# BMJ Open <br> The contribution of primary prevention medication and dietary change in coronary mortality reduction in England between 2000 and 2007: a modelling study 

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## ABSTRACT

Objective: To analyse the falls in coronary heart disease (CHD) mortality in England between 2000 and 2007 and quantify the relative contributions from preventive medications and population-wide changes in blood pressure (BP) and cholesterol levels, particularly by exploring socioeconomic inequalities.
Design: A modelling study.
Setting: Sources of data included controlled trials and meta-analyses, national surveys and official statistics.
Participants: English population aged $25+$ in 20002007.

Main outcome measures: Number of deaths prevented or postponed (DPPs) in 2007 by socioeconomic status. We used the IMPACT ${ }_{\text {sEc }}$ model which applies the relative risk reduction quantified in previous randomised controlled trials and metaanalyses to partition the mortality reduction among specific treatments and risk factor changes.
Results: Between 2000 and 2007, approximately 20400 DPPs were attributable to reductions in BP and cholesterol in the English population. The substantial decline in BP was responsible for approximately 13000 DPPs. Approximately 1800 DPPs came from medications and some 11200 DPPs from populationwide changes. Reduction in population BP prevented almost twoofold more deaths in the most deprived quintile compared with the most affluent. Reduction in cholesterol resulted in approximately 7400 DPPs; approximately 5300 DPPs were attributable to statin use and approximately 2100 DPPs to population-wide changes. Statins prevented almost $50 \%$ more deaths in the most affluent quintile compared with the most deprived. Conversely, population-wide changes in cholesterol prevented threefold more deaths in the most deprived quintile compared with the most affluent.
Conclusions: Population-wide secular changes in systolic blood pressure (SBP) and cholesterol levels helped to substantially reduce CHD mortality and the associated socioeconomic disparities. Mortality reductions were, in absolute terms, greatest in the most deprived quintiles, mainly reflecting their bigger

## Strengths and limitations of this study

- This is the first IMPACT model to quantify the contributions of population risk factors and primary prevention treatments to recent changes in CHD mortality rates by socioeconomic quintiles.
- The datasets used for the model are representative of the English population and used deprivation scores for area of residence as an acceptable proxy indicator for socioeconomic status.
- Unlike the previous IMPACTSEC models, our study stratifies the analysis and results by gender. This allowed us to gain valuable new insights, for example the change in uptake levels for women in the least deprived quintile was almost as effective as the population-wide changes in SBP and cholesterol.
- We assumed that changes in the risk factors and treatment uptakes have equal effect across socioeconomic groups.
- The model was not able to explain around $14 \%$ of the total CHD mortality fall. One possible contributor might be the exclusion of other 'upstream' cardiovascular risk factors, which might affect SES groups differentially.
initial burden of disease. Statins for high-risk individuals also made an important contribution but maintained socioeconomic inequalities. Our results strengthen the case for greater emphasis on preventive approaches, particularly population-based policies to reduce SBP and cholesterol.


## INTRODUCTION

The UK, as many other industrialised countries, has experienced a remarkable $60 \%$ reduction in coronary heart disease (CHD) mortality since the 1970s. However, CHD remains the leading cause of premature death. ${ }^{1}$

Approximately one-third of this initial CHD mortality reduction was attributable to treatments, and two-thirds to reductions in major risk factors. The biggest contributions came from a large decline in smoking prevalence since the 1960 s and more recent reductions in blood pressure and cholesterol. ${ }^{2}{ }^{3}$

The CHD mortality declines have demonstrated a changing relationship with socioeconomic status (SES). ${ }^{4-6}$ Initially it demonstrated a positive relationship with SES (ie, with affluence). ${ }^{7}$ However, this has now reversed in more recent studies in the UK, US, New Zealand, Australia and Scandinavia. ${ }^{8-10}$

Risk factors have also demonstrated strong socioeconomic patterning. Substantial positive associations between lower SES and higher smoking prevalence and higher blood pressure levels have been reported in several studies. ${ }^{11-13}$ However, for cholesterol, the evidence has been less dramatic, with a higher intake of saturated fats among the more deprived populations reported in most studies, ${ }^{14-16}$ but not all. ${ }^{17-19}$ Socioeconomic differences in both risk factors may thus explain some of the CHD mortality gradients. Thus, any attempt to reduce the CHD burden and tackle the associated socioeconomic inequalities should explicitly consider these major risk factors. ${ }^{20}$

Primary prevention medications to lower blood pressure and cholesterol therefore, have been a standard UK health policy for almost two decades. However, while their quantitative benefits to whole populations are accepted, their potential contributions to reduce inequalities are less clear. ${ }^{79}$ 21-25

The aim of this study was, therefore, to analyse the recent falls in CHD mortality and quantify the relative contributions from preventive medications and from population-wide changes in blood pressure and cholesterol levels, particularly exploring the potential effects on different socioeconomic groups.

