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Surgical care for the aged: a retrospective cross-sectional study of a national surgical mortality audit

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ABSTRACT

Objective: It is assumed that increased age signifies increased surgical care. Few surgical studies describe the differences in care provided to older patients compared with younger patients. We aimed to examine the relationships between increasing age, preoperative factors and markers of postoperative care in adults who died in-hospital after surgery in Australia.

Design: This retrospective cross-sectional study extracted data from a national surgical mortality audit—an independent, peer-reviewed process.

Setting: From January 2009 to December 2012, 111 public and 61 private Australian hospitals notified the audit of in-hospital deaths after general anaesthetic surgery or if the patient was admitted under a surgeon.

Participants: Notified deaths totalled 19 723. We excluded deaths if patients were brain dead, younger than 17 years or never had an operation (n=11 376). From this baseline population, we divided 11 201 deaths into three patient age groups: youngest (17–64 years), medium (65–79 years) and oldest (≥80 years).

Outcome measures: Univariable and multivariable logistic regression analyses determined the relationships between increasing age and the measured preoperative factors and postoperative variables.

Results: The baseline population’s median age was 78 years (IQR 66–85), 43.7% (4892/11 201) were 80 years or older and 83.4% (9319/11 723) had emergency admissions. The oldest group had increased trauma and emergency admissions than the medium and youngest age groups. Seven of the eight measured markers of postoperative care demonstrated strong and significant relationships with increasing age. The oldest group compared with the medium group had decreased rates of: unplanned returns to theatre (11.2% (526/4709) vs 20.2% (726/3586)), unplanned intensive care admissions (16.3% (545/3350) vs 24.0% (601/2504)) and treatment in intensive care units (59.7% (2689/4507) vs 76.7% (2754/3590)).

Conclusions: The oldest patients received lower levels of care than the medium and youngest age groups.

INTRODUCTION

Older patients generally require more surgical services relative to younger patients1 as a consequence of the ageing process and accumulation of chronic disease.2–3 As life expectancy and population numbers increase,4,5 so will the requirement for elective and emergency surgical services.6–8 But surgical care for the older patient is complex. The presence of comorbidities in older patients complicates surgical decision-making,9 and is reported to be associated with increases in postoperative complications10,11 due to the patients functional, physiological, psychological and social effects.2,8,12

There is a scarcity of data about overall surgical management in older patients. Few surgical studies8,13 have described the differences in care provided to older patients compared with that provided to younger patients. Boumendil et al14 found that elderly patients in France had fewer admissions to intensive care than younger patients. Most studies discuss technical surgical procedures on elderly patients, rather than postoperative care for this age group.15

Surgical literature is inconsistent in defining old age, with ≥65,14 ≥758 and ≥80 years4
all being used. Old age is sometimes used as an independent predictor for surgical morbidity and mortality, but it is reported as having both minimal and increased effects—possibly because researchers assess the impact of chronological age in a dichotomous fashion using an arbitrary age.

This study aimed to examine the relationships between increasing age, particularly old age, preoperative factors and markers of postoperative care in adult surgical patients who died in-hospital after surgery.

To do this effectively, the audit data were divided into three age groups: youngest (17–64 years), medium (65–79 years) and oldest (≥80 years). We describe the differences in preoperative factors and postoperative care provided to these patients who died.

**METHODS**

This retrospective cross-sectional study analysed data collected in the Australian and New Zealand Audit of Surgical Mortality (ANZASM), between 1 January 2009 and 31 December 2012. This study covers all of Australia, which has a population of 22.9 million people. During the study period, there were ~8.4 million patient separations from hospital after surgery. No baseline patient population data were available for patients discharged alive from hospital after surgery. The audit data are provided by treating surgeons and 94% (4268/4540) of Australian surgeons participate in the audit.

