BMJ Open Effects of maternal smoking on body size and proportions at birth: a register-based cohort study of 1.4 million births

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ABSTRACT

Objectives The aim of our work was to analyse the effect of maternal smoking on body size and body proportions of newborns when the mother had smoked only during the first trimester, in comparison with continued smoking after the first trimester. Furthermore, we have evaluated how growth restriction associated with maternal smoking contributes to changes in body proportions.

Design Register-based cohort study

Setting Maternal Exposure (MATEX) cohort identified from the Finnish Medical Birth Register.

Participants Singleton births without congenital anomalies and missing data (1.38 million) from 1 January 1991 to 31 December 2016.

Methods Logistic regression was used to quantify the effect of maternal smoking, stratified by the maternal smoking status.

Outcome measures Body proportions indicated by low brain-to-body ratio (defined as <10th percentile); high ponderal index and high head-to-length ratio (defined as >90th percentile); small body size for gestational age at birth (defined as weight, length or head circumference <10th percentile) and preterm birth (<37 weeks) and low birth weight (2500 g).

Results Continued smoking after the first trimester was associated with high ponderal index (OR 1.26, 95% CI 1.23 to 1.28), low brain-to-body ratio (1.11, 1.07–1.15) and high head-to-length ratio (1.22, 1.19–1.26), corresponding with absolute risks of 22%, 10% and 19%, respectively. The effects were slightly lower when smoking had been quit during the first trimester. Similar effects were seen for the body size variables and low birth weight. Preterm birth was not associated with smoking only during first trimester.

Conclusions Maternal smoking, independent of smoking duration during pregnancy, was associated with abnormal body proportions resulting from larger reduction of length and head circumference in comparison to weight. The effects of having quit smoking during the first trimester and having continued smoking after the first trimester were similar, suggesting the importance of early pregnancy as a sensitive exposure window.

INTRODUCTION

Smoking during pregnancy increases the risk for adverse pregnancy outcomes. Adverse pregnancy outcomes are not only associated

Strengths and limitations of this study

- The register-based design of this study provided a big study size to detect small risks and minimises risks for recall bias.
- The register-based design of this study allowed for sensitivity analyses including stratification by socio-economic status and birth year, as well as testing of additional adjustment models for sociodemographic factors and comorbidities.
- The data content of the Finnish Medical Birth Register has been validated for accuracy and completeness.
- Smoking status was self-reported during antenatal visits, leading to possible reporting bias.
- The register-based design restricted availability of information on confounders. Thus, lifestyle-related confounders, such as alcohol consumption and exposure to secondhand tobacco smoke, could not be adjusted for.

with complications in the neonatal period, but also much later, potentially in late adulthood.¹ Tobacco smoke contains thousands of chemicals, which can cross the placenta and enter fetal circulation. Among them, nicotine has a multitude of adverse effects on the development of organs including brain.² Other well-known toxic chemicals in tobacco smoke include carbon monoxide, which can interfere with oxygen supply of the unborn child, as well as genotoxic and carcinogenic polycyclic aromatic hydrocarbons and tobacco-specific nitrosamines, which are teratogenic in animal studies.³

The association between maternal smoking and low birth weight, commonly defined as weight below 2500 g, is well established. In addition, data on the susceptibility of anthropometric indices, such as body length, head size and abdominal circumference to maternal smoking is emerging.⁴ Low birth weight as such does not hold information whether the reduction of weight is due to loss of lean or fat body mass. Similarly, reduction in any anthropometric index alone fails to

identify abnormal body proportions. Symmetrical growth restriction in utero is a stronger risk factor for later life morbidity and mortality than asymmetrical growth restriction with decreased amount of fat tissue.⁵ Small for gestational age, used as a substitute for in utero growth restriction, is not an optimal proxy.⁶ This clearly demonstrates the importance of body proportions for future health of the newborn.

The effects of early smoking cessation on body size are not well understood. Smoking only during early pregnancy has been shown to be less harmful on body size than continued smoking during late pregnancy. Previous small studies have indicated that anthropometric indices in newborns exposed only during early pregnancy are similar to those in newborns of non-smoking mothers.⁷ On the other hand, current bigger studies report increased risk for growth restriction even in fetuses exposed only during the first trimester.⁸ There are insufficient data about anthropometric indices, other than birth weight, in newborns exposed only during early pregnancy in comparison with newborns of non-smokers.⁹

The aim of our work was to analyse the effect of maternal smoking on body size and body proportions of newborns when the mother had smoked only during the first trimester, in comparison with newborns of mothers having continued smoking after the first trimester. Furthermore, we studied how growth restriction associated with maternal smoking contributes to abnormal body proportions. For this, we used the MATEX cohort identified from the Finnish Medical Birth Register.¹⁰ Additionally, we discuss the possibility of mechanistic interpretations of differences in body proportions in newborns of smoking mothers compared with non-exposed newborns.