## METHODS

We used an extended version of the well-known IMPACT model to estimate the contributions of population-level risk factor changes and changes in treatment uptake on the CHD mortality decline in England between 2000 and 2007 for adults aged 25 and above, for two major risk factors-blood pressure and cholesterol. ${ }^{10}$

The IMPACT model applies the relative risk reduction quantified in previous randomised controlled trials (RCT) and meta-analyses to estimate the mortality reduction attributable to (1) temporal change in risk factor prevalence and (2) net change over the period in the uptake of specific treatments in patients with each specific form of CHD. This previously validated deterministic cell-based model has been described in detail elsewhere. ${ }^{21} 26$

The extended version IMPACT SEC model $^{2}$ includes all the major CHD risk factors: smoking, systolic blood pressure (SBP), total cholesterol, body mass index (BMI),
diabetes, physical inactivity and fruit and vegetable consumption. It also includes 45 medical and surgical treatments employed in nine different patient groups. Additionally, the model allows exploring the variation in CHD mortality trends by socioeconomic circumstances. Model inputs and outputs are stratified by the Index of Multiple Deprivation (IMD) quintiles as a proxy indicator of SES. ${ }^{14}$

Our primary outcome measure was the mortality fall or more specifically, the total number of deaths prevented or postponed (DPPs), for each deprivation quintile, that can be attributed to either population-level risk factor changes in SBP and cholesterol, or changes in the uptake of antihypertensive and dyslipidaemia treatments. The DPPs in 2007 relative to 2000 are defined as the difference between the number of CHD expected deaths in 2007 (had age, sex and SES quintile-specific CHD mortality rates in 2000 remained unchanged) and the observed figures.

To calculate the expected number of CHD deaths in 2007, we multiplied the age-sex-IMD quintile specific mortality rates from CHD in 2000 by the population counts for 2007 in that age-sex-IMD quintile stratum. Summing over all strata then yielded the expected number of deaths in 2007 had mortality rates that remained unchanged. Population counts, CHD mortality rates and observed number of deaths used in this step, along with sources, are enlisted in sections 3.1 and 3.2 of the online supplementary technical appendix.

The first part of the IMPACT SEC model calculates the net benefit of statins and antihypertensive treatment in 2007. First, we calculated the expected number of DPPs if statin and antihypertensive uptake rates in 2000 remained constant by multiplying the 2000 age-sex-IMD quintile specific treatment uptake levels by the population counts for 2000 in that age-sex-IMD quintile stratum, the 1 -year case fatality rate and the relative reduction in the case fatality rate as a result of the administered treatment. We did the same for the expected number of DPPs in 2007 by using the 2007 age-sex-IMD quintile specific treatment uptake levels. The difference between the expected number of DPPs (ie, using the treatment uptake rates in 2000) and the estimated number DPPs (ie, using the 2007 uptake rates) is the net benefit of treatments in 2007.

The uptake levels for antihypertensives and statins were defined as the prevalence of never having had angina or heart attack, and currently taking medication specifically prescribed to treat high-blood pressure or lipid-lowering treatment. Treatment uptake values, estimates of treatment efficacy (relative risk reductions) and age-sex specific case fatality rates, along with their sources, are presented in sections 3.3-3.6 of the online supplementary technical appendix.

The second part of the IMPACT SEC model estimates the number of DPPs related to changes in SBP and cholesterol levels in the population. To calculate DPPs from changes in risk factors we used the regression approach,
where the number of CHD deaths in 2000 were multiplied by the absolute change in risk factor level (absolute difference in the risk factors levels between 2000 and 2007) and by a regression $\beta$-coefficient quantifying the estimated relative change in CHD mortality that would result from a one unit change in risk factor level. Risk factors mean levels and $\beta$-coefficients, along with their sources, are presented in sections 3.7-3.9 of the online supplementary technical appendix.

Recent reductions in CHD mortality have been the result of simultaneous change in multiple risk factors. Hence, part of the effect of one risk factor may be mediated through another. In this regard, we used a cumulative risk reduction adjustment factor (AF) to adjust downwards the DPPs attributed to multiple risk factors acting additively or separately, more details can be found in section 2.5 of the online supplementary technical appendix.

Also we considered that some overlap between pharmacological and non-pharmacological contributions to risk factor DPPs occur. Therefore, to estimate the impact of population-wide reduction in total cholesterol due to non-pharmacological change only, we subtracted the estimated effect of cholesterol-lowering treatments uptake levels change from the overall number of DPPs due to change in mean total cholesterol. A similar procedure was carried out for SBP and antihypertensive treatments. For more details see section 2.6 of the online supplementary technical appendix.

