwartz Center Compassionate Care Scale (SCCCS), the revised Edmonton Symptom Assessment Scale (ESAS-r) and Picker Patient Experience Questionnaire (PPEQ) were also administered in phase 2 to assess convergent and divergent validity.

Setting and participants  633 participants were recruited over two study phases. In the EFA phase, a 54-item version of the measure was administered to 303 participants, with 330 participants being administered the final 15-item measure in the CFA phase.

Results  Both EFA and CFA confirmed compassion as a single factor construct with factor loadings for the 15-item measure ranging from 0.76 to 0.86, with excellent test–retest reliability (intraclass correlation coefficient range: 0.74–0.89) and excellent internal reliability (Cronbach’s alpha of 0.96). The measure was positively correlated with the SCCCS (r=0.75, p<0.001) and PPEQ (r=0.60, p<0.001). Participants reporting higher experiences of compassion had significantly greater well-being and lower depression on the ESAS-r. Patients in acute care and hospice reported significantly greater experiences of compassion than LTC residents.

Conclusions  There is strong initial psychometric evidence for the Sinclair Compassion Questionnaire (SCQ) as a valid and reliable patient-reported compassion measure. The SCQ provides healthcare providers, settings and administrators the means to routinely measure patients’ experiences of compassion, while providing researchers a robust measure to conduct high-quality research.

BACKGROUND

Compassion is ‘a virtuous response that seeks to address the suffering and needs of a person through relational understanding and action’.1 Increasingly, compassion is considered a key indicator of quality care by patients, families, medical associations,2,3 policy makers, healthcare organisations and governments.4–9 In addition to addressing each of the goals of the Framework for Healthcare Improvement,10 recent research has reported that compassion is positively associated with a variety of patient-reported outcomes, the alleviation of healthcare provider (HCP) burnout, positive health
outcomes and improved quality of care ratings.\textsuperscript{4, 6, 11–13} Deficiencies in compassion are associated with increased patient complaints, malpractice suits, healthcare costs, non-disclosure of health information, adverse medical events and patient mortality.\textsuperscript{4, 6, 12–15} Notably, a systemic lack of compassion has been identified as a common and central factor in a number of high-profile national reviews of healthcare failures.\textsuperscript{4–7, 12–15}

Although compassion in healthcare is increasingly mandated by governments and given credence by healthcare organisations, a persistent barrier to improving compassion cited in the literature, has been the lack of a valid and reliable patient-reported measure of compassion.\textsuperscript{5, 11, 15–17} While measures have begun to emerge to address this gap, a recent systematic review revealed that existing measures have significant limitations, including but not limited to: insufficient evidence of internal consistency; validity; test accuracy; reliability and sensitivity.\textsuperscript{11, 18, 19} Currently, there are no patient-reported compassion measures that: (a) adhere to measure development guidelines\textsuperscript{20–22}; (b) were developed from an empirical model of the construct of interest; (c) adequately assess the credibility and transferability of the measure across patient populations and importantly (d) engaged patients across each stage of measure development.\textsuperscript{11} The objective of the current study was to develop and validate a psychometrically rigorous, patient-reported measure of compassion for use in clinical practice and research.

Before embarking on the current study we: conducted a scoping review of compassion\textsuperscript{1}; defined and developed an empirical model of compassion from the perspective of patients\textsuperscript{1} (see online supplemental figure S1) demarcated compassion from sympathy and empathy,\textsuperscript{23} and validated the model among HCPs\textsuperscript{16} (see online supplemental figure S2). The current study began with qualitative interviews with patients to establish the transferability of the model across our study populations and focus groups with HCP, educators and administrators (n=24) to determine the feasibility, challenges, facilitators and clinical utility of the proposed measure (see online supplemental table S1). The results of this first study stage,\textsuperscript{24} along with the findings of our afore-mentioned literature review and model development directly informed the item generation stage of the study,\textsuperscript{25} in accordance with development guidelines.\textsuperscript{20–22} Finally, the content validity of the draft measure was established through a Delphi process with international subject matter experts and patient advisors, along with cognitive interviews with patients.\textsuperscript{26} In addition to revising the initial item pool, this initial validation phase established the: relevancy of each item; representativeness of items to the construct and associated domain of the compassion model; clarity and readability of each item; and the relevancy of the recall period and the response scale. This article reports on the results of the exploratory and confirmatory factor analysis (CFA) stages of this large multi-centred study, producing the final version of the Sinclair Compassion Questionnaire (SCQ).

