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Risk of Long-Term Benzodiazepine and Z-Drug Use Following the First Prescription among Community-Dwelling Adults with Anxiety/Mood and Sleep Disorders: A retrospective cohort study

Running Title: Risk of long-term benzodiazepine receptor agonist use

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ABSTRACT

Objective: To measure the incidence of long-term benzodiazepine receptor agonist (BZRA) use among individuals with anxiety, mood and/or sleep disorders. To identify factors associated with long-term use following the first prescription.

Methods: This was a population-based retrospective cohort study using administrative databases in Manitoba, Canada. Individuals with anxiety/mood or sleep disorder who received their first BZRA between April 1, 2001 to March 31, 2015 were included. Long-term use was defined as ≥180 days.

Results: Among 206,933 individuals included, long-term BZRA use in the first episode of use ranged from 4.5% (≥180 days) following their first prescription. Factors associated with ≥180 days of use included male sex (adjusted odds ratio (aOR) 1.33, 95% CI 1.27 to 1.39), age ≥65 (aOR 5.15, 95% CI 4.81 to 5.52), income assistance (aOR 1.68, 95% CI 1.55 to 1.81), previous non-BZRA psychotropic (aOR 1.93 (95% CI 1.83 to 2.02) or opioid use (aOR 1.16, 95% CI 1.11 to 1.22), high comorbidity (aOR 1.43, 95% CI 1.32 to 1.55), high healthcare use (aOR 1.46 (95% CI 1.33 to 1.60), and psychiatrist prescriber (aOR 2.11, 95% CI 1.93 to 2.32).

Conclusions: Less than one in ten patients use BZRAs ≥180 days in their first treatment episode. Several factors were associated with long-term use following the first prescription and further investigation into whether these factors need to be considered at the point of prescribing is warranted. In light of these findings, future research should examine the predictors of cumulative BZRA exposure.

Key Words: benzodiazepine, anxiety, insomnia, pharmacoepidemiology, clinical practice guidelines, z-drug hypnotics

Strengths and Limitations of Study

- This study used administrative data from the Manitoba Centre for Health Policy, which is one of the most comprehensive datasets in North America containing >140 de-identified linked datasets on healthcare, education, social/families, justice and registries for all residents of Manitoba (population of 1.4 million people) not restricted by age or income
- All diagnoses are identified through physician claims data or hospitalizations, which are
 dependent on people seeking treatment and may be prone to some misclassification. Drug
 information is also based on dispensing records from community pharmacies and does not
 confirm the patient actually took the drug. However, we performed multiple sensitivity
 analyses to address this.
- The databases do not capture participation in psychological interventions such as cognitive behavioral therapy.

Introduction

The use of benzodiazepine receptor agonists (BZRAs)*, benzodiazepines (BZD) and Z-Drugs, in the treatment of anxiety and insomnia has shifted based on the evolving safety and efficacy data on long-term use in the literature. ¹⁻⁴ Upon their initial introduction into the clinical practice in the late 1950's, benzodiazepines were considered to be a safer alternative to barbiturates.⁵ However, safety concerns such as psychomotor impaired accidents (i.e., falls and motor-vehicle accidents), dependency and misuse/abuse are now well known. ⁶⁻⁸ Benzodiazepine in combination with opioid prescription has also been reported to increase the risk of opioid-related death by 1.5 to 3.9-fold. ¹⁰⁻¹³ Opioid-related hospitalizations have ranged from 17.1% in British Columbia to 35.6% in Manitoba among individuals with a co-prescription of opioid and benzodiazepine. ¹² Recent studies have raised concerns proposing possible links to dementia, recurrence of mood episode, respiratory disease exacerbation and suicide. ¹³⁻¹⁷ However, the association of BZRA use for these newer harms is uncertain given conflicting evidence and confounding in previous studies. ^{18,19}

In spite of ongoing adverse effect concerns, justification for less restrictive BZRA use have stemmed from their reputation as rapidly effective anxiolytic sedatives.²⁰ Some view that withholding BZRA is at times impractical and may increase psychiatric symptom burden and patient distress.²¹ Moreover, the use of alternative pharmacotherapy, including trazodone, atypical antipsychotics, barbiturates, and tricyclic antidepressants are not without harm. Nevertheless, a patient-centered approach which carefully accounts for the benefits and risks of BZRA use is expected to yield the best outcomes for the patient.²² It should also be noted that the difficulties

^{*} Abbreviations: BZRA, benzodiazepine receptor agonist; BZD, benzodiazepine

with de-prescribing these agents reported in the literature and have added caution to the initiation of these agents in practice.

Clinical practice guidelines have attempted to provide general direction to practitioners and pharmacists on how these medications should be managed according to the best available evidence. There are a small number of population-wide prescribing practice evaluations to determine the extent of adherence to guideline recommendations and only one considered data on duration of use. As such, this study sought to i) measure the incidence of long-term BZRA use among a cohort of community-dwelling Canadian adults with anxiety, mood and/or sleep disorders. ii) To determine factors associated with progression to long-term BZD use following the first prescription in this population.

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Methods

Study Design and Data Sources

This was a retrospective, cohort study using routinely collected administrative healthcare data pertaining to prescription drug dispensations, outpatient physician claims, hospitalization discharge abstracts, income assistance records and prescriber demographics (Table 1). All data used was extracted from the Manitoba Centre for Health Policy's Population Research Data Repository. The Repository provides comprehensive coverage of all Manitoba residents contact with the primary healthcare system. The Drug Program Information Network provides information on outpatient prescription drugs dispensed in Manitoba with the exception of medications dispensed in hospital and nursing stations. In Manitoba, eligible outpatient prescriptions are 100% covered for residents after an income-based deductible is paid for each fiscal year. Merging of the

various data sources was facilitated via linkage of unique de-identified Personal Health Information Numbers. The Charlson Comorbidity score [0 (lowest risk), 1, \geq 2 (high risk)] was determined based on 17 categories of comorbidities using ICD-9-CM or ICD-10-CA equivalent codes in administrative data to provide the weight-based adjusted risk of death or resource use.²⁹ All data was cleaned and analyzed using Base SAS v9.4©.

Cohort Inclusion/Exclusion Criteria

Eligible patients were adults age 18 years and older who initiated a new benzodiazepine or Z-Drug prescription (defined as no use in the one year prior to the first prescription^{30,31}) between April 1, 2001 to March 31, 2015, with no preceding dispensations from April 1, 2000 to March 31, 2001 (first year of the dataset) to avoid prevalent user bias (Figure 1). A ≥1-year of follow-up prior to and after the first prescription, as determined by insurance registry coverage, was required for cohort inclusion.

Eligibility was also based on diagnostic criteria for anxiety/mood related disorders and/or insomnia based on ICD-9-CM or ICD-10-CA medical claims, either at outpatient physician visits or hospitalizations, occurring within a 5-year period prior to the first prescription. The ICD diagnostic criteria chosen are a combination of the definitions from two sources; the Canadian Public Health Association on mental health surveillance and the MCHP concept dictionary, which listed the various past-case definitions employed in previous research within Manitoba for mood and anxiety disorders (Table A1).³²⁻³⁶ Lastly, because reliance on ICD codes is expected (and has been previously shown) to underestimate capture of sleep disorder cases, we also accepted receipt of a Z-Drug in the definition for insomnia as this was their sole approved indication.³⁷

To reduce confounding, we established cohort exclusion criteria that otherwise may have justified long-term use of BZDs in clinical scenarios beyond the scope of general guideline recommendations for anxiety and insomnia. Namely, patients were excluded if they had ≥1 ICD code for cancer, a seizure disorder or if there was placement in the Manitoba palliative care drug program at any point in the 5 years preceding their first prescription for a BZRA (Table A2). Where patients became palliative ≥1 year after the initial BZRA dispensation, their ongoing use of BZRA was censored beginning from the date of their placement, but all use prior to their palliative status was retained. Clobazam use was excluded entirely from the evaluated drug claims because it is approved only as an adjunctive agent for epilepsy in Canada. Finally, patients were excluded if they lacked at least 1-year of registry coverage from their first-prescription index date. This was to eliminate any biasing effect from early mortality, moving out of province or other loss to follow-0/0 up.

Main Outcome Measures

Long-term use was defined as ≥180 days based on the recommendation from a previous systematic review of similar studies.³² This duration is longer than clinical practice guideline duration recommendations and is believed to be of sufficient length, with repeated dosing, for some degree of dependency to arise in many users.³⁸ One-third of individuals who use BZDs for longer than six months have been previously reported to be unable to stop completely due to withdrawal symptoms (e.g., anxiety, insomnia, muscle spasms).³⁸ A sensitivity analysis, ranging from 60 to 365 days, was also used in our study to account for variances in dispensing patterns and to allow for a period long enough to develop tolerance.³²

Patients were followed forward in time from the date of their first BZRA prescription. BZRA 'use episodes' were determined according to consecutive prescription overlap based on dispensation dates and coded day supply values. The allowable gap between prescriptions was the greater of either 30 days or 50% of the last prescription day supply after the prescription end date (end date = dispensation date + day-supply) of the prior prescription. This gap was chosen because we believed it was an acceptable compromise, in the absence of prescription use directions, because it allowed for clinically significant, but persistent, 'as needed' BZRA use while preventing more infrequent 'as needed' prescription fills from contributing to 'use episode' duration. The episode end date was calculated as the date of the last prescription in a given 'use episode' plus its associated day-supply. To account for immeasurable time bias, hospitalization time was assumed to be a continuation of BZD use given that in-patient drug use data was limited.³⁹ The provincial drug program subsidizes dispensations of up to a 100 day-supply.

Individuals were able to have multiple use episodes over the entire study duration. First episode duration and average episode duration were calculated for each user. If patients only had one use episode both of these values were the same. Patients were allowed to switch from one BZRA to another without it interrupting their 'use episodes'. This included switching from a BZD to a Z-drug and vice versa.

Independent Variables

Variables used for statistical prediction of long-term use included age, sex, geographic residence, residential mobility, socioeconomic status, marriage, concurrent opioid or prescription psychotropic use, comorbidity burden, healthcare usage, time period of first prescription and prescriber characteristics (Table A3 and Table A4). Variables were assessed at baseline; either

within 1-year before the index date, at the index date or up to 6-months past the index date (in the case of prescription opioids and other psychotropics, such as antidepressants, antipsychotics, and mood stabilizers).

Logistic Regression Modelling

Standard reporting criteria were followed in the approach to logistic regression modelling. Univariate analysis was performed first in the form of simple logistic regression. The multi-variable model was constructed to determine the most parsimonious model for prediction of long-term BZRA use defined as \geq 180 days in the first episode of use with adjustment of clinically relevant covariates based on previous literature. Differences between models in their maximum log-likelihood estimation, likelihood ratios and other goodness-of-fit statistics enabled model discrimination. Multicollinearity and effect-measure modification (i.e., interaction effects) were assessed when it was suspected that variables may be either correlated or non-independent. In order to perform these diagnostics, the binary dependent variable was first substituted for a linear variable (first-episode duration in days) to conduct a multiple *linear* regression. Specifically, collinearity was determined to be a model threat if any correlation coefficient in the independent variable correlation matrix was \geq |0.8| or if any variance inflation factor was unreasonably high (\geq 10) while the corresponding tolerance factor was miniscule (\leq 0.1). Analyses were assessed at p<0.01 threshold set a priori for statistical significance.

For the multiple logistic regression, 'complete case-analysis' was used because the extent of missing data was too small to justify the need for multiple imputation procedures. ⁴³ In this study, no claims were excluded on the basis of missing data fields. Only 1,568 claims (<0.01%) were excluded for being spurious (i.e '0' day/quantity supply or incredibly high dispensed quantity to day-supply ratio) Furthermore, observed missing data was believed to be missing at random. The

only variable with significant missing data was that of 'prescriber type' (~38,000 missing observations or 17.5% of final sample).

Sensitivity Analysis

To assess the robustness of the primary outcome, 6 sensitivity analyses (Tables A7 and Table A8) were conducted to determine how the proportion of long-term use changed under differing parameter assumptions.⁴³ The threshold duration for long-term use was adjusted to values ranging from 60 days to 365 days. Additionally, the episode lapse criteria (i.e., prescription gap rule) was changed. While the analysis was not exhaustive for every conceivable combination of these key parameters, the selected values were chosen because they were judged to be representative of how peers in the international clinical community may have defined or measured 'long-term use' of BZRA.

Ethical Approval

Access to the data for this project was approved by the University's Health Research Ethics Board (registration number H2017:052 (HS20498) and the Health Information Privacy Committee (no. 2016/2017-62) of the provincial government.

Results

Episodic BZD/Z-Drug Use

There were 206,933 patients in our cohort representing 931,271 unique BZRA dispensations over the 15-year study duration, accounting for a total of 337,341 person-years of BZRA use based upon our use-duration measurement method. Over the study period, cohort

individuals had a median of three and average of 4.5 BZRA use episodes, respectively. First-episodes of use were of a median duration of 20 days (IQR = 10-30 days). For all use-episodes, the median average use duration was 30 days (IQR = 15-111 days). Evaluation of long-term use revealed that 4.51% of patients used a BZRA for ≥180-days in their 'first' episode of use. At most, this proportion increased to 9.64% when a sensitivity analysis of 60 days or greater was used for the definition of 'long-term use' for the first episode of use. However, the proportion of long-term users increased considerably after averaging for all episodes for each user (sensitivity analysis range: 15.6%-35.1%) (Table A7).

To evaluate treatment duration for insomnia, a sensitivity analysis was performed on only Z-Drugs (*n*=110,663). This was done to mitigate any effects of concurrent BZD use and to get a more specific estimate for insomnia treatment duration; however, the results were similar. All results for the Z-Drug cohort are provided in the supplemental appendix (Tables A8-A11).

Factors Predicting Long-term First Episode Use

Logistic regression analysis revealed that male sex (adjusted odds ratio (OR) 1.33, 95% CI 1.27 to 1.39), older age (adjusted OR 2.24 (95% CI 2.11 to 2.38) and 5.15 (95% CI 4.81 to 5.52) for aged 45-64 years and ≥65 years, respectively, compared to <45 years), receipt of income assistance (adjusted OR 1.68, 95% CI 1.55 to 1.81), previous non-BZRA psychotropic (adjusted OR 1.93 (95% CI 1.83 to 2.02) or opioid use (adjusted OR 1.16, 95% CI 1.11 to 1.22), high comorbidity (Charlson Comorbidity Index 1 and ≥2, adjusted OR 1.11 (95% CI 1.04 to 1.17) and 1.43 (95% CI 1.32 to 1.55), respectively), high healthcare resource use (resource utilization band of 4 and 5, adjusted OR 1.15 (95% CI 1.07 to 1.23) and 1.46 (95% CI 1.33 to 1.60), respectively), first prescription from psychiatrist (adjusted OR 2.11, 95% CI 1.93 to 2.32) and receipt of first prescription after 2006 (2006-2011, adjusted OR 1.74, 95% CI 1.64 to 1.85; 2011-2015, adjusted

OR 2.99, 95% CI 2.80 to 3.18), were all predictive of long-term use of ≥180 days in the first episode. Rural residence (adjusted OR 1.10, 95% CI 1.04 to 1.15) and high residential mobility (adjusted OR 1.14, 95% CI 1.08 to 1.21) were also associated with a higher risk of long-term use in the first episode. Married status was associated with a lower risk of meeting the long-term use definition (adjusted OR 0.79, 95% CI 0.76 to 0.83). These findings were also replicated in the sensitivity analysis restricted to Z-Drug users. Both the crude and adjusted odds ratios are presented for the full cohort in Table 2.

A sub-analysis of the higher comorbidity scores in the long-term user groups shows that this relationship was mainly driven by cardiovascular diseases, diabetes and dementia (though nearly all diagnoses had statistically significant differences) (Table 3). Proportions for these particular diagnoses were 2 to 5 times higher in the long-term user group, with the greatest difference existing for dementia (long-term; 8.5% vs. short-term; 1.5%).

Discussion

This study demonstrates that 'first-episode' use appears to be largely consistent with current practice guideline recommendations in regards to usage duration among those with prior anxiety, depression, or insomnia. Approximately 4.5% of the full cohort and 7.4% of the Z-Drug cohort were 'long-term' first-episode users according to the best available evidence-based consensus definition of 180 days.³² Restricting the analysis to Z-Drug use showed that the frequency of long-term use was higher than that of the main cohort. However, strictly in terms of practice guideline recommendations, the duration of use advocated for Z-Drugs in the treatment of primary insomnia is often shorter (range of ≤4-6 weeks) than that allowed for BZD in anxiety states.⁴⁵ Therefore, these results suggest greater disparity from practice guidelines in the case of

Z-drug use for insomnia. Of note, more recent insomnia guidelines have recognized that while non-drug alternatives have a favourable safety profile, these interventions may be difficult to achieve for certain populations, which could explain the deviation between practice recommendations and real-world use of these agents.⁴⁶

The proportion of patients who met criteria for 'long-term' use after accounting for all of their use-episodes was approximately 3.5 times higher than the proportion of patients meeting criteria after only their first episode of use. These results indicate that repeated episodes of BZRA use are associated with progression to longer-term use episodes. Though, the majority of repeat users still only take BZRAs for intermittent, short-term periods. Furthermore, confounding variables such as age and accrued comorbidity over time suggest a potentially legitimate requirement for future long-term use in some patients. Nonetheless, these results support the observed difficulty in de-prescribing once BZRA use has become chronic, which has also been reported in previous literature.^{4,46,47} Lastly, other clinical considerations such as risk of protracted withdrawal symptoms, risk of rebound insomnia and/or anxiety, patient dissatisfaction, limited alternate drug and non-drug interventions, or interference with another prescriber's decisions likely undermine potential de-prescribing efforts.

Older age and female sex have also been identified in previous studies as being associated with long-term use. 48–55 While we found females to have greater representation in all patterns of BZRA use, we found males were more specifically predictive of long-term use after the first episode of use. 56–58 As with almost all of the previously published studies, older age was strongly associated with long-term BZRA use. 55-59 It should be noted that older individuals may have had a greater opportunity to be exposed to BZRA use. Therefore, it is possible that age could be a

confounder if increased BZRA exposure is associated with decreased likelihood for BZRA cessation.

As supported by previous evidence, income assistance was associated with long-term BZRA use ^{51,60}. Our study also found frequent moving, unmarried status, and rural residence to be associated with increased odds of long-term use. Frequency of moving, income assistance, and marriage status could be a proxy for social or general life stability^{53,65,62}. Rural residence may have a small effect on longer-term BZRA use due to the relative unavailability of timely scheduled follow-up, which may necessitate prescriptions of greater quantity or for longer periods. Another study also found rural adults to be at higher odds of inappropriate BZD use .⁶¹

Healthcare consumption and the presence of various physical illnesses have been consistent predictors of long-term BZRA use 50,52,53,58,63. In this study, as both of these variables increased, so did the odds of longer-term use. We speculate that the positive relationship between these two indices and long-term use may be partially explained by unmeasured 'health' anxiety or associated mental health issues arising secondary to physical comorbidities or by additional disruptive effects of physical illness on sleep. Investigation of this link in future studies may better inform clinicians on prescribing of BZRA for such 'atypical' anxiety states.

The Charlson comorbidity score findings were not surprising given the relatively higher proportion of older adults in the long-term user group. Nonetheless, the greater degree of BZRA exposure among those patients with dementia is alarming given the ongoing controversy between dementia and BZD use ^{9,19}. This concern is echoed by a previous European study that found higher prevalence rates of long-term use of BZD in community dwelling elderly with Alzheimer's disease.⁶⁴

In concordance with previous studies, prescriptions for an opioid or a psychotropic agent, such as antidepressants, antipsychotics or mood stabilisers, during the baseline period were modestly predictive for future long-term use. 51,54,56,58,60,65 Those having received a non-BZD prescription agent for a psychiatric disorder could be expected to have had greater disease severity on average than those BZRA users who did not receive such treatment early on. Furthermore, certain antidepressants, namely SSRIs, may stimulate a greater need for a BZD due to their adverse pharmacology resulting in what has been termed "anxiety/jitteriness syndrome". 64 Therefore, undetected anxiogenic or sleep disrupting effects of other psychotropic medications may, in some cases, result in persisting BZD use.

An unexpected finding was the increased odds of long-term use associated with the more recent time period of the first prescription. This is contrary to what may be expected from cumulative knowledge on BZRA and the long-standing emphasis on short-term use advised in guidelines and clinical literature. Nevertheless, this trend may be partially explained by changes in the clinical selection of BZRA over the course of the 15-year study period and the corresponding evidence for the popularity of certain agents.⁶⁷ This finding may reflect the growing awareness that BZRAs should not be used as a first-line treatment resulting in only those with greater risk factors and fewer coping strategies to be more likely to receive BZRAs and who may be less likely to respond to other alternatives.

In regards to zopiclone, the relative absence of preferred alternative first-line pharmacotherapies in the Canadian prescriptive armamentarium may have resulted in the default selection of this agent by many prescribers to treat insomnia. Furthermore, a perception of lesser risk (compared to BZD) coupled with increases in population prevalence of insomnia over time (due to various factors such as population aging, increased technological screen time etc.) may

account for why the incidence of long-term use has increased. Lastly, long-term clonazepam usage was also observed in previous studies.^{68,69} Some studies have shown greater abuse liability with clonazepam over other BZD.^{70,71}

The present study has a number of strengths. This study used a large administrative data sources that were near complete in their coverage of the study population's prescription drug dispensations and healthcare contact. Application of cohort inclusion and exclusion criteria in a carefully constructed new user longitudinal design limited confounding and bias to the extent possible. Multiple sensitivity analyses on the main outcome measure, the duration of BZD/Z-Drug use measurement method and the association between the independent and dependent variables for two cohorts reduced quantitative bias to increase confidence in the results.

A few important limitations should be acknowledged. Firstly, administrative data is prone to some misclassification of variables. For instance, diagnostic criteria for cohort case inclusion and exclusion will differ in their true sensitivity and specificity, regardless of prior validation of case definitions. Drugs used during any hospitalizations were not available and was assumed to be continued BZD exposure. As all independent variables were only measured cross-sectionally before or at the time of the first prescription of the first use-episode, the logistic regression model was only predictively valid for the first use episode duration and not users' average episode duration. Since DPIN only captures the days supply provided, it is possible that not all of the medication was actually taken by the patient. However, this study was able to provide insight into the prescribing practices of benzodiazepines that are filled in the pharmacy in this population. Our study did not evaluate the extent of concurrent use of multiple BZD or other psychiatric diagnoses such as substance use disorder. The databases also do not capture participation in psychological interventions such as cognitive behavioral therapy. This study was done in a setting where there is

a universal healthcare system and medication costs are covered for all Manitobans after an incomebased deductible is met every year. As a result, findings may be generalizable to similar settings. Future research should aim to examine the association of repeat exposure to BZRA and risk of chronic use.

Conclusion

Prescribing of BZRAs was in accordance with clinical practice guideline recommendations on use duration for the majority of individuals with a prior history of anxiety, depression, or insomnia. However, the proportion of long-term use among new users was up to one in three based on the average of all episodes of use, warranting future research in this area. Patients who are male, of older age, are socially or financially deprived, have poor physical health, use opioids or other psychotropic agents and are frequent consumers of healthcare resources are more likely to use BZRA long-term after their first prescription. Future research could be done to explore whether these factors need to be considered at the point of prescribing in clinical practice.

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Availability of Data and Materials

Data used in this article was derived from administrative health and social data as a secondary use. The data was provided under specific data sharing agreements only for approved use at MCHP. The original source data is not owned by the researchers or MCHP and as such cannot be provided to a public repository.

Author Statement

All authors contributed to the original design, analysis, interpretation, and writing of this study and manuscript.

Conflict of Interest Disclosure

The authors declare no conflicts of interest related to any aspects of this work.

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Patient and Public Involvement

We have a patient advisory group who provided feedback on the dissemination of research findings.