Finally, we implemented sensitivity analysis using the EXCEL add-in Ersatz software which allows Monte Carlo simulation. This allows us to calculate $95 \%$ uncertainty intervals ( $95 \%$ UI) for all outputs, based on 5000 draws from specified probabilistic distributions for the model input variables. The probabilistic distributions and their parameters used for the each of the input variables can be found in section 2.8 of the online supplementary technical appendix.

More details on the methodology and worked examples can be found in the online supplementary technical appendix.

## RESULTS

## SBP and cholesterol population levels

Figure 1 depicts the trends in population SBP and cholesterol levels between 2000 and 2007, stratified by IMD quintiles and sex. SBP fell substantially between 2000 and 2007, by an average of 5.4 mm Hg in women and by 2.5 mm Hg in men. Total cholesterol also fell substantially (by approximately $0.20 \mathrm{mmol} / \mathrm{L}$ ), but equally in men and women.

There was no evidence of a social gradient, since the population factors levels were similar across IMD quintiles with no statistically significant difference between them.

## Antihypertensive and statin treatment uptakes

Figure 2 depicts treatments uptakes between 2000 and 2007: there was a substantial increase in both treatment
uptakes, especially statins. Uptake levels of antihypertensive treatments and statins were remarkably equitable across quintiles for men and women, with no statistically significant differences between them.

## Deaths prevented or postponed

There were approximately 38000 fewer CHD deaths in 2007 than if 2000 mortality rates had persisted and been applied to 2007 population estimates for England. Our model was able to explain approximately 32800 ( $86.3 \%$ ) of these fewer deaths (see table 1). Approximately 7100 ( $95 \%$ UI, $3500-14200$ ) fewer deaths ( $19 \%$ of the total mortality reduction) were attributed to increases in the uptake levels of treatments for high-blood pressure and raised cholesterol. Approximately 13300 (8500-17 400) DPPs ( $35 \%$ of the mortality reduction) were attributed to population falls in blood pressure and cholesterol in asymptomatic individuals after subtracting the estimated effect of increases in treatment uptakes. The remaining $32 \%$ of the deaths prevented or postponed in our model were attributed to other risk factors and treatments.

Figure 3 shows the number of deaths prevented or postponed from changes in the population mean levels of SBP and cholesterol (figure 3A, left panel) and from changes in the treatments uptake levels (figure 3B, right panel). We can highlight some key aspects:
(1) Population falls in SBP and cholesterol resulted in more DPPs than increases in uptake levels changes of antihypertensives and statins; (2) Most of the mortality reduction through population changes reflected falls in SBP rather than in cholesterol; (3) By contrast, most of the effect of treatment uptake levels changes was through increments in the uptake levels in statin use rather than antihypertensive use, reflecting the larger increase in statins use during the period of study (e.g., statin uptake rate in 2000 was around $1 \%$ compared to $12 \%$ in 2007); (4) Substantial numbers of DPPs were observed in all social class groups; (5) The absolute effect of population changes on DPPs was larger among persons residing in the most deprived quintiles; and (6) by contrast, the number of DPPs attributable to increases in treatment uptake levels was remarkably equitable across SES groups. However, statin uptakes apparently postponed or prevented slightly more deaths in the most affluent quintile than in the most deprived quintile (figure 3B).

## Systolic blood pressure

Overall, SBP falls between 2000 and 2007 prevented or postponed approximately 13000 ( $8100-17500$ ) deaths ( $34.2 \%$ of the total mortality reduction). Approximately 1800 (700-3900) of those were attributable to antihypertensive treatments ( $4.7 \%$ of the total mortality reduction) and some 11200 DPPs (6500-15 100), over sixfold more, were attributable to population-wide SBP changes (29.5\% of the total mortality reduction). Substantially more DPPs through population-wide changes occurred in the most deprived quintile: 2400 (1600-3100)


Figure 1 Mean values of systolic blood pressure (SBP) and cholesterol (with 95\% UI) between 2000 and 2007 for England stratified by deprivation quintiles and sex.
compared with the most affluent quintiles: 1800 (1000-2600). Thus population-wide changes apparently helped to reduce inequalities in absolute terms.

Anti-hypertensive uptake levels for Men


Statins uptake levels for Men


Conversely, changes in treatment uptake levels demonstrated the opposite effect, since more deaths were prevented in the most affluent quintile (360 DPPs)

## Anti-hypertensive uptake levels for Women



Statins uptake levels for Women


Figure 2 Uptake levels and proportion change in treatment uptake (with $95 \%$ UI) between 2000 and 2007 for England stratified by deprivation quintiles and sex.