Hospitals included in the audit are teaching hospitals that perform surgery that requires an anaesthetist. The full audit process is published (see online supplementary 1) and briefly described here. Hospitals report in-hospital deaths to the audit, independent of the surgeon, if patients were under the care of a surgeon at the time of death. Reported deaths also include patients palliated postoperatively either at the same hospital or an associated facility. Audit data are systematically collected using a standard data collection form (designed by surgeons) which covers all aspects of surgical care (see online supplementary 2—not for publication). The treating surgeon completes the form, which is de-identified and assessed by external peer surgeons. When required, forms were sent for secondary assessment. Some variables have minor denominator variation as not every question was always answered. Preoperative and postoperative variable answers are based on hospital medical record notes. Surgeon assessors use their professional judgement and clinical evidence when answering subjective questions on the form. Clinical incidents in the audit are defined as areas in healthcare that the surgeon believed could be improved or different; or should have been better; or adverse events.

For the analysis, cases were excluded if forms were pending from the treating surgeon and if the patient never had surgery, had an American Society of Anaesthesiologists (ASA) class of VI (organ donor), was admitted for non-operative terminal care and was younger than 17 years of age. We extracted preoperative and postoperative variables from the audit database.

Preoperative variables included: patient age (categorised as youngest, 17–64 years; medium, 65–79 years; and oldest, ≥80 years), gender, type of admission (emergency or elective), the presence of one or more comorbidities, malignancy status, involvement of trauma, ASA class, patient transfer, delay in diagnosis and surgical specialty.

Postoperative markers of care variables included: fluid balance problems, unplanned return to theatre, unplanned intensive care unit (ICU) admission, treated in ICU, clinical incidents, postoperative complication, infection present at death and if retrospectively the surgeon would have done anything differently when managing the patient.

We divided the included audit data into three age groups for several reasons: the discrepant definitions of old age in the literature; our opinion that 65 years is not a true cut-off mark for old age and the biological differences found with increasing age.

Selection bias for reporting of deaths is not present, as deaths are reported independent of the surgeons. Classification bias is not present as death is an end point. Reporting bias may be present for clinical incidents (these are reported retrospectively by the surgeon) and malignancy (a result of missing data, which have been theoretically corrected using multiple imputation).

**External validation of the audit data**

The sensitivity of the audit reporting process was externally validated. The deaths reported to the audit were compared with the total number of postoperative mortalities in Australia using Australian Institute of Health and Welfare (AIHW) data.

**Ethical approval**

No ethical approval was sought as ANZASM is a protected quality assurance activity under Part VC of the Health Insurance Act 1973 (gazetted August 2011). New Zealand data are excluded in this analysis.

“According to the policy activities that constitute research at the Australian and New Zealand Audit of Surgical Mortality, this work met criteria for operational improvement activities exempt from ethics review.”

**Statistical analysis**

We calculated ORs for each variable, comparing by age group (ie, youngest vs medium, youngest vs oldest and medium vs oldest). Summary statistics are presented as median (IQR) for continuous variables and frequency (percentage) for categorical variables. The association between continuous variables and age was calculated using the Mann-Whitney U test. The association between categorical variables and age was calculated with Fisher’s exact test.

We calculated the association between age and both preoperative factors and surgical complications using
both univariable and multivariable logistic regression. In all multivariable models, gender, ASA class, type of admission and presence of malignancy were included as covariables except for the outcomes gender (where gender is excluded as a covariable), admission type (where admission type is excluded as a covariable) and presence of malignancy (where presence of malignancy is excluded as a covariable). Owing to the relatively high amount of missing data for the variable malignancy, multiple imputations using logistic regression was used to impute 20 sets of values. The covariables used in the imputation were gender, type of admission (emergency/elective), presence of comorbidities (yes/no), whether patient was transferred (yes/no) and whether diagnosis was delayed (yes/no). The univariable association between presence of malignancy and age was assessed using imputed malignancy data. In all multivariable logistic regressions, imputed malignancy data were used. For all other variables, missing items were excluded from analyses. Regression results are presented as OR and 95% CIs. Significance values were based on two-tailed tests, with p<0.05 considered statistically significant. We performed all analyses using SPSS V.19 (IBM, Armonk, New York, USA) and Stata V.12.1 (StataCorp, College Station, Texas, USA).