Table 1 -Raw Data Sources and Relevant Corresponding Data Elements

Database	Date Range	Relevant Data Elements
	of Data	
Drug Program	Apr. 1/2000 –	Prescriptions for benzodiazepines (ATC codes
Information Network	Mar. 31/2016	N03AE, N05BA, N05CD), Z-Drugs (N05CF),
(DPIN)		Antidepressants, Antipsychotics, Mood stabilisers,
		Lithium and Opioids
		-Drug, dosage strength, dosage type, metric
		quantity dispensed, day supply, date of
		dispensation
Manitoba Health	Apr. 1/1996 –	Birth date/age of patient; sex; location of
Insurance Registry	Mar. 31/2016	residence, marital status, date of Manitoba Health
		coverage, date of coverage end, reason for
		coverage end (i.e death, emigration etc.)
Medical Claims	Apr. 1/1996 –	Services - type of physician (e.g., psychiatrist);
(Physician Billings)	Mar. 31/2016	dates of services, specific diagnoses (ICD-9 or
		ICD-10 equivalent)
Hospital Separations	Apr. 1/1996 –	Diagnoses (ICD-9 or ICD-10 equivalent), length
Abstracts	Mar. 31/2016	of stay, admission dates, discharge dates,
Provider	Apr. 1/1996 –	Physician Age, Sex, Specialty
Registry/Physician	Mar. 31/2016	
Master File		
Social Allowances	Apr. 1/2001-	Receipt of income assistance
Management Information	Mar. 31/2013	
Network (SAMIN)		

BMJ Open

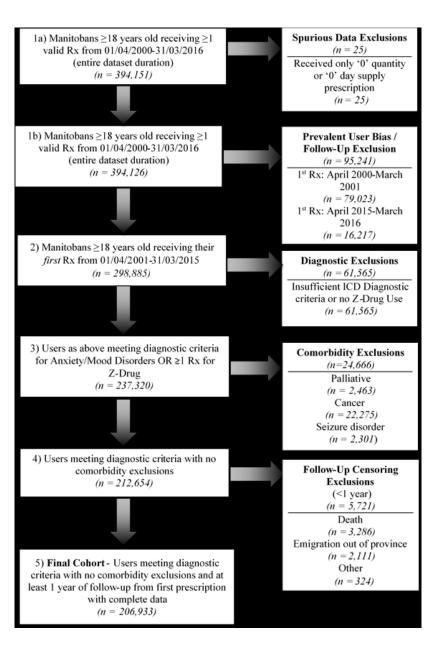
Table 2 – Statistical Associations between Predictor Variables and Long-term Use of BZRAs

Table 2 – Statistical Associations between Predictor Variables and Long-term Use of BZRAs							
				Use D	uration	916	
Independent Variable		≥180 Days		≥90 Days		2 ≥60 Days	
		Crude OR (95% CI)	Adjusted OR (95% CI)	Crude OR (95% CI)	Adjusted OR (95% CI)	Crixde OR (95% CI)	Adjusted OR (95% CI)
Male		1.41 (1.35-1.47)	1.33 (1.27-1.39)	1.40 (1.35-1.45)	1.34 (1.29-1.40)	12830 (1.26-1.34)	1.27 (1.23-1.31)
_	18-44	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 g ref)	1 (ref)
Age	45-64	1.82 (1.73-1.92)	2.24 (2.11-2.38)	1.77 (1.70-1.85)	2.00 (1.91-2.10)	ਜ਼ੁਿਲ1 (1.7%-1.86)	1.89 (1.82-1.97)
	65+	4.06 (3.86-4.28)	5.15 (4.81-5.52)	3.56 (3.41-3.72)	4.11 (3.88-4.36)	(3.2 ½ -3.47)	3.52 (3.36-3.70)
Rural Residenc	e	1.07 (1.02-1.11)	1.10 (1.04-1.15)	0.97 (0.93-1.00)	0.97 (0.94-1.02)	(0.8 3 -0.92)	0.92 (0.88-0.95)
High Residential Mo	obility	1.52 (1.45-1.60)	1.14 (1.08-1.21)	1.35 (1.29-1.40)	1.06 (1.01-1.11)	(1.1 <mark>2</mark> -1.18)	1.01 (0.97-1.06)
Income Assistan	се	1.46 (1.37-1.55)	1.68 (1.55-1.81)	1.14 (1.08-1.21)	1.35 (1.26-1.45)	(0.8 3 -0.93)	1.12 (1.06-1.20)
_	<-1	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 g (ref)	1 (ref)
Socio-Economic Factor	-1 to 0	1.08 (1.00-1.15)	0.99 (0.92-1.07)	0.96 (0.91-1.02)	0.91 (0.86-0.97)	<u></u> €90 (0.87-0.95)	0.89 (0.85-0.94)
Index-2 (SEFI-2) – Score –	0 to 1	1.16 (1.07-1.24)	1.02 (0.94-1.10)	0.98 (0.93-1.04)	0.92 (0.87-0.98)	687 (0.82-0.91)	0.89 (0.84-0.94)
	>1	1 (0.92-1.09)	0.93 (0.84-1.03)	0.78 (0.73-0.84)	0.80 (0.74-0.87)	(0.5 9 -0.67)	0.73 (0.68-0.78)
Married		0.91 (0.87-0.95)	0.79 (0.76-0.83)	1.01 (0.98-1.05)	0.89 (0.85-0.92)	1.13 (1.1 2 -1.16)	0.95 (0.92-0.99)
Opioid Use		1.19 (1.14-1.27)	1.16 (1.11-1.22)	1.08 (1.04-1.12)	1.09 (1.05-1.14)	<u>€</u> .99 (0.9€-1.02)	1.05 (1.01-1.09)

			BMJ Open			36/bmjopen-2020		
		Use Duration				04 916 ≥60 Days		
Independent Variable			Days		Days		Days	
		Crude OR (95% CI)	Adjusted OR (95% CI)	Crude OR (95% CI)	Adjusted OR (95% CI)	Cri&de OR (95⅔ CI)	Adjusted OF (95% CI)	
		1.82	1.93	1.62	1.75	€.34	1.49	
Psychotropic Rx Use (non-BZRA)	(1.75-1.90)	(1.83-2.02)	(1.56-1.67)	(1.69-1.83)	(1.3\$\overline{2}\overline{2}\overline{1}\overline{2}\o	(1.44-1.54)	
	0	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1gref)	1 (ref)	
Chamban Camanhidit		1.44	1.11	1.33	1.08	₽24	1.04	
Charlson Comorbidity Index Score		(1.36-1.51)	(1.04-1.17)	(1.27-1.39)	(1.02-1.13)	(1.1 9 -1.29)	(1.00-1.08)	
Thuex Score	2+	2.96	1.43	2.41	1.33	<u>≥</u> 01	1.23	
	2+	(2.79 - 3.15)	(1.32-1.55)	(2.29-2.54)	(1.24-1.42)	(1.92 - 2.11)	(1.15-1.31)	
	0-3	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1⊈ref)	1 (ref)	
Dagarra III:li-niian	4	1.84	1.15	1.58	1.08	<u>₿</u> 37	1.00	
Resource Utilization Band		(1.73-1.95)	(1.07-1.23)	(1.50-1.66)	(1.01-1.14)	$(1.3\overline{\$}-1.43)$	(0.94-1.05)	
Бина	5	3.48	1.46	2.73	1.31	2 .21	1.17	
		(3.24-3.73)	(1.33-1.60)	(2.56-2.92)	(1.20-1.42)	(2.08 - 2.35)	(1.09-1.27)	
Male Prescriber of First	t Dragarintian	1.07	1.03	1.07	1.04	£ .01	0.98	
Male Trescriber of First	i i rescripiion	(1.02-1.12)	(0.98-1.09)	(1.02-1.11)	(0.99-1.09)	(0.98-1.05)	(0.94-1.02)	
Prescriber Age ≥5	O Years	1.08	0.98	1.08	0.99	<u>§</u> 15	1.08	
Trescriber Age 25	o rears	(1.03-1.12)	(0.94-1.03)	(1.04-1.12)	(0.95-1.03)	(1.1g-1.18)	(1.04-1.11)	
	GP	1 (ref)	1 (ref)	1 (ref)	1 (ref)	<u></u> ≥1	1 (ref)	
Type of Prescriber of	Psychiatrist	2.06	2.11	1.85	1.89	₹54	1.63	
First Prescription	1 sychiairisi	(1.89-2.25)	(1.93-2.32)	(1.72-2.00)	(1.75-2.05)	(1.43-1.65	(1.51-1.75)	
Tirsi Trescription	 Other	1.09	0.92	1.07	0.92	1816	1.03	
	Oinei	(0.98-1.21)	(0.82-1.03)	(0.98-1.17)	(0.84-1.01)	(1.47-1.24)	(0.96-1.11)	
	2001-2006	1 (ref)	1 (ref)	1 (ref)	1 (ref)	<u><u></u> <u><u></u> <u> </u></u></u>	1 (ref)	
Period of First	2006-2011	1.66	1.74	1.58	1.65		1.48	
Prescription		(1.58-1.75)	(1.64-1.85)	(1.51-1.65)	(1.57-1.7)	(1.3\$\overline{6}-1.46)	(1.42-1.54)	
Γιεσειιριισιι	2011-2015	2.93	2.99	2.59	2.71	₿97	2.07	
		(2.78-3.08)	(2.80-3.18)	(2.48-2.71)	(2.57-2.8)	(1.96-2.05)	(1.98-2.16)	

Table 3 – Frequency of Charlson Comorbidity Group Diagnoses by First Use Episode **Cohort**

Charlson Diagnosis	Short-Term 'First-Episode'	Long-Term 'First- Episode' Users	Z-Test of Two	916 on 1 November 2021. Downloaded from http://bmjopen.bmj.com/ on April 18, 2024
	Users	(n=9,327)	Proportions	ve m
1: 1 x 0	(n=197,606)	201 (2.00/)	.0.01	ber
Myocardial Infarction	2,474 (1.3%)	281 (3.0%)	p < 0.01	202
Congestive Heart Failure	3,943 (2.0%)	628 (6.7%)	p < 0.01	13
Peripheral Vascular Disease	2,367 (1.2%)	256 (2.7%)	p < 0.01	Jownk
Cerebrovascular Disease	3,690 (1.9%)	544 (5.8%)	p < 0.01	oad
Dementia	2,928 (1.5%)	796 (8.5%)	p < 0.01	ed f
COPD	23,064 (11.7%)	1,163 (12.5%)	p = 0.02	ro m
Connective Tissue/Rheumatic Disease	2,793 (1.4%)	222 (2.4%)	p < 0.01	http://
Peptic Ulcer Disease	2,140 (1.1%)	114 (1.2%)	p = 0.20	Э <u>.</u>
Mild Liver Disease	2,406 (1.2%)	135 (1.4%)	p = 0.05	оре
Moderate/Severe Liver Disease	341 (0.1%)	28 (0.0%)	p < 0.01	n.bmj.c
Uncomplicated Diabetes	14,131 (7.2%)	1,099 (11.8%)	p < 0.01) Ö B
Complicated Diabetes	1,611 (0.8%)	252 (2.7%)	p < 0.01	on /
Paraplegia and Hemiplegia	794 (0.4%)	136 (1.5%)	p < 0.01	Αp
Renal Disease	1,858 (0.9%)	238 (2.6%)	p < 0.01	<u>→</u>
Cancer	829 (0.4%)	64 (0.1%)	p < 0.01	8, 2
Metastatic Carcinoma	64 (0.0%)	13 (0.0%)	p < 0.01	024
HIV/AIDS	50 (0.0%)	10 (0.0%)	p < 0.01	by guest. Protected by copyright.



Flowchart of study population

48x70mm (300 x 300 DPI)

Supplemental Appendix Tables

Table A1 – International Classification for Disease Coding for Mood/Anxiety/Sleep Disorders (Cohort Inclusion)

	Source 1 - CPHA	Source 2 - MCHP	Study Algorithm
ICD Codes	All Mental Health Disorders: 9-CM: 290-319 10-CA: F00-F99	Mood Disorders: 296 and 311 (ICD-9-CM) or F30-F34, F39 (ICD 10-CA) Anxiety Disorders: 300 (ICD-9-CM) or F40-F42	Mood disorders: 296 and 311 (ICD-9-CM) or F30-F34, F39 (ICD 10-CA) Anxiety disorders: 300 (ICD-9-CM) or F40-F43 (ICD-10- CA) Sleep disorders: 307, 780 or F51, G47 ICD-10-CA)
Case Definition	≥1 hospitalization or	≥1 hospitalization or	≥1 hospitalization or
	outpatient medical	≥1-3 outpatient	≥3 outpatient
	claim within 1 year	medical claims within	medical claims within
	,	3-5 years*	5 years**

^{*}Range of similar definitions between studies from 2000 to 2016

^{**}The decision to use a 5-year pre-exposure window was based on the fact that all patients received a BZRA, which itself increases specificity for anxiety/sleep disorder diagnoses.

Table A2 – International Classification for Disease Coding Algorithms for Seizure, Cancer and Palliation (Cohort Exclusion)

	Seizure	Cancer and other Neoplasms	Palliation
ICD Codes	9-CM: 345 10-CA: G40	9-CM: 140-165, 170- 176,179-195, 200-208	N/A*
Case Definition	≥1 hospitalization or ≥3 outpatient medical claim within 5 years before index date	10-CA: C00-C99 ≥1 hospitalization or ≥3 outpatient medical claims within 5 years before index date	Carrier code indicating palliative drug program enrollment in DPIN

^{*}While ICD codes do exist for palliation, the DPIN carrier code '04' is expected to be a reliable indicator of when patients become ill enough that community use of medication is required for symptom management.

Table A3 – Independent 'Patient' Variables for Prediction of Long-Term BZRA Use

Baseline Patient Characteristics		
Age	3 age groups; 18-44, 45-64, 65+ (Ordinal)	Index Date
Sex	Male or Female (Dichotomous Categorical)	Index Date
Region	Urban; Winnipeg or Brandon postal-codes Rural; Any other Manitoba postal-code (Dichotomous Categorical)	Census Period closest in time to the index date
Socioeconomic Status	Socio-Economic Factor Index – Version 2 (SEFI-2) score composite of four variables based on geography; i) unemployment rate ii) average household income iii) proportion of single-parent households iv) proportion of population without high school education. Scores <0 indicate more favourable socioeconomic conditions Scores >0 indicate less ideal socioeconomic conditions (Ordinal Scale)	Census Period closest in time to the index date
Income Assistance	Record of income assistance (Dichotomous Categorical)	Up to 1-year before the Index Date
Marriage Record	Record of Marriage (Dichotomous Categorical)	Entire available registry period up to the Index Date
Residential Mobility (i.e frequent mover)	Average of 1 move every 3 years from beginning of registry coverage to index date (Dichotomous)	Entire available registry period up to the Index Date
Comorbidity Burden	Charlson Comorbidity Index (CCI) Score; 0, 1, 2+ (Ordinal Scale)	Up to 1-year before the Index Date
Healthcare Resource Use	Johns Hopkins Adjusted Clinical Groups Resource Utilization Band (Ordinal Scale); placement into a band (0 to 5) based on grouping of	Up to 1-year before the Index Date

	ICD	
Prescription Psychotropic Use (non-BZRA)	Receipt of Prescription (Dichotomous Categorical)	Up to 1-year before the Index Date and 6 months after the Index Date
Prescription Opioid Use	Receipt of Prescription (Dichotomous Categorical)	Up to 1-year before the Index Date and 6 months after the Index Date

Table A4 - Independent 'First-Prescription' Variables for Prediction of Long-Term BZRA
Use

Characteristics of First Consultation and Subsequent Prescription	Definition	Measurement Period
Fiscal Year Period	Fiscal year of first prescription Assigned to 3 five-year intervals; 2001-2005, 2006- 2010, 2011-2015 (Ordinal)	Index Date
Prescriber	10 Years or More (Dichotomous)	Index Date
Sex of Prescriber	Male or Female (Dichotomous)	Index Date
Prescriber Specialty	General Practitioner, Psychiatry or Other (Categorical)	Index Date

Table A5 – Logistic Regression Methodology

Criteria	Approach			
Variable Selection	-Informal selection via published literature			
variable Selection	-Simple logistic regression; β values (p < 0.25)			
	-Dichotomous Categorical; 0 or 1			
	-Ordinal; discrete number scale starting at 1			
Variable Coding				
without couring	-Polychotomous Categorical; 0 or 1 with auto-			
	generated dummy variables			
	-No continuous variables retained			
Events-per-Variable	-Minimum 10 events per independent variable rule			
Conformity of Linear Gradient	-Ordered categorical variables assessed for conformity of linear gradient; nonconformity handled by variable transformation or separation into additional (design) variables (i.e fiscal year was shown to be linear with respect to outcome so condensed variable into 5-year increments)			
Interaction effects	-Assessed at p < 0.01. Suspected interactions included; age*sex, residential mobility*SEFI*income assistance, psychotropic use*opioid use, RUB*CCI			
Collinearity	-Analysis of variance inflation factor, correlation coefficients, eigenvalues			

	-Significant collinearity; combine variables or removal of inferior explanatory variable
Statistical Significance	-Wald 95% CI for β and OR's
Goodness-of-Fit Measures	-C-statistic, Log-Likelihood Ratio, Hosmer- Lemeshow Statistic
Fitting Procedure	-Stepwise addition/subtraction of variables -Assessment of clinical significance



Table A6 – Goodness of Fit for Final Logistic Regression Models Predicting Fong-Term Use of **BZRA**

Model	Model Type	Independent Variables	Likelihood Ratio (higher is better)	C statistic	Hosmer- Lemeshow Chi-Square Statistic
1	Main-Effects	9 Variables; Age-Sex Category, Period of First Rx, Psychotropic Use, Opioid Use, Income Assistance, Marriage, RUB CCI Score, Residential Mobility	6932 (p < 0.001)	l. Downloaded from http://bmjopen.bmj.c 738	$ \begin{array}{c} 10.78 \\ (p = 0.215) \end{array} $
2	Main-Effects + Interaction Effects	10 Variables: All from Model 1 + Residential Mobility*Income Assistance	6945 (p < 0.001)	0.739 0024 by gue	$ \begin{array}{c} 11.02 \\ (p = 0.20) \end{array} $

Table A7 – Proportion of Long-Term BZRA Use by Differing Parameters and Duration Thresholds

Scenario*	Long-Term Use Parameter	Prescription Lapse Criteria	Patients (n)	Proportion of Cohort
A1**	First-Use Episode ≥ 180 days	30 days or 50% of previous Day Supply	9,327	4.51%
A2	First-Use Episode ≥ 90 days	30 days or 50% of previous Day Supply	13,745	6.64%
A3	First-Use Episode ≥ 60 days	30 days or 50% of previous Day Supply	19,948	9.64%
A4	First-Use Episode ≥ 180 days	60 Days or 50% of previous Day Supply	13,050	6.31%
A5	First-Use Episode ≥ 180 days	90 Days	16,831	8.13%
A6	First-Use Episode ≥ 270 days	90 Days	15,214	7.35%
A7	First-Use Episode ≥ 365 days	90 Days	14,219	6.87%
B1	Mean Episode Duration ≥ 180 days	30 days or 50% of previous Day Supply	38,853	18.78%
B2	Mean Episode Duration ≥ 90 days	30 days or 50% of previous Day Supply	58,442	28.24%
В3	Mean Episode Duration ≥ 60 days	30 days or 50% of previous Day Supply	72,639	35.10%
B4	Mean Episode Duration ≥ 180 days	60 Days or 50% of previous Day Supply	44,593	21.55%
B5	Mean Episode Duration ≥ 180 days	90 Days	50,142	24.23%
В6	User Mean Episode Duration ≥ 270 days	90 Days	39,395	19.04%
В7	User Mean	90 Days	32,200	15.56%

Episode Duration		
≥ 365 days		

^{*}A=First Episode Scenario; B=Mean Episode Duration Scenario

 ${\bf Table~A8-Proportion~of~Long-Term~Z-Drug~Use~by~Differing~Parameters~and~Duration~Thresholds}$

Scenario	Long-Term Use Parameter	Prescription Lapse Criteria	Patients (n)	Proportion of Sub-Cohort
A1	First-Use Episode ≥ 180 days	30 days or 50% of previous Day Supply	8,206	7.41%
A2	First-Use Episode ≥ 90 days	30 days or 50% of previous Day Supply	12,155	11.0%
A3	First-Use Episode ≥ 60 days	30 days or 50% of previous Day Supply	17,126	15.5%
A4	First-Use Episode ≥ 180 days	60 Days or 50% of previous Day Supply	10,437	9.43%
A5	First-Use Episode ≥ 180 days	90 Days	12,719	11.49%
A6	First-Use Episode ≥ 270 days	90 Days	11,117	10.04%
A7	First-Use Episode ≥ 365 days	90 Days	10,045	9.07%
B1	User Mean Episode Duration ≥ 180 days	30 days or 50% of previous Day Supply	21,859	19.75%
B2	User Mean Episode Duration ≥ 90 days	30 days or 50% of previous Day Supply	32,020	28.92%
В3	User Mean Episode Duration ≥ 60 days	30 days or 50% of previous Day Supply	39,690	35.85%
B4	User Mean Episode Duration ≥ 180 days	60 Days or 50% of previous Day Supply	24,098	21.77%
B5	User Mean	90 Days	26,477	23.92%

^{**}Primary Scenario Used for Logistic Regression

	Episode Duration			
	≥ 180 days			
	User Mean	90 Days		
B6	Episode Duration		21,040	19.01%
	≥ 270 days		·	
	User Mean	90 Days		
B7	Episode Duration		17,358	15.68%
	≥ 365 days			

Table A9 – Patient Characteristics of Z-Drug Users by First Use Episode Duration

	9	Short-term	Long-term	Total
Number of U	Users	102,459 (100%)	8,204 (100%)	110,663 (100%)
Sex Distribution	Male	40,516 (39.5%)	3,473 (42.3%)	43,989 (39.8%)
Sex Distribution	Female	61,943 (60.5%)	4,731 (57.7%)	66,674 (60.2%)
	18-44	42,663 (41.6%)	1,795 (21.9%)	44,458 (40.2%)
Age Category	45-64	39,817 (38.9%)	3,184 (38.8%)	43,001 (38.9%)
	65+	20,011 (19.5%)	3,227 (39.3%)	23,238 (21.0%)
	<-1	13,678 (13.3%)	981 (12.0%)	14,659 (13.2%)
CEEL 2 Cooks	-1 to 0	45,136 (44.1%)	3,674 (44.8%)	48,810 (44.1%)
SEFI-2 Score	0 to 1	33,719 (32.9%)	2,885 (35.2%)	36,604 (33.1%)
	>1	9,958 (9.7%)	666 (8.1%)	10,624 (9.6%)
Residence	Urban	63,207 (61.7%)	3,313 (40.4%)	66,520 (60.1%)
Distribution	Rural	39,284 (38.3%)	4,893 (59.6%)	44,177 (39.9%)
High Residential	Mobility	22,408 (21.9%)	2,523 (30.8%)	24,931 (22.5%)
Receipt of Income	Receipt of Income Assistance		758 (9.2%)	9,109 (8.2%)
Marriage Record		57,308 (55.9%)	4,595 (56.0%)	61,903 (55.9%)

			I	I
	0 (no utilization)	1,771 (1.7%)	234 (2.9%)	2,005 (1.8%)
	1	3,205 (3.1%)	175 (2.1%)	3,380 (3.1%)
Johns Hopkins Healthcare	2	17,523 (17.1%)	1,012 (12.3%)	18,535 (16.7%)
Resource Utilization	3	65,067 (63.5%)	4,699 (57.3%)	69,766 (63.0%)
Band	4	10,810 (10.6%)	1,259 (15.3%)	12,069 (10.9%)
	5 (high- utilization)	4,083 (4.0%)	825 (10.1%)	4,908 (4.4%)
		Short-term	Long-term	Total
Number of U	Jsers	102,459 (100%)	8,204 (100%)	110,663 (100%)
Charlson	0	72,490 (70.8%)	4,528 (55.2%)	77,018 (69.6%)
Comorbidity index	1	19,495 (19.0%)	1,905 (23.2%)	21,400 (19.3%)
Score	2+	10,506 (10.3%)	1,773 (21.6%)	12,279 (11.1%)
Non-BZRA	O	27,797 (27.1%)	1,784 (21.7%)	29,581 (26.7%)
Psychotropic Prescription	1	36,939 (36.1%)	2,156 (26.3%)	39,095 (35.3%)
Dispensations	2+	37,755 (36.8%)	4,266 (52.0%)	42,021 (38.0%)
	0	47,427 (46.3%)	3,298 (40.2%)	50,725 (45.8%)
Opioid Prescription Dispensations	1	34,505 (33.7%)	2,772 (33.8%)	37,277 (33.7%)
	2+	20,559 (20.1%)	2,136 (26.0%)	22,695 (20.5%)
Sex of Prescriber Issuing First Prescription	Male	71,485 (69.8%)	5,627 (68.6%)	77,112 (69.7%)
	Female	28,485 (27.8%)	2,273 (27.7%)	30,758 (27.8%)
Age of Prescriber Issuing First Prescription	50+ Years	47,871 (46.7%)	4,014 (48.9%)	51,885 (46.9%)
	<50 Years	49,257 (48.1%)	3,758 (45.8%)	53,015 (47.9%)

Type of Prescriber	General Practitioner	78,610 (76.7%)	6,366 (77.6%)	84,976 (76.8%)
Issuing First Prescription	Psychiatry	3,912 (3.8%)	475 (5.8%)	4,387 (4.0%)
Prescription	Other	3,881 (3.8%)	381 (4.6%)	4,262 (3.9%)
Period of First Prescription	2001-2006	34,360 (33.5%)	1,526 (18.6%)	35,886 (32.4%)
	2006-2011	37,752 (36.8%)	2,808 (34.2%)	40,560 (36.7%)
	2011-2016	30,379 (29.6%)	3,872 (47.2%)	34,251 (31.0%)

Table A10 – Frequency of Charlson Comorbidity Group Diagnoses by First Use Episode Duration for Z-Drug Cohort