MATERIALS AND METHODS

Study design

To study the effects of maternal smoking on body size and proportions at birth, we conducted a register-based cohort study using the Finnish MATEX cohort. The MATEX cohort was identified from the Finnish Medical Birth Register. It is described in more detail elsewhere.¹⁰ The Finnish Medical Birth register contains perinatal outcomes, pregnancy characteristics and sociodemographic information for all live births and stillbirths after the 22nd gestational week or with a birth weight of at least 500 g. The Medical Birth Register receives information from standardised forms filled out by nurses and midwives during antenatal care visits and after the delivery of the baby.

This work focused on the effects of maternal smoking on singleton pregnancies delivered between the 1 January 1991 and the 31 December 2016. From initial 1.75 million mother–child pairs, 1.38 million were included in the analyses after exclusion of multiple births, newborns with congenital anomalies and newborns with missing information on maternal smoking status or the covariates (see online supplementary material, figure S1).

Within the MATEX birth cohort, information on head circumference and maternal weight and height, as well as maternal comorbidities was available only for the subpopulation born between 2004 and 2016. The subpopulation included 659 157 mother–child pairs (see online supplementary material, figure S1). Thus, the analyses of endpoints related to head size (small head circumference, brain-to-body ratio and head-to-length ratio), as well as sensitivity analyses, were limited to the smaller subpopulation.

Exposure

Maternal smoking data are recorded in the Medical Birth Register during antenatal care visits as reported by the pregnant women. In the MATEX cohort, smoking status during pregnancy was assigned as three categories: (1) non-smoker; (2) quit smoking during the first trimester and (3) continued smoking after the first trimester. The trends in smoking during the study period have been described in detail elsewhere.¹¹

Outcomes

Four groups of outcomes were included: (1) preterm birth (PTB); (2) low birth weight (LBW) (as a crude measure of small body size); (3) small body size for gestational age and (4) body proportions (table 1).

PTB was defined as a birth before gestational week 37, and term birth as birth during gestational week 37 or later.

LBW was categorised in accordance with the International Classification of Diseases 10th Revision (ICD-10) diagnostic criteria as weight below 2500 g. As the reference category normal weight was defined as 2500–4500 g, excluding high birth weight according to ICD-10 definition (>4500 g).

In this work, we use SGA as a general expression to describe the small size of the newborn. It was defined as measurement of body weight, body length or head circumference below a cut-off at 10th percentile, while the normal range defined as above the 10th percentile. Percentiles were defined based on sex-specific and parity-specific mean and SD for the corresponding gestational age (in weeks) at birth as reported in the Finnish standard reference population.¹² It was included as an endpoint to take into account the impact of gestational age on body size.

Body proportionality was assessed by three anthropometric indices in relation with each other: ponderal index, brain: body ratio and head: length ratio. Percentiles of each ratio have been separately estimated for each gestational age (in weeks) at births in the study population. The 10–90th percentiles were categorised as normal and used as reference, while the tails of the distribution (<10th percentile and >90th percentile) were categorised as abnormal.

Ponderal index was calculated using birth weight and body length (Equation 1). It was categorised normal (10–90th percentile of the study population, used as the

Table 1 Summary of the definitions of endpoints and adjustments in the regression model used to study the association of maternal smoking and preterm birth, low birth weight, small for gestational age and abnormal body proportions at birth