RESULTS
This Australia-wide surgical mortality audit included 61 private and 111 public teaching hospitals that notified the audit of their surgical deaths (this covered 20% of private hospitals and 99% of public teaching hospitals). Treating surgeons at the participating hospitals completed forms for their own cases, with a loss to follow-up rate of 2.4% (482 of 19 723). Seventy-five percent (74.6%, 11 376/15 021) of the included deaths were considered to be the baseline population of which 98.5% (11 201 of 15 021) were categorised by age group. The sensitivity of the reporting process was externally validated by comparison with AIHW data and this showed a 97.6% correlation—14 659 deaths reported to AIHW compared with 15 021 in the audit. To

Figure 1  Flow chart of included mortality data.
†Cases from non-participating surgeons.
‡Cases not returned after 2 years from notification, despite regular reminders, were considered “lots to follow-up”.
§Cases not returned, but <2 years, therefore not “lost of follow-up”.
||Patients classed as American Society of Anaesthesiologists (ASA) 6 are brain-dead.
Table 1  Patients’ characteristics of the baseline population and by age group

<table>
<thead>
<tr>
<th>Patient characteristics</th>
<th>Baseline population (n=11 376)*</th>
<th>Youngest Age 17–64 (n=2514)†</th>
<th>Medium Age 65–79 (n=3795)†</th>
<th>Oldest Age 80+ (n=4892)†</th>
</tr>
</thead>
<tbody>
<tr>
<td>Median age (years) (IQR)</td>
<td>78.0 (66–85)</td>
<td>54.0 (43–60)</td>
<td>74.0 (69–77)</td>
<td>86.0 (83–90)</td>
</tr>
<tr>
<td>Sex‡</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Female</td>
<td>5081/11 373 (44.7)</td>
<td>952/2514 (37.9)</td>
<td>1526/3795 (40.2)</td>
<td>2531/4892 (51.7)</td>
</tr>
<tr>
<td>Male</td>
<td>6292/11 373 (55.3)</td>
<td>1560/2514 (62.1)</td>
<td>2269/3795 (59.8)</td>
<td>2361/4892 (48.3)</td>
</tr>
<tr>
<td>Type of admission‡</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Elective</td>
<td>1854/11 173 (16.6)</td>
<td>382/2467 (15.5)</td>
<td>807/3717 (21.7)</td>
<td>637/4822 (13.2)</td>
</tr>
<tr>
<td>Emergency</td>
<td>9319/11 173 (83.4)</td>
<td>2084/2467 (84.5)</td>
<td>2909/3717 (78.3)</td>
<td>4185/4822 (86.8)</td>
</tr>
<tr>
<td>Malignancy‡</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Malignancy not present</td>
<td>5255/7755 (67.8)</td>
<td>1220/1698 (71.8)</td>
<td>1588/2566 (61.9)</td>
<td>2301/3264 (70.5)</td>
</tr>
<tr>
<td>Malignancy present</td>
<td>2287/7755 (29.5)</td>
<td>455/1612 (28.2)</td>
<td>938/2493 (37.6)</td>
<td>882/3296 (26.8)</td>
</tr>
<tr>
<td>Malignancy unknown</td>
<td>213/7755 (2.8)</td>
<td>42/1698 (2.4)</td>
<td>50/2566 (1.9)</td>
<td>86/3264 (2.6)</td>
</tr>
<tr>
<td>ASA class‡</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1</td>
<td>262 (2.4)</td>
<td>223 (8.9)</td>
<td>341 (0.9)</td>
<td>147 (0.3)</td>
</tr>
<tr>
<td>2</td>
<td>886 (8.1)</td>
<td>219 (8.7)</td>
<td>319 (8.4)</td>
<td>264 (5.4)</td>
</tr>
<tr>
<td>3</td>
<td>3644 (33.4)</td>
<td>458 (18.2)</td>
<td>1203 (31.7)</td>
<td>1722 (35.2)</td>
</tr>
<tr>
<td>4</td>
<td>4889 (44.9)</td>
<td>978 (38.9)</td>
<td>1730 (45.6)</td>
<td>2284 (46.7)</td>
</tr>
<tr>
<td>5</td>
<td>1218 (11.2)</td>
<td>561 (22.3)</td>
<td>501 (13.2)</td>
<td>592 (12.1)</td>
</tr>
<tr>
<td>Comorbidity‡</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Comorbidity present</td>
<td>9687/10 768 (90.0)</td>
<td>1774/2435 (72.9)</td>
<td>3427/3740 (91.6)</td>
<td>4801/4866 (98.7)</td>
</tr>
<tr>
<td>Length of hospital stay—median days (IQR)‡</td>
<td>8 (3–17)</td>
<td>8 (3–20)</td>
<td>11 (5–23)</td>
<td>9 (4–19)</td>
</tr>
<tr>
<td>Length of time from operation to death median days (IQR)‡</td>
<td>7 (2–16)</td>
<td>6 (1–14)</td>
<td>8 (2–18)</td>
<td>7 (2–15)</td>
</tr>
<tr>
<td>Surgical specialty§</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>General</td>
<td>4496/11 376 (39.5)</td>
<td>965/2514 (38.4)</td>
<td>1662/3795 (43.8)</td>
<td>1869/4892 (38.2)</td>
</tr>
<tr>
<td>Neurosurgery</td>
<td>1484/11 376 (13.0)</td>
<td>787/2514 (31.3)</td>
<td>501/3795 (13.2)</td>
<td>196/4892 (4.0)</td>
</tr>
<tr>
<td>Orthopaedics</td>
<td>2282/11 376 (20.0)</td>
<td>131/2514 (5.2)</td>
<td>429/3795 (11.3)</td>
<td>172/4892 (35.2)</td>
</tr>
<tr>
<td>Cardiothoracic</td>
<td>699/11 376 (6.1)</td>
<td>214/2514 (8.5)</td>
<td>338/3795 (8.9)</td>
<td>147/4892 (3.0)</td>
</tr>
</tbody>
</table>