Charlson Diagnosis	Short-Term	Long-Term 'First-		
	'First-Episode'	Episode' Users	Z-Test of Two	
	Users	(n=8,204)	Proportions	
	(n=102,459)			
Myocardial Infarction	1,836 (1.8%)	306 (3.7%)	p < 0.01	
Congestive Heart Failure	3,174 (3.1%)	700 (8.5%)	p < 0.01	
Peripheral Vascular Disease	1,772 (1.7%)	284 (3.5%)	p < 0.01	
Cerebrovascular Disease	2,321 (2.3%)	550 (6.7%)	p < 0.01	
Dementia	1,925 (1.9%)	865 (10.5%)	p < 0.01	
COPD	12,357 (12.1%)	1,171 (14.3%)	p < 0.01	
Connective	1 006 (1 00/)	243 (3.0%)	p < 0.01	
Tissue/Rheumatic Disease	1,906 (1.9%)	243 (3.0%)	p < 0.01	
Peptic Ulcer Disease	1,111 (1.1%)	123 (1.5%)	p < 0.01	
Mild Liver Disease	1,672 (1.6%)	139 (1.7%)	p = 0.33	
Moderate/Severe Liver Disease	275 (0.2%)	38 (0.4%)	p < 0.01	
Uncomplicated Diabetes	9,317 (9.1%)	1,150 (14.0%)	p < 0.01	
Complicated Diabetes	1,639 (1.6%)	328 (4.0%)	p < 0.01	
Paraplegia and Hemiplegia	508 (0.5%)	136 (1.7%)	p < 0.01	
Renal Disease	1,543 (1.5%)	293 (3.6%)	p < 0.01	
Cancer	2,109 (2.1%)	247 (3.0%)	p < 0.01	
Metastatic Carcinoma	429 (0.4%)	45 (0.5%)	p = 0.04	
HIV/AIDS	118 (0.1%)	16 (0.2%)	p = 0.02	

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Table A11 – Statistical Associations between Predictor Variables and Long-term Use of Z-Drugs

				Use D	uration	916	
Independent Variable		≥180 days		≥90 days		⁹ ≥ 60 days	
		Crude OR (95% CI)	Adjusted OR (95% CI)	Crude OR (95% CI)	Adjusted OR (95% CI)	Cru & OR (95% CI)	Adjusted OR (95% CI)
Мс	ale	1.12 (1.07-1.18)	1.04 (0.99-1.09)	1.13 (1.08-1.17)	1.05 (1.01-1.10)	1∯8 (1.05≋1.12)	1.04 (1.00-1.08)
	18-44	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 (<u>r</u> ef)	1 (ref)
Age	45-64	1.90 (1.79-2.02)	2.02 (1.89-2.17)	1.74 (1.66-1.82)	1.78 (1.68-1.88)	1ৠ1 (1.6세 1.78)	1.68 (1.60-1.76)
	<i>65</i> +	3.83 (3.61-4.07)	3.71 (3.44-4.00)	3.24 (3.08-3.40)	3.08 (2.90-3.28)	2899 (2.87₹3.12)	2.78 (2.64-2.93)
Rural Re	Rural Residence		1.13 (1.07-1.19)	0.99 (0.96-1.03)	1.02 (0.98-1.07)	1308 (1.04=1.11)	0.95 (0.91-0.99)
High Residen	High Residential Mobility		1.26 (1.19-1.33)	1.53 (1.46-1.59)	1.21 (1.15-1.27)	1 <u>3</u> 0 (1.261.35)	1.12 (1.07-1.17)
Income A	Income Assistance		1.47 (1.34-1.61)	1.02 (0.95-1.09)	1.29 (1.19-1.40)	032 (0.77=0.87)	1.08 (1.00-1.17)
	<-1	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 (% ef)	1 (ref)
SEFI-2 Score	-1 to 0	1.14 (1.06-1.22)	1.07 (0.99-1.16)	1.03 (0.97-1.09)	0.98 (0.92-1.04)	0.55 (0.9 \(\bar{b} \)1.00)	0.94 (0.89-0.99)
	0 to 1	1.19 (1.11-1.29)	1.08 (0.99-1.17)	1.04 (0.98-1.11)	0.99 (0.93-1.06)	0. 3 2 (0.8₹0.97)	0.93 (0.88-0.99)
	>1	0.93 (0.84-1.03)	0.84 (0.75-0.94)	0.80 (0.73-0.87)	0.77 (0.70-0.85)	0 <u>6</u> 8 (0.63 <u>-</u> 0.73)	0.72 (0.66-0.78)
Married		1.00 (0.96-1.05)	0.86 (0.82-0.91)	1.07 (1.03-1.10)	0.93 (0.89-0.98)	1┋3 (1.10월1.17)	0.98 (0.94-1.01)
		1.28	1.15	1.26	1.15	1.48	1.11

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Opioid Use		(1.22-1.34)	(1.09-1.21)	(1.21-1.31)	(1.11-1.20)	(1.1461.21)	(1.07-1.15)	
		Use Duration ♀						
Indonandar	nt Variabla	≥180	O days	≥90 days			days	
<u>Independent Variable</u>		Crude OR (95% CI)	Adjusted OR (95% CI)	Crude OR (95% CI)	Adjusted OR (95% CI)	Crude OR (95% CI)	Adjusted OR (95% CI)	
Psychotropic Rx	Use (Non-BZRA)	1.34 (1.27-1.41)	1.24 (1.17-1.32)	1.35 (1.29-1.41)	1.27 (1.20-1.33)	1 <u>8</u> 2 (1.17-1.27)	1.19 (1.14-1.24)	
	0	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 (g ef)	1 (ref)	
Charlson Comorbidity	1	1.56 (1.48-1.65)	1.25 (1.18-1.33)	1.45 (1.39-1.52)	1.21 (1.15-1.27)	1 🗟 3 (1.28 1.38)	1.13 (1.08-1.19)	
Index Score	2+	2.70 (2.55-2.87)	1.46 (1.36-1.58)	2.34 (2.22-2.46)	1.38 (1.29-1.47)	2 2 2 (1.93 2.12)	1.30 (1.22-1.37)	
	0-3 (≤Moderate)	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 (₹ef)	1 (ref)	
Resource Utilization Band	4 (High)	1.67 (1.56-1.78)	1.16 (1.08-1.25)	1.47 (1.39-1.56)	1.09 (1.01-1.16)	1 <u>3</u> 0 (1.24 <u>6</u> 1.37)	1.00 (0.95-1.07)	
Unitzation Bana	5 (Very High)	2.89 (2.67-3.13)	1.55 (1.41-1.70)	2.43 (2.26-2.61)	1.42 (1.30-1.55)	1.85 (1.85 2.11)	1.22 (1.12-1.32)	
Male Prescriber of	Male Prescriber of First Prescription		0.97 (0.92-1.03)	0.98 (0.94-1.02)	0.98 (0.93-1.02)	0.94 (0.90±0.97)	0.93 (0.90-0.97)	
Prescriber Aş	Prescriber Age ≥50 Years		0.98 (0.93-1.03)	1.10 (1.06-1.15)	0.98 (0.94-1.02)	1월5 (1.1년1.19)	1.05 (1.01-1.09)	
	GP	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 (<u>ke</u> f)	1 (ref)	
Prescriber of First Prescription	Psychiatrist	1.50 (1.36-1.66)	1.96 (1.76-2.17)	1.36 (1.25-1.49)	1.72 (1.57-1.89)	1 ½ 1 (1.02 1.20)	1.38 (1.27-1.51)	
τ ποι τ τεοστιμιίση	Other	1.21 (1.09-1.35)	0.92 (0.82-1.03)	1.18 (1.07-1.29)	0.91 (0.83-1.00)	159 (1.10-1.29)	0.98 (0.91-1.07)	
Davied of First	2001-2006	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 (F ef)	1 (ref)	
Period of First Prescription	2006-2011	1.68 (1.57-1.79)	1.57 (1.46-1.68)	1.67 (1.59-1.76)	1.56 (1.47-1.66)	1863 (1.4691.60)	1.46 (1.39-1.54)	

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2011-2015	2.87 (2.70-3.05)	2.45 (2.28-2.65)	2.83 (2.69-2.97)	2.44 (2.30-2.59)	2 2 0 (2.10 <u>2</u> 2.29)	1.96 (1.86-2.07)
	, , , , , ,		<u> </u>		on 1 7	(1.80-2.07)

STROBE Statement—checklist of items that should be included in reports of observational studies

Item No	Recommendation	Page No
1	(a) Indicate the study's design with a commonly used term in the title or	1
	the abstract	
	(b) Provide in the abstract an informative and balanced summary of what	3
	was done and what was found	
2	Explain the scientific background and rationale for the investigation being reported	5
3	•	6
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4	Present key elements of study design early in the paper	6
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	(c) Explain now missing data were addressed	10-
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		11
	controls was addressed	
	<i>Cross-sectional study</i> —If applicable, describe analytical methods taking account of sampling strategy	
	No 1	(a) Indicate the study's design with a commonly used term in the title or the abstract (b) Provide in the abstract an informative and balanced summary of what was done and what was found

Continued on next page



Results			1
Participants	13*	(a) Report numbers of individuals at each stage of study—eg numbers potentially	11-
		eligible, examined for eligibility, confirmed eligible, included in the study,	12
		completing follow-up, and analysed	
		(b) Give reasons for non-participation at each stage	7
		(c) Consider use of a flow diagram	7,
			fig1
Descriptive	14*	(a) Give characteristics of study participants (eg demographic, clinical, social) and	11,12
data		information on exposures and potential confounders	
		(b) Indicate number of participants with missing data for each variable of interest	Fig 1
		(c) Cohort study—Summarise follow-up time (eg, average and total amount)	11
Outcome data	15*	Cohort study—Report numbers of outcome events or summary measures over time	12-
			13
		Case-control study—Report numbers in each exposure category, or summary	n/a
		measures of exposure	
		Cross-sectional study—Report numbers of outcome events or summary measures	n/a
Main results	16	(a) Give unadjusted estimates and, if applicable, confounder-adjusted estimates and	Table
		their precision (eg, 95% confidence interval). Make clear which confounders were	2
		adjusted for and why they were included	
		(b) Report category boundaries when continuous variables were categorized	Table
			2
		(c) If relevant, consider translating estimates of relative risk into absolute risk for a	Table
		meaningful time period	2
Other analyses	17	Report other analyses done—eg analyses of subgroups and interactions, and	Table
		sensitivity analyses	2-3
Discussion			
Key results	18	Summarise key results with reference to study objectives	13
Limitations	19	Discuss limitations of the study, taking into account sources of potential bias or	17
		imprecision. Discuss both direction and magnitude of any potential bias	
Interpretation	20	Give a cautious overall interpretation of results considering objectives, limitations,	13-
		multiplicity of analyses, results from similar studies, and other relevant evidence	17
Generalisability	21	Discuss the generalisability (external validity) of the study results	17-
			18
Other informati	ion		
Funding	22	Give the source of funding and the role of the funders for the present study and, if	25
		applicable, for the original study on which the present article is based	

^{*}Give information separately for cases and controls in case-control studies and, if applicable, for exposed and unexposed groups in cohort and cross-sectional studies.

Note: An Explanation and Elaboration article discusses each checklist item and gives methodological background and published examples of transparent reporting. The STROBE checklist is best used in conjunction with this article (freely available on the Web sites of PLoS Medicine at http://www.plosmedicine.org/, Annals of Internal Medicine at http://www.annals.org/, and Epidemiology at http://www.epidem.com/). Information on the STROBE Initiative is available at www.strobe-statement.org.

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Risk of Long-Term Benzodiazepine and Z-Drug Use Following the First Prescription among Community-Dwelling Adults with Anxiety/Mood and Sleep Disorders: A retrospective cohort study

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Risk of Long-Term Benzodiazepine and Z-Drug Use Following the First Prescription among Community-Dwelling Adults with Anxiety/Mood and Sleep Disorders: A retrospective cohort study

Running Title: Risk of long-term benzodiazepine receptor agonist use

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ABSTRACT

Objective: To measure the incidence of long-term benzodiazepine receptor agonist (BZRA) use among individuals with anxiety, mood and/or sleep disorders. To identify factors associated with long-term use following the first prescription.

Methods: This was a population-based retrospective cohort study using administrative databases in Manitoba, Canada. Individuals with anxiety/mood or sleep disorder who received their first BZRA between April 1, 2001 to March 31, 2015 were included. Long-term use was defined as ≥180 days. Logistic regression modelling was used to examine predictors of long-term use.

Results: Among 206,933 individuals included, long-term BZRA use in the first episode of use was 4.5% (≥180 days) following their first prescription. Factors associated with ≥180 days of use included male sex (adjusted odds ratio (aOR) 1.33, 95% CI 1.27 to 1.39), age ≥65 (aOR 5.15, 95% CI 4.81 to 5.52), income assistance (aOR 1.68, 95% CI 1.55 to 1.81), previous non-BZRA psychotropic (aOR 1.93 (95% CI 1.83 to 2.02) or opioid use (aOR 1.16, 95% CI 1.11 to 1.22), high comorbidity (aOR 1.43, 95% CI 1.32 to 1.55), high healthcare use (aOR 1.46 (95% CI 1.33 to 1.60), and psychiatrist prescriber (aOR 2.11, 95% CI 1.93 to 2.32).

Conclusions: Less than one in twenty patients use BZRAs \geq 180 days in their first treatment episode. Several factors were associated with long-term use following the first prescription and further investigation into whether these factors need to be considered at the point of prescribing is warranted. In light of these findings, future research should examine the predictors of cumulative repeat episodes of BZRA exposure.

Key Words: benzodiazepine, anxiety, insomnia, pharmacoepidemiology, clinical practice guidelines, z-drug hypnotics

Strengths and Limitations of Study

- This study used administrative data from the Manitoba Centre for Health Policy, which is one of the most comprehensive datasets in North America containing >140 de-identified linked datasets on healthcare, education, social/families, justice and registries for all residents of Manitoba (population of 1.4 million people) not restricted by age or income
- All diagnoses are identified through physician claims data or hospitalizations, which are
 dependent on people seeking treatment and may be prone to some misclassification. Drug
 information is also based on dispensing records from community pharmacies and does not
 confirm the patient actually took the drug. However, we performed multiple sensitivity
 analyses to address this.
- The databases do not capture participation in psychological interventions such as cognitive behavioral therapy.

Introduction

The use of benzodiazepine receptor agonists (BZRAs)*, benzodiazepines (BZD) and Z-Drugs, in the treatment of anxiety and insomnia has shifted based on the evolving data on safety risks and limited efficacy on long-term use in the literature. ¹⁻⁴ Upon their initial introduction into clinical practice in the late 1960s, benzodiazepines were considered to be a safer alternative to barbiturates.⁵ However, safety concerns such as psychomotor impaired accidents (i.e., falls and motor-vehicle accidents), dependency and misuse/abuse are now well known. ⁶⁻⁸ Recent studies have also raised concerns proposing possible links to dementia, recurrence of mood episode, respiratory disease exacerbation and suicide with long-term BZRA use. ⁹⁻¹³ However, the association of BZRA use for these newer harms is uncertain given conflicting evidence and confounding in previous studies.¹⁴

In spite of ongoing adverse effect concerns, justification for less restrictive BZRA use have stemmed from their clinical utility as rapidly effective anxiolytic sedatives.¹⁵ Some view that limiting BZRA use is at times impractical.¹⁶ Moreover, the use of alternative pharmacotherapy, including trazodone, atypical antipsychotics, barbiturates, and tricyclic antidepressants are not without adverse effects. It should also be noted that the difficulties with de-prescribing BZRAs reported in the literature have added caution to the initiation of these agents in practice.^{4,17}

Previous studies examining the pattern of BZRA use have found a decline in benzodiazepine (particularly lorazepam) incident use and an increase in the incidence of Z-drug use. ^{18,19} Limited studies have examined predictors of long-term use after a first prescription. ^{20,21} As such, this study sought i) to measure the incidence of long-term BZRA use among a cohort of

^{*} Abbreviations: BZRA, benzodiazepine receptor agonist; BZD, benzodiazepine

community-dwelling Canadian adults with anxiety, mood and/or sleep disorders, and ii) to determine factors associated with progression to long-term BZD use following the first prescription in this population.

Methods

Study Design and Data Sources

This was a retrospective, cohort study using routinely collected administrative healthcare data pertaining to prescription drug dispensations, outpatient physician claims, hospitalization discharge abstracts, income assistance records and prescriber demographics (Table 1). All data used was extracted from the Manitoba Centre for Health Policy Population Research Data Repository. The Repository provides comprehensive coverage of all Manitoba residents contact with the primary healthcare system. The Drug Program Information Network (DPIN) provides information on outpatient prescription drugs dispensed in Manitoba with the exception of medications dispensed in hospital and nursing stations. In Manitoba, eligible outpatient prescriptions are 100% covered for residents after an income-based deductible is paid for each fiscal year. DPIN captures information on the drug name, strength, quantity, day-supply, and date of all outpatient prescriptions dispensed regardless of coverage. Merging of the various data sources was facilitated via linkage of unique de-identified Personal Health Information Numbers. The Charlson Comorbidity score $[0 \text{ (lowest risk)}, 1, \ge 2 \text{ (high risk)}]$ was also determined to examine the effects of comorbidity of duration of use. This was determined based on 17 categories of comorbidities using ICD-9-CM or ICD-10-CA equivalent codes in administrative data to provide the weight-based adjusted risk of death or resource use.²²

Cohort Inclusion/Exclusion Criteria

Eligible patients were adults age 18 years and older who initiated a new benzodiazepine or Z-Drug prescription (defined as no use in the one year prior to the first prescription^{20,21}) between April 1, 2001 to March 31, 2015, with no preceding dispensations from April 1, 2000 to March 31, 2001 (first year of the dataset) to avoid prevalent user bias (Figure 1). All individuals with at least one year of registry coverage prior to and after the first prescription was required for cohort inclusion. As such, individuals who received a benzodiazepine in the distant past could be included in the cohort as a new user, provided that the benzodiazepine was not used in the past one year. A sensitivity analysis was also performed in which incident use was defined as no prescription for a BZRA was received in the three years prior to the first prescription.²³

Eligibility was also based on diagnostic criteria for anxiety/mood related disorders and/or insomnia based on ICD-9-CM or ICD-10-CA medical claims, either at outpatient physician visits or hospitalizations, occurring within a 5-year period prior to the first prescription. The ICD diagnostic criteria chosen are a combination of the definitions from two sources; the Canadian Public Health Association on mental health surveillance and the MCHP concept dictionary, which listed the various past-case definitions employed in previous research within Manitoba for mood and anxiety disorders (Table A1).²⁴⁻²⁸ Lastly, because reliance on ICD codes is expected (and has been previously shown) to underestimate capture of sleep disorder cases, we also accepted receipt of a Z-Drug in the definition for insomnia as this was their sole approved indication.²⁹

To reduce confounding, we established cohort exclusion criteria that otherwise may have justified long-term use of BZDs in clinical scenarios beyond the scope of general guideline recommendations for anxiety and insomnia. Namely, patients were excluded if they had ≥1 ICD

code for cancer, a seizure disorder or if there was placement in the Manitoba palliative care drug program at any point in the 5 years preceding their first prescription for a BZRA (Table A2). Where patients became palliative ≥1 year after the initial BZRA dispensation, their ongoing use of BZRA was censored beginning from the date of their placement, but all use prior to their palliative status was retained. Clobazam use was excluded entirely from the evaluated drug claims because it is approved only as an adjunctive agent for epilepsy in Canada. Finally, patients were excluded if they lacked at least 1-year of registry coverage from their first-prescription index date. This was to eliminate any biasing effect from early mortality, moving out of province or other loss to follow-up.

Main Outcome Measures

Long-term use was defined as ≥180 days based on the recommendation from a previous systematic review of similar studies (Figure 2).²⁴ This duration is longer than clinical practice guideline duration recommendations and is believed to be of sufficient length for risk of dependence to occur.³⁰ One-third of individuals who use BZDs for longer than six months have been previously reported to be unable to stop completely due to withdrawal symptoms (e.g., anxiety, insomnia, muscle spasms).³⁰ A sensitivity analysis, ranging from 60 to 365 days, was also used in our study to account for varying definitions of long-term use reported in the literature.²⁴

Patients were followed forward in time from the date of their first BZRA prescription. BZRA 'use episodes' were determined according to consecutive prescription overlap based on dispensation dates and coded day supply values. The allowable gap between prescriptions was the greater of either 30 days or 50% of the last prescription day supply after the prescription end date (end date = dispensation date + day-supply) of the prior prescription. This gap was chosen to account for those who regularly or frequently used "as needed" BZRA in the 'use episode' duration

(Figure 3). The episode end date was calculated as the date of the last prescription in a given 'use episode' plus its associated day-supply. To account for immeasurable time bias, hospitalization time was assumed to be a continuation of BZD use given that in-patient drug use data was limited.³¹ The provincial drug program subsidizes dispensations of up to a 100 day-supply.

Individuals were able to have multiple use episodes over the entire study duration. First episode duration and average episode duration were calculated for each user. If patients only had one use episode both of these values were the same. Patients were allowed to switch from one BZRA to another without it interrupting their 'use episodes'. This included switching from a BZD to a Z-drug and vice versa.

Independent Variables

Variables used for statistical prediction of long-term use were determined a priori and included age, sex, geographic residence, residential mobility, socioeconomic status, marriage, concurrent opioid or prescription psychotropic use, comorbidity burden, healthcare usage, time period of first prescription and prescriber characteristics (Table A3 and Table A4). Variables were assessed at baseline; either within 1-year before the index date, at the index date or up to 6-months past the index date (in the case of prescription opioids and other psychotropics, such as antidepressants, antipsychotics, and mood stabilizers).

Statistical Analysis

Standard reporting criteria were followed in the approach to logistic regression modelling (Table A5 and A6).³² Univariate analysis was performed first in the form of simple logistic regression. The multi-variable model was constructed to determine the most parsimonious model for prediction of long-term BZRA use defined as \geq 180 days in the first episode of use with

adjustment of clinically relevant covariates based on previous literature.²⁴ Differences between models in their maximum log-likelihood estimation, likelihood ratios and other goodness-of-fit statistics enabled model discrimination.³² Multicollinearity and effect-measure modification (i.e., interaction effects) were assessed when it was suspected that variables may be either correlated or non-independent.³² In order to perform these diagnostics, the binary dependent variable was first substituted for a linear variable (first-episode duration in days) to conduct a multiple *linear* regression. Specifically, collinearity was determined to be a model threat if any correlation coefficient in the independent variable correlation matrix was $\geq |0.8|$ or if any variance inflation factor was unreasonably high (≥ 10) while the corresponding tolerance factor was miniscule (≤ 0.1).³³ Analyses were assessed at p ≤ 0.01 threshold set a priori for statistical significance.

For the multiple logistic regression, 'complete case-analysis' was used because the extent of missing data was too small to justify the need for multiple imputation procedures.³⁴ In this study, no claims were excluded on the basis of missing data fields. Only 1,568 claims (<0.01%) were excluded for being spurious (i.e '0' day/quantity supply or incredibly high dispensed quantity to day-supply ratio) Furthermore, observed missing data was believed to be missing at random.³⁵ The only variable with significant missing data was that of 'prescriber type' (~38,000 missing observations or 17.5% of final sample).

A subgroup analysis of each of the 17 categories of the Charlson Comorbidity Score was also performed using Z-test of two proportions to describe the specific comorbidities that may contribute to the relationship between Charlson Comorbidity Score and long-term use.

Sensitivity Analysis

To assess the robustness of the primary outcome, 6 sensitivity analyses (Tables A7 and Table A8) were conducted to determine how the proportion of long-term use changed under

differing parameter assumptions.³⁶ The threshold duration for long-term use was adjusted to values ranging from 60 days to 365 days. Additionally, the episode lapse criteria (i.e., prescription gap rule) was changed. While the analysis was not exhaustive for every conceivable combination of these key parameters, the selected values were chosen because they were judged to be representative of how peers in the international clinical community may have defined or measured 'long-term use' of BZRA. All data was cleaned and analyzed using SAS v9.4©.

Ethical Approval

Access to the data for this project was approved by the University's Health Research Ethics Board (HREB, registration number H2017:052 (HS20498) and the Health Information Privacy Committee (HIPC, no. 2016/2017-62) of the provincial government. Consent for this study was not required by HREB given the retrospective nature of the study and data agreements in place through HIPC.

Results

Episodic BZD/Z-Drug Use

Study population demographics are presented in Table 2. There were 206,933 patients in our cohort representing 931,271 unique BZRA dispensations over the 15-year study duration. Over the study period, cohort individuals had a median of three and average of 4.5 BZRA use episodes, respectively. First-episodes of use were of a median duration of 20 days (IQR = 10-30 days). For all use-episodes, the median average use duration was 30 days (IQR = 15-111 days). Evaluation of long-term use revealed that 4.51% of patients used a BZRA for ≥180-days in their 'first' episode

of use. At most, this proportion increased to 9.64% when a sensitivity analysis of 60 days or greater was used for the definition of 'long-term use' for the first episode of use. However, the proportion of long-term users increased considerably after averaging for all episodes for each user (sensitivity analysis range: 15.6%-35.1%) (Table A7).