METHODS

Patient and public involvement

Patients and former patients, who comprised the studies patient advisory group, were involved in the design, or conduct, or reporting, or dissemination plans of this study—including assessing the transferability of the patient model, item development, cognitive interviews and measure development.\textsuperscript{1, 23–26}

METHODS: PHASE 1—EXPLORATORY FACTOR ANALYSIS

Phase 1: study design and study population

To assess the structure and reliability of the draft 54-item compassion measure that emerged from the content validity stage\textsuperscript{26} using EFA, 303 participants were recruited between February 2018 and September 2018 from four care settings in two Canadian cities: acute care (n=105), home care (n=13), long term care (LTC) (n=130) and hospice (n=55) (see online supplemental table S2). Eligibility criteria included: being age 18 years or older; being able to read and speak English; living with a life-limiting illness (eg, cancer, chronic obstructive pulmonary disease, coronary heart disease, dementia); having the cognitive capacity to complete the study (as assessed by the healthcare team) and being able to provide informed consent. Eligible patients/residents (n=713) were identified by a member of the healthcare team and informed of the study. Patients/residents were not referred to the study if they were cognitively impaired, unable to provide informed consent, or too ill to participate. Of the 713 eligible participants informed about the study, 209 (29.3%) chose not to or could not participate for various reasons (too tired n=43; not interested n=92; language barrier n=8; discharged/deceased n=14; other n=13; undisclosed n=39). Of the remaining 504 patients/care residents who agreed to speak to a Research Assistant (RA) about participating, 201 (39.9%) chose not to or could not participate. The remaining 303 participants (response rate=42.5%) were provided with further study details by the RA, who was not a member of the healthcare team and assured patients that their responses would not be shared with members of their healthcare team.

The protocol was administered within 24 hours after obtaining written informed consent (see online supplemental table S2). Participants were asked to self-administer the questionnaire via tablets (adjustable font size) using REDCap survey software, or on paper if preferred. If a participant had difficulty completing the questionnaire (eg, poor eyesight or weakness), the RA administered the questionnaire.

Phase 1: study procedures

Study participants were administered the draft 54-item compassion measure asking them to rate their experience of compassion from their HCPs over the past 7 days using a 5-point Likert scale of agreement (1=strongly disagree, 2=disagree, 3=neutral, 4=agree, 5=strongly agree). Additionally, participants were asked to rate the importance of
each of the 54 items on a 5-point Likert scale (1=not at all important, 2=not very important, 3=somewhat important, 4=very important, 5=extremely important). Test–retest reliability was assessed by having the same RA re-administer the 54-item measure to 65 consenting participants within a 24-hour period.22

Phase 1: data analysis
All data analyses were completed using SPSS Statistical Software (V.24). In an effort to achieve maximum measurement stability with the fewest number of items, the set of 54 initial items emerging from initial validity testing,26 was further reduced based on the test–retest reliability results. Items below an intraclass correlation coefficient (ICC) threshold of 0.70, the upper end of ‘good’ ICC values,27 were considered less stable across testing sessions and were not considered optimal items. They were therefore eliminated, a strategy that has been used in other scale development work.25 26 Exploratory factor analysis (EFA) using a principal axis factor (PAF) extraction with direct oblimin rotation was used to explore the underlying structure of the data.26 Pairwise deletion was used for any missing values. Parallel analysis was used to assess the number of retained factors.34 Refinement of the compassion measure occurred by examining factor loadings, domain coverage within the Patient Compassion Model (see online supplemental figure S1),1 and internal consistency.

RESULTS: PHASE 1—EFA
Participants took approximately 30min to complete the draft 54-item SCQ and importance questions, regardless of whether it was self-administered (n=68) or RA facilitated (n=234), with 71.9% (n=217) of the questionnaires being administered by tablet (see online supplemental table S2). Initial test–retest assessment of the 54 items, resulted in five items achieving an ICC below 0.70 and were therefore discarded (see online supplemental table S3). In order to make an objective decision regarding item retention, the remaining 49 items were subjected to an EFA, using PAF to identify which of a number of alternate worded items to retain (eg, my HCP showed genuine concern vs genuine interest), while removing the lower loaded factor alternate. To first ensure that the data were suitable for factor analysis we conducted the Kaiser-Meyer-Olkin (KMO) measure of sampling adequacy (0.976) and Bartlett’s test of sphericity ($\chi^2=137.797.78$, df=1176, p<0.001). The minimum threshold for acceptability for KMO values in conducting factor analysis is \(<0.50,32\) with a significant Bartlett’s test of sphericity.35 These values indicated that our data were optimal for EFA. The PAF and subsequent analysis of factor loadings between alternate worded items resulted in the removal of 11 redundant items (see online supplemental table S4).