*Baseline population (includes 175 patients not categorised by age).
†Patients categorised by age.
‡Missing items (denominator variation) applicable to the baseline population: sex=3, type of admission=203, malignancy=3621, ASA=477, length of stay=3414.
§Smaller specialties not listed.
ASA, American Society of Anaesthesiologists.
demonstrate the differences in care by age, 22.4% (2514) were categorised as youngest, 33.9% (3795) as medium and 43.7% (4892) as oldest (figure 1).

Patient characteristics for the baseline population and by age group are shown in table 1. The baseline population’s median age was 78 years (IQR 66–85) and less than a quarter of the patients had an elective admission. Nearly half the patients (44.9%, 4892/10 899) had pre-operative incapacitating systemic disease that posed a constant threat to life (ie, ASA4 class). Preoperatively surgeons expected 13% of the patients would die and 87% had less than expected risk of death regardless of age.

Patients in the oldest group (table 1) were more likely to be female (51.7% (2531/4892)) compared with the baseline population (44.7% (5081/11 376)) as well as the medium (37.9% (952/2514)) and youngest groups (40.2% (1526/3795)). The prevalence of comorbidities increased with increasing age (72.9% (1774/2435) vs 91.6% (3427/3740) vs 98.7% (4801/4866)). The median length of hospital stay was longest for the medium group (11 vs 9 days in the oldest and 8 days in the youngest group).

Analysed preoperative variables (table 2) demonstrated a significant influence of age on gender, emergency admission, presence of malignancy, presence of at least one or more comorbidity and whether the patient was transferred. There was no significant influence of age on evidence of trauma. Adequate discrimination and calibration of all logistic regression models was attained.