Table 1 CHD deaths prevented or postponed between 2000 and 2007 in England, stratified by deprivation quintiles
Deaths prevented or postponed (DPP)

|  | England | IMD quintile 1 <br> affluent | IMD quintile 2 | IMD quintile 3 | IMD quintile 4 | IMD quintile 5 <br> deprived |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Mean | 32770 | 5775 | 6745 | 7015 | 6870 | 6370 |
| $95 \%$ LL | 25990 | 4430 | 5320 | 5420 | 5400 | 5100 |
| $95 \%$ UL | 41550 | 7705 | 8515 | 9360 | 8765 | 7830 |

compared to the most deprived (280 DPPs). However, in both cases, SES differences were not statistically significant. Detailed outputs with uncertainty intervals can be found in section 4 of the online supplementary technical appendix.

## Cholesterol

Overall, cholesterol falls between 2000 and 2007 resulted in approximately 7400 (3900-14500) fewer deaths ( $19.5 \%$ of the total mortality reduction). This total comprised some 5300 (2100-12 300) fewer deaths ( $13.9 \%$ of the total mortality reduction) attributable to statin medications and approximately 2100 (1000-3200) fewer deaths ( $5.5 \%$ of the total mortality reduction) attributable to population-wide falls in cholesterol. Statin medications prevented some $1100(400-2700)$ deaths in the most affluent quintile compared to approximately 800 (300-1900) DPPs in the most deprived quintile. Conversely, population changes in cholesterol resulted in approximately 700 (500-1000) DPPs in the most deprived quintile and some 200 (40-400) DPPs in the
most affluent quintile. However, like SBP, there was no clear SES gradient. Section 4 of the online supplementary technical appendix provides detailed outputs with uncertainty intervals.

## Gender differences

Figure 4 shows the number of deaths prevented or postponed in men and women from falls in the population mean levels of SBP and cholesterol (figure 4A, left panels) and from increases in the treatment uptake levels (figure 4B, right panels). For men, although most of the mortality reduction came from population falls in SBP, cholesterol reductions also had a considerable larger effect in reducing mortality compared to women (four times higher). By contrast, the number of DPPs due to increases in treatment uptakes in men appeared remarkably equitable across SES groups.

For women, the impressive reduction in SBP mean level between 2000 and 2007 contributed the most to the total mortality reduction and in all quintiles, whereas population level reductions of cholesterol had a

A


| DPPs | England | IMD <br> $\mathbf{1}$ | IMD <br> $\mathbf{2}$ | IMD <br> $\mathbf{3}$ | IMD <br> $\mathbf{4}$ | IMD <br> $\mathbf{5}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Total | 13253 | 2093 | 2666 | 2608 | 2742 | 3143 |
| SBP | 11162 | 1861 | 2168 | 2321 | 2391 | 2421 |
| Cholesterol | 2091 | 232 | 498 | 287 | 351 | 722 |

B DPPS through changes in the treatments uptakes
$\square$ SBP Cholesterol


| DPPs | England | IMD <br> $\mathbf{1}$ | IMD <br> $\mathbf{2}$ | IMD <br> $\mathbf{3}$ | IMD <br> $\mathbf{4}$ | IMD <br> $\mathbf{5}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Total | 7098 | 1411 | 1386 | 1713 | 1539 | 1049 |
| SBP | 1798 | 357 | 411 | 408 | 345 | 277 |
| Cholesterol | 5300 | 1054 | 975 | 1305 | 1194 | 772 |

Figure 3 Number of deaths prevented or postponed (DPPs) between 2000 and 2007 in England, attributable to changes in the population in SBP and cholesterol (A, left panel), changes in uptakes levels for antihypertensive treatments and statins (B, right panel); stratified by deprivation quintiles


Figure 4 ( $A$ and $B$ ) Number of deaths prevented or postponed (DPPs) from changes in the population in systolic blood pressure (SBP) and cholesterol, changes in uptakes levels for antihypertension and statins between 2000 and 2007 in England, stratified by deprivation quintiles and sex.
smaller benefit. Moreover, the joint benefit of increasing treatment uptakes (antihypertensive and statins) in women appeared to have an important effect: for example, in the most affluent quintile (IMDQ1) the reduction in DPPs due to the increase in treatment uptakes for women was almost as effective as the population-wide falls in both sexes for that quintile.

However, in terms of differences between men and women, the results of the uncertainty analysis suggest that these are not significant in statistical terms. More detailed outputs split by gender can be found in section 5 of the online supplementary technical appendix.

## DISCUSSION

CHD mortality in England fell by a remarkable 34\% between 2000 and 2007. This represents an impressive 38000 fewer deaths from CHD in 2007 than if the 2000 mortality rates had persisted. Reductions in major cardiovascular risk factors of blood pressure and cholesterol explained for almost two-thirds of this large mortality fall.