Endpoint	Definition case	Definition control	No of mother-child pairs included in the regression *		Adjustment in stratified multiple logistic regression
			Quit during first trimester †	Continued after first trimester ‡	
Preterm birth	Gestational age <37 weeks	Gestational age ≥37 weeks	1 210 410	1 286 667	Maternal age, sex, parity, SES
Low birth weight	Birth weight <2500 g	Birth weight 2500–4500 g	1 170 187	1 328 221	Maternal age, sex, parity, gestational weeks, SES
Small for gestational age					
Weight	<10th percentile of weight for gestational age (weeks) birth	≥10th percentile of weight for gestational age (weeks) birth	1 210 048	1 327 783	Maternal age, sex, parity, SES
Body length	<10th percentile of length for gestational age (weeks) birth	≥10th percentile of length for gestational age (weeks) birth	1 210 048	1 327 783	Maternal age, sex, parity, SES
Head circumference	<10th percentile of head circumference for gestational age (weeks) birth	≥10th percentile of head circumference for gestational age (weeks) birth	573 343	601 407	Maternal age, sex, parity, SES
Abnormal body proportions					
Ponderal index	>90th percentile of weight-to-length ratio for gestational age (weeks) birth	10–90th percentile of weight-to-length ratio for gestational age (weeks) birth	1 088 451	1 196 479	Maternal age, sex, parity, SES, weight z-score
Brain-to-body ratio §	<10th percentile of weight-to-head circumference ratio for gestational age (weeks) birth	10–90th percentile of weight-to-head circumference ratio for gestational age (weeks) birth	518 704	541 549	Maternal age, sex, parity, SES, weight z-score
Head-to-length ratio §	>90th percentile of head circumference-to-length ratio for gestational age (weeks) birth	10–90th percentile of head circumference-to-length ratio for gestational age (weeks) birth	521 420	547 597	Maternal age, sex, parity, SES

*The number of mother-child pairs is the sum of non-exposed and exposed pregnancies; the number varies for outcomes due to exclusion of mother-child pairs with missing information on the exposure, outcome or any of the confounding factors included in the multivariate regression.

†Comparison quit smoking during the first trimester with no smoking during pregnancy.

‡Comparison continued smoking after the first trimester with no smoking during pregnancy.

§Available for the years 2004–2016.

SES, socioeconomic status.

reference) and high (>90th percentile). Newborns below the 10th percentile were excluded.

$$\text{Ponderal Index} = 100 \times \frac{\text{birth weight (g)}}{\text{body length (cm)}^3} \quad (1)$$

Brain-to-body ratio was calculated based on head circumference and birth weight (Equation 2). It was categorised as low (<10th percentile of the study population) and normal (10-90th percentile, reference). Newborns above the 90th percentile were excluded.

$$\text{Brain - to - Body Ratio} = 100 \times \frac{0.037 \times \text{head circumference (cm)}^{2.57}}{\text{birth weight (g)}} \quad (2)$$

The nominator of the formula is the estimation of the brain weight according to the National Institute of Neurological and Communicative Disorders and Stroke's Collaborative Perinatal Project.¹³

Head-to-length ratio was calculated using head circumference and body length (Equation 3). It was categorised normal (10-90th percentile of the study population, reference) and high (>90th percentile). Newborns below the 10th percentile were excluded.

$$\text{Head - to - Length Ratio} = \frac{\text{head circumference (cm)}}{\text{body length (cm)}} \quad (3)$$

Covariates

Maternal age and gestational age in weeks were used as continuous variables in the regression models. Parity was defined as nulliparous or multiparous. Sex was defined as male or female. Socioeconomic status (SES) was categorised as upper white collar (upper level employees with administrative, managerial, professional and related occupations), lower white collar (lower level employees with administrative and clerical occupations), blue collar (manual workers) and others (farmers, self-employed, students, pensioners), based on the Finnish national classification of occupations.¹⁴ An additional category (information missing) was added to this classification.

Statistical analyses

Multiple logistic regressions were performed to estimate ORs with 95% CIs. The regressions were stratified by exposure status with no smoking as reference, that is, smoking after the first trimester was compared with no smoking, and separately smoking only during the first trimester was compared with no smoking. The regression models were adjusted for potential confounders (table 1; online supplementary material, chapter 3.2). The potential confounding factors were selected based on a combination of available data and previously published factors that could affect both maternal smoking and the outcome measures. The data were analysed using R Statistical software V.3.4.3. The study power was estimated as the smallest detectable risk ratio. The calculations were done using R Statistical Software epiR package with assumed 95% CI and a study power of 90% (see online supplementary material, table S1).

Register data permit

Due to the full register-based design of the study, non-informed consent was required from the study participants according to the Finnish Personal Data Act 1050/2018.

Patient and public involvement

No patients were involved in the design, recruitment or conduct of the study. The used register data are routinely collected in Finland.

RESULTS

Of all women with singleton births included in this study (n=1 376 778), 84.5% were non-smokers, 3.5% quit smoking during the first trimester and 12.0% continued smoking after the first trimester. Smoking pregnant women were younger and more likely to be nulliparous (table 2, supplementary material, table S2).

Any maternal smoking was associated with an increased risk for SGA and abnormal body proportions, while PTB was only associated with smoking continued after the first trimester (figure 1; online supplementary material, table S3).