Figure 2 illustrates the trends of the analysed markers of postoperative care reported in table 3. More than 90% of these markers of care showed a statistically significant and strong relationship with increasing age. These findings were evident in both univariable and multivariable analyses. The oldest patients received less care than the medium age group. The oldest patients were less likely to have postoperative complications reported (33.6%, 1621/4824) than the middle (40.5%, 1504/3711) but more likely than the youngest age group (30.3%, 737/2433).

**DISCUSSION**

We found that decreased patterns of postoperative care occurred in the oldest patients when we divided the patients into three age groups. The differences in care between these groups were statistically significant.

Most of the postoperative markers of care analysed (seven of the eight) demonstrated lower levels of aggressive intervention and resource use. These interventions included admissions to intensive care, when the patients were over 80 years of age. The oldest patients were treated differently from the younger patient groups. Despite being admitted with increased incidence of trauma and more extensive comorbidities, they had fewer postoperative complications diagnosed and shorter length of hospital stay before death than the middle-aged group.
The medium age group (65–79 years) died in a setting of high resource use: highest rates of interhospital transfers, postoperative complications, clinical incidents, unplanned returns to theatre, unplanned intensive care admissions and they had the longest length of hospital stay. Interestingly, this group had the lowest rate of trauma but the highest rate of malignancy, which may have influenced the decisions for ongoing aggressive treatment.

The levels of care in the youngest age group (17–64 years) parallel the levels of care in the oldest age group (80+ years); however, the level of care for both of these age groups was lower than the care given to the medium age group (65–79 years).

Only a few publications have assessed the use of postoperative intensive care for older surgical patients. Our findings are similar to Boumendil et al. who reported decreased admission rates to ICU in patients 80 years and older (10% compared with 29% for patients aged 65–79 years). Recently, however, Nguyen et al. reported that the expected benefits of medical or unplanned surgical ICU admissions of patients aged 80 years and older are particularly weak and that the admission of these categories of patients to ICU is questionable. Admission to an ICU is often rationalised due to high operational costs whatever the age group, and may be influenced by the presence of malignancy and its staging. Because our data showed that the oldest group had the lowest prevalence of malignancy and the lowest admission to ICU, it is clear that factors other than the diagnosis influenced the decision for ICU admission.

Comorbidities in older patients are reported as being associated with increases in postoperative complications. The management decisions with these patients may be influenced by the patients’ functional, physiological, psychological and social conditions. But our data showed that detection of postoperative complications was lowest in the oldest age group, despite virtually all having multiple comorbidities. Our data suggest that there may be a culture of less intensive investigation, monitoring and possible failure to intervene in the elderly group. From our data, it is difficult to determine the appropriateness of these levels of care because the study is based on after death self-reporting. Guidelines and scoring systems have been developed to ascertain the point at which ‘intervention’ is appropriate, that is, when surgery or aggressive intervention is futile. The use of scoring systems is rarely quoted by self-reporting surgeons in the data of this mortality audit. This suggests that such systems should be used in the mostly elderly surgical population on an everyday basis.

The greatest strength of this study is its large sample size. This ensured the robustness of the data and the ability to make broad statements concerning the findings. It encompassed all surgical specialties, is multi-centred covering approximately 70% of Australian surgical hospitals—both private and public. Over 90% of Australian surgeons participated as part of the Continuing Professional Development Program of the Royal Australasian College of Surgeons. Data were systematically collected using a standard data tool.
### Table 3: Comparison of postoperative factors of oldest against youngest and medium age patients