## Blood pressure trends

Declines in the population blood pressure level made the largest contribution to the overall fall in CHD mortality. In contrast, antihypertensive treatments produced only modest benefits. First, because the baseline CHD event rate was low in asymptomatic individuals ( $\leq 1 \%$ per year) yielding only a small reduction of the attributable risk during the period of study. ${ }^{27}$ Second, treatment efficacy is low; and third, blood pressure control is still poor (adherence levels to medication are around $60 \%$ ), ${ }^{7}$ leading in conjunction to a substantial residual risk. ${ }^{21} 28$

## Cholesterol trends

Population-wide falls in cholesterol levels averted more deaths in the most deprived quintiles, reflecting similar absolute falls but much higher baseline mortality rates. The increase in the uptake of statins between 2000 and 2007 made an even greater contribution to the overall mortality fall: twofold greater than the change in population cholesterol ( $16 \%$ vs $6 \%$ ), and with equitable benefits across all five SES groups.

## Comparisons with other studies

Our results are consistent with previous analyses in the UK and around the world, supporting the importance of this study beyond England. Using the IMPACT model to examine contributions to the overall reductions in CHD mortality in England and Wales population between 1981 and 2000, Unal, Critchley ${ }^{3}$ reported a higher contribution from blood pressure changes (compared to cholesterol). Some $76 \%$ of this contribution was attributable to population-wide changes rather than antihypertensive medications. IMPACT analyses carried out in the USA and Irish populations between 19802000 and 1985-2000 likewise observed substantially greater benefits attributable to secular changes in risk factors rather than treatments. ${ }^{26} 27$

The analysis by DeWilde, Carey ${ }^{29}$ suggested that reported blood pressure treatments were responsible for $25 \%$ of 5 mm Hg reduction in SBP during the period 1994-2009 for England.

Emberson et al $e^{\beta 0}$ applied a very different methodology using evidence from randomised control trials and cohort studies to analyse the effectiveness of populationwide changes in risk factor levels against the high-risk individual approach. Their findings were entirely consistent with ours. They concluded that a mere $10 \%$ reduction in population-wide blood pressure and cholesterol levels might achieve a $45 \%$ reduction in cardiac events in the long term. However, it would be needed to provide treatment to approximately $26 \%$ of the UK population at high risk to achieve only a $34 \%$ reduction in cardiac events. The US CHD policy model likewise reported that population-wide reductions of salt intake $(3 \mathrm{~g} /$ day $)$ might prevent between 44000 and 90000 deaths. ${ }^{31}$

## Strengths and limitations

This is the first IMPACT model to quantify the contributions of population risk factors and primary prevention treatments to recent changes in CHD mortality rates by socioeconomic quintiles.

The data sets used for the model are representative of the English population and used deprivation scores for area of residence as an acceptable proxy indicator for socioeconomic status. This allowed a sufficient sample size to quantify the effect of risk factor modification through changes in population-wide risk factor levels and treatment uptakes.

Unlike, the previous IMPACT ${ }_{\text {SEC }}$ models (Bajekal et al ${ }^{2}$ and Scholes et $a l^{22}$ ), our study stratifies the analysis and results by gender. This allowed us to gain valuable new insights. For example, changes in SBP and cholesterol population levels for women led to the highest number of DPPs for all quintiles. More surprisingly, the change in uptake levels for women in the least deprived quintile was almost as effective as the population-wide changes in SBP and cholesterol. This suggests that any attempt to tackle the socioeconomic inequalities in CHD mortality should explicitly consider these gender differences.

However, our study has limitations that should also be acknowledged. First, the area-level categorisation may not be representative of individual circumstances. Furthermore, observed SES differences in CHD mortality might reflect not material deprivation but other confounding and mediator factors such as alcohol consumption, obesity or ethnicity. However, the IMD is a comprehensive multidimensional construct of socioeconomic status made up of seven domains, and based on small geographical areas (less than 1500 residents) called lower level super output areas (LSOAs). The advantage of using LSOAs is that their smaller geographical size also allows for a more detailed knowledge of deprived areas.

Our risk factor effect data might still have some residual confounding. Statins and antihypertensive medication data are from the surveys; therefore, some misclassification bias might be present.

We assumed that treatments and lifestyle changes have an immediate effect on CHD mortality, which might not be entirely true. However, Capewell and O'Flaherty ${ }^{23} 32$ pointed out evidence from clinical trials and policy interventions which consistently suggest that changes in diet and lifestyle across entire populations can be rapidly followed by dramatic declines in mortality.

We assumed that changes in the risk factors and treatment uptakes have equal effect across socioeconomic groups. However, the benefits of falls in risk factors or increases in treatment uptakes may be higher in more affluent groups. ${ }^{2}$ This may partly explain the faster rates of CHD mortality decline in the most affluent quintiles as Bajekal et al ${ }^{10}$ pointed out. Likewise, we assumed that the relative risk reduction due to treatments remained constant from 2000 to 2007.