To ensure that the remaining 38 items were adequate for factor analysis, the KMO measure of sampling adequacy (0.974) and Bartlett’s test of sphericity ($\chi^2=10023.92$, df=703, p<0.001) were again assessed and deemed appropriate. The second PAF was conducted (see online supplemental table S5), which resulted in the extraction of a single factor that explained approximately 57% of the variance, yielding an overall Cronbach’s alpha of 0.979. The decision to extract a single factor was determined through the use of parallel analysis, rather than eigenvalues greater than one, or scree plots, due to the tendency of the latter to result in over-extraction.31

To determine the optimal number of measure items, balancing maximal reliability and clinical feasibility, the measure was further refined based on factor loadings, internal reliability and domain coverage within the Patient Compassion Model (see online supplemental figure S1).1 Adhering to methodological guidelines,30 the Cronbach’s alpha for the first four items that loaded highest onto this single factor was calculated and additional items were loaded sequentially, one-by-one, to determine the reliability of the items collectively until all 38 items were loaded to assess the diminishing returns of additional items (figure 1). Through an iterative consensus process among the research team, a cut-off of 15 items with the highest factor loadings was determined as optimal, as it yielded a Cronbach’s alpha of 0.96, while also providing sufficient domain coverage (see online supplemental figure S1), such that at least two items for each domain were included. These methodological decisions were based on an array of considerations, including psychometric soundness, the pragmatics of scale length and existing theoretical considerations, and are consistent with healthcare measure development recommendations.30 Although not explicitly a part of the selection process for these cuts, our decision was further supported by patients’ mean importance ratings for the 15-items, which remained high, ranging from (SD) 4.08 (0.75) to 4.30 (0.60) (table 1). Descriptive statistics on the 15-item SCQ, including means, SD, skewedness, kurtosis, range, test–retest reliabilities and theoretical domain coverage were also assessed (tables 1 and 2).

METHODS: PHASE 2—CFA
Phase 2: study design and study population
Between February 2019 and September 2019, a new sample of 330 participants were recruited to confirm the factor structure obtained at the EFA stage and test convergent and divergent validity, using valid and reliable measures (see online supplemental table S6). Participants were recruited from acute care (n=109), hospice (n=82) and LTC settings (n=139), across two Canadian cities (see online supplemental table S2). Homemare patients were excluded in this phase. The 7-day recall period in the question stem proved to be problematic, as many homemare patients did not require or receive an in-person visitation from a member of the homemare team over the course of a week. Of the 710 eligible (same criteria as phase 1) individuals, 176 (24.8%) chose not to, or could not, participate for various reasons (too
Compassion was also assessed using the 12-item Schwartz Center Compassionate Care Scale (SCCCS). The SCCCS question stem (‘During your recent hospitalization, how successfully did your doctor …’) was modified to align with the current study’s recall period and multi-provider focus (‘In thinking about your Healthcare Providers over the past seven days, how successfully did they …’). Next, symptom distress was assessed using the 9-item revised Edmonton Symptom Assessment Scale (ESAS-r). Finally, patient satisfaction was assessed using the 13-item PICKER Patient Experience Questionnaire (PPEQ). It was hypothesised that scores on the SCCCS would be

Table 1: Descriptive statistics for the final 15-item Sinclair Compassion Questionnaire based on phase 1 (exploratory factor analysis) sample