<table>
<thead>
<tr>
<th>Postoperative factors</th>
<th>Youngest Age 17-64</th>
<th>Medium Age 65-79</th>
<th>Oldest Age 80+</th>
<th>Youngest versus medium</th>
<th>Youngest versus oldest</th>
<th>Medium versus oldest</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>n=2514</td>
<td>n=3795</td>
<td>n=4892</td>
<td>Univariable OR</td>
<td>p Value</td>
<td>Univariable OR</td>
</tr>
<tr>
<td>Fluid balance problems (%)</td>
<td>136/2308 (5.9%)</td>
<td>325/3543 (9.2%)</td>
<td>548/4645 (11.8%)</td>
<td>1.61 (1.31 to 1.98)</td>
<td>&lt;0.001*</td>
<td>1.53 (1.18 to 1.98)</td>
</tr>
<tr>
<td>Unplanned return to theatre (%)</td>
<td>439/2331 (18.8%)</td>
<td>726/3586 (20.2%)</td>
<td>526/4709 (11.2%)</td>
<td>1.09 (0.96 to 1.25)</td>
<td>0.182</td>
<td>1.12 (0.94 to 1.33)</td>
</tr>
<tr>
<td>Unplanned ICU admission (%)</td>
<td>329/1618 (20.3%)</td>
<td>601/2504 (24.0%)</td>
<td>454/3350 (16.3%)</td>
<td>1.24 (1.06 to 1.44)</td>
<td>0.006*</td>
<td>1.02 (0.86 to 1.21)</td>
</tr>
<tr>
<td>Treated in ICU (%)</td>
<td>1973/2410 (81.9%)</td>
<td>2754/3590 (76.7%)</td>
<td>2689/4507 (59.7%)</td>
<td>0.73 (0.64 to 0.83)</td>
<td>0.001*</td>
<td>0.73 (0.61 to 0.88)</td>
</tr>
<tr>
<td>Different action by surgeon (%)</td>
<td>295/1586 (18.6%)</td>
<td>518/2466 (21.0%)</td>
<td>490/3295 (14.9%)</td>
<td>1.16 (0.99 to 1.36)</td>
<td>0.062</td>
<td>1.04 (0.88 to 1.24)</td>
</tr>
<tr>
<td>Clinical incidents (%)</td>
<td>446/1719 (26.0%)</td>
<td>728/2594 (28.1%)</td>
<td>745/3436 (21.7%)</td>
<td>1.12 (0.97 to 1.28)</td>
<td>0.119</td>
<td>1.04 (0.89 to 1.21)</td>
</tr>
<tr>
<td>Postoperative complication (%)</td>
<td>737/2433 (30.3%)</td>
<td>1504/3711 (40.5%)</td>
<td>1621/4824 (33.6%)</td>
<td>1.57 (1.41 to 1.75)</td>
<td>&lt;0.001*</td>
<td>1.45 (1.25 to 1.68)</td>
</tr>
<tr>
<td>Infection present at death† (%)</td>
<td>155/474 (32.6%)</td>
<td>251/699 (35.9%)</td>
<td>347/935 (37.1%)</td>
<td>1.15 (0.90 to 1.48)</td>
<td>0.257</td>
<td>1.08 (0.82 to 1.40)</td>
</tr>
</tbody>
</table>

Reference group: the youngest age group is always the point of comparison (youngest vs medium; youngest vs oldest and medium vs oldest).

*Statistically significant (p<0.05).

†New question in data collection tool (denominators are lower).

ICU, intensive care unit.

No patient selection bias was present, as the cohort consisted of all patients who died during hospital surgery and survived. Moreover, there were no frailty assessment scores or other methodological differences in the surgical subspecialties. These findings show that the variable ‘care was given or withheld’ in the analysis by the surgeons, especially in the presence of malignancy.

The collection of comorbidities prevalent in each surgical subspecialty was not included in the analysis. The variable ‘care was given or withheld’ in the analysis by the surgeons, especially in the presence of malignancy.

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CONCLUSION
This national mortality audit data uniquely identified complex relationships between advancing age and post-operative care in adults who died following surgery. We showed strong decreases in resource use in the oldest group. We demonstrated a decrease in levels of post-operative care in patients 80 years or older. These findings may indicate a willingness to offer an operation on presentation, but early withdrawal of treatment if complications occur—rather than initial instigation of palliative care. As a result, surgical costs may increase at a lower rate than expected if older people continue to have fewer complex postoperative interventions.

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