To evaluate treatment duration for insomnia, a sensitivity analysis was performed on only Z-Drugs (*n*=110,663), which found similar results (Tables A8-A11).

Factors Predicting Long-term First Episode Use

Logistic regression analysis revealed that male sex (adjusted odds ratio (OR) 1.33, 95% CI 1.27 to 1.39), older age (adjusted OR 2.24 (95% CI 2.11 to 2.38) and 5.15 (95% CI 4.81 to 5.52) for aged 45-64 years and \geq 65 years, respectively, compared to <45 years), receipt of income assistance (adjusted OR 1.68, 95% CI 1.55 to 1.81), previous non-BZRA psychotropic (adjusted OR 1.93 (95% CI 1.83 to 2.02) or opioid use (adjusted OR 1.16, 95% CI 1.11 to 1.22), high comorbidity (Charlson Comorbidity Index 1 and ≥2, adjusted OR 1.11 (95% CI 1.04 to 1.17) and 1.43 (95% CI 1.32 to 1.55), respectively), high healthcare resource use (resource utilization band of 4 and 5, adjusted OR 1.15 (95% CI 1.07 to 1.23) and 1.46 (95% CI 1.33 to 1.60), respectively), first prescription from psychiatrist (adjusted OR 2.11, 95% CI 1.93 to 2.32) and receipt of first prescription after 2006 (2006-2011, adjusted OR 1.74, 95% CI 1.64 to 1.85; 2011-2015, adjusted OR 2.99, 95% CI 2.80 to 3.18), were all predictive of long-term use of \geq 180 days in the first episode. Rural residence (adjusted OR 1.10, 95% CI 1.04 to 1.15) and high residential mobility (adjusted OR 1.14, 95% CI 1.08 to 1.21) were also associated with a higher risk of long-term use in the first episode. Married status was associated with a lower risk of meeting the long-term use definition (adjusted OR 0.79, 95% CI 0.76 to 0.83). These findings were also replicated in the sensitivity analysis restricted to Z-Drug users. Both the crude and adjusted odds ratios are presented for the full cohort in Table 3.

A sub-analysis of the higher comorbidity scores in the long-term user groups shows that this relationship was mainly driven by cardiovascular diseases, diabetes and dementia (Table 4). Proportions for these particular diagnoses were 2 to 5 times higher in the long-term user group, with the greatest difference existing for dementia (long-term; 8.5% vs. short-term; 1.5%). A sensitivity analysis was performed changing the definition of incident user to no receipt of BZRA prescription in the three years prior to the first BZRA prescription. No change in results were found.

Discussion

This study found approximately 4.5% of the full cohort and 7.4% of the Z-Drug cohort were 'long-term' first-episode users according to the best available evidence-based consensus definition of 180 days. ²⁴ Restricting the analysis to Z-Drug use showed that the frequency of long-term use was higher than that of the main cohort. Practice guidelines typically recommend a shorter duration of use for Z-Drugs in the treatment of insomnia (range of ≤2-6 weeks)³⁷⁻³⁹ compared to BZD for anxiety disorder (up to ≤12 weeks depending on indication). ⁴⁰⁻⁴² Therefore, these results suggest greater disparity from practice guidelines in the case of Z-drug use for insomnia. Of note, more recent insomnia guidelines have recognized that while non-drug alternatives have a favourable safety profile, these interventions may be difficult to achieve for certain populations, which could explain the deviation between practice recommendations and real-world use of these agents. ³⁸

The proportion of patients who met criteria for 'long-term' use after accounting for all of their use-episodes (i.e., rather than just the first episode of use) was approximately 3.5 times higher than the proportion of patients meeting criteria after only their first episode of use. These results may indicate that repeated episodes of BZRA use may be associated with a higher risk of being exposed to a BZRA for a duration of ≥ 180 days in one episode. An area of future research is to examine whether repeated episodes of BZRA use is associated with progression to long-term use as demonstrated in a previous study that observed the number of episodes of dispensing in the first month was a significant predictor of the total duration of dispensing in the later period.⁴³ Of note, the majority of people with repeated use still only take BZRAs for intermittent, short-term periods. Furthermore, confounding variables such as age and accrued comorbidity over time may influence the risk of future long-term use in some patients. Nonetheless, these results support the observed difficulty in de-prescribing once BZRA use has become chronic, which has also been reported in previous literature. 4,44 Lastly, other clinical considerations such as risk of protracted withdrawal symptoms, risk of rebound insomnia and/or anxiety, severity of indication, patient dissatisfaction, limited alternate drug and non-drug interventions, or interference with another prescriber's decisions likely undermine potential de-prescribing efforts.

Older age and female sex have also been identified in previous studies as being associated with long-term use.^{45–51} While we found females to have greater representation in all patterns of BZRA use, we found males were more specifically predictive of long-term use after the first episode of use. ^{52–54} As with almost all of the previously published studies, older age was strongly associated with long-term BZRA use.⁵¹⁻⁵⁵ It should be noted that older individuals may have had a greater opportunity to be exposed to BZRA use.

As supported by previous evidence, income assistance was associated with long-term BZRA use ^{48,56}. Our study also found frequent moving, unmarried status, and rural residence to be associated with increased odds of long-term use. Frequency of moving and income assistance could be a proxy for general life stability ^{50,57,58}. Rural residence may have a small effect on longer-term BZRA use due to the relative limitations of timely scheduled follow-up, which may necessitate prescriptions of greater quantity or for longer periods. Another study also found rural adults to be at higher odds of inappropriate BZD use .⁵⁹

Healthcare use and the presence of various physical illnesses have been consistent predictors of long-term BZRA use ^{47,49,50,60}. In this study, as both of these variables increased, so did the odds of long-term use. We speculate that the positive relationship between these two indices and long-term use may be partially explained by unmeasured 'health' anxiety or associated mental health issues arising secondary to physical comorbidities or by additional disruptive effects of physical illness on sleep.

The Charlson comorbidity score findings were not surprising given the relatively higher proportion of older adults in the long-term use group. Nonetheless, the greater degree of BZRA exposure among those patients with dementia is of concern given the risk of BZD use in this population. Similar to previous studies, prescriptions for an opioid or a psychotropic agent, such as antidepressants, antipsychotics or mood stabilisers, during the baseline period were modestly predictive for future long-term use. 48,52,54,56,58,61 Those having received a non-BZD prescription agent for a psychiatric disorder could be expected to have had greater disease severity on average than those BZRA users who did not receive such treatment early on.

An unexpected finding was the increased odds of long-term use associated with the more recent time period of the first prescription. This is contrary to what may be expected from

cumulative knowledge on BZRA and the long-standing emphasis on short-term use advised in guidelines and clinical literature. This finding may reflect the growing awareness that BZRAs should not be used as a first-line treatment resulting in only those who have not responded to other alternatives to be more likely to receive BZRAs long-term.

The present study has a number of strengths. This study used a large administrative data source that were near complete in their coverage of the study population's prescription drug dispensations and healthcare contact. Application of cohort inclusion and exclusion criteria in a carefully constructed new user longitudinal design limited confounding and bias to the extent possible. Multiple sensitivity analyses on the main outcome measure, the duration of BZRA use measurement method and the association between the independent and dependent variables for two cohorts reduced quantitative bias to increase confidence in the results.

A few important limitations should be acknowledged. Firstly, administrative data is prone to some misclassification of variables. For instance, diagnostic criteria for cohort case inclusion and exclusion will differ in their true sensitivity and specificity, regardless of prior validation of case definitions. Drugs used during any hospitalizations were not available and was assumed to be continued BZD exposure. As all independent variables were only measured cross-sectionally before or at the time of the first prescription of the first use-episode, the logistic regression model was only predictively valid for the first use episode duration and not users' average episode duration. Since DPIN only captures the days supply provided, it is possible that not all of the medication was actually taken by the patient. However, this study was able to provide insight into the prescribing practices of benzodiazepines that are filled in the pharmacy in this population. Our study did not evaluate the extent of concurrent use of multiple BZD or other psychiatric diagnoses such as substance use disorder. The databases also do not capture participation in psychological

interventions such as cognitive behavioral therapy. Moreover, while the databases are able to link several data on health information regardless of age and coverage, they do not capture other potential confounding factors such as education status and ethnicity. This study was done in a setting where there is a universal healthcare system and medication costs are covered for all Manitobans after an income-based deductible is met every year. As a result, findings may be generalizable to similar settings. Future research should aim to examine the association of repeat exposure to BZRA and risk of chronic use. Future research could also examine specific benzodiazepine type and formulations on risk of long-term use.

Conclusion

Prescribing of BZRAs was used for less than six months duration for the majority of individuals with a prior history of anxiety, depression, or insomnia. However, the proportion of long-term use among new users was up to one in three based on the average of all episodes of use, warranting future research in this area. Patients who are male, of older age, are socially or financially deprived, have poor physical health, use opioids or other psychotropic agents and are frequent consumers of healthcare resources are more likely to use BZRA long-term after their first prescription. Future research could be done to explore whether these factors need to be considered at the point of prescribing in clinical practice.

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Availability of Data and Materials

Data used in this article was derived from administrative health and social data as a secondary use. The data was provided under specific data sharing agreements only for approved use at MCHP. The original source data is not owned by the researchers or MCHP and as such cannot be provided to a public repository.

Author Statement

Jaden Brandt contributed to the conception, design, acquisition of data, analysis, and writing of the manuscript. Donica Janzen contributed to the analysis, interpretation, and writing of the manuscript. Silvia Alessi-Severini contributed to the interpretation and writing of the manuscript. Dan Chateau contributed to the interpretation and analysis of the study. Murray Enns contributed to the interpretation and writing of the study. Alexander Singer contributed to the interpretation and writing of the study. Christine Leong contributed to the conception, design, interpretation, and writing of the manuscript.

Conflict of Interest Disclosure

The authors declare no conflicts of interest related to any aspects of this work.

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Patient and Public Involvement

We have a patient advisory group who provided feedback on the dissemination of research findings.

Table 1 –Raw Data Sources and Relevant Corresponding Data Elements

Database	Date Range of Data	Relevant Data Elements
Drug Program	Apr. 1/2000 –	Prescriptions for benzodiazepines (ATC codes
Information Network	Mar. 31/2016	N03AE, N05BA, N05CD), Z-Drugs (N05CF),
(DPIN)		Antidepressants, Antipsychotics, Mood stabilisers,
		Lithium and Opioids
		-Drug, dosage strength, dosage type, metric
		quantity dispensed, day supply, date of
		dispensation
Manitoba Health	Apr. 1/1996 –	Birth date/age of patient; sex; location of
Insurance Registry	Mar. 31/2016	residence, marital status, date of Manitoba Health
		coverage, date of coverage end, reason for
		coverage end (i.e death, emigration etc.)
Medical Claims	Apr. 1/1996 –	Services - type of physician (e.g., psychiatrist);
(Physician Billings)	Mar. 31/2016	dates of services, specific diagnoses (ICD-9 or
		ICD-10 equivalent)
Hospital Separations	Apr. 1/1996 –	Diagnoses (ICD-9 or ICD-10 equivalent), length
Abstracts	Mar. 31/2016	of stay, admission dates, discharge dates,

Provider Registry/Physician Master File	Apr. 1/1996 – Mar. 31/2016	Physician Age, Sex, Specialty
Social Allowances	1	Receipt of income assistance
Management Information Network (SAMIN)	Mar. 31/2013	



 Table 2. Characteristics of BZRA Users by First Use Episode Duration

	36/bmjoper			
le 2. Characteristics of BZ	RA Users by First Use	Episode Duration		36/bmjopen-2020-046916 on 1
		Short-term	Long-term	
Number of Users		197,606 (100%)	9,327 (100%)	206,933 (100%)
Sex Distribution*	Male	74,487 (37.7%)	4,295 (46.1%)	Vove 206,933 (100%) er 20,78,782 (38.1%)
	Female	123,057 (62.3%)	5,029 (53.9%)	
Age Category	18-44	101,709 (51.5%)	2,776 (29.8%)	D 128,086 (61.9%)
	45-64	66,752 (33.8%)	3,320 (35.6%)	ਰੂੰ 70,072 (33.9%)
	65+	29,143 (14.7%)	3,231 (34.6%)	32,374 (15.6%)
SEFI-2 Score	<-1	24,955 (12.6%)	1,089 (11.7%)	26,044 (12.6%)
	-1 to 0	81,718 (41.4%)	3,835 (41.1%)	32,374 (15.6%) 26,044 (12.6%) 85,553 (41.3%) Ppril 68,241 (33.0%)
	0 to 1	64,967 (32.9%)	3,274 (35.1%)	9 68,241 (33.0%)
	>1	25,966 (13.1%)	1,129 (12.1%)	% 27,095 (13.1%)
esidence Distribution	Urban	125,950 (63.7%)	5,802 (62.2%)	1371,752 (63.7%)
	Rural	71,656 (36.3%)	3,525 (37.8%)	7.5,181 (36.3%)
High Residential I	Mobility	36,392 (18.4%)	2,385 (25.6%)	2

		BMJ Open		36/bmjopen-2020-046916
Receipt of Income A	ssistance	18,530 (9.4%)	1,222 (13.1%)	20 20 04 46 46 19,752 (9.5%)
Marriage Rec	ord	102,461 (51.9%)	4,618 (49.5%)	9 107,079 (51.8%)
	0 (no utilization)	3,001 (1.5%)	349 (3.7%)	3,350 (1.6%) 5,980 (2.9%) 35,166 (17.0)
Johns Hopkins Healthcare	1	5,798 (2.9%)	182 (2.0%)	5,980 (2.9%)
Resource Utilization	2	33,974 (17.2%)	1,192 (12.8%)	
Band**	3	127,824 (64.7%)	5,151 (55.2%)	§ 132,975 (64.3%)
	4	20,065 (10.2%)	1,486 (15.9%)	21,551 (10.4%)
	5 (high-utilization)	6,882 (3.5%)	964 (10.3%)	fo 7,846 (3.8%)
		Short-term	Long-term	from 7,846 (3.8%) http://bmjop
Number of Us	sers	197,606 (100%)	9,327 (100%)	206,933 (100%)
Charles Consorbidity in day	0	148,257 (75.0%)	5,783 (62.0%)	8 154,040 (74.4%)
Charlson Comorbidity index Score	1	36,261 (18.4%)	2,031 (21.8%)	38,292 (18.5%)
30016	2+	13,088 (6.6%)	1,513 (16.2%)	38,292 (18.5%) 3 14,601 (7.1%)
	0	111,216 (56.3%)	3,862 (41.4%)	⇒ 115,078 (55.6%) ∞ 115,078 (55.6%)
Non-BZRA Psychotropic Prescription Dispensations	1	17,661 (8.9%)	518 (5.6%)	₹ 18,179 (8.8%)
	2+	68,729 (34.8%)	4,947 (53.0%)	Pr 73,676 (35.6%)
Opioid Prescription Dispensations	0	132,027 (66.8%)	5,855 (62.8%)	137,882 (66.6%)

	1	30,530 (15.5%)	1,011 (10.8%)	26 26 27 28 29 20 20 20 20 20 20 20 20 20 20 20 20 20
	2+	35,049 (17.7%)	2,461 (26.4%)	37,510 (18.2%)
Sex of Prescriber Issuing	Male	143,619 (75.3%)	6,928 (76.5%)	150,547 (75.3%)
First Prescription***	Female	47,128 (24.7%)	2,126 (23.5%)	No. 1 49,254 (24.7%)
Age of Prescriber Issuing	50+ Years	95,629 (52.1%)	4,775 (53.9%)	100,404 (52.2%)
First Prescription†	<50 Years	87,833 (47.9%)	4.076 (46.1%)	91,909 (47.8%)
Type of Prescriber Issuing	General Practitioner	146,823 (91.6%)	7,013 (87.5%)	153,836 (91.4%)
First Prescription‡	Psychiatry	6,338 (4.1%)	624 (7.8%)	6,962 (4.1%)
	Other	7,183 (4.5%)	375 (4.7%)	7,558 (4.5%)
	2001-2006	90,008 (45.5%)	2,608 (28.0%)	92,616 (44.8%)
Period of First Prescription	2006-2011	65,750 (33.3%)	3,170 (34.0%)	68,920 (33.3%)
	2011-2016	41,848 (21.2%)	3,549 (38.1%)	³ 45,397 (21.9%)
*N=197,544 (short-term use	rs); N=9,324 (long-te	erm users); N=206,868 (total	users)	pril
**N=197,544 (short-term us	ers); N=9,324 (long-	term users); N=206,868 (tota	al users)	18,
***N=190,747 (short-term u	sers); N=9,054 (long	-term users); N=199,801 (to	tal users)	2024
†N=183,462 (short-term use	rs); N=8,851 (long-te	erm users); N=192,313 (total	users)	4 b)
‡N=160,344 (short-term use	rs); N=8,012 (long-te	erm users); N=168,356 (total	users)	by gu

0,344 (short-term users); N=8,012 (long-term users); N=168,356 (total users)

Table 3 – Statistical Associations between Predictor Variables and Long-term Use of BZRAs

			BMJ Open			36/bmj	
						36/bmjopen-2020	
				Use D	uration	0.046	
Independent Var	<u>iable</u>	≥180	Days	≥90	Days	916	Days
		Crude OR (95% CI)	Adjusted OR (95% CI)	Crude OR (95% CI)	Adjusted OR (95% CI)	Cru y le OR (95% CI)	Adjusted OR (95% CI)
Male		1.41 (1.35-1.47)	1.33 (1.27-1.39)	1.40 (1.35-1.45)	1.34 (1.29-1.40)	हूँ 30 (1.2 % -1.34)	1.27 (1.23-1.31)
_	18-44	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 .(ref)	1 (ref)
Age	45-64	1.82 (1.73-1.92)	2.24 (2.11-2.38)	1.77 (1.70-1.85)	2.00 (1.91-2.10)	₹.81 (1.7 3 -1.86)	1.89 (1.82-1.97)
_	65+	4.06 (3.86-4.28)	5.15 (4.81-5.52)	3.56 (3.41-3.72)	4.11 (3.88-4.36)	\$34 (3.2₹-3.47)	3.52 (3.36-3.70)
Rural Residenc	re	1.07 (1.02-1.11)	1.10 (1.04-1.15)	0.97 (0.93-1.00)	0.97 (0.94-1.02)	9.90 (0.8 7 -0.92)	0.92 (0.88-0.95)
High Residential M	obility	1.52 (1.45-1.60)	1.14 (1.08-1.21)	1.35 (1.29-1.40)	1.06 (1.01-1.11)	1 1 1 1 1 1 1 1 1 1	1.01 (0.97-1.06)
Income Assistan	ice	1.46 (1.37-1.55)	1.68 (1.55-1.81)	1.14 (1.08-1.21)	1.35 (1.26-1.45)	© .88 (0.8 ₹ -0.93)	1.12 (1.06-1.20)
_	<-1	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 <mark>3</mark> ref)	1 (ref)
Socio-Economic Factor	-1 to 0	1.08 (1.00-1.15)	0.99 (0.92-1.07)	0.96 (0.91-1.02)	0.91 (0.86-0.97)	9 .90 (0. 智 -0.95)	0.89 (0.85-0.94)
Index-2 (SEFI-2) - Score	0 to 1	1.16 (1.07-1.24)	1.02 (0.94-1.10)	0.98 (0.93-1.04)	0.92 (0.87-0.98)	(0.83-0.91)	0.89 (0.84-0.94)
_	>1	1 (0.92-1.09)	0.93 (0.84-1.03)	0.78 (0.73-0.84)	0.80 (0.74-0.87)	0263 (0.59-0.67)	0.73 (0.68-0.78)
Married		0.91 (0.87-0.95)	0.79 (0.76-0.83)	1.01 (0.98-1.05)	0.89 (0.85-0.92)	€.13 (1.1⊉-1.16)	0.95 (0.92-0.99)
Opioid Use		1.19 (1.14-1.27)	1.16 (1.11-1.22)	1.08 (1.04-1.12)	1.09 (1.05-1.14)	(0.9%-1.02)	1.05 (1.01-1.09)
Independent Var	iabl <u>e</u>	,	, , ,	Use D	uration	<u> </u>	, ,
					29	by copyright	

			BMJ Open			36/bmjopen-2020	
		≥180	Days	≥90	Days		Days
		Crude OR (95% CI)	Adjusted OR (95% CI)	Crude OR (95% CI)	Adjusted OR (95% CI)	Crugle OR (95% CI)	Adjusted OR (95% CI)
Psychotropic Rx Use (non-BZRA)	1.82 (1.75-1.90)	1.93 (1.83-2.02)	1.62 (1.56-1.67)	1.75 (1.69-1.83)	1.38 (1.38-1.38)	1.49 (1.44-1.54)
	0	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 gref)	1 (ref)
Charlson Comorbidity Index Score	1	1.44 (1.36-1.51)	1.11 (1.04-1.17)	1.33 (1.27-1.39)	1.08 (1.02-1.13)	§24 (1.1 ÿ -1.29)	1.04 (1.00-1.08)
Thatex Score	2+	2.96 (2.79-3.15)	1.43 (1.32-1.55)	2.41 (2.29-2.54)	1.33 (1.24-1.42)	汉 01 (1.9 美 -2.11)	1.23 (1.15-1.31)
	0-3	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 g ref)	1 (ref)
Resource Utilization Band	4	1.84 (1.73-1.95)	1.15 (1.07-1.23)	1.58 (1.50-1.66)	1.08 (1.01-1.14)	(1.3 3 -1.43)	1.00 (0.94-1.05)
Бини	5	3.48 (3.24-3.73)	1.46 (1.33-1.60)	2.73 (2.56-2.92)	1.31 (1.20-1.42)	2.21 (2.08-2.35)	1.17 (1.09-1.27)
Male Prescriber of First	Prescription	1.07 (1.02-1.12)	1.03 (0.98-1.09)	1.07 (1.02-1.11)	1.04 (0.99-1.09)	(0.9 3 -1.05)	0.98 (0.94-1.02)
Prescriber Age ≥5	0 Years	1.08 (1.03-1.12)	0.98 (0.94-1.03)	1.08 (1.04-1.12)	0.99 (0.95-1.03)	1.15 (1.13-1.18)	1.08 (1.04-1.11)
	GP	1 (ref)	1 (ref)	1 (ref)	1 (ref)	<u>o</u> 1	1 (ref)
Type of Prescriber of	Psychiatrist	2.06 (1.89-2.25)	2.11 (1.93-2.32)	1.85 (1.72-2.00)	1.89 (1.75-2.05)		1.63 (1.51-1.75)
First Prescription	Other	1.09 (0.98-1.21)	0.92 (0.82-1.03)	1.07 (0.98-1.17)	0.92 (0.84-1.01)	16 (1.02-1.24)	1.03 (0.96-1.11)
	2001-2006	1 (ref)	1 (ref)	1 (ref)	1 (ref)	\$1	1 (ref)
Period of First Prescription	2006-2011	1.66 (1.58-1.75)	1.74 (1.64-1.85)	1.58 (1.51-1.65)	1.65 (1.57-1.7)	\$.41 (1.3\frac{6}{2}-1.46)	1.48 (1.42-1.54)
r rescription	2011-2015	2.93 (2.78-3.08)	2.99 (2.80-3.18)	2.59 (2.48-2.71)	2.71 (2.57-2.8)	夏97 (1.9 9 -2.05)	2.07 (1.98-2.16)

Table 4 – Frequency of Charlson Comorbidity Group Diagnoses by First Use Episode **Cohort**

Charlson Diagnosis	Short-Term 'First-Episode'	Long-Term 'First- Episode' Users	Z-Test of Two	916 on 1 November 2021. Downloaded from http://bmjopen.bmj.com/ on April 18, 2024
	Users	(n=9,327)	Proportions	verr
	(n=197,606)			nbei
Myocardial Infarction	2,474 (1.3%)	281 (3.0%)	p < 0.01	r 20
Congestive Heart Failure	3,943 (2.0%)	628 (6.7%)	p < 0.01	21.
Peripheral Vascular Disease	2,367 (1.2%)	256 (2.7%)	p < 0.01	Downli
Cerebrovascular Disease	3,690 (1.9%)	544 (5.8%)	p < 0.01	oad
Dementia	2,928 (1.5%)	796 (8.5%)	p < 0.01	ed f
COPD	23,064 (11.7%)	1,163 (12.5%)	p = 0.02	rom
Connective Tissue/Rheumatic Disease	2,793 (1.4%)	222 (2.4%)	p < 0.01	http://
Peptic Ulcer Disease	2,140 (1.1%)	114 (1.2%)	p = 0.20	omj
Mild Liver Disease	2,406 (1.2%)	135 (1.4%)	p = 0.05	ope
Moderate/Severe Liver Disease	341 (0.1%)	28 (0.0%)	p < 0.01	n.bmj.c
Uncomplicated Diabetes	14,131 (7.2%)	1,099 (11.8%)	p < 0.01	Ö
Complicated Diabetes	1,611 (0.8%)	252 (2.7%)	p < 0.01	/on
Paraplegia and Hemiplegia	794 (0.4%)	136 (1.5%)	p < 0.01	Ар
Renal Disease	1,858 (0.9%)	238 (2.6%)	p < 0.01	<u>1</u>
Cancer	829 (0.4%)	64 (0.1%)	p < 0.01	8, 2
Metastatic Carcinoma	64 (0.0%)	13 (0.0%)	p < 0.01	024
HIV/AIDS	50 (0.0%)	10 (0.0%)	p < 0.01	by guest. Protected by copyright