We simply subtracted the mortality gains from increasing uptake levels of statins from the overall gains due to reductions in total cholesterol to estimate the impact of population-wide reduction in total cholesterol due to non-pharmacological change only. This mutually exclusive adjudication of cause adjustment might overestimate medication benefit.

Given the background of higher mortality and morbidity in the more deprived quintiles, DPPs might overestimate the actual health gain, as we do not know the additional life span gained by preventing a specific death at a specific time. This might result in a lesser reduction in inequalities than DPPs alone would suggest.

Finally, our model was not able to explain around $14 \%$ of the total CHD mortality fall between 2000 and 2007. One possible contributor might be the exclusion of other 'upstream' cardiovascular risk factors, which might affect SES groups differentially, for example, psychosocial stress. ${ }^{33}$

## Implications for public health and clinical care

This study shows that population-wide secular falls in blood pressure and cholesterol have substantially helped to decrease CHD mortality and reduce the associated
socioeconomic disparities in absolute terms. Furthermore, as we discussed earlier, there is an increasing body of evidence to support the use of population-wide approaches to reduce CHD risk factors. Mackenbach et al ${ }^{\beta^{4}}$ recently evaluated 22 successful preventive interventions in the Netherlands. Approximately $75 \%$ of the health gains during the period 1970-2010 were achieved by a population approach and just $25 \%$ by a high-risk individual approach.

In the UK, the population-wide fall in blood pressure is consistent with the recent successful implementation of policies to reduce salt intake. Similar trends have been reported in other developed countries. ${ }^{21}{ }^{28}$ There are also several international examples where policy interventions have proven to be effective at achieving significant reductions in saturated fats, trans-fats and calories in processed foods and takeaway meals. ${ }^{24} 3135 \quad 36$ However, policies to reduce saturated fats and trans-fats have so far been neglected in the UK. ${ }^{25}$

Conversely, targeting high-risk individuals with medication appears less effective and may also widen socioeconomic inequalities in CHD mortality. ${ }^{37}{ }^{38}$ Any intervention that requires people to mobilise their own resources (material and psychological) will understandably favour those who have greater resources ${ }^{37}$ and thus, widen social inequalities. Thus, those with the poorest health will benefit the least from such interventions. ${ }^{38}$

However, there is no simple choice between either population-based or high-risk strategies to reduce CHD mortality. The approaches are complementary in delivering the greatest public health benefit. ${ }^{39} 40$ It is, however, clear that individual-based treatment strategies can afford only modest reductions in mortality compared with addressing risk factors population wide.

Severely limited healthcare budgets are now forcing planning systems to consider how best to allocate future resources. Our results strengthen the case for greater emphasis on preventive approaches, particularly population-based policies to reduce blood pressure and cholesterol. Such strategies might be more powerful, rapid, cost-effective and equitable than additional preventive medications. ${ }^{25}$

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Contributors MGC drafted the manuscript, analysed the results and conducted the uncertainty analysis in collaboration with SC, MOF and RA. RA conducted the initial literature review and initial set up of the model in collaboration with SC. NH, SS, EW and JL contributed to the interpretation of the results and to the drafting and finalisation of the manuscript.

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## Correction

Guzman-Castillo M, Ahmed R, Hawkins N, et al. The contribution of primary prevention medication and dietary change in coronary mortality reduction in England between 2000 and 2007: a modelling study. BMJ Open 2015;5:e006070.
(1) The Acknowledgement section should read:

We would like to thank Rosalind Raine and Madhavi Bajekal for their work in developing the original IMPACTSEC model. The English IMPACTsec project was conducted by Madhavi Bajekal (MB) and Shaun Scholes (SS). Hande Love set up the worksheet template; SS populated the model and was its custodian. MB ensured the integrity of inputs and outputs and provided SEC-related methodological solutions. Martin O'Flaherty \& Nat Hawkins provided support, clinical expertise and generated the therapeutic input. The UCL team was led by Rosalind Raine. Simon Capewell co-ordinated the overall project.
(2) The design section should read:

Design: Retrospective analysis using the IMPACTsec policy model.
(3) 'Strengths and limitations' and 'Summary' sections should read:

IMPACTsec is the first model to quantify the contributions of population risk factors and primary prevention treatments to recent changes in CHD mortality rates by socioeconomic quintiles.
(4) Appendix/References section has been replaced online to include a version that references Bajekal; no other changes have been made.