<table>
<thead>
<tr>
<th>Item</th>
<th>N</th>
<th>Min</th>
<th>Max</th>
<th>Mean</th>
<th>SD</th>
<th>Skewness</th>
<th>Kurtosis</th>
<th>Mean importance rating (SD)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Feel cared for</td>
<td>303</td>
<td>1</td>
<td>5</td>
<td>4.15</td>
<td>0.76</td>
<td>−0.95</td>
<td>1.70</td>
<td>4.24 (0.68)</td>
</tr>
<tr>
<td>Genuine concern</td>
<td>302</td>
<td>1</td>
<td>5</td>
<td>4.12</td>
<td>0.75</td>
<td>−0.82</td>
<td>1.10</td>
<td>4.20 (0.68)</td>
</tr>
<tr>
<td>Communicated sensitive</td>
<td>303</td>
<td>1</td>
<td>5</td>
<td>4.14</td>
<td>0.74</td>
<td>−0.77</td>
<td>0.98</td>
<td>4.17 (0.64)</td>
</tr>
<tr>
<td>Attentive</td>
<td>303</td>
<td>1</td>
<td>5</td>
<td>4.14</td>
<td>0.72</td>
<td>−0.70</td>
<td>1.08</td>
<td>4.30 (0.60)</td>
</tr>
<tr>
<td>Provided comfort</td>
<td>303</td>
<td>1</td>
<td>5</td>
<td>4.07</td>
<td>0.73</td>
<td>−0.89</td>
<td>1.68</td>
<td>4.23 (0.62)</td>
</tr>
<tr>
<td>Very supportive</td>
<td>301</td>
<td>1</td>
<td>5</td>
<td>4.18</td>
<td>0.66</td>
<td>−0.63</td>
<td>1.46</td>
<td>4.19 (0.62)</td>
</tr>
<tr>
<td>Provided care</td>
<td>302</td>
<td>2</td>
<td>5</td>
<td>4.19</td>
<td>0.65</td>
<td>−0.65</td>
<td>1.21</td>
<td>4.25 (0.58)</td>
</tr>
<tr>
<td>Spoke with kindness</td>
<td>303</td>
<td>1</td>
<td>5</td>
<td>4.24</td>
<td>0.69</td>
<td>−1.02</td>
<td>2.79</td>
<td>4.27 (0.60)</td>
</tr>
<tr>
<td>Saw as person</td>
<td>302</td>
<td>1</td>
<td>5</td>
<td>4.09</td>
<td>0.78</td>
<td>−0.87</td>
<td>1.03</td>
<td>4.23 (0.64)</td>
</tr>
<tr>
<td>Behaved in caring way</td>
<td>300</td>
<td>2</td>
<td>5</td>
<td>4.20</td>
<td>0.66</td>
<td>−0.59</td>
<td>0.77</td>
<td>4.28 (0.57)</td>
</tr>
<tr>
<td>Really understood needs</td>
<td>303</td>
<td>1</td>
<td>5</td>
<td>3.91</td>
<td>0.88</td>
<td>−0.76</td>
<td>0.48</td>
<td>4.24 (0.66)</td>
</tr>
<tr>
<td>Good relationship</td>
<td>303</td>
<td>1</td>
<td>5</td>
<td>4.17</td>
<td>0.75</td>
<td>−0.88</td>
<td>1.24</td>
<td>4.28 (0.56)</td>
</tr>
<tr>
<td>See my perspective</td>
<td>303</td>
<td>1</td>
<td>5</td>
<td>3.85</td>
<td>0.89</td>
<td>−0.55</td>
<td>0.05</td>
<td>4.08 (0.75)</td>
</tr>
<tr>
<td>Warm presence</td>
<td>302</td>
<td>2</td>
<td>5</td>
<td>4.11</td>
<td>0.70</td>
<td>−0.51</td>
<td>0.34</td>
<td>4.15 (0.65)</td>
</tr>
<tr>
<td>Sincere</td>
<td>303</td>
<td>2</td>
<td>5</td>
<td>4.15</td>
<td>0.71</td>
<td>−0.61</td>
<td>0.40</td>
<td>4.21 (0.55)</td>
</tr>
</tbody>
</table>
strongly and positively associated with the SCQ scores, whereas scores from the discriminant constructs such as symptom distress (ESAS-r) and patient satisfaction (PPEQ) would be positively associated with the SCQ, but less so.