Any maternal smoking increased the risk for LBW (<2500 g). Smoking continued after the first trimester in comparison to no smoking was associated with twice as high a risk for LBW (OR 2.22, 95% CI 2.14 to 2.30). Smoking only during the first trimester was also associated with an increased risk for LBW compared with non-smokers (OR 1.10, 95% CI 1.02 to 1.19) although not as strong as with continued smoking (figure 1A).

Any smoking during pregnancy was associated with an increased risk for weight or body length at birth being below the 10th percentile. Mothers who quit smoking during the first trimester were at elevated but not statistically significant risk for giving birth to a child with a small head circumference. Among the newborns of the mothers who continued smoking after the first trimester the risk for small head circumference was clearly increased with an OR of 1.64 (95% CI 1.60 to 1.68) (figure 1A).

The risk for abnormal body proportions of newborns was significantly increased by any maternal smoking. A stronger increase in risk was observed for high ponderal index and high head-to-length ratio than for low brain-to-body ratio. ORs were consistently higher if smoking was continued after the first trimester than when the mothers quit smoking during the first trimester. Nevertheless, smoking only during the first trimester was associated with statistically significantly increased risks for abnormal body proportions in newborns. Especially the risk for brain-to-body ratio was almost similar in those exposed only during the first trimester and those exposed also after the first trimester (figure 1B).

We stratified the analysis by SES to investigate the influence of lifestyle factors and health behaviours that correlate with SES. Stratification by SES did not result in statistically significant differences in the risk estimates between the SES groups (see online supplementary

Table 2 Baseline characteristics by smoking status for singleton births in the MATEX cohort (1991–2016)

	All 1376778	Non-smoker 84.5% (1 163 225)	Quit smoking during first trimester 3.5% (47 819)	Continued smoking after first trimester 12% (165 734)
Mother				
	Mean (SD)	Mean (SD)	Mean (SD)	Mean (SD)
Age (years)	29.39 (5.34)	29.78 (5.16)	27.18 (5.43)	27.3 (5.83)
Parity (nulliparous)	59.5 (818 294)	39.4 (458 010)	58.8 (28 097)	43.2 (71 629)
Marital status (married or partnership)	62 (846 557)	66.8 (771 065)	35.9 (17 039)	35.7 (58 453)
Socioeconomic status	% (count)	% (count)	% (count)	% (count)
Upper white collar worker	14.9 (205 770)	16.8 (195 402)	6.4 (3081)	4.4 (7287)
Lower white collar worker	36.4 (501 780)	37.4 (435 623)	30.9 (14 786)	31 (51 371)
Blue collar worker	15.1 (207 962)	13.2 (153 567)	19.2 (9189)	27.3 (45 206)
Other	16.8 (231 481)	16.3 (189 157)	17.2 (8229)	20.6 (34 095)
Missing	16.7 (229 785)	16.3 (189 476)	26.2 (12 534)	16.8 (27 775)
Child				
	Mean (SD)	Mean (SD)	Mean (SD)	Mean (SD)
Gestational age (days)	278.66 (12.05)	278.74 (11.85)	279.34 (12.11)	277.87 (13.29)
Birth weight (g)	3549.45 (542.2)	3573.19 (536.38)	3540.62 (535.37)	3385.36 (556.1)
Body length (cm)	50.21 (2.43)	50.32 (2.4)	50.12 (2.39)	49.5 (2.59)
Head circumference (cm)*	34.94 (1.64)	34.99 (1.62)	34.91 (1.65)	34.55 (1.71)
Preterm birth	% (count)	% (count)	% (count)	% (count)
Preterm birth (<37 weeks)	4.3 (58 828)	4.1 (47 775)	4.3 (2040)	5.4 (9013)
Low/High birth weight	% (count)	% (count)	% (count)	% (count)
Low birth weight (<2500 g)	3.0 (40 006)	2.7 (30 443)	3.0 (1387)	5.0 (8176)
High birth weight (>4500 g)	3.1 (41 054)	3.3 (37 165)	2.8 (1313)	1.6 (2576)
Small for gestational age (<10th percentile)	% (count)	% (count)	% (count)	% (count)
Weight	11.5 (158 817)	10.4 (120 800)	13.6 (6483)	19.1 (31 534)
Body length	6.0 (81 869)	5.2 (60 003)	7.0 (3345)	11.2 (18 521)
Head circumference (cm)*	10.9 (69 350)	10.1 (54 321)	13.3 (4675)	16.4 (10 354)
Abnormal body proportions	% (count)	% (count)	% (count)	% (count)
High ponderal index	11.0 (136 918)	11.1 (116 902)	11.8 (5111)	10.1 (14 905)
Low brain-to-body ratio*	11.1 (63 520)	11.2 (54 554)	12.5 (3968)	9.2 (4998)
High head-to-length ratio*	9.6 (60 358)	9.8 (49 830)	8.9 (3560)	8.3 (6960)