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TECHNICAL APPENDIX FOR THE IMPACT $_{\text {SEC }}$ MODEL
Contents
1 Overview of the IMPACT SEC model ..... 3
2 METHOD AND EXAMPLES OF DEATHS PREVENTED OR POSTPONED (DPP) CALCULATIONS ..... 4
2.1. Changes in mortality rates from CHD, England 2000 to 2007 ..... 4
2.2. Expected and observed number of deaths from CHD ..... 4
2.3. Treatment component of IMPACT SEC model ..... 4
2.4. Risk factor component of IMPACT SEC model ..... 6
2.5. Cumulative risk-reduction ..... 7
2.5.1. Background ..... 7
2.5.2. 1.3.2 Implementation ..... 7
2.5.3. Calculating aggregate change in risk factors over 2000 and 2007 ..... 8
2.5.4. Regression models to estimate risk factor change, 2000-2007 ..... 8
2.5.5. Adjustment factors by age-sex-IMD ..... 9
2.6. Overlap between pharmacological and non-pharmacological contributions to risk factor DPPs 10
2.7. Net effects ..... 10
2.8. Uncertainty analyses ..... 11
2.8.1. Allocating areas to socioeconomic quintiles using the Index of Multiple Deprivation, 2007 ..... 11
3 Data sources ..... 12
3.1. Population and patient data sources used in the IMPACTSEC model ..... 12
3.2. Demographic data 2000 and 2007 by sex and deprivation quintiles ..... 13
3.3. Data sources for treatment uptake levels ..... 14
3.4. Treatment uptake in 2000 and 2007 ..... 15
3.5. Clinical efficacy of interventions: relative risk reductions obtained from meta-analyses, and randomised clinical trials ..... 16
3.6. Case fatality rates for each patient group ..... 17
3.7. Risk factors - variable definitions and source ..... 18
3.8. Risk factor levels in 2000 and 2007 by sex and deprivation quintiles ..... 19
3.9. Beta coefficients for major risk factors ..... 20
3.10. Cumulative benefit: Adjustment factors by age, sex and IMD quintile ..... 22
3.11. Uncertainty analysis: parameter distributions, functions and sources ..... 23
4 Tables ..... 26
5 Tables by gender ..... 28
5.1. Men ..... 28
5.2. Women ..... 31
5.3. Percentage difference in men relative to women ..................................................... 33

DPPS through changes in the population ................................................................................... 33
6 Reference List.................................................................................................................... 37

## 1 Overview of the IMPACT SEC model

The IMPACT model accommodates sub-national variation in CHD mortality trends by socioeconomic circumstances (IMPACT SEC model). We used the Index of Multiple Deprivation 2007 (IMD) quintiles as a proxy indicator of socioeconomic circumstances. This model examines the effects of changes in treatment uptake and risk factor trends on changes in mortality from coronary heart disease (CHD) among adults in England aged 25 years and over, stratified into equal quintiles by population size. The tables included in this Technical Appendix provide details about the sources and methods that were used.

## 2 METHOD AND EXAMPLES OF DEATHS PREVENTED OR POSTPONED (DPP) CALCULATIONS

### 2.1. Changes in mortality rates from CHD, England 2000 to 2007

Data sources used in examining the changes in CHD mortality rates over 2000 to 2007 are shown in Table A. Mortality rates from CHD were calculated using the underlying cause of death (2000: ICD9 410-414; 2007: ICD10 I20-I25). Both unadjusted and age-adjusted mortality rates were calculated. The direct method of age-standardisation was used with the European Union reference population as standard.

### 2.2. Expected and observed number of deaths from CHD

Data sources used to estimate the observed and expected number of deaths from CHD for 2000 and 2007 are shown in Table A. The expected number of CHD deaths in 2007 was calculated by multiplying the age-sex-IMD quintile specific mortality rates from CHD in 2000 by the population counts for 2007 in that age-sex-IMD quintile stratum. Summing over all strata then yielded the expected number of deaths in 2007 had mortality rates remained unchanged. The difference between the number of expected and observed deaths from CHD represented the mortality fall, or the total DPPs in 2007 relative to 2000 . Population counts, CHD mortality rates, observed and expected numbers of deaths are shown in sections 3.1and 3.2

### 2.3. Treatment component of IMPACT SEC model

The treatment component of the IMPACT SEC model included nine mutually exclusive CHD patient groups (see below). However, for the purposes of our model, we just take into account groups 8 and 9

1. Patients treated in hospital for acute myocardial infarction (ST-elevation myocardial infarction and non-ST elevation acute coronary syndrome)
2. Patients admitted to hospital with unstable angina
3. Community-dwelling patients who have survived a myocardial infarction for over a year
4. Patients who have undergone a revascularisation procedure up to and including the years 2000 and 2007: Coronary Artery Bypass Grafting (CABG), or a Percutaneous Coronary Intervention (PCI)
5. Community-dwelling patients with stable coronary artery disease
6. Patients admitted to hospital with heart failure (due to CHD)
7. Community-dwelling patients with heart failure (due to CHD)
8. Hypercholesterolaemic subjects without CHD eligible for cholesterol lowering therapy such as statins
9. Hypertensive individuals without CHD eligible for anti-hypertensive therapy