Phase 2: data analysis

The initial factor structure identified via EFA was tested through CFA, using AMOS (V.24) with missing data excluded listwise. Factor loadings were explored, and global model fit was evaluated by examining model $\chi^2$, comparative fit index (CFI), root-mean-square-error of approximation (RMSEA) and standardised root-mean-squared residual (SRMR). Ideal fit criteria for these indices include non-significant $\chi^2$ tests, CFI above 0.95, SRMR values less than 0.08, and RMSEA values below 0.08. It is important to note, that model $\chi^2$ tests are highly sensitive to minor deviations from perfect model fit and sample size, and are often found to be statistically significant despite an excellent fitting model. Potential areas of misfit were therefore improved by including several covariances between residuals in the measurement structure. Item response theory (IRT) was also used to evaluate the relationships between the latent construct of compassion and the items intended to measure the construct. Missing values for observation were excluded from the analysis and did not contribute to the likelihood estimates in the IRT analysis.

Validity evidence between overall scores from the SCQ and those from the measures of convergent and discriminant validity were explored using Pearson correlations. For exploratory purposes, potential differences in overall SCQ scores as a function of study sites (acute care, hospice and LTC) and sociodemographic variables were examined using an analysis of variance (ANOVA). Any missing values for these analyses were treated with pairwise deletion.

RESULTS: PHASE 2—CFA AND RELIABILITY

Participants took approximately 20 min to complete the protocol, with 259 (78.5%) of the participants requiring the assistance of an RA. The 15-item SCQ took participants between 3 and 5 min to complete, with 224 (67.9%) being completed via tablet and 21.6% being self-administered (see online supplemental table S2). Only participants with complete data on the SCQ compassion items were included in the CFA (N=327).

Cronbach’s alpha for the 15-item measure in phase 2 was 0.96. Given the single factor solution revealed at the EFA stage, a single latent compassion factor was specified with loadings from each of the 15 items. Initial model estimation revealed strong standardised factor loadings ranging between 0.75 and 0.86, with some global fit indices initially being unsatisfactory. Global fit improved by adding covariances to the model, yielding standardised factor loadings that remained strong, between 0.76 and 0.86 (figure 2). Overall, these results support a single factor of compassion.

IRT analyses indicated that the SCQ precisely measures compassion across the wide range of patients’ experiences with their HCPs. The average marginal reliability of the SCQ was 0.85. The parameter estimates were reasonable and their standard errors were small. For each of the SCQ items, the item discrimination parameter was statistically significant, implying that each item can differentiate between individuals with different ratings of compassion. Further analysis indicated that participants who felt that

| Table 2 Test-retest reliability and theoretical domain coverage for the final 15-item Sinclair Compassion Questionnaire |
|-----------------|-----------------|--------------------------------------|
| Item            | Intraclass correlation | Patient Compassion Model domain      |
| Feel cared for  | 0.74***          | Relational communication             |
| Genuine concern | 0.78***          | Relational communication             |
| Communicated sensitive | 0.78***   | Relational communication             |
| Attentive       | 0.77***          | Virtuous response                    |
| Provided comfort| 0.74***          | Attending to needs                   |
| Very supportive | 0.86***          | Relational communication             |
| Provided care   | 0.83***          | Attended to needs                    |
| Spoke with kindness | 0.83***    | Relational communication             |
| Saw as person   | 0.89***          | Seeking to understand                |
| Behaved in caring way | 0.79***   | Relational communication             |
| Really understood needs | 0.87*** | Seeking to understand                |
| Good relationship| 0.81***          | Relational space                     |
| See my perspective | 0.82***    | Relational communication             |
| Warm presence   | 0.86***          | Relational space                     |
| Sincere         | 0.87***          | Virtuous response                    |

***p<0.001, **p<0.01, *p<0.05.
Mean time interval=24 hours.
their HCPs were very compassionate were highly likely to select ‘strongly agree’ to relevant SCQ items, while those who believed that their HCPs were less compassionate were likely to select the appropriate response; providing further confidence in the reliability and precision of the SCQ. Descriptive statistics for the final 15-item SCQ based on phase 2 (CFA) sample can be referred to in table 3.