Figure Legend/Caption

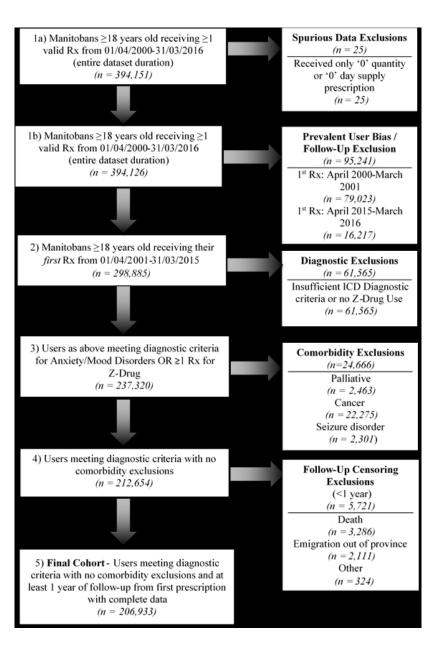
Figure 1. Flowchart of study population

Figure 2. Definition of long-term use (\geq 180 days)

Figure 3a. Duration of use determination

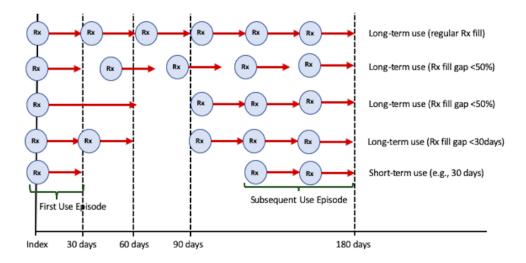
Figure 3b. Legend





Flowchart of study population

48x70mm (300 x 300 DPI)



Definition of long-term use (≥180 days)

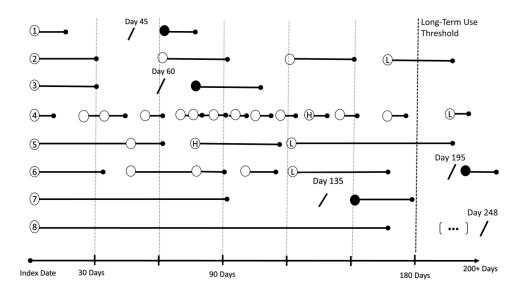


Figure 3a. Duration of use determination $338 \times 190 \text{mm} (300 \times 300 \text{ DPI})$

Legend

= Initial Prescription at Start of Use Episode (# denotes distinct individuals)

= Prescription During Use Episode

(H) = Hospitalization (BZD use assumed to continue)

= Last prescription at end of use episode (end date is end of prescription day supply)

= Lapse date (episode expiration)*

= Prescription after lapse date (not part of use episode)

*Lapse date determined as the greater of either 30 days or 50% of the previous prescription day supply

Figure 3b. Legend

338x190mm (300 x 300 DPI)

Supplemental Appendix Tables

Table A1 – International Classification for Disease Coding for Mood/Anxiety/Sleep Disorders (Cohort Inclusion)

	Source 1 - CPHA	Source 2 - MCHP	Study Algorithm
ICD Codes	All Mental Health Disorders: 9-CM: 290-319 10-CA: F00-F99	Mood Disorders: 296 and 311 (ICD-9-CM) or F30-F34, F39 (ICD 10-CA) Anxiety Disorders: 300 (ICD-9-CM) or F40-F42	Mood disorders: 296 and 311 (ICD-9-CM) or F30-F34, F39 (ICD 10-CA) Anxiety disorders: 300 (ICD-9-CM) or F40-F43 (ICD-10- CA) Sleep disorders: 307, 780 or F51, G47 ICD-10-CA)
Case Definition	≥1 hospitalization or	≥1 hospitalization or	≥1 hospitalization or
	outpatient medical	≥1-3 outpatient	≥3 outpatient
	claim within 1 year	medical claims within	medical claims within
		3-5 years*	5 years**

^{*}Range of similar definitions between studies from 2000 to 2016

^{**}The decision to use a 5-year pre-exposure window was based on the fact that all patients received a BZRA, which itself increases specificity for anxiety/sleep disorder diagnoses.

Table A2 – International Classification for Disease Coding Algorithms for Seizure, Cancer and Palliation (Cohort Exclusion)

	Seizure	Cancer and other Neoplasms	Palliation
ICD Codes	9-CM: 345 10-CA: G40	9-CM: 140-165, 170- 176,179-195, 200-208 10-CA: C00-C99	N/A*
Case Definition	≥1 hospitalization or ≥3 outpatient medical claim within 5 years before index date	≥1 hospitalization or ≥3 outpatient medical claims within 5 years before index date	Carrier code indicating palliative drug program enrollment in DPIN

^{*}While ICD codes do exist for palliation, the DPIN carrier code '04' is expected to be a reliable indicator of when patients become ill enough that community use of medication is required for symptom management.

Table A3 – Independent 'Patient' Variables for Prediction of Long-Term BZRA Use

Baseline Patient Characteristics		
Age	3 age groups; 18-44, 45-64, 65+ (Ordinal)	Index Date
Sex	Male or Female (Dichotomous Categorical)	Index Date
Region	Urban; Winnipeg or Brandon postal-codes Rural; Any other Manitoba postal-code (Dichotomous Categorical)	Census Period closest in time to the index date
Socioeconomic Status	Socio-Economic Factor Index – Version 2 (SEFI-2) score composite of four variables based on geography; i) unemployment rate ii) average household income iii) proportion of single-parent households iv) proportion of population without high school education. Scores <0 indicate more favourable socioeconomic conditions Scores >0 indicate less ideal socioeconomic conditions (Ordinal Scale)	Census Period closest in time to the index date
Income Assistance	Record of income assistance (Dichotomous Categorical)	Up to 1-year before the Index Date
Marriage Record	Record of Marriage (Dichotomous Categorical)	Entire available registry period up to the Index Date
Residential Mobility (i.e frequent mover)	Average of 1 move every 3 years from beginning of registry coverage to index date (Dichotomous)	Entire available registry period up to the Index Date
Comorbidity Burden	Charlson Comorbidity Index (CCI) Score; 0, 1, 2+ (Ordinal Scale)	Up to 1-year before the Index Date
Healthcare Resource Use	Johns Hopkins Adjusted Clinical Groups Resource Utilization Band (Ordinal Scale); placement into a band (0 to 5) based on grouping of	Up to 1-year before the Index Date

	ICD	
Prescription Psychotropic Use (non-BZRA)	Receipt of Prescription (Dichotomous Categorical)	Up to 1-year before the Index Date and 6 months after the Index Date
Prescription Opioid Use	Receipt of Prescription (Dichotomous Categorical)	Up to 1-year before the Index Date and 6 months after the Index Date

Table A4 - Independent 'First-Prescription' Variables for Prediction of Long-Term BZRA
Use

Characteristics of First Consultation and Subsequent Prescription	Definition	Measurement Period
Fiscal Year Period	Fiscal year of first prescription Assigned to 3 five-year intervals; 2001-2005, 2006- 2010, 2011-2015 (Ordinal)	Index Date
Prescriber	10 Years or More (Dichotomous)	Index Date
Sex of Prescriber	Male or Female (Dichotomous)	Index Date
Prescriber Specialty	General Practitioner, Psychiatry or Other (Categorical)	Index Date

Table A5 - Logistic Regression Methodology

Criteria Variable Selection	Approach -Informal selection via published literature
Variable Selection	-Informal selection via published literature
variable Selection	
	-Simple logistic regression; β values (p < 0.25)
	-Dichotomous Categorical; 0 or 1
	-Ordinal; discrete number scale starting at 1
Variable Coding	-Polychotomous Categorical; 0 or 1 with auto-
	generated dummy variables
	generated dammy variables
	-No continuous variables retained
Events-per-Variable	-Minimum 10 events per independent variable rule
Conformity of Linear Gradient	-Ordered categorical variables assessed for conformity of linear gradient; nonconformity handled by variable transformation or separation into additional (design) variables (i.e fiscal year was shown to be linear with respect to outcome so condensed variable into 5-year increments)
Interaction effects	-Assessed at p < 0.01. Suspected interactions included; age*sex, residential mobility*SEFI*income assistance, psychotropic use*opioid use, RUB*CCI
Collinearity	-Analysis of variance inflation factor, correlation coefficients, eigenvalues -Significant collinearity; combine variables or

	removal of inferior explanatory variable
Statistical Significance	-Wald 95% CI for β and OR's
Goodness-of-Fit Measures	-C-statistic, Log-Likelihood Ratio, Hosmer-
Goodiless-of-Fit Measures	Lemeshow Statistic
Eitting Dragadyra	-Stepwise addition/subtraction of variables
Fitting Procedure	-Assessment of clinical significance



Table A6 – Goodness of Fit for Final Logistic Regression Models Predicting Fong-Term Use of **BZRA**

Model	Model Type	Independent Variables	Likelihood Ratio (higher is better)	C statistic	Hosmer- Lemeshow Chi-Square Statistic
1	Main-Effects	9 Variables; Age-Sex Category, Period of First Rx, Psychotropic Use, Opioid Use, Income Assistance, Marriage, RUB CCI Score, Residential Mobility	6932 (p < 0.001)	l. Downloaded from http://bmjopen.bmj.c 73	$ \begin{array}{c} 10.78 \\ (p = 0.215) \end{array} $
2	Main-Effects + Interaction Effects	10 Variables: All from Model 1 + Residential Mobility*Income Assistance	6945 (p < 0.001)	om/ on April 18, 2024 by gue	11.02 (p = 0.20)

 $\begin{tabular}{ll} Table A7-Proportion of Long-Term BZRA Use by Differing Parameters and Duration Thresholds \\ \end{tabular}$

Scenario*	Long-Term Use Parameter	Prescription Lapse Criteria	Patients (n)	Proportion of Cohort
A1**	First-Use Episode ≥ 180 days	Greater of either 30 days or 50% of previous Day Supply	9,327	4.51%
A2	First-Use Episode ≥ 90 days	Greater of either 30 days or 50% of previous Day Supply	13,745	6.64%
A3	First-Use Episode ≥ 60 days	Greater of either 30 days or 50% of previous Day Supply	19,948	9.64%
A4	First-Use Episode ≥ 180 days	Greater of either 60 Days or 50% of previous Day Supply	13,050	6.31%
A5	First-Use Episode ≥ 180 days	90 Days	16,831	8.13%
A6	First-Use Episode ≥ 270 days	90 Days	15,214	7.35%
A7	First-Use Episode ≥ 365 days	90 Days	14,219	6.87%
В1	Mean Episode Duration ≥ 180 days	Greater of either 30 days or 50% of previous Day Supply	38,853	18.78%
B2	Mean Episode Duration ≥ 90 days	Greater of either 30 days or 50% of previous Day Supply	58,442	28.24%
В3	Mean Episode Duration ≥ 60 days Mean Episode Or 50% of previous Day Supply		72,639	35.10%
B4	Mean Episode Greater of either 60 Days		44,593	21.55%
В5	Mean Episode Duration ≥ 180 days	90 Days	50,142	24.23%
В6	User Mean Episode Duration ≥ 270 days	90 Days	39,395	19.04%
В7	User Mean Episode Duration	90 Days	32,200	15.56%

≥ 365 days

*A=First Episode Scenario; B=Mean Episode Duration Scenario

Table A8 - Proportion of Long-Term Z-Drug Use by Differing Parameters and Duration Thresholds

Scenario	Long-Term Use Parameter	Prescription Lapse Criteria	Patients (n)	Proportion of Sub-Cohort
A1	First-Use Episode ≥ 180 days	Greater of either 30 days or 50% of previous Day Supply	8,206	7.41%
A2	First-Use Episode ≥ 90 days	Greater of either 30 days or 50% of previous Day Supply	12,155	11.0%
A3	First-Use Episode ≥ 60 days	Greater of either 30 days or 50% of previous Day Supply	17,126	15.5%
A4	First-Use Episode ≥ 180 days	Greater of either 60 Days or 50% of previous Day Supply	10,437	9.43%
A5	First-Use Episode ≥ 180 days	90 Days	12,719	11.49%
A6	First-Use Episode ≥ 270 days	90 Days	11,117	10.04%
A7	First-Use Episode ≥ 365 days	90 Days	10,045	9.07%
B1	User Mean Episode Duration ≥ 180 days	Greater of either 30 days or 50% of previous Day Supply	21,859	19.75%
В2	User Mean Episode Duration ≥ 90 days	Greater of either 30 days or 50% of previous Day Supply	32,020	28.92%
В3	User Mean Episode Duration ≥ 60 days	Greater of either 30 days or 50% of previous Day Supply	39,690	35.85%
B4	User Mean Episode Duration	Greater of either 60 Days or 50% of	24,098	21.77%

^{**}Primary Scenario Used for Logistic Regression

	≥ 180 days	previous Day Supply		
B5	User Mean Episode Duration ≥ 180 days	90 Days	26,477	23.92%
В6	User Mean Episode Duration ≥ 270 days	90 Days	21,040	19.01%
В7	User Mean Episode Duration ≥ 365 days	90 Days	17,358	15.68%

Table A9 – Patient Characteristics of Z-Drug Users by First Use Episode Duration

	0	Short-term	Long-term	Total
Number of U	Number of Users		102,459 (100%) 8,204 (100%)	
	Male	40,516 (39.5%)	3,473 (42.3%)	43,989 (39.8%)
Sex Distribution	Female	61,943 (60.5%)	4,731 (57.7%)	66,674 (60.2%)
	18-44	42,663 (41.6%)	1,795 (21.9%)	44,458 (40.2%)
Age Category	45-64	39,817 (38.9%)	3,184 (38.8%)	43,001 (38.9%)
	65+	20,011 (19.5%)	3,227 (39.3%)	23,238 (21.0%)
	<-1	13,678 (13.3%)	981 (12.0%)	14,659 (13.2%)
	-1 to 0	45,136 (44.1%)	3,674 (44.8%)	48,810 (44.1%)
SEFI-2 Score	0 to 1	33,719 (32.9%)	2,885 (35.2%)	36,604 (33.1%)
	>1	9,958 (9.7%)	666 (8.1%)	10,624 (9.6%)
Residence	Urban	63,207 (61.7%)	3,313 (40.4%)	66,520 (60.1%)
Distribution	Rural	39,284 (38.3%)	4,893 (59.6%)	44,177 (39.9%)
High Residential	High Residential Mobility		2,523 (30.8%)	24,931 (22.5%)
Receipt of Income Assistance		8,351 (8.2%)	758 (9.2%)	9,109 (8.2%)

Marriage Re	ecord	57,308 (55.9%)	4,595 (56.0%)	61,903 (55.9%)
	O (no utilization)	1,771 (1.7%)	234 (2.9%)	2,005 (1.8%)
	1	3,205 (3.1%)	175 (2.1%)	3,380 (3.1%)
Johns Hopkins Healthcare	2	17,523 (17.1%)	1,012 (12.3%)	18,535 (16.7%)
Resource Utilization	3	65,067 (63.5%)	4,699 (57.3%)	69,766 (63.0%)
Band	4	10,810 (10.6%)	1,259 (15.3%)	12,069 (10.9%)
	5 (high- utilization)	4,083 (4.0%)	825 (10.1%)	4,908 (4.4%)
	0	Short-term	Long-term	Total
Number of U	Jsers	102,459 (100%)	8,204 (100%)	110,663 (100%)
Charlson	0	72,490 (70.8%)	4,528 (55.2%)	77,018 (69.6%)
Comorbidity index	1	19,495 (19.0%)	1,905 (23.2%)	21,400 (19.3%)
Score	2+	10,506 (10.3%)	1,773 (21.6%)	12,279 (11.1%)
Non-BZRA	0	27,797 (27.1%)	1,784 (21.7%)	29,581 (26.7%)
Psychotropic Prescription	1	36,939 (36.1%)	2,156 (26.3%)	39,095 (35.3%)
Dispensations	2+	37,755 (36.8%)	4,266 (52.0%)	42,021 (38.0%)
	0	47,427 (46.3%)	3,298 (40.2%)	50,725 (45.8%)
Opioid Prescription Dispensations	1	34,505 (33.7%)	2,772 (33.8%)	37,277 (33.7%)
	2+	20,559 (20.1%)	2,136 (26.0%)	22,695 (20.5%)
Sex of Prescriber Issuing First Prescription	Male	71,485 (69.8%)	5,627 (68.6%)	77,112 (69.7%)
	Female	28,485 (27.8%)	2,273 (27.7%)	30,758 (27.8%)
Age of Prescriber Issuing First	50+ Years	47,871 (46.7%)	4,014 (48.9%)	51,885 (46.9%)

Prescription	<50 Years	49,257 (48.1%)	3,758 (45.8%)	53,015 (47.9%)
Type of Prescriber	General Practitioner	78,610 (76.7%)	6,366 (77.6%)	84,976 (76.8%)
Issuing First	Psychiatry	3,912 (3.8%)	475 (5.8%)	4,387 (4.0%)
Prescription	Other	3,881 (3.8%)	381 (4.6%)	4,262 (3.9%)
Period of First	2001-2006	34,360 (33.5%)	1,526 (18.6%)	35,886 (32.4%)
	2006-2011	37,752 (36.8%)	2,808 (34.2%)	40,560 (36.7%)
Prescription	2011-2016	30,379 (29.6%)	3,872 (47.2%)	34,251 (31.0%)

Table A10 – Frequency of Charlson Comorbidity Group Diagnoses by First Use Episode Duration for Z-Drug Cohort

CI I D:	CI . TD	T	
Charlson Diagnosis	Short-Term	Long-Term 'First-	
	'First-Episode'	Episode' Users	Z-Test of Two
	Users	(n=8,204)	Proportions
	(n=102,459)		
Myocardial Infarction	1,836 (1.8%)	306 (3.7%)	p < 0.01
Congestive Heart Failure	3,174 (3.1%)	700 (8.5%)	p < 0.01
Peripheral Vascular Disease	1,772 (1.7%)	284 (3.5%)	p < 0.01
Cerebrovascular Disease	2,321 (2.3%)	550 (6.7%)	p < 0.01
Dementia	1,925 (1.9%)	865 (10.5%)	p < 0.01
COPD	12,357 (12.1%)	1,171 (14.3%)	p < 0.01
Connective	1 007 (1 00/)	242 (2.00/)	< 0.01
Tissue/Rheumatic Disease	1,906 (1.9%)	243 (3.0%)	p < 0.01
Peptic Ulcer Disease	1,111 (1.1%)	123 (1.5%)	p < 0.01
Mild Liver Disease	1,672 (1.6%)	139 (1.7%)	p = 0.33
Moderate/Severe Liver	275 (0.2%)	38 (0.4%)	p < 0.01
Disease	273 (0.270)	36 (0.470)	p < 0.01
Uncomplicated Diabetes	9,317 (9.1%)	1,150 (14.0%)	p < 0.01
Complicated Diabetes	1,639 (1.6%)	328 (4.0%)	p < 0.01
Paraplegia and Hemiplegia	508 (0.5%)	136 (1.7%)	p < 0.01
Renal Disease	1,543 (1.5%)	293 (3.6%)	p < 0.01
Cancer	2,109 (2.1%)	247 (3.0%)	p < 0.01
Metastatic Carcinoma	429 (0.4%)	45 (0.5%)	p = 0.04
HIV/AIDS	118 (0.1%)	16 (0.2%)	p = 0.02



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Table A11 – Statistical Associations between Predictor Variables and Long-term Use of Z-Drugs

	- Statistical Ass	Use Duration						
		≥180) days	≥90	O days	on ≥ 60	° ≥ 60 days	
<u>Independent Variable</u>		Crude OR (95% CI)	Adjusted OR (95% CI)	Crude OR (95% CI)	Adjusted OR (95% CI)	Cru z e OR (95§ CI)	Adjusted OR (95% CI)	
Ма	ule	1.12 (1.07-1.18)	1.04 (0.99-1.09)	1.13 (1.08-1.17)	1.05 (1.01-1.10)	1∯8 (1.05€1.12)	1.04 (1.00-1.08)	
	18-44	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 (<u>ref)</u> 1≱1	1 (ref)	
Age	45-64	1.90 (1.79-2.02)	2.02 (1.89-2.17)	1.74 (1.66-1.82)	1.78 (1.68-1.88)	1ৠ1 (1.6ৄ1.78)	1.68 (1.60-1.76)	
-	65+	3.83 (3.61-4.07)	3.71 (3.44-4.00)	3.24 (3.08-3.40)	3.08 (2.90-3.28)	299 (2.87 5 3.12)	2.78 (2.64-2.93)	
Rural Re	Rural Residence		1.13 (1.07-1.19)	0.99 (0.96-1.03)	1.02 (0.98-1.07)	1308 $(1.041.11)$	0.95 (0.91-0.99)	
High Residen	tial Mobility	1.59 (1.51-1.67)	1.26 (1.19-1.33)	1.53 (1.46-1.59)	1.21 (1.15-1.27)	1 <u>3</u> 0 (1.261.35)	1.12 (1.07-1.17)	
Income A	ssistance	1.15 (1.06-1.24)	1.47 (1.34-1.61)	1.02 (0.95-1.09)	1.29 (1.19-1.40)	0.32 (0.77=0.87)	1.08 (1.00-1.17)	
	<-1	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 gef)	1 (ref)	
SEFI-2 Score	-1 to 0	1.14 (1.06-1.22)	1.07 (0.99-1.16)	1.03 (0.97-1.09)	0.98 (0.92-1.04)	0 <u></u> 3 5 (0.9 <u>1.00)</u>	0.94 (0.89-0.99)	
SEF1-2 Score	0 to 1	1.19 (1.11-1.29)	1.08 (0.99-1.17)	1.04 (0.98-1.11)	0.99 (0.93-1.06)	0 <u>9</u> 2 (0.87 <u>2</u> 0.97)	0.93 (0.88-0.99)	
	>1	0.93 (0.84-1.03)	0.84 (0.75-0.94)	0.80 (0.73-0.87)	0.77 (0.70-0.85)	058 (0.6350.73)	0.72 (0.66-0.78)	
Married		1.00 (0.96-1.05)	0.86 (0.82-0.91)	1.07 (1.03-1.10)	0.93 (0.89-0.98)	1	0.98 (0.94-1.01)	
Opioid	d Use	1.28 (1.22-1.34)	1.15 (1.09-1.21)	1.26 (1.21-1.31)	1.15 (1.11-1.20)	1 1 2 8 (1.14 2 1.21)	1.11 (1.07-1.15)	

			BMJ Open			36/bmjopen-2020		
						n-2020-(
		Use Duration $\overset{\circ}{4}$						
Independent Variable		≥180 days		≥90 days		⁹ / ₆ ≥60 days		
		Crude OR (95% CI)	Adjusted OR (95% CI)	Crude OR (95% CI)	Adjusted OR (95% CI)	Crude OR (95½ CI)	Adjusted OR (95% CI)	
Psychotropic Rx	Use (Non-BZRA)	1.34 (1.27-1.41)	1.24 (1.17-1.32)	1.35 (1.29-1.41)	1.27 (1.20-1.33)	1 2 1.19 (1.1 2 1.27) (1.14-1.24)		
	0	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 (gef)	1 (ref)	
Charlson Comorbidity Index Score	1	1.56 (1.48-1.65)	1.25 (1.18-1.33)	1.45 (1.39-1.52)	1.21 (1.15-1.27)	1 (gef) 1:33 (1.28 1.38)	1.13 (1.08-1.19)	
	2+	2.70 (2.55-2.87)	1.46 (1.36-1.58)	2.34 (2.22-2.46)	1.38 (1.29-1.47)	2∯2 (1.93€2.12)	1.30 (1.22-1.37)	
Resource Utilization Band	0-3 (≤Moderate)	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 (g ef)	1 (ref)	
	4 (High)	1.67 (1.56-1.78)	1.16 (1.08-1.25)	1.47 (1.39-1.56)	1.09 (1.01-1.16)	130 (1.241.37)	1.00 (0.95-1.07)	
	5 (Very High)	2.89 (2.67-3.13)	1.55 (1.41-1.70)	2.43 (2.26-2.61)	1.42 (1.30-1.55)	1 <u>9</u> 7 (1.8 5 2.11)	1.22 (1.12-1.32)	
Male Prescriber of	Male Prescriber of First Prescription		0.97 (0.92-1.03)	0.98 (0.94-1.02)	0.98 (0.93-1.02)	0.994 (0.99-0.97)	0.93 (0.90-0.97)	
Prescriber Aş	ge ≥50 Years	1.10 (1.05-1.15)	0.98 (0.93-1.03)	1.10 (1.06-1.15)	0.98 (0.94-1.02)	15.5 (1.11.19)	1.05 (1.01-1.09)	
	GP	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 @ ef)	1 (ref)	
Prescriber of First Prescription	Psychiatrist	1.50 (1.36-1.66)	1.96 (1.76-2.17)	1.36 (1.25-1.49)	1.72 (1.57-1.89)	1 = 1 (1.02 = 1.20)	1.38 (1.27-1.51)	
	Other	1.21 (1.09-1.35)	0.92 (0.82-1.03)	1.18 (1.07-1.29)	0.91 (0.83-1.00)	1¥9 (1.1%1.29)	0.98 (0.91-1.07)	
Period of First Prescription	2001-2006	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 (g ef)	1 (ref)	
	2006-2011	1.68 (1.57-1.79)	1.57 (1.46-1.68)	1.67 (1.59-1.76)	1.56 (1.47-1.66)	1. 5 3 (1.4 6 1.60)	1.46 (1.39-1.54)	
	2011-2015	2.87 (2.70-3.05)	2.45 (2.28-2.65)	2.83 (2.69-2.97)	2.44 (2.30-2.59)	220 (2.1022.29)	1.96 (1.86-2.07)	