The general approach to calculating the number of DPPs from an intervention among a particular patient group was first to stratify by age, sex and IMD quintile; then to multiply the estimated number of patients in 2007 in turn by: the proportion of these patients receiving a particular treatment; the one-year case fatality rate; and the relative reduction in the case fatality rate due to the administered treatment. Sources for treatment uptake are shown in sections 3.3 and 3.4. Sources for estimates of treatment efficacy (relative risk reductions) are shown in section 3.5. We obtained the relative risks based on the most recent published systematic reviews and metaanalyses of epidemiological studies. Each treatment relative risk value in the model was based on a meta-analysis comparison with an older therapy, or in some cases with a placebo if relevant. Age-sex specific case fatality rates for each patient group are presented in section 3.6

It was assumed that compliance (adherence), i.e. the proportion of treated patients actually taking therapeutically effective levels of medication, was $100 \%$ among hospital patients, $70 \%$ among symptomatic community patients, and $50 \%$ among asymptomatic community patients taking lipid-lowering drugs or anti-hypertensive medication for primary prevention. An adjustment was also made in certain cases for sub-optimal dose.

## Example 1: Estimation of DPPs from a specific treatment

## Mortality fall as a result of taking statins in men aged 55-64 in the most affluent quintile

For example, in 2007 , about 685,000 men aged $55-64$ were classified as the most affluent quintile. Uptake of statins in primary prevention was estimated to be approximately $15 \%$ with $100 \%$ assumed to comply. Statins in primary prevention reduces case fatality in patients by approximately $35 \%$. The underlying one-year case fatality rate in these men was approximately $0.6 \%$. The DPPs for at least a year were therefore calculated as:

## Patient numbers $\times$ treatment uptake $\times$ compliance $\times$ relative mortality reduction $\times$ one year case fatality

$=685,000 \times 15 \% \times 50 \% \times 35 \% \times 0.6 \% \approx 108$ DPPs

This calculation was then repeated for each age-sex-IMD quintile group.

### 2.4. Risk factor component of IMPACT SEC model

The second part of the IMPACT SEC model estimated the number of DPPs related to changes in cardiovascular risk factor levels in the population. The risk factors considered were total cholesterol and systolic blood pressure. The Health Survey for England was used to calculate trends in the prevalence (or mean values) of each risk factor (section 3.7). For the purposes of this paper, we used the regression approach to calculate DPPs from changes in risk factors.

In this approach regression approach the number of CHD deaths in 2000 (the start year) after adjusting for population change between 2000 and 2007 were multiplied by the absolute change in risk factor level, and by a regression coefficient ('beta') quantifying the estimated relative change in CHD mortality that would result from a one-unit change in risk factor level (see section 3.9). Natural logarithms were used, as is conventional, in order to best describe the log-linear relationship between absolute changes in risk factor levels and relative change in mortality. Levels of risk factors in 2000 and 2007 by sex and IMD quintile are shown in section 3.8.

## Example 2: Estimation of DPPs from risk factor changes using regression method

Mortality fall due to reduction in SBP in women aged 55-64 in the most affluent quintile
For example, in 2000, there were 227 CHD deaths among 573,291 women aged 55-64 years in the most affluent quintile. The population total had increased to 714,111 in 2007. Applying the CHD death rate from $2000(39.6$ per 100,000$)$ to the 2007 population gives an (adjusted) total of 283 expected deaths in 2007.

Mean SBP in this group fell by an estimated 4.28 millimetres of mercury ( mmHg ) (from 133.8 in 2000 to 129.5 in 2007). The largest meta-analysis reports an estimated age-sex specific reduction in mortality of $50 \%$ for every 20 mmHg reduction in SBP, generating a logarithmic coefficient of -0.035 (i.e. natural logarithm of 0.5 divided by 20). The subsequent reduction in CHD deaths between 2000 and 2007 was then estimated as the product of three variables:

DPPs = expected CHD deaths in 2007 (had mortality rates in 2000 remained constant) $\times$ absolute risk factor reduction between 2000 and $2007 \times$ regression coefficient exponentiated

DPPs $=(1-($ exponential $($ regression coefficient $\times$ absolute change $))) \times$ expected deaths in 2007
DPPs $=(1-($ exponential $(-0.035 \times 4.28))) \times 283 \approx 39$

This calculation was then repeated for each age-sex-quintile group.

The regression coefficients were assumed equal across deprivation quintiles. A 'fixed gradient' approach was used to stabilise estimates of risk factor change across the quintiles; this method is discussed in 2.5.5

### 2.5. Cumulative risk-reduction

### 2.5.1. Background

CHD deaths are usually caused by multiple risk factors acting simultaneously. Hence