When assessing convergent validity, the SCCCS ($\alpha=0.97$) and the SCQ were strongly positively correlated, $r=0.75$, $p<0.001$, providing support that the SCQ taps into the construct of compassion. When assessing for divergent validity, the SCQ and the PPEQ ($\alpha=0.88$) were positively correlated with moderately high strength ($r=0.60$, $p<0.001$). This provides evidence that the SCQ is related to, but sufficiently distinct from, patient satisfaction. Finally, the relationships between the SCQ and each individual ESAS-r symptom were assessed for evidence of divergent validity. The SCQ was significantly associated with depression and well-being, with individuals who reported lower depression or greater well-being having higher compassion scores. Compassion was, however, weakly and

<table>
<thead>
<tr>
<th>Item</th>
<th>N</th>
<th>Min</th>
<th>Max</th>
<th>Mean</th>
<th>SD</th>
<th>Skewness</th>
<th>Kurtosis</th>
</tr>
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<tbody>
<tr>
<td>Feel cared for</td>
<td>330</td>
<td>1</td>
<td>5</td>
<td>4.31</td>
<td>0.87</td>
<td>-1.53</td>
<td>2.62</td>
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<td>Genuine concern</td>
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<td>4.28</td>
<td>0.79</td>
<td>-1.03</td>
<td>0.90</td>
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<tr>
<td>Communicated sensitive</td>
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<td>4.17</td>
<td>0.87</td>
<td>-1.20</td>
<td>1.63</td>
</tr>
<tr>
<td>Attentive</td>
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<td>5</td>
<td>4.21</td>
<td>0.83</td>
<td>-1.02</td>
<td>0.84</td>
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<tr>
<td>Provided comfort</td>
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<td>-1.17</td>
<td>1.44</td>
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<tr>
<td>Very supportive</td>
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<td>5</td>
<td>4.23</td>
<td>0.79</td>
<td>-1.03</td>
<td>1.30</td>
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<tr>
<td>Provided care</td>
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<td>4.23</td>
<td>0.82</td>
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<td>Spoke with kindness</td>
<td>330</td>
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<td>5</td>
<td>4.31</td>
<td>0.72</td>
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<td>Saw as person</td>
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<td>5</td>
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<td>0.83</td>
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<td>Really understood needs</td>
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<td>3.98</td>
<td>0.89</td>
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<td>0.68</td>
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<td>Good relationship</td>
<td>330</td>
<td>1</td>
<td>5</td>
<td>4.31</td>
<td>0.72</td>
<td>-1.17</td>
<td>2.18</td>
</tr>
<tr>
<td>See my perspective</td>
<td>330</td>
<td>1</td>
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<td>3.95</td>
<td>0.88</td>
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<td>Warm presence</td>
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<td>4.21</td>
<td>0.77</td>
<td>-1.03</td>
<td>1.69</td>
</tr>
<tr>
<td>Sincere</td>
<td>328</td>
<td>1</td>
<td>5</td>
<td>4.25</td>
<td>0.79</td>
<td>-1.23</td>
<td>2.04</td>
</tr>
</tbody>
</table>
negatively associated with depression (r = −0.13, p = 0.02), and well-being (r = −0.17, p = 0.002), with higher depression scores on the ESAS-r indicating greater depression and higher well-being scores indicating worse well-being. No additional relationships between the SCQ and other symptoms measured by the ESAS-r (eg, pain, fatigue, anxiety, sleep, etc) reached statistical significance. This provides support that the SCQ construct of compassion is distinct from symptom distress.

To further validate the SCQ, we assessed whether SCQ scores varied as a function of demographic variables or case setting. ANOVAs yielded significant mean differences on the SCQ by gender (p = 0.784), marital status (p = 0.403), education (p = 0.240), ethnicity (p = 0.551) or spirituality/religiousness (p = 0.589). A weak, negative correlation between age and reported SCQ scores was identified (r = −0.13, p = 0.021). Significant differences emerged between compassion scores and care location, F(2,327) = 16.62, p < 0.001. Post-hoc Tukey’s tests revealed that individuals in acute care (M = 4.39) reported significantly higher (p < 0.001) levels of compassion than those in LTC (M = 3.97). Similarly, participants in hospice (M = 4.37) reported significantly greater levels of compassion (p < 0.001) than those in LTC. No differences were found between those in acute care and hospice settings (p = 0.975). Given that compassion scores were influenced by age and care location, exploratory follow-up analyses were conducted to examine (a) potential age differences by care location and (b) differences in compassion by care location when statistically controlling for age. Age differed by care location, F(2,317) = 27.55, p < 0.001, with acute care participants (M = 66.89) being significantly younger than LTC residents (M = 79.37, p < 0.001) and hospice patients (M = 75.61, p < 0.001). However, when examining compassion by care location statistically controlling for age, results remained consistent with those reported earlier, F(3,316) = 11.65, p < 0.001, with compassion being higher in acute care and hospice settings versus LTC, p < 0.001. A separate secondary CFA analysis was also conducted to assess the validity and reliability of a 5-item short-form version (SCQ-SF), in order to provide further flexibility and utility to clinicians and survey administrators wanting to embed a measure of compassion in their clinical assessments and patient/family surveys, without compromising psychometric rigour (see online supplemental table S7 and figure S8).