*Available since 1 January 2004.

material figure S2). We stratified the data by birth year to investigate potential influence of changes in the composition of tobacco (eg, use of additives and changes in tar and nicotine content) and decreased social acceptability of smoking during the study period on the risk estimates. Stratification by birth year did not indicate a clear temporal pattern in risk estimates for any of the analysed endpoints (see online supplementary material, figure S3). Sensitivity analysis was performed to test the sensitivity of risk estimates to choice of adjustment factors in the regression model. Alternative multivariate adjustment models, including comorbidities or additional socioeconomic factors, did not significantly alter the risk

estimates for any of the reported endpoints (see online supplementary material figure S4).

DISCUSSION

In this work, we investigated the effect of maternal smoking during pregnancy, categorised as quit during the first trimester and continued after the first trimester on preterm birth, on body size and body proportions at birth. The most important finding of our study is that although the risk for low birth weight decreases by smoking cessation during the first trimester, brain size and body length in relation to body weight seem not to catch up. Among

Figure 1 Adjusted ORs for association of maternal smoking with (A) traditional birth outcomes (preterm birth, low birth weight and small for gestational age); and with (B) abnormal body proportions; adjusted for maternal age, sex, parity, socioeconomic status and gestational week (for low birth weight) and weight z-score (in regressions with weight included in the dependent variable); marker+error bar: OR (95% CI).

the newborns exposed to maternal smoking only during the first trimester all three measurements of body size (birth weight, body length and head circumference) showed signs of growth restriction. In addition, body proportions were abnormal.

Our work indicates a difference in susceptibility for growth restriction between weight, body length and head circumference. The observed positive association of maternal smoking with ponderal index suggests a stronger reduction in length than in weight. Similarly, the association with low brain-to-body ratio suggests reduction rather in brain size than in weight. However, the association of maternal smoking with high head-to-length ratio suggests a stronger reduction in length than in head size. It is in line with previous research showing that

smoking during pregnancy predominantly affects lean body mass and not fat tissue.⁴ While the associations with reduced body size and abnormal body proportion were stronger in newborns of mothers who continued smoking after the first trimester, there was a clear association also in newborns exposed only during the first trimester. This can be interpreted as an effect of maternal smoking on cell proliferation during organogenesis in early prenatal development. Insults during this period have been shown to persist throughout life.⁵ This stresses the importance of smoking cessation before pregnancy since even smoking only during early pregnancy has potentially devastating effects on long-term health of the unborn child.

The importance of body proportions at birth has been summarised by Zanelli and coworkers.¹⁵ Children born

small and thin were shown to be more likely to develop coronary heart disease as obese adults than their peers who were born small but not thin. It is not possible to infer from our study whether the high ponderal index and higher risk for shorter body length is comparable to stunting due to malnutrition and infections. Mechanistic studies of the observed effects are needed to extrapolate the risks to later life.

Additionally, smaller head circumference has been shown to directly translate into a smaller brain.^{16 17} Insults during early development of the brain, such as maternal smoking, have been shown to result in differences in DNA methylation, altered expression of genes regulating brain structure and function¹⁸ and reduction in neuronal content of the brain.¹⁹ Also, neurophysiological functions and overall brain functions are altered due to prenatal smoking.²⁰ The smaller brain volume observed in newborns has been shown to persist into young adulthood.²¹

Overall, our results are in line with previously reported studies by other groups (see online supplementary material, table S4).^{22–25} Smoking quit during the first trimester had a weaker effect on reduction in weight or length measures, whereas smoking especially at the end of pregnancy reduced femur length, abdominal circumference and biparietal diameter.²⁶ A clear dose response of smoking (number of cigarettes per day) on reduction of birth weight and increase in ponderal index has been demonstrated.^{17 27} Previous studies examining the effect of smoking cessation during pregnancy have consistently reported a reduction in harm compared with continued smoking.⁴

There is increasing evidence from animal studies of nicotine as a causative agent for reproductive toxicity, including detrimental effects on brain.^{28 29} In a large epidemiological study, among the few existing ones, aberrations in lung development due to nicotine replacement products used during pregnancy has already been suggested.³⁰ Epidemiological studies on the effects of nicotine products, other than cigarettes, are still needed. We recommend the inclusion of information on the use of nicotine products in the Medical Birth Register. This would allow detecting pregnancies at risk more reliably and facilitating epidemiological research on nicotine exposure during pregnancy beyond maternal smoking.