STROBE Statement—checklist of items that should be included in reports of observational studies

	Item No	Recommendation	Page No	
Title and abstract	1	(a) Indicate the study's design with a commonly used term in the title or		
		the abstract		
		(b) Provide in the abstract an informative and balanced summary of what	3	
		was done and what was found		
Introduction				
Background/rationale	2	Explain the scientific background and rationale for the investigation being reported		
Objectives	3	State specific objectives, including any prespecified hypotheses	6	
Methods			I	
Study design	4	Present key elements of study design early in the paper		
Setting	5	Describe the setting, locations, and relevant dates, including periods of	7-9	
Setting		recruitment, exposure, follow-up, and data collection		
Participants	6	(a) Cohort study—Give the eligibility criteria, and the sources and	7-9	
i articipants	O	methods of selection of participants. Describe methods of follow-up	/-/	
		Case-control study—Give the eligibility criteria, and the sources and		
		methods of case ascertainment and control selection. Give the rationale		
		for the choice of cases and controls		
		Cross-sectional study—Give the eligibility criteria, and the sources and		
		methods of selection of participants	7.0	
		(b) Cohort study—For matched studies, give matching criteria and	7-9	
		number of exposed and unexposed		
		Case-control study—For matched studies, give matching criteria and the		
		number of controls per case		
Variables	7	Clearly define all outcomes, exposures, predictors, potential confounders,	8-10	
Data sources/	8*	For each variable of interest, give sources of data and details of methods		
measurement		of assessment (measurement). Describe comparability of assessment		
		methods if there is more than one group		
Bias	9	roll effect modifiers. Give diagnostic criteria, if applicable For each variable of interest, give sources of data and details of methods of assessment (measurement). Describe comparability of assessment methods if there is more than one group Describe any efforts to address potential sources of bias Explain how the study size was arrived at Explain how quantitative variables were handled in the analyses. If		
Study size	10	Explain how the study size was arrived at		
Quantitative variables	11	Explain how quantitative variables were handled in the analyses. If	7-10	
		applicable, describe which groupings were chosen and why		
Statistical methods	12	(a) Describe all statistical methods, including those used to control for	10	
		confounding		
		(b) Describe any methods used to examine subgroups and interactions	10	
		(c) Explain how missing data were addressed	10-	
		•	11	
		(d) Cohort study—If applicable, explain how loss to follow-up was	10-	
		addressed	11	
		Case-control study—If applicable, explain how matching of cases and		
		Case-control study—If applicable, explain how matching of cases and controls was addressed		
		controls was addressed		

Continued on next page

Results			
Participants	13*	(a) Report numbers of individuals at each stage of study—eg numbers potentially eligible, examined for eligibility, confirmed eligible, included in the study, completing follow-up, and analysed	11-12
		(b) Give reasons for non-participation at each stage	7
		(c) Consider use of a flow diagram	7, fig1
Descriptive data	14*	(a) Give characteristics of study participants (eg demographic, clinical, social) and information on exposures and potential confounders	11,12
		(b) Indicate number of participants with missing data for each variable of interest	Fig 1
		(c) <i>Cohort study</i> —Summarise follow-up time (eg, average and total amount)	11
Outcome data	15*	Cohort study—Report numbers of outcome events or summary measures over time	12- 13
		Case-control study—Report numbers in each exposure category, or summary measures of exposure	n/a
		Cross-sectional study—Report numbers of outcome events or summary measures	n/a
Main results	16	(a) Give unadjusted estimates and, if applicable, confounder-adjusted estimates and	Table
		their precision (eg, 95% confidence interval). Make clear which confounders were adjusted for and why they were included	2
		(b) Report category boundaries when continuous variables were categorized	Table 2
		(c) If relevant, consider translating estimates of relative risk into absolute risk for a meaningful time period	Table 2
Other analyses	17	Report other analyses done—eg analyses of subgroups and interactions, and sensitivity analyses	Table 2-3
Discussion			
Key results	18	Summarise key results with reference to study objectives	13
Limitations	19	Discuss limitations of the study, taking into account sources of potential bias or imprecision. Discuss both direction and magnitude of any potential bias	17
Interpretation	20	Give a cautious overall interpretation of results considering objectives, limitations,	13-
•		multiplicity of analyses, results from similar studies, and other relevant evidence	17
Generalisability	21	Discuss the generalisability (external validity) of the study results	17-
			18
Other informati	on		
Funding	22	Give the source of funding and the role of the funders for the present study and, if applicable, for the original study on which the present article is based	25

^{*}Give information separately for cases and controls in case-control studies and, if applicable, for exposed and unexposed groups in cohort and cross-sectional studies.

Note: An Explanation and Elaboration article discusses each checklist item and gives methodological background and published examples of transparent reporting. The STROBE checklist is best used in conjunction with this article (freely available on the Web sites of PLoS Medicine at http://www.plosmedicine.org/, Annals of Internal Medicine at http://www.annals.org/, and Epidemiology at http://www.epidem.com/). Information on the STROBE Initiative is available at www.strobe-statement.org.

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Risk of Long-Term Benzodiazepine and Z-Drug Use Following the First Prescription among Community-Dwelling Adults with Anxiety/Mood and Sleep Disorders: A retrospective cohort study

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Risk of Long-Term Benzodiazepine and Z-Drug Use Following the First Prescription among Community-Dwelling Adults with Anxiety/Mood and Sleep Disorders: A retrospective cohort study

Running Title: Risk of long-term benzodiazepine receptor agonist use

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ABSTRACT

Objective: To measure the incidence of long-term benzodiazepine receptor agonist (BZRA) use among individuals with anxiety, mood and/or sleep disorders. To identify factors associated with long-term use following the first prescription.

Methods: This was a population-based retrospective cohort study using administrative databases in Manitoba, Canada. Individuals with anxiety/mood or sleep disorder who received their first BZRA between April 1, 2001 to March 31, 2015 were included. Long-term use was defined as ≥180 days. Logistic regression modelling was used to examine predictors of long-term use.

Results: Among 206,933 individuals included, long-term BZRA use in the first episode of use was 4.5% (≥180 days) following their first prescription. Factors associated with ≥180 days of use included male sex (adjusted odds ratio (aOR) 1.33, 95% CI 1.27 to 1.39), age ≥65 (aOR 5.15, 95% CI 4.81 to 5.52), income assistance (aOR 1.68, 95% CI 1.55 to 1.81), previous non-BZRA psychotropic (aOR 1.93 (95% CI 1.83 to 2.02) or opioid use (aOR 1.16, 95% CI 1.11 to 1.22), high comorbidity (aOR 1.43, 95% CI 1.32 to 1.55), high healthcare use (aOR 1.46 (95% CI 1.33 to 1.60), and psychiatrist prescriber (aOR 2.11, 95% CI 1.93 to 2.32).

Conclusions: Less than one in twenty patients use BZRAs \geq 180 days in their first treatment episode. Several factors were associated with long-term use following the first prescription and further investigation into whether these factors need to be considered at the point of prescribing is warranted. In light of these findings, future research should examine the predictors of cumulative repeat episodes of BZRA exposure.

Key Words: benzodiazepine, anxiety, insomnia, pharmacoepidemiology, clinical practice guidelines, z-drug hypnotics

Strengths and Limitations of Study

- This study used administrative data from the Manitoba Centre for Health Policy, which is one of the most comprehensive datasets in North America containing >140 de-identified linked datasets on healthcare, education, social/families, justice and registries for all residents of Manitoba (population of 1.4 million people) not restricted by age or income
- All diagnoses are identified through physician claims data or hospitalizations, which are
 dependent on people seeking treatment and may be prone to some misclassification. Drug
 information is also based on dispensing records from community pharmacies and does not
 confirm the patient actually took the drug. However, we performed multiple sensitivity
 analyses to address this.
- The databases do not capture participation in psychological interventions such as cognitive behavioral therapy.

Introduction

The use of benzodiazepine receptor agonists (BZRAs)*, benzodiazepines (BZD) and Z-Drugs, in the treatment of anxiety and insomnia has shifted based on the evolving data on safety risks and limited efficacy on long-term use in the literature. ^{1–4} Upon their initial introduction into clinical practice in the late 1960s, benzodiazepines were considered to be a safer alternative to barbiturates.⁵ However, safety concerns such as psychomotor impaired accidents (i.e., falls and motor-vehicle accidents), dependency and misuse/abuse are now well known. ^{6–8} Recent studies have also raised concerns proposing possible links to dementia, recurrence of mood episode, respiratory disease exacerbation and suicide with long-term BZRA use. ^{9–13} However, the association of BZRA use for these newer harms is uncertain given conflicting evidence and confounding in previous studies.¹⁴

In spite of ongoing adverse effect concerns, justification for less restrictive BZRA use have stemmed from their clinical utility as rapidly effective anxiolytic sedatives.¹⁵ Some view that limiting BZRA use is at times impractical.¹⁶ Moreover, the use of alternative pharmacotherapy, including trazodone, atypical antipsychotics, barbiturates, and tricyclic antidepressants are not without adverse effects. It should also be noted that the difficulties with de-prescribing BZRAs reported in the literature have added caution to the initiation of these agents in practice.^{4,17}

Previous studies examining the pattern of BZRA use have found a decline in benzodiazepine (particularly lorazepam) incident use and an increase in the incidence of Z-drug use. Limited studies have examined predictors of long-term use after a first prescription. As such, this study sought i) to measure the incidence of long-term BZRA use among a cohort of

^{*} Abbreviations: BZRA, benzodiazepine receptor agonist; BZD, benzodiazepine

community-dwelling Canadian adults with anxiety, mood and/or sleep disorders, and ii) to determine factors associated with progression to long-term BZD use following the first prescription in this population.

Methods

Study Design and Data Sources

This was a retrospective, cohort study using routinely collected administrative healthcare data pertaining to prescription drug dispensations, outpatient physician claims, hospitalization discharge abstracts, income assistance records and prescriber demographics (Table A1). All data used was extracted from the Manitoba Centre for Health Policy Population Research Data Repository. The Repository provides comprehensive coverage of all Manitoba residents contact with the primary healthcare system. The Drug Program Information Network (DPIN) provides information on outpatient prescription drugs dispensed in Manitoba with the exception of medications dispensed in hospital and nursing stations. In Manitoba, eligible outpatient prescriptions are 100% covered for residents after an income-based deductible is paid for each fiscal year. DPIN captures information on the drug name, strength, quantity, day-supply, and date of all outpatient prescriptions dispensed regardless of coverage. Merging of the various data sources was facilitated via linkage of unique de-identified Personal Health Information Numbers. The Charlson Comorbidity score $[0 \text{ (lowest risk)}, 1, \ge 2 \text{ (high risk)}]$ was also determined to examine the effects of comorbidity of duration of use. This was determined based on 17 categories of comorbidities using ICD-9-CM or ICD-10-CA equivalent codes in administrative data to provide the weight-based adjusted risk of death or resource use.²²

Cohort Inclusion/Exclusion Criteria

Eligible patients were adults age 18 years and older who initiated a new benzodiazepine or Z-Drug prescription (defined as no use in the one year prior to the first prescription^{20,21}) between April 1, 2001 to March 31, 2015, with no preceding dispensations from April 1, 2000 to March 31, 2001 (first year of the dataset) to avoid prevalent user bias (Figure 1). All individuals with at least one year of registry coverage prior to and after the first prescription was required for cohort inclusion. As such, individuals who received a benzodiazepine in the distant past could be included in the cohort as a new user, provided that the benzodiazepine was not used in the past one year. A sensitivity analysis was also performed in which incident use was defined as no prescription for a BZRA was received in the three years prior to the first prescription.²³

Eligibility was also based on diagnostic criteria for anxiety/mood related disorders and/or insomnia based on ICD-9-CM or ICD-10-CA medical claims, either at outpatient physician visits or hospitalizations, occurring within a 5-year period prior to the first prescription. The ICD diagnostic criteria chosen are a combination of the definitions from two sources; the Canadian Public Health Association on mental health surveillance and the MCHP concept dictionary, which listed the various past-case definitions employed in previous research within Manitoba for mood and anxiety disorders (Table A2).²⁴⁻²⁸ Lastly, because reliance on ICD codes is expected (and has been previously shown) to underestimate capture of sleep disorder cases, we also accepted receipt of a Z-Drug in the definition for insomnia as this was their sole approved indication.²⁹

To reduce confounding, we established cohort exclusion criteria that otherwise may have justified long-term use of BZDs in clinical scenarios beyond the scope of general guideline recommendations for anxiety and insomnia. Namely, patients were excluded if they had ≥1 ICD

code for cancer, a seizure disorder or if there was placement in the Manitoba palliative care drug program at any point in the 5 years preceding their first prescription for a BZRA (Table A3). Where patients became palliative ≥1 year after the initial BZRA dispensation, their ongoing use of BZRA was censored beginning from the date of their placement, but all use prior to their palliative status was retained. Clobazam use was excluded entirely from the evaluated drug claims because it is approved only as an adjunctive agent for epilepsy in Canada. Finally, patients were excluded if they lacked at least 1-year of registry coverage from their first-prescription index date. This was to eliminate any biasing effect from early mortality, moving out of province or other loss to follow-up.

Main Outcome Measures

Long-term use was defined as ≥180 days based on the recommendation from a previous systematic review of similar studies.²⁴ This duration is longer than clinical practice guideline duration recommendations and is believed to be of sufficient length for risk of dependence to occur.³⁰ One-third of individuals who use BZDs for longer than six months have been previously reported to be unable to stop completely due to withdrawal symptoms (e.g., anxiety, insomnia, muscle spasms).³⁰ A sensitivity analysis, ranging from 60 to 365 days, was also used in our study to account for varying definitions of long-term use reported in the literature.²⁴

Patients were followed forward in time from the date of their first BZRA prescription. BZRA 'use episodes' were determined according to consecutive prescription overlap based on dispensation dates and coded day supply values. The allowable gap between prescriptions was the greater of either 30 days or 50% of the last prescription day supply after the prescription end date (end date = dispensation date + day-supply) of the prior prescription. This gap was chosen to account for those who regularly or frequently used "as needed" BZRA in the 'use episode'

duration. The episode end date was calculated as the date of the last prescription in a given 'use episode' plus its associated day-supply. To account for immeasurable time bias, hospitalization time was assumed to be a continuation of BZD use given that in-patient drug use data was limited.³¹ The provincial drug program subsidizes dispensations of up to a 100 day-supply.

Individuals were able to have multiple use episodes over the entire study duration. First episode duration and average episode duration were calculated for each user. If patients only had one use episode both of these values were the same. Patients were allowed to switch from one BZRA to another without it interrupting their 'use episodes'. This included switching from a BZD to a Z-drug and vice versa.

Independent Variables

Variables used for statistical prediction of long-term use were determined a priori and included age, sex, geographic residence, residential mobility, socioeconomic status, marriage, concurrent opioid or prescription psychotropic use, comorbidity burden, healthcare usage, time period of first prescription and prescriber characteristics (Table A4 and Table A5). Variables were assessed at baseline; either within 1-year before the index date, at the index date or up to 6-months past the index date (in the case of prescription opioids and other psychotropics, such as antidepressants, antipsychotics, and mood stabilizers).

Statistical Analysis

Standard reporting criteria were followed in the approach to logistic regression modelling (Table A6 and A7).³² Univariate analysis was performed first in the form of simple logistic regression. The multi-variable model was constructed to determine the most parsimonious model for prediction of long-term BZRA use defined as \geq 180 days in the first episode of use with

adjustment of clinically relevant covariates based on previous literature.²⁴ Differences between models in their maximum log-likelihood estimation, likelihood ratios and other goodness-of-fit statistics enabled model discrimination.³² Multicollinearity and effect-measure modification (i.e., interaction effects) were assessed when it was suspected that variables may be either correlated or non-independent.³² In order to perform these diagnostics, the binary dependent variable was first substituted for a linear variable (first-episode duration in days) to conduct a multiple *linear* regression. Specifically, collinearity was determined to be a model threat if any correlation coefficient in the independent variable correlation matrix was $\geq |0.8|$ or if any variance inflation factor was unreasonably high (≥ 10) while the corresponding tolerance factor was miniscule (≤ 0.1).³³ Analyses were assessed at p ≤ 0.01 threshold set a priori for statistical significance.

For the multiple logistic regression, 'complete case-analysis' was used because the extent of missing data was too small to justify the need for multiple imputation procedures.³⁴ In this study, no claims were excluded on the basis of missing data fields. Only 1,568 claims (<0.01%) were excluded for being spurious (i.e '0' day/quantity supply or incredibly high dispensed quantity to day-supply ratio) Furthermore, observed missing data was believed to be missing at random.³⁵ The only variable with significant missing data was that of 'prescriber type' (~38,000 missing observations or 17.5% of final sample).

A subgroup analysis of each of the 17 categories of the Charlson Comorbidity Score was also performed using Z-test of two proportions to describe the specific comorbidities that may contribute to the relationship between Charlson Comorbidity Score and long-term use.

Sensitivity Analysis

To assess the robustness of the primary outcome, 6 sensitivity analyses (Tables A8 and Table A9) were conducted to determine how the proportion of long-term use changed under

differing parameter assumptions.³⁶ The threshold duration for long-term use was adjusted to values ranging from 60 days to 365 days. Additionally, the episode lapse criteria (i.e., prescription gap rule) was changed. While the analysis was not exhaustive for every conceivable combination of these key parameters, the selected values were chosen because they were judged to be representative of how peers in the international clinical community may have defined or measured 'long-term use' of BZRA. All data was cleaned and analyzed using SAS v9.4©.

Ethical Approval

Access to the data for this project was approved by the University's Health Research Ethics Board (HREB, registration number H2017:052 (HS20498) and the Health Information Privacy Committee (HIPC, no. 2016/2017-62) of the provincial government. Consent for this study was not required by HREB given the retrospective nature of the study and data agreements in place through HIPC.

Results

Episodic BZD/Z-Drug Use

Study population demographics are presented in Table 1. There were 206,933 patients in our cohort representing 931,271 unique BZRA dispensations over the 15-year study duration. Over the study period, cohort individuals had a median of three and average of 4.5 BZRA use episodes, respectively. First-episodes of use were of a median duration of 20 days (IQR = 10-30 days). For all use-episodes, the median average use duration was 30 days (IQR = 15-111 days). Evaluation of long-term use revealed that 4.51% of patients used a BZRA for ≥180-days in their 'first' episode

of use. At most, this proportion increased to 9.64% when a sensitivity analysis of 60 days or greater was used for the definition of 'long-term use' for the first episode of use. However, the proportion of long-term users increased considerably after averaging for all episodes for each user (sensitivity analysis range: 15.6%-35.1%) (Table A7).

To evaluate treatment duration for insomnia, a sensitivity analysis was performed on only Z-Drugs (n=110,663), which found similar results (Tables A9-A12).

Factors Predicting Long-term First Episode Use

Logistic regression analysis revealed that male sex (adjusted odds ratio (OR) 1.33, 95% CI 1.27 to 1.39), older age (adjusted OR 2.24 (95% CI 2.11 to 2.38) and 5.15 (95% CI 4.81 to 5.52) for aged 45-64 years and \geq 65 years, respectively, compared to <45 years), receipt of income assistance (adjusted OR 1.68, 95% CI 1.55 to 1.81), previous non-BZRA psychotropic (adjusted OR 1.93 (95% CI 1.83 to 2.02) or opioid use (adjusted OR 1.16, 95% CI 1.11 to 1.22), high comorbidity (Charlson Comorbidity Index 1 and ≥2, adjusted OR 1.11 (95% CI 1.04 to 1.17) and 1.43 (95% CI 1.32 to 1.55), respectively), high healthcare resource use (resource utilization band of 4 and 5, adjusted OR 1.15 (95% CI 1.07 to 1.23) and 1.46 (95% CI 1.33 to 1.60), respectively), first prescription from psychiatrist (adjusted OR 2.11, 95% CI 1.93 to 2.32) and receipt of first prescription after 2006 (2006-2011, adjusted OR 1.74, 95% CI 1.64 to 1.85; 2011-2015, adjusted OR 2.99, 95% CI 2.80 to 3.18), were all predictive of long-term use of \geq 180 days in the first episode. Rural residence (adjusted OR 1.10, 95% CI 1.04 to 1.15) and high residential mobility (adjusted OR 1.14, 95% CI 1.08 to 1.21) were also associated with a higher risk of long-term use in the first episode. Married status was associated with a lower risk of meeting the long-term use definition (adjusted OR 0.79, 95% CI 0.76 to 0.83). These findings were also replicated in the sensitivity analysis restricted to Z-Drug users. Both the crude and adjusted odds ratios are presented for the full cohort in Table 2.

A sub-analysis of the higher comorbidity scores in the long-term user groups shows that this relationship was mainly driven by cardiovascular diseases, diabetes and dementia (Table 3). Proportions for these particular diagnoses were 2 to 5 times higher in the long-term user group, with the greatest difference existing for dementia (long-term; 8.5% vs. short-term; 1.5%). A sensitivity analysis was performed changing the definition of incident user to no receipt of BZRA prescription in the three years prior to the first BZRA prescription. No change in results were found.

Discussion

This study found approximately 4.5% of the full cohort and 7.4% of the Z-Drug cohort were 'long-term' first-episode users according to the best available evidence-based consensus definition of 180 days. ²⁴ Restricting the analysis to Z-Drug use showed that the frequency of long-term use was higher than that of the main cohort. Practice guidelines typically recommend a shorter duration of use for Z-Drugs in the treatment of insomnia (range of ≤2-6 weeks)³⁷⁻³⁹ compared to BZD for anxiety disorder (up to ≤12 weeks depending on indication). ⁴⁰⁻⁴² Therefore, these results suggest greater disparity from practice guidelines in the case of Z-drug use for insomnia. Of note, more recent insomnia guidelines have recognized that while non-drug alternatives have a favourable safety profile, these interventions may be difficult to achieve for certain populations, which could explain the deviation between practice recommendations and real-world use of these agents. ³⁸

The proportion of patients who met criteria for 'long-term' use after accounting for all of their use-episodes (i.e., rather than just the first episode of use) was approximately 3.5 times higher than the proportion of patients meeting criteria after only their first episode of use. These results may indicate that repeated episodes of BZRA use may be associated with a higher risk of being exposed to a BZRA for a duration of ≥ 180 days in one episode. An area of future research is to examine whether repeated episodes of BZRA use is associated with progression to long-term use as demonstrated in a previous study that observed the number of episodes of dispensing in the first month was a significant predictor of the total duration of dispensing in the later period.⁴³ Of note, the majority of people with repeated use still only take BZRAs for intermittent, short-term periods. Furthermore, confounding variables such as age and accrued comorbidity over time may influence the risk of future long-term use in some patients. Nonetheless, these results support the observed difficulty in de-prescribing once BZRA use has become chronic, which has also been reported in previous literature. 4,44 Lastly, other clinical considerations such as risk of protracted withdrawal symptoms, risk of rebound insomnia and/or anxiety, severity of indication, patient dissatisfaction, limited alternate drug and non-drug interventions, or interference with another prescriber's decisions likely undermine potential de-prescribing efforts.

Older age and female sex have also been identified in previous studies as being associated with long-term use.^{45–51} While we found females to have greater representation in all patterns of BZRA use, we found males were more specifically predictive of long-term use after the first episode of use. ^{52–54} As with almost all of the previously published studies, older age was strongly associated with long-term BZRA use.⁵¹⁻⁵⁵ It should be noted that older individuals may have had a greater opportunity to be exposed to BZRA use.

As supported by previous evidence, income assistance was associated with long-term BZRA use ^{48,56}. Our study also found frequent moving, unmarried status, and rural residence to be associated with increased odds of long-term use. Frequency of moving and income assistance could be a proxy for general life stability^{50,57,58}. Rural residence may have a small effect on longer-term BZRA use due to the relative limitations of timely scheduled follow-up, which may necessitate prescriptions of greater quantity or for longer periods. Another study also found rural adults to be at higher odds of inappropriate BZD use .⁵⁹

Healthcare use and the presence of various physical illnesses have been consistent predictors of long-term BZRA use ^{47,49,50,60}. In this study, as both of these variables increased, so did the odds of long-term use. We speculate that the positive relationship between these two indices and long-term use may be partially explained by unmeasured 'health' anxiety or associated mental health issues arising secondary to physical comorbidities or by additional disruptive effects of physical illness on sleep.