**INSTRUCTION MANUAL, SCORING GUIDELINES AND VERSIONS OF THE SCQ**

Scoring of the SCQ is performed by calculating the mean score of all the items, with an overall score being indicative of greater reported compassion. The SCQ, a detailed instruction manual including administration, data entry, response coding and scoring instructions is available from the authors or at www.compassionmeasure.com. The SCQ French version (Questionnaire de Compassion Sinclair) is also available, with a Spanish version forthcoming.

**DISCUSSION**

This study validated a patient-reported measure of compassion, the SCQ, providing researchers, clinicians, survey administrators and healthcare organisations a clinically informed, patient-reported measure of compassion that has strong initial evidence supporting its validity and reliability. The SCQ contains items that cover patients’ experiences of compassion within each of the theoretical domains of the Patient Compassion Model with our results showing that these domains are subsumed under a single latent construct of compassion. These results are a defining feature of reflective measures, whereby individual items each reflect the underlying construct, underscoring the necessity of conducting foundational research and initial validation studies to establish construct validity—an essential, but overlooked stage in the development of compassion measures and measure in general. As a result, the SCQ has excellent internal consistency (Cronbach’s alpha of 0.96) and test-retest reliability (ranging from 0.74 to 0.89).

Results demonstrate that the SCQ overcomes the inherent limitations of previous compassion measures in healthcare. Notably, the SCQ is a patient-reported measure of compassion (vs a clinician or proxy assessment) that rigorously adheres to measure development guidelines and uses appropriate response scales for the construct of interest (agreement vs frequency or satisfaction). Further, the perspectives of patients were incorporated across each stage of this study, fortifying our foundational patient-centred research that defined the construct of interest and its associated domains, including how compassion is delineated and preferred over similar constructs such as sympathy and empathy—which are often conflated within existing compassion measures. This is critical, as while compassion subsumes and enhances elements of empathy, patient-centred care and clinical communication, it is a separate construct that includes action and the virtues of HCPs—providing a multimodal and optimal effect on various patient outcomes that do not come at the cost of HCP well-being.

In addition, the SCQ provides healthcare organisations with the means to assess patient experiences of compassion alongside other quality care indicators, patient experience measures and satisfaction surveys such as the PPEQ. Reporting patients’ compassion scores at the unit, institutional and systems-level allows researchers and health administrators to identify benchmarks and analyse associations between compassion and other key variables (eg, patient satisfaction, workplace well-being, adverse medical events and between patient groups). Further, a valid patient-reported compassion measure provides investigators the opportunity to study distributions of compassion preferences across various populations and care settings. While the 15-item SCQ is brief and easy to complete, the 5-item SCQ-SF provides additional flexibility for clinicians who want to assess and improve compassion in their professional practice.
Secondary analysis revealed compassion scores were significantly lower among those in LTC and were weakly and negatively correlated with age. After controlling for age, a negative statistically significant difference remained between SCQ scores in LTC and other settings. Possible explanations for lower SCQ scores in LTC may be due to differing staff–patient ratios, levels of acuity or suffering, or differences in practice culture; however, further research is required to examine this difference.