This work is solely based on routinely collected register data, which dictates the data availability. We tested our results for sensitivity to different adjustment models. Our results were robust against inclusion of maternal comorbidities, maternal anthropometric indices, social background and reproductive history as confounders in the statistical model. Smoking and socioeconomic status have been shown to correlate well with other lifestyle related factors and they are reliable markers for the unaccounted factors.³¹ The smoking information was self-reported by the mother during antenatal care visits. Thus, reporting bias cannot be excluded. It was not possible to analyse the impact of timing of smoking cessation in more detail or possible dose–response relationships due to lack of data. In addition, we lack information on paternal and household

smoking. Secondhand tobacco smoke exposure during pregnancy has been shown to increase the risk for low birth weight and growth restriction.³² We cannot exclude the possibility that some observed effects are partly attributable to secondhand tobacco smoke exposure, especially in those women who ceased smoking during their first trimester. Our definition of indicators for body proportions were constrained by available data in the register. Ideally, the outcomes in this work would be supplemented with clinical criteria collected during for example prenatal ultrasound scans, such as femur length, abdominal circumference and skinfold thickness. Further, the cut-off at the 10th percentile for the definition of small body size and abnormal body proportions was arbitrary due to a lack of clear data on a threshold for increased risk for complications later in life. Socioeconomic status was assigned here solely on maternal occupation and no information about the father's occupation was available. For a high proportion of mothers (18%) the occupation was not available. Previous studies applied the same SES categorisation and showed that the missing information did not bias the proportions in the other SES categories.²⁴

Study strengths include the register-based design with a large study size and practically complete population representativeness. Earlier analyses on the occupation codes available in the Medical Birth Register have shown that the socioeconomic confounding is reasonably well taken into account. Overall, occupation is well correlated with education and income in Finland, and it can be used as an indicator for socioeconomic health differences.^{33 34} An earlier study showed a reasonable match between serum cotinine and self-reported smoking status as applied in the Medical Birth Register.³⁵

CONCLUSIONS

This study showed that maternal smoking is associated with a stronger reduction in body length and head circumference than in birth weight, leading to changed body proportions. The effects on body proportions of having quit smoking during the first trimester or having continued smoking after the first trimester were similar, stressing the importance of early pregnancy as a sensitive exposure window. Furthermore, it suggests a limited potential to repair fetal damage induced in early pregnancy. Lower brain-to-body ratio suggests that any smoking during the pregnancy may lead to losses in the development of central nervous system. The effects on body size (weight, length and head circumference) were more pronounced in newborns of mothers who continued smoking after the first trimester.

It seems important to study the association of growth restriction and other adverse effects with the use of nicotine therapy products, as already demonstrated in animal studies. Until their safety has been proven, caution should be taken when advising pregnant women. Routine collection of information on the use of nicotine replacement products in the Medical Birth Register is needed for more careful follow-up of risk pregnancies and to facilitate

scientific research on specific effects associated with nicotine replacement products.

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Contributors IR, OH, KV and MV conceived the research question and designed the study. IR conducted statistical analysis, interpreted the results and wrote the first and subsequent drafts of the manuscript with support of OH, KV, MV, MG and HdR contributed to data interpretation and revisions of the manuscript. MV, OH, KV and IR obtained funding. All authors approved the final version of the submitted manuscript.

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Competing interests Kirsi Vähäkangas was the chair in 2018–2019 of the scientific advisory committee of the Savuton Suomi (Smoke-free Finland) 2030 initiative.

Patient consent for publication Not required.

Ethics approval The study was approved by the ethics committee of the Northern Ostrobothnia Hospital District (EETM 44/2016; issued 18th April 2016). The right to use of register data held by the Finnish Institute for Health and Welfare was granted under the document number THL/838/6.02.00/2016 (issued 22nd June 2016).

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Data availability statement Data may be obtained from a third party and are not publicly available. The Finnish Institute of Health and Welfare is the controller of the Medical Birth Register. Data may be obtained from the register controller (<https://thl.fi/en/web/thlfi-en/statistics/information-on-statistics/register-descriptions/newborns> accessed 29 May 2019).

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