The Charlson comorbidity score findings were not surprising given the relatively higher proportion of older adults in the long-term use group. Nonetheless, the greater degree of BZRA exposure among those patients with dementia is of concern given the risk of BZD use in this population. Similar to previous studies, prescriptions for an opioid or a psychotropic agent, such as antidepressants, antipsychotics or mood stabilisers, during the baseline period were modestly predictive for future long-term use. 48,52,54,56,58,61 Those having received a non-BZD prescription agent for a psychiatric disorder could be expected to have had greater disease severity on average than those BZRA users who did not receive such treatment early on.

An unexpected finding was the increased odds of long-term use associated with the more recent time period of the first prescription. This is contrary to what may be expected from

cumulative knowledge on BZRA and the long-standing emphasis on short-term use advised in guidelines and clinical literature. This finding may reflect the growing awareness that BZRAs should not be used as a first-line treatment resulting in only those who have not responded to other alternatives to be more likely to receive BZRAs long-term.

The present study has a number of strengths. This study used a large administrative data source that were near complete in their coverage of the study population's prescription drug dispensations and healthcare contact. Application of cohort inclusion and exclusion criteria in a carefully constructed new user longitudinal design limited confounding and bias to the extent possible. Multiple sensitivity analyses on the main outcome measure, the duration of BZRA use measurement method and the association between the independent and dependent variables for two cohorts reduced quantitative bias to increase confidence in the results.

A few important limitations should be acknowledged. Firstly, administrative data is prone to some misclassification of variables. For instance, diagnostic criteria for cohort case inclusion and exclusion will differ in their true sensitivity and specificity, regardless of prior validation of case definitions. Drugs used during any hospitalizations were not available and was assumed to be continued BZD exposure. As all independent variables were only measured cross-sectionally before or at the time of the first prescription of the first use-episode, the logistic regression model was only predictively valid for the first use episode duration and not users' average episode duration. Since DPIN only captures the days supply provided, it is possible that not all of the medication was actually taken by the patient. However, this study was able to provide insight into the prescribing practices of benzodiazepines that are filled in the pharmacy in this population. Our study did not evaluate the extent of concurrent use of multiple BZD or other psychiatric diagnoses such as substance use disorder. The databases also do not capture participation in psychological

interventions such as cognitive behavioral therapy. Moreover, while the databases are able to link several data on health information regardless of age and coverage, they do not capture other potential confounding factors such as education status and ethnicity. This study was done in a setting where there is a universal healthcare system and medication costs are covered for all Manitobans after an income-based deductible is met every year. As a result, findings may be generalizable to similar settings. Future research should aim to examine the association of repeat exposure to BZRA and risk of chronic use. Future research could also examine specific benzodiazepine type and formulations on risk of long-term use.

Conclusion

Prescribing of BZRAs was used for less than six months duration for the majority of individuals with a prior history of anxiety, depression, or insomnia. However, the proportion of long-term use among new users was up to one in three based on the average of all episodes of use, warranting future research in this area. Patients who are male, of older age, are socially or financially deprived, have poor physical health, use opioids or other psychotropic agents and are frequent consumers of healthcare resources are more likely to use BZRA long-term after their first prescription. Future research could be done to explore whether these factors need to be considered at the point of prescribing in clinical practice.

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Availability of Data and Materials

Data used in this article was derived from administrative health and social data as a secondary use. The data was provided under specific data sharing agreements only for approved use at MCHP. The original source data is not owned by the researchers or MCHP and as such cannot be provided to a public repository.

Author Statement

Jaden Brandt contributed to the conception, design, acquisition of data, analysis, and writing of the manuscript. Donica Janzen contributed to the analysis, interpretation, and writing of the manuscript. Silvia Alessi-Severini contributed to the interpretation and writing of the manuscript. Dan Chateau contributed to the interpretation and analysis of the study. Murray Enns contributed to the interpretation and writing of the study. Alexander Singer contributed to the interpretation and writing of the study. Christine Leong contributed to the conception, design, interpretation, and writing of the manuscript.

Conflict of Interest Disclosure

The authors declare no conflicts of interest related to any aspects of this work.

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Patient and Public Involvement

We have a patient advisory group who provided feedback on the dissemination of research findings.

 Table 1. Characteristics of BZRA Users by First Use Episode Duration

		BMJ Open		i6/bmjoper	
le 1. Characteristics of BZ	RA Users by First Use	Episode Duration		36/bmjopen-2020-046916 on 1	
		Short-term	Long-term		
Number of U	sers	197,606 (100%)	9,327 (100%)	206,933 (100%)	
Sex Distribution*	Male	74,487 (37.7%)	4,295 (46.1%)	Vove 206,933 (100%) er 20,78,782 (38.1%)	
	Female	123,057 (62.3%)	5,029 (53.9%)		
Age Category	18-44	101,709 (51.5%)	2,776 (29.8%)	D 128,086 (61.9%)	
	45-64	66,752 (33.8%)	3,320 (35.6%)	र्के र्के 70,072 (33.9%)	
	65+	29,143 (14.7%)	3,231 (34.6%)	32,374 (15.6%)	
SEFI-2 Score	<-1	24,955 (12.6%)	1,089 (11.7%)	26,044 (12.6%)	
	-1 to 0	81,718 (41.4%)	3,835 (41.1%)	32,374 (15.6%) 26,044 (12.6%) 85,553 (41.3%) Pprii 68,241 (33.0%)	
	0 to 1	64,967 (32.9%)	3,274 (35.1%)	9 68,241 (33.0%)	
	>1	25,966 (13.1%)	1,129 (12.1%)	27,095 (13.1%)	
esidence Distribution	Urban	125,950 (63.7%)	5,802 (62.2%)	131,752 (63.7%)	
	Rural	71,656 (36.3%)	3,525 (37.8%)	7.5,181 (36.3%)	
High Residential I	Mobility	36,392 (18.4%)	2,385 (25.6%)	9 38,777 (18.7%)	

			ijoper
sistance	18 530 (9 4%)	1 222 (13 1%)	36/bmjopen-2020-046916
Sistance	18,330 (3.470)	1,222 (13.1/0)	59 13,732 (3.370)
rd	102,461 (51.9%)	4,618 (49.5%)	€ 107,079 (51.8%)
0 (no utilization)	3,001 (1.5%)	349 (3.7%)	3,350 (1.6%) 5,980 (2.9%) 35,166 (17.0)
1	5,798 (2.9%)	182 (2.0%)	5,980 (2.9%)
2	33,974 (17.2%)	1,192 (12.8%)	35,166 (17.0)
3	127,824 (64.7%)	5,151 (55.2%)	§ 132,975 (64.3%)
4	20,065 (10.2%)	1,486 (15.9%)	21,551 (10.4%)
5 (high-utilization)	6,882 (3.5%)	964 (10.3%)	ਰੂ 7,846 (3.8%)
	Short-term	Long-term	http://bmjopen.bm
ers .	197,606 (100%)	9,327 (100%)	206,933 (100%)
0	148,257 (75.0%)	5,783 (62.0%)	8 154,040 (74.4%)
1	36,261 (18.4%)	2,031 (21.8%)	38,292 (18.5%)
2+	13,088 (6.6%)	1,513 (16.2%)	14,601 (7.1%)
0	111,216 (56.3%)	3,862 (41.4%)	38,292 (18.5%) 3 14,601 (7.1%) 3 115,078 (55.6%)
1	17,661 (8.9%)	518 (5.6%)	₹ 18,179 (8.8%)
2+	68,729 (34.8%)	4,947 (53.0%)	Pr 73,676 (35.6%)
0	132,027 (66.8%)	5,855 (62.8%)	137,882 (66.6%) opyright
	(no utilization) 1 2 3 4 5 (high-utilization) 2rs 0 1 2+ 0 1 2+	102,461 (51.9%) 0 (no utilization) 1	102,461 (51.9%) 4,618 (49.5%) 0 (no utilization) 3,001 (1.5%) 349 (3.7%) 1 5,798 (2.9%) 182 (2.0%) 2 33,974 (17.2%) 1,192 (12.8%) 3 127,824 (64.7%) 5,151 (55.2%) 4 20,065 (10.2%) 1,486 (15.9%) 5 (high-utilization) 6,882 (3.5%) 964 (10.3%) Short-term Long-term ers 197,606 (100%) 9,327 (100%) 0 148,257 (75.0%) 5,783 (62.0%) 1 36,261 (18.4%) 2,031 (21.8%) 2+ 13,088 (6.6%) 1,513 (16.2%) 0 111,216 (56.3%) 3,862 (41.4%) 1 17,661 (8.9%) 518 (5.6%) 2+ 68,729 (34.8%) 4,947 (53.0%)

				<u> </u>
	1	30,530 (15.5%)	1,011 (10.8%)	6 6 169,423 (15.2%)
	2+	35,049 (17.7%)	2,461 (26.4%)	37,510 (18.2%)
Sex of Prescriber Issuing	Male	143,619 (75.3%)	6,928 (76.5%)	고 항 하는 150,547 (75.3%)
First Prescription***	Female	47,128 (24.7%)	2,126 (23.5%)	20 21 21 49,254 (24.7%) □
Age of Prescriber Issuing	50+ Years	95,629 (52.1%)	4,775 (53.9%)	0 8 100,404 (52.2%)
First Prescription†	<50 Years	87,833 (47.9%)	4.076 (46.1%)	91,909 (47.8%)
Type of Prescriber Issuing	General Practitioner	146,823 (91.6%)	7,013 (87.5%)	153,836 (91.4%) 6,962 (4.1%)
First Prescription‡	Psychiatry	6,338 (4.1%)	624 (7.8%)	6,962 (4.1%)
	Other	7,183 (4.5%)	375 (4.7%)	7,558 (4.5%)
	2001-2006	90,008 (45.5%)	2,608 (28.0%)	92,616 (44.8%)
Period of First Prescription	2006-2011	65,750 (33.3%)	3,170 (34.0%)	68,920 (33.3%)
	2011-2016	41,848 (21.2%)	3,549 (38.1%)	³ 45,397 (21.9%)
*N=197,544 (short-term use	rs); N=9,324 (long-te	erm users); N=206,868 (total	users)	pril .
**N=197,544 (short-term use	ers); N=9,324 (long-	term users); N=206,868 (tota	al users)	18,
***N=190,747 (short-term u	sers); N=9,054 (long	g-term users); N=199,801 (to	tal users)	2024 by gu
†N=183,462 (short-term use				4 by
‡N=160,344 (short-term use	rs); N=8,012 (long-te	erm users); N=168,356 (total	users)	gu

Table 2 – Statistical Associations between Predictor Variables and Long-term Use of BZRAs

			BMJ Open			36/bmjopen-2020	
				Use D	uration	046	
Independent Variable		<u>≥180</u> . Crude OR	Days		Days	16 on ≥60	Days
			Adjusted OR (95% CI)	Crude OR (95% CI)	Adjusted OR (95% CI)	Cru z le OR (95% CI)	Adjusted OR (95% CI)
Male		1.41 (1.35-1.47)	1.33 (1.27-1.39)	1.40 (1.35-1.45)	1.34 (1.29-1.40)	ਛੂ30 (1.2 g -1.34)	1.27 (1.23-1.31)
_	18-44	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 <u>-(</u> ref)	1 (ref)
- Age	45-64	1.82 (1.73-1.92)	2.24 (2.11-2.38)	1.77 (1.70-1.85)	2.00 (1.91-2.10)	€.81 (1.7 3 -1.86)	1.89 (1.82-1.97)
_	<i>65</i> +	4.06 (3.86-4.28)	5.15 (4.81-5.52)	3.56 (3.41-3.72)	4.11 (3.88-4.36)	\$.34 (3.2 2 -3.47)	3.52 (3.36-3.70)
Rural Residenc	e	1.07 (1.02-1.11)	1.10 (1.04-1.15)	0.97 (0.93-1.00)	0.97 (0.94-1.02)	9.90 (0.8 3 -0.92)	0.92 (0.88-0.95)
High Residential Mo	obility	1.52 (1.45-1.60)	1.14 (1.08-1.21)	1.35 (1.29-1.40)	1.06 (1.01-1.11)	1 4 (1. 19 -1.18)	1.01 (0.97-1.06)
Income Assistan	ce	1.46 (1.37-1.55)	1.68 (1.55-1.81)	1.14 (1.08-1.21)	1.35 (1.26-1.45)	©.88 (0.8 4 -0.93)	1.12 (1.06-1.20)
_	<-1	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 <mark>3</mark> ref)	1 (ref)
Socio-Economic Factor	-1 to 0	1.08 (1.00-1.15)	0.99 (0.92-1.07)	0.96 (0.91-1.02)	0.91 (0.86-0.97)	9 .90 (0. ₹ -0.95)	0.89 (0.85-0.94)
Index-2 (SEFI-2) – Score	0 to 1	1.16 (1.07-1.24)	1.02 (0.94-1.10)	0.98 (0.93-1.04)	0.92 (0.87-0.98)	(0.8 3 -0.91)	0.89 (0.84-0.94)
_	>1	1 (0.92-1.09)	0.93 (0.84-1.03)	0.78 (0.73-0.84)	0.80 (0.74-0.87)	0263 (0.5 9 -0.67)	0.73 (0.68-0.78)
Married		0.91 (0.87-0.95)	0.79 (0.76-0.83)	1.01 (0.98-1.05)	0.89 (0.85-0.92)	5 .13 (1.1 以 -1.16)	0.95 (0.92-0.99)
Opioid Use		1.19 (1.14-1.27)	1.16 (1.11-1.22)	1.08 (1.04-1.12)	1.09 (1.05-1.14)	(0.9%-1.02)	1.05 (1.01-1.09)
<u>Independent Variable</u>		,	·	Use D	uration	by c	,

36/bmjopen-2020

		≥180 Days		≥90 Days		≥60 Days		
		Crude OR (95% CI)	Adjusted OR (95% CI)	Crude OR (95% CI)	Adjusted OR (95% CI)	Crudle OR (95% CI)	Adjusted OR (95% CI)	
Psychotropic Rx Use ((non-BZRA)	1.82 (1.75-1.90)	1.93 (1.83-2.02)	1.62 (1.56-1.67)	1.75 (1.69-1.83)	1.38 (1.38)	1.49 (1.44-1.54)	
	0	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1∰ref)	1 (ref)	
Charlson Comorbidity	1	1.44 (1.36-1.51)	1.11 (1.04-1.17)	1.33 (1.27-1.39)	1.08 (1.02-1.13)	ម្លី24 (1.1 ÿ -1.29)	1.04 (1.00-1.08)	
Index Score	2+	2.96 (2.79-3.15)	1.43 (1.32-1.55)	2.41 (2.29-2.54)	1.33 (1.24-1.42)	\$01 (1.9\$-2.11)	1.23 (1.15-1.31)	
	0-3	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1gref)	1 (ref)	
Resource Utilization Band	4	1.84 (1.73-1.95)	1.15 (1.07-1.23)	1.58 (1.50-1.66)	1.08 (1.01-1.14)	1.3 3 -1.43)	1.00 (0.94-1.05)	
	5	3.48 (3.24-3.73)	1.46 (1.33-1.60)	2.73 (2.56-2.92)	1.31 (1.20-1.42)	2 21 (2.0 8 -2.35)	1.17 (1.09-1.27)	
Male Prescriber of First Prescription		1.07 (1.02-1.12)	1.03 (0.98-1.09)	1.07 (1.02-1.11)	1.04 (0.99-1.09)	(0.9 % -1.05)	0.98 (0.94-1.02)	
Prescriber Age ≥50 Years		1.08 (1.03-1.12)	0.98 (0.94-1.03)	1.08 (1.04-1.12)	0.99 (0.95-1.03)	15 (1.13-1.18)	1.08 (1.04-1.11)	
	GP	1 (ref)	1 (ref)	1 (ref)	1 (ref)	<u>2</u> 1	1 (ref)	
Type of Prescriber of	Psychiatrist	2.06 (1.89-2.25)	2.11 (1.93-2.32)	1.85 (1.72-2.00)	1.89 (1.75-2.05)	<u>\$</u> 54 (1.44-1.65	1.63 (1.51-1.75)	
First Prescription	Other	1.09 (0.98-1.21)	0.92 (0.82-1.03)	1.07 (0.98-1.17)	0.92 (0.84-1.01)	1516 (1.07-1.24)	1.03 (0.96-1.11)	
	2001-2006	1 (ref)	1 (ref)	1 (ref)	1 (ref)	\$1	1 (ref)	
Period of First	2006-2011	1.66 (1.58-1.75)	1.74 (1.64-1.85)	1.58 (1.51-1.65)	1.65 (1.57-1.7)	\$.41 (1.3 <u>6</u> -1.46)	1.48 (1.42-1.54)	
Prescription	2011-2015	2.93 (2.78-3.08)	2.99 (2.80-3.18)	2.59 (2.48-2.71)	2.71 (2.57-2.8)	臣97 (1.9 % -2.05)	2.07 (1.98-2.16)	

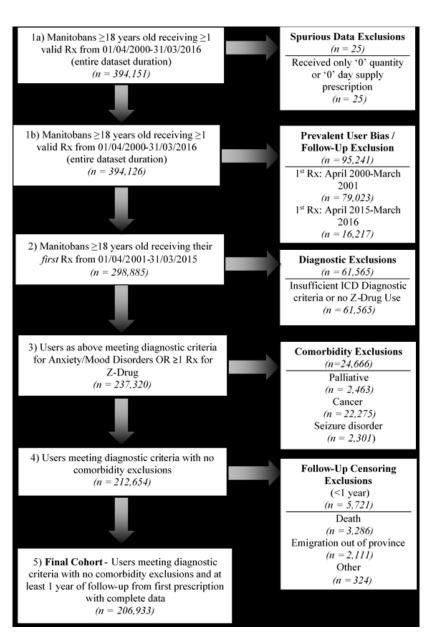
Table 3 – Frequency of Charlson Comorbidity Group Diagnoses by First Use Episode **Cohort**

Cohort Charlson Diagnosis	Short-Term Long-Term 'First- 'First-Episode' Episode' Users		Z-Test of Two	916 on 1 November 2021. Downloaded from http://bmjopen.bmj.com/ on April 18, 2024
	Users	(n=9,327)	Proportions	ver
	(n=197,606)			nber
Myocardial Infarction	2,474 (1.3%)	281 (3.0%)	p < 0.01	. 20
Congestive Heart Failure	3,943 (2.0%)	628 (6.7%)	p < 0.01	21.
Peripheral Vascular Disease	2,367 (1.2%)	256 (2.7%)	p < 0.01	Downk
Cerebrovascular Disease	3,690 (1.9%)	544 (5.8%)	p < 0.01	oade
Dementia	2,928 (1.5%)	796 (8.5%)	p < 0.01	ed f
COPD	23,064 (11.7%)	1,163 (12.5%)	p = 0.02	rom
Connective Tissue/Rheumatic Disease	2,793 (1.4%)	222 (2.4%)	p < 0.01	http:///
Peptic Ulcer Disease	2,140 (1.1%)	114 (1.2%)	p = 0.20	omj
Mild Liver Disease	2,406 (1.2%)	135 (1.4%)	p = 0.05	эре
Moderate/Severe Liver Disease	341 (0.1%)	28 (0.0%)	p < 0.01	n.bmj.c
Uncomplicated Diabetes	14,131 (7.2%)	1,099 (11.8%)	p < 0.01	om,
Complicated Diabetes	1,611 (0.8%)	252 (2.7%)	p < 0.01	/on
Paraplegia and Hemiplegia	794 (0.4%)	136 (1.5%)	p < 0.01	Ар
Renal Disease	1,858 (0.9%)	238 (2.6%)	p < 0.01	7 5
Cancer	829 (0.4%)	64 (0.1%)	p < 0.01	8, 2
Metastatic Carcinoma	64 (0.0%)	13 (0.0%)	p < 0.01	024
HIV/AIDS	50 (0.0%)	10 (0.0%)	p < 0.01	by guest. Protected by copyright

Figure Legend/Caption

Figure 1. Flowchart of study population





Flowchart of study population

48x70mm (300 x 300 DPI)

Supplemental Appendix Tables

Table A1 -Raw Data Sources and Relevant Corresponding Data Elements

Database	Date Range of Data	Relevant Data Elements
Drug Program	Apr. 1/2000 –	Prescriptions for benzodiazepines (ATC codes
Information Network	Mar. 31/2016	N03AE, N05BA, N05CD), Z-Drugs (N05CF),
(DPIN)		Antidepressants, Antipsychotics, Mood stabilisers,
		Lithium and Opioids
		-Drug, dosage strength, dosage type, metric
		quantity dispensed, day supply, date of
		dispensation
Manitoba Health	Apr. 1/1996 –	Birth date/age of patient; sex; location of
Insurance Registry	Mar. 31/2016	residence, marital status, date of Manitoba Health
		coverage, date of coverage end, reason for
		coverage end (i.e death, emigration etc.)
Medical Claims	Apr. 1/1996 –	
(Physician Billings)	Mar. 31/2016	dates of services, specific diagnoses (ICD-9 or
		ICD-10 equivalent)
Hospital Separations	Apr. 1/1996 –	Diagnoses (ICD-9 or ICD-10 equivalent), length
Abstracts	Mar. 31/2016	of stay, admission dates, discharge dates,
Provider	Apr. 1/1996 –	Physician Age, Sex, Specialty
Registry/Physician	Mar. 31/2016	
Master File		
Social Allowances	Apr. 1/2001-	Receipt of income assistance
Management Information	Mar. 31/2013	
Network (SAMIN)		

Table A2 – International Classification for Disease Coding for Mood/Anxiety/Sleep Disorders (Cohort Inclusion)

Source 1 - CPHA	Source 2 - MCHP	Study Algorithm

-	T	T	
ICD Codes	All Mental Health	Mood Disorders: 296	Mood disorders: 296
	<u>Disorders:</u>	and 311 (ICD-9-CM)	and 311 (ICD-9-CM)
	9-CM: 290-319	or F30-F34, F39	or F30-F34, F39
	10-CA: F00-F99	(ICD 10-CA)	(ICD 10-CA)
			, , ,
			Anxiety disorders:
		Anxiety Disorders:	300 (ICD-9-CM) or
		300 (ICD-9-CM) or	F40-F43 (ICD-10-
		F40-F42	CA)
			Sleep disorders: 307,
			780 or F51, G47
			ICD-10-CA)
Case Definition	≥1 hospitalization or	≥1 hospitalization or	≥1 hospitalization or
	outpatient medical	≥1-3 outpatient	≥3 outpatient
	claim within 1 year	medical claims within	medical claims within
		3-5 years*	5 years**
		1	- 1

^{*}Range of similar definitions between studies from 2000 to 2016

Table A3 – International Classification for Disease Coding Algorithms for Seizure, Cancer and Palliation (Cohort Exclusion)

	Seizure	Cancer and other	Palliation
		Neoplasms	
ICD Codes	9-CM: 345	9-CM: 140-165, 170-	N/A*
	10-CA: G40	176,179-195, 200-208	
		10-CA: C00-C99	
Case Definition	≥1 hospitalization or	≥1 hospitalization or ≥3	Carrier code
	≥3 outpatient	outpatient medical	indicating palliative
	medical claim within	claims within 5 years	drug program
	5 years before index	before index date	enrollment in DPIN
	date		

^{*}While ICD codes do exist for palliation, the DPIN carrier code '04' is expected to be a reliable indicator of when patients become ill enough that community use of medication is required for symptom management.

^{**}The decision to use a 5-year pre-exposure window was based on the fact that all patients received a BZRA, which itself increases specificity for anxiety/sleep disorder diagnoses.