**Strengths and limitations**

This study is not without limitations. First, while the establishment of the recall period was determined by the construct being measured; measure development guidelines, and was validated through a Delphi process with international SMEs and cognitive interviews with patients, the SCQ may have lost some specificity, as we chose a 7-day recall period that asked respondents to consider all interactions with their HCPs during that period. This decision was intentional and evidence-informed, as patients, members of our HCP focus groups, subject matter experts and members of our patient advisory group felt that a 7-day period was most appropriate for our study populations. The same data informed our decision to measure patients overall experience of compassion from all HCPs they interacted with, rather than care from a single HCP at a specific timepoint—a decision that is further supported by guidelines for measuring quality care within the patient experience. Second, while we anticipate that the SCQ is transferable to other patient populations and while we validated the measure across diverse care settings, in sampling patients with a life limiting illness, the generalisability of the measure to other patient populations requires further research, for example, in the general population, younger patients and other healthcare contexts. Third, while we attempted to validate the measure within a homecare setting, the question stem proved to be problematic, resulting in this cohort being removed from the CFA stage of the study. While it is anticipated that the SCQ is adaptable to other settings, as the wording of individual items is not context or disease specific, further research is needed to assess the validity of adapted versions of the SCQ with modified question stems. Fourth, although the strength of the SCQ as a brief, patient-reported measure addresses many of the limitations of previous measures, due to the acuity of our patient population, most patients required the aid of an RA to complete the protocol (78% in both study phases). As such, when patients are unable to self-administer the SCQ, we recommend that it be administered by someone other than the patients’ HCPs (eg, a patient care manager or RA) in order to mitigate response and social desirability bias. Fifth, in slightly modifying the recall period and the question stem of the measures we used for convergent and divergent validity with the SCQ, this may have influenced the fidelity and results from these measures (see online supplemental table S6).

While the initial evidence for construct validity of the SCQ is encouraging, additional evidence is needed to establish its responsiveness, interpretability and criterion validity.

**CONCLUSIONS**

The SCQ is a reliable and valid patient-reported compassion measure for research and practice. The SCQ will allow HCPs, practice settings, institutions and healthcare systems to routinely assess and improve compassion, while providing researchers the means to conduct empirical research on this important care construct.

**Author affiliations**

1Faculty of Nursing, University of Calgary, Calgary, Alberta, Canada
2Compassion Research Lab, University of Calgary, Calgary, Alberta, Canada
3Department of Oncology, Cumming School of Medicine, University of Calgary, Calgary, Alberta, Canada
4College of Nursing, Rady Faculty of Health Sciences, University of Manitoba, Winnipeg, Manitoba, Canada
5Psychosocial Oncology and Cancer Nursing Research, IH Asper Clinical Research Institute, Winnipeg, Manitoba, Canada
6Department of Psychology, University of Calgary, Calgary, Alberta, Canada

**Twitter** Shane Sinclair @ShaneASinclair, Aynharan Sinnarajah @DrASinnarajah, Genevieve Thompson @DRThompson_GN and The COMPASS Research Team @CRLinaction

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**Collaborators** The COMPASS Research Team: Harvey Chochinov, Shelly Cory, Neil Hagen, Max Jajszczak, Leah Lechelt, Carlo Leget, Christina Puchalski, Patrick Quail, Lucy Selman.

**Contributors** SS conceptualised the study, offered funding, developed, refined and determined the final wording of the SCQ items, oversaw data collection and analysis across sites in all stages of measure development, wrote the final manuscript and was responsible for all stages of the study. TFH conceptualised the study, obtained funding, developed, refined and determined the final wording of the SCQ items, oversaw data collection at the Winnipeg sites, contributed to data analysis and the final manuscript. CCM oversaw data analysis and interpretation in the EFA and CFA stages and contributed to the final manuscript. PJ coordinated the entire study, managed data across all sites, contributed to the development and refinement of the final wording of the SCQ items, contributed to data analysis and the final manuscript. HB conducted the data analysis and interpretation in the EFA and CFA stages and contributed to the final manuscript. SM contributed to the conceptualisation of the study, supported data collection at the Winnipeg sites and contributed to the final manuscript. GT contributed to the conceptualisation of the study, supported data collection at the Winnipeg sites and contributed to the final manuscript. AS contributed to the conceptualisation of the study, supported data analysis.
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**ORCID iDs**
- Sinclair S http://orcid.org/0000-0003-4542-9911
- Thomas F Hack http://orcid.org/0000-0002-6913-8732
- Cara C MacInnis http://orcid.org/0000-0001-5418-1626
- Harrison Boss http://orcid.org/0000-0003-4472-1048
- Aynharan Sinnarajah http://orcid.org/0000-0002-7367-159X
- Genewieve Thompson http://orcid.org/0000-0003-2559-9926

**REFERENCES**

39. Hu Li-tze, Bentler PM. Cutoff criteria for fit indexes in covariance structure analysis: conventional criteria versus new alternatives.
Structural Equation Modeling: A Multidisciplinary Journal