Table A4 – Independent 'Patient' Variables for Prediction of Long-Term BZRA Use

Baseline Patient Characteristics	Definition (Variable Type)	Measurement Period
Age	3 age groups; 18-44, 45-64, 65+ (Ordinal)	Index Date
Sex	Male or Female (Dichotomous Categorical)	Index Date
Region	Urban; Winnipeg or Brandon postal-codes Rural; Any other Manitoba postal-code (Dichotomous Categorical)	Census Period closest in time to the index date
Socioeconomic Status	Socio-Economic Factor Index – Version 2 (SEFI-2) score composite of four variables based on geography; i) unemployment rate ii) average household income iii) proportion of single-parent households iv) proportion of population without high school education. Scores <0 indicate more favourable socioeconomic conditions Scores >0 indicate less ideal socioeconomic conditions (Ordinal Scale)	Census Period closest in time to the index date
Income Assistance	Record of income assistance (Dichotomous Categorical)	Up to 1-year before the Index Date
Marriage Record	Record of Marriage (Dichotomous Categorical)	Entire available registry period up to the Index Date
Residential Mobility (i.e frequent mover)	Average of 1 move every 3 years from beginning of registry coverage to index date (Dichotomous)	Entire available registry period up to the Index Date
Comorbidity Burden	Charlson Comorbidity Index (CCI) Score; 0, 1, 2+ (Ordinal Scale)	Up to 1-year before the Index Date
Healthcare Resource Use	Johns Hopkins Adjusted Clinical Groups Resource Utilization Band (Ordinal	Up to 1-year before the Index Date

	Scale); placement into a band (0 to 5) based on grouping of ICD	
Prescription Psychotropic Use (non-BZRA)	Receipt of Prescription (Dichotomous Categorical)	Up to 1-year before the Index Date and 6 months after the Index Date
Prescription Opioid Use	Receipt of Prescription (Dichotomous Categorical)	Up to 1-year before the Index Date and 6 months after the Index Date

Table A5 - Independent 'First-Prescription' Variables for Prediction of Long-Term BZRA
Use

Characteristics of First Consultation and Subsequent Prescription	Definition	Measurement Period
Fiscal Year Period	Fiscal year of first prescription Assigned to 3 five-year intervals; 2001-2005, 2006- 2010, 2011-2015 (Ordinal)	Index Date
Prescriber	10 Years or More (Dichotomous)	Index Date
Sex of Prescriber	Male or Female (Dichotomous)	Index Date
Prescriber Specialty	General Practitioner, Psychiatry or Other (Categorical)	Index Date

Table A6 - Logistic Regression Methodology

Criteria	Approach
	-Informal selection via published literature
Variable Selection	-Simple logistic regression; β values (p <
	0.25)
	-Dichotomous Categorical; 0 or 1
Variable Coding	-Ordinal; discrete number scale starting at 1 -Polychotomous Categorical; 0 or 1 with auto-generated dummy variables
	-No continuous variables retained
Events-per-Variable	-Minimum 10 events per independent variable rule
Conformity of Linear Gradient	-Ordered categorical variables assessed for conformity of linear gradient; nonconformity handled by variable transformation or separation into additional (design) variables (i.e fiscal year was shown to be linear with respect to outcome so condensed variable into 5-year increments)
Interaction effects	-Assessed at p < 0.01. Suspected interactions included; age*sex, residential mobility*SEFI*income assistance, psychotropic use*opioid use, RUB*CCI
Collinearity	-Analysis of variance inflation factor, correlation coefficients, eigenvalues -Significant collinearity; combine variables or removal of inferior explanatory variable
Statistical Significance	-Wald 95% CI for β and OR's
Goodness-of-Fit Measures	-C-statistic, Log-Likelihood Ratio, Hosmer- Lemeshow Statistic
Fitting Procedure	-Stepwise addition/subtraction of variables -Assessment of clinical significance

Table A7 – Goodness of Fit for Final Logistic Regression Models Predicting Gong-Term Use of **BZRA**

Model	Model Type	Independent Variables	Ratio (higher is better)	C statistic	Hosmer- Lemeshow Chi-Square Statistic
1	Main-Effects	9 Variables; Age-Sex Category, Period of First Rx, Psychotropic Use, Opioid Use, Income Assistance, Marriage, RUB CCI Score, Residential Mobility	6932 (p < 0.001)	l. Downloaded from attp://bmjopen.bmj.c	$ \begin{array}{c} 10.78 \\ (p = 0.215) \end{array} $
2	Main-Effects + Interaction Effects	10 Variables: All from Model 1 + Residential Mobility*Income Assistance	6945 (p < 0.001)	om/ on April 9 , 2024 by gues	$ \begin{array}{c} 11.02 \\ (p = 0.20) \end{array} $

 $\begin{tabular}{ll} Table \ A8-Proportion \ of \ Long-Term \ BZRA \ Use \ by \ Differing \ Parameters \ and \ Duration \ Thresholds \end{tabular}$

Scenario*	Long-Term Use Parameter	Prescription Lapse Criteria	Patients (n)	Proportion of Cohort
A1**	First-Use Episode ≥ 180 days	30 days or 50% of previous Day Supply	9,327	4.51%
A2	First-Use Episode ≥ 90 days	30 days or 50% of previous Day Supply	13,745	6.64%
A3	First-Use Episode ≥ 60 days	30 days or 50% of previous Day Supply	19,948	9.64%
A4	First-Use Episode ≥ 180 days	60 Days or 50% of previous Day Supply	13,050	6.31%
A5	First-Use Episode ≥ 180 days	90 Days	16,831	8.13%
A6	First-Use Episode ≥ 270 days	90 Days	15,214	7.35%
A7	First-Use Episode ≥ 365 days	90 Days	14,219	6.87%
B1	Mean Episode Duration ≥ 180 days	30 days or 50% of previous Day Supply	38,853	18.78%
B2	Mean Episode Duration ≥ 90 days	30 days or 50% of previous Day Supply	58,442	28.24%
В3	Mean Episode Duration ≥ 60 days	30 days or 50% of previous Day Supply	72,639	35.10%
B4	Mean Episode Duration ≥ 180 days	60 Days or 50% of previous Day Supply	44,593	21.55%
B5	Mean Episode Duration ≥ 180 days	90 Days	50,142	24.23%
В6	User Mean Episode Duration ≥ 270 days	90 Days	39,395	19.04%
В7	User Mean Episode Duration ≥ 365 days	90 Days	32,200	15.56%

^{*}A=First Episode Scenario; B=Mean Episode Duration Scenario

^{**}Primary Scenario Used for Logistic Regression

 $\begin{tabular}{ll} Table A9 - Proportion of Long-Term Z-Drug Use by Differing Parameters and Duration Thresholds \\ \end{tabular}$

Scenario	Long-Term Use Parameter	Prescription Lapse Criteria	Patients (n)	Proportion of Sub-Cohort
A1	First-Use Episode ≥ 180 days	30 days or 50% of previous Day Supply	8,206	7.41%
A2	First-Use Episode ≥ 90 days	30 days or 50% of previous Day Supply	12,155	11.0%
A3	First-Use Episode ≥ 60 days	30 days or 50% of previous Day Supply	17,126	15.5%
A4	First-Use Episode ≥ 180 days	60 Days or 50% of previous Day Supply	10,437	9.43%
A5	First-Use Episode ≥ 180 days	90 Days	12,719	11.49%
A6	First-Use Episode ≥ 270 days	90 Days	11,117	10.04%
A7	First-Use Episode ≥ 365 days	90 Days	10,045	9.07%
B1	User Mean Episode Duration ≥ 180 days	30 days or 50% of previous Day Supply	21,859	19.75%
B2	User Mean Episode Duration ≥ 90 days	30 days or 50% of previous Day Supply	32,020	28.92%
В3	User Mean Episode Duration ≥ 60 days	30 days or 50% of previous Day Supply	39,690	35.85%
B4	User Mean Episode Duration ≥ 180 days	60 Days or 50% of previous Day Supply	24,098	21.77%
В5	User Mean Episode Duration ≥ 180 days	90 Days	26,477	23.92%
В6	User Mean Episode Duration ≥ 270 days	90 Days	21,040	19.01%
В7	User Mean Episode Duration ≥ 365 days	90 Days	17,358	15.68%

Table A10 – Patient Characteristics of Z-Drug Users by First Use Episode Duration

		Short-term	Long-term	Total
			20.18 10.111	. 0 0 0
Number of Users		102,459 (100%)	8,204 (100%)	110,663 (100%)
	Male	40,516 (39.5%)	3,473 (42.3%)	43,989 (39.8%)
Sex Distribution	Female	61,943 (60.5%)	4,731 (57.7%)	66,674 (60.2%)
	18-44	42,663 (41.6%)	1,795 (21.9%)	44,458 (40.2%)
Age Category	45-64	39,817 (38.9%)	3,184 (38.8%)	43,001 (38.9%)
	65+	20,011 (19.5%)	3,227 (39.3%)	23,238 (21.0%)
	<-1	13,678 (13.3%)	981 (12.0%)	14,659 (13.2%)
6551.3.6	-1 to 0	45,136 (44.1%)	3,674 (44.8%)	48,810 (44.1%)
SEFI-2 Score	0 to 1	33,719 (32.9%)	2,885 (35.2%)	36,604 (33.1%)
	>1	9,958 (9.7%)	666 (8.1%)	10,624 (9.6%)
Residence	Urban	63,207 (61.7%)	3,313 (40.4%)	66,520 (60.1%)
Distribution	Rural	39,284 (38.3%)	4,893 (59.6%)	44,177 (39.9%)
High Residential Mobility		22,408 (21.9%)	2,523 (30.8%)	24,931 (22.5%)
Receipt of Income	Assistance	8,351 (8.2%)	758 (9.2%)	9,109 (8.2%)
Marriage Re	ecord	57,308 (55.9%)	4,595 (56.0%)	61,903 (55.9%)
	0 (no utilization)	1,771 (1.7%)	234 (2.9%)	2,005 (1.8%)
	1	3,205 (3.1%)	175 (2.1%)	3,380 (3.1%)
Johns Hopkins Healthcare	2	17,523 (17.1%)	1,012 (12.3%)	18,535 (16.7%)
Resource	3	65,067 (63.5%)	4,699 (57.3%)	69,766 (63.0%)
Utilization Band	4	10,810 (10.6%)	1,259 (15.3%)	12,069 (10.9%)
	5 (high- utilization)	4,083 (4.0%)	825 (10.1%)	4,908 (4.4%)

		Short-term	Long-term	Total	
Number of U	Jsers	102,459 (100%)	8,204 (100%)	110,663 (100%)	
Charlson	0	72,490 (70.8%)	4,528 (55.2%)	77,018 (69.6%)	
Comorbidity index	1	19,495 (19.0%)	1,905 (23.2%)	21,400 (19.3%)	
Score	2+	10,506 (10.3%)	1,773 (21.6%)	12,279 (11.1%)	
Non-BZRA	0	27,797 (27.1%)	1,784 (21.7%)	29,581 (26.7%)	
Psychotropic Prescription	1	36,939 (36.1%)	2,156 (26.3%)	39,095 (35.3%)	
Dispensations	2+	37,755 (36.8%)	4,266 (52.0%)	42,021 (38.0%)	
	0	47,427 (46.3%)	3,298 (40.2%)	50,725 (45.8%)	
Opioid Prescription Dispensations	1	34,505 (33.7%)	2,772 (33.8%)	37,277 (33.7%)	
	2+	20,559 (20.1%)	2,136 (26.0%)	22,695 (20.5%)	
Sex of Prescriber Issuing First	Male	71,485 (69.8%)	5,627 (68.6%)	77,112 (69.7%)	
Prescription	Female	28,485 (27.8%)	2,273 (27.7%)	30,758 (27.8%)	
Age of Prescriber Issuing First	50+ Years	47,871 (46.7%)	4,014 (48.9%)	51,885 (46.9%)	
Prescription	<50 Years	49,257 (48.1%)	3,758 (45.8%)	53,015 (47.9%)	
Type of Prescriber Issuing First	General Practitioner	78,610 (76.7%)	6,366 (77.6%)	84,976 (76.8%)	
Prescription	Psychiatry	3,912 (3.8%)	475 (5.8%)	4,387 (4.0%)	
1	Other	3,881 (3.8%)	381 (4.6%)	4,262 (3.9%)	
Period of First	2001-2006	34,360 (33.5%)	1,526 (18.6%)	35,886 (32.4%)	
Prescription	2006-2011	37,752 (36.8%)	2,808 (34.2%)	40,560 (36.7%)	
	2011-2016	30,379 (29.6%)	3,872 (47.2%)	34,251 (31.0%)	

Table A11 – Frequency of Charlson Comorbidity Group Diagnoses by First Use Episode Duration for Z-Drug Cohort

Charlson Diagnosis	Short-Term	Long-Term 'First-	
	'First-Episode'	Episode' Users	Z-Test of Two
	Users	(n=8,204)	Proportions
	(n=102,459)		
Myocardial Infarction	1,836 (1.8%)	306 (3.7%)	p < 0.01
Congestive Heart Failure	3,174 (3.1%)	700 (8.5%)	p < 0.01
Peripheral Vascular Disease	1,772 (1.7%)	284 (3.5%)	p < 0.01
Cerebrovascular Disease	2,321 (2.3%)	550 (6.7%)	p < 0.01
Dementia	1,925 (1.9%)	865 (10.5%)	p < 0.01
COPD	12,357 (12.1%)	1,171 (14.3%)	p < 0.01
Connective	1,906 (1.9%)	243 (3.0%)	p < 0.01
Tissue/Rheumatic Disease	1,900 (1.970)	243 (3.070)	p < 0.01
Peptic Ulcer Disease	1,111 (1.1%)	123 (1.5%)	p < 0.01
Mild Liver Disease	1,672 (1.6%)	139 (1.7%)	p = 0.33
Moderate/Severe Liver	275 (0.2%)	38 (0.4%)	p < 0.01
Disease	273 (0.270)	38 (0.470)	p < 0.01
Uncomplicated Diabetes	9,317 (9.1%)	1,150 (14.0%)	p < 0.01
Complicated Diabetes	1,639 (1.6%)	328 (4.0%)	p < 0.01
Paraplegia and Hemiplegia	508 (0.5%)	136 (1.7%)	p < 0.01
Renal Disease	1,543 (1.5%)	293 (3.6%)	p < 0.01
Cancer	2,109 (2.1%)	247 (3.0%)	p < 0.01
Metastatic Carcinoma	429 (0.4%)	45 (0.5%)	p = 0.04
HIV/AIDS	118 (0.1%)	16 (0.2%)	p = 0.02

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Table A12 – Statistical Associations between Predictor Variables and Long-term see of Z-Drugs

	– Statisticai Ass				uration	6916	- -
		≥180 days		≥90 days		° ≥60 days	
Independent Variable		Crude OR (95% CI)	Adjusted OR (95% CI)	Crude OR (95% CI)	Adjusted OR (95% CI)	Crı ₹ le OR (9 . % CI)	Adjusted OR (95% CI)
Male		1.12 (1.07-1.18)	1.04 (0.99-1.09)	1.13 (1.08-1.17)	1.05 (1.01-1.10)	₹.08 (1.0₹-1.12)	1.04 (1.00-1.08)
	18-44	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 (ref)
Age	45-64	1.90 (1.79-2.02)	2.02 (1.89-2.17)	1.74 (1.66-1.82)	1.78 (1.68-1.88)	2 .71 (1. ₫ -1.78)	1.68 (1.60-1.76)
-	<i>65</i> +	3.83 (3.61-4.07)	3.71 (3.44-4.00)	3.24 (3.08-3.40)	3.08 (2.90-3.28)	\$.99 (2.8 3 -3.12)	2.78 (2.64-2.93)
Rural Re	esidence	0.92 (0.88-0.96)	1.13 (1.07-1.19)	0.99 (0.96-1.03)	1.02 (0.98-1.07)	1.08 (1.04-1.11)	0.95 (0.91-0.99)
High Resider	ntial Mobility	1.59 (1.51-1.67)	1.26 (1.19-1.33)	1.53 (1.46-1.59)	1.21 (1.15-1.27)	(1.26-1.35)	1.12 (1.07-1.17)
Income A	ssistance	1.15 (1.06-1.24)	1.47 (1.34-1.61)	1.02 (0.95-1.09)	1.29 (1.19-1.40)	6 .82 (0.7∄-0.87)	1.08 (1.00-1.17)
	<-1	1 (ref)	1 (ref)	1 (ref)	1 (ref)	l (ref)	1 (ref)
SEFI-2 Score	-1 to 0	1.14 (1.06-1.22)	1.07 (0.99-1.16)	1.03 (0.97-1.09)	0.98 (0.92-1.04)	€ .95 (0. 9 1.00)	0.94 (0.89-0.99)
SEF1-2 Score	0 to 1	1.19 (1.11-1.29)	1.08 (0.99-1.17)	1.04 (0.98-1.11)	0.99 (0.93-1.06)	6 92 (0.8 2 -0.97)	0.93 (0.88-0.99)
	>1	0.93 (0.84-1.03)	0.84 (0.75-0.94)	0.80 (0.73-0.87)	0.77 (0.70-0.85)	(0.6%) -0.73)	0.72 (0.66-0.78)
Married		1.00 (0.96-1.05)	0.86 (0.82-0.91)	1.07 (1.03-1.10)	0.93 (0.89-0.98)	1.13 (1.13)-1.17)	0.98 (0.94-1.01)
Opioid Use		1.28 (1.22-1.34)	1.15 (1.09-1.21)	1.26 (1.21-1.31)	1.15 (1.11-1.20)	1.18 (1.14-1.21)	1.11 (1.07-1.15)

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				Use Di	uration		
r 1 1	4 \$7 . 11	≥180 days		≥90 days		04 69 16 ≥60 days	
<u>Independent Variable</u>		Crude OR (95% CI)	Adjusted OR (95% CI)	Crude OR (95% CI)	Adjusted OR (95% CI)	Crude OR (95½/6 CI)	Adjusted OR (95% CI)
		1.34	1.24	1.35	1.27	(95€ 6 CI) § .22	1.19
Psychotropic Rx U	Use (Non-BZRA)	(1.27-1.41)	(1.17-1.32)	(1.29-1.41)	(1.20-1.33)	(1. 1 .27)	(1.14-1.24)
	0	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1gref)	1 (ref)
Charlson Comorbidity	1	1.56 (1.48-1.65)	1.25 (1.18-1.33)	1.45 (1.39-1.52)	1.21 (1.15-1.27)	1.33 (1.2 3 -1.38)	1.13 (1.08-1.19)
Index Score	2.	2.70	1.46	2.34	1.38	₹.02	1.30
	2+	(2.55-2.87)	(1.36-1.58)	(2.22-2.46)	(1.29-1.47)	(1.93 - 2.12)	(1.22-1.37)
	0-3 (≤Moderate)	1 (ref)	1 (ref)	1 (ref)	1 (ref)	1 ≩ ref)	1 (ref)
Resource	1 (High)	1.67	1.16	1.47	1.09	2 30	1.00
Utilization Band	4 (High)	(1.56-1.78)	(1.08-1.25)	(1.39-1.56)	(1.01-1.16)	(1.24-1.37)	(0.95-1.07)
Cittization Bana	5 (Very High)	2.89	1.55	2.43	1.42	<u>\$</u>.97	1.22
	J (very migh)	(2.67-3.13)	(1.41-1.70)	(2.26-2.61)	(1.30-1.55)	(1.85-2.11)	(1.12-1.32)
Male Prescriber of	First Prescription	0.99	0.97	0.98	0.98	g .94	0.93
mare i reservoer of	1 trist 1 resertpiton	(0.94-1.04)	(0.92-1.03)	(0.94-1.02)	(0.93-1.02)	(0.90-0.97)	(0.90-0.97)
Prescriber Age ≥50 Years		1.10	0.98	1.10	0.98	§ .15	1.05
		(1.05-1.15)	(0.93-1.03)	(1.06-1.15)	(0.94-1.02)	(1.19-1.19)	(1.01-1.09)
	GP	1 (ref)	1 (ref)	1 (ref)	1 (ref)	l₫(ref)	1 (ref)
Prescriber of	Psychiatrist	1.50	1.96	1.36	1.72	<u><u>=</u> 11</u>	1.38
First Prescription	Other	(1.36-1.66)	(1.76-2.17)	(1.25-1.49)	(1.57-1.89)	(1.02-1.20)	(1.27-1.51)
-		1.21 (1.09-1.35)	0.92 (0.82-1.03)	1.18 (1.07-1.29)	0.91 (0.83-1.00)	1.19 (1.19-1.29)	0.98 (0.91-1.07)
	2001-2006	1 (ref)	1 (ref)	1 (ref)	1 (ref)	la (ref)	1 (ref)
D 1 CD	2006-2011	1.68	1.57	1.67	1.56	<u>1</u> 53	1.46
Period of First Prescription		(1.57-1.79)	(1.46-1.68)	(1.59-1.76)	(1.47-1.66)	$(1.4\overline{\$}-1.60)$	(1.39-1.54)
Ττεκετιριίση	2011-2015	2.87	2.45	2.83	2.44	क्रे.20	1.96
	2011-2013	(2.70-3.05)	(2.28-2.65)	(2.69-2.97)	(2.30-2.59)	(2.19-2.29)	(1.86-2.07)

STROBE Statement—checklist of items that should be included in reports of observational studies

	Item No	Recommendation	Page No
Title and abstract	1	(a) Indicate the study's design with a commonly used term in the title or	1
		the abstract	
		(b) Provide in the abstract an informative and balanced summary of what	3
		was done and what was found	
Introduction			
Background/rationale	2	Explain the scientific background and rationale for the investigation being	5
		reported	
Objectives	3	State specific objectives, including any prespecified hypotheses	6
Methods			
Study design	4	Present key elements of study design early in the paper	6
Setting	5	Describe the setting, locations, and relevant dates, including periods of	7-9
s • · · · · · · · · · · · · · · · · · ·		recruitment, exposure, follow-up, and data collection	'
Participants	6	(a) Cohort study—Give the eligibility criteria, and the sources and	7-9
1 with pulled	Ü	methods of selection of participants. Describe methods of follow-up	'
		Case-control study—Give the eligibility criteria, and the sources and	
		methods of case ascertainment and control selection. Give the rationale	
		for the choice of cases and controls	
		Cross-sectional study—Give the eligibility criteria, and the sources and	
		methods of selection of participants	7.0
		(b) Cohort study—For matched studies, give matching criteria and	7-9
		number of exposed and unexposed	
		Case-control study—For matched studies, give matching criteria and the	
		number of controls per case	
Variables	7	Clearly define all outcomes, exposures, predictors, potential confounders,	8-10
		and effect modifiers. Give diagnostic criteria, if applicable	
Data sources/	8*	For each variable of interest, give sources of data and details of methods	7-8
measurement		of assessment (measurement). Describe comparability of assessment	
		methods if there is more than one group	
Bias	9	Describe any efforts to address potential sources of bias	n/a
Study size	10	Explain how the study size was arrived at	7
Quantitative variables	11	Explain how quantitative variables were handled in the analyses. If	7-10
		applicable, describe which groupings were chosen and why	
Statistical methods	12	(a) Describe all statistical methods, including those used to control for	10
		confounding	
		(b) Describe any methods used to examine subgroups and interactions	10
		(c) Explain how missing data were addressed	10-
			11
		(d) Cohort study—If applicable, explain how loss to follow-up was	10-
		addressed	111
		Case-control study—If applicable, explain how matching of cases and	1.1
		controls was addressed	
		Cross-sectional study—If applicable, describe analytical methods taking	
		account of sampling strategy	
		(\underline{e}) Describe any sensitivity analyses	1

Continued on next page



Results			
Participants	13*	(a) Report numbers of individuals at each stage of study—eg numbers potentially eligible, examined for eligibility, confirmed eligible, included in the study, completing follow-up, and analysed	11- 12
		(b) Give reasons for non-participation at each stage	7
		(c) Consider use of a flow diagram	7, fig1
Descriptive data	14*	(a) Give characteristics of study participants (eg demographic, clinical, social) and information on exposures and potential confounders	11,12
		(b) Indicate number of participants with missing data for each variable of interest	Fig 1
		(c) <i>Cohort study</i> —Summarise follow-up time (eg, average and total amount)	11
Outcome data	15*	Cohort study—Report numbers of outcome events or summary measures over time	12- 13
		Case-control study—Report numbers in each exposure category, or summary measures of exposure	n/a
		Cross-sectional study—Report numbers of outcome events or summary measures	n/a
Main results	16	(a) Give unadjusted estimates and, if applicable, confounder-adjusted estimates and	Table
		their precision (eg, 95% confidence interval). Make clear which confounders were adjusted for and why they were included	2
		(b) Report category boundaries when continuous variables were categorized	Table 2
		(c) If relevant, consider translating estimates of relative risk into absolute risk for a meaningful time period	Table 2
Other analyses	17	Report other analyses done—eg analyses of subgroups and interactions, and sensitivity analyses	Table 2-3
Discussion			•
Key results	18	Summarise key results with reference to study objectives	13
Limitations	19	Discuss limitations of the study, taking into account sources of potential bias or imprecision. Discuss both direction and magnitude of any potential bias	
Interpretation	20	Give a cautious overall interpretation of results considering objectives, limitations,	13-
C 1: - 1:11:4	21	multiplicity of analyses, results from similar studies, and other relevant evidence	17
Generalisability	21	Discuss the generalisability (external validity) of the study results	17- 18
Other informati	on		1
Funding	22	Give the source of funding and the role of the funders for the present study and, if applicable, for the original study on which the present article is based	25

^{*}Give information separately for cases and controls in case-control studies and, if applicable, for exposed and unexposed groups in cohort and cross-sectional studies.

Note: An Explanation and Elaboration article discusses each checklist item and gives methodological background and published examples of transparent reporting. The STROBE checklist is best used in conjunction with this article (freely available on the Web sites of PLoS Medicine at http://www.plosmedicine.org/, Annals of Internal Medicine at http://www.annals.org/, and Epidemiology at http://www.epidem.com/). Information on the STROBE Initiative is available at www.strobe-statement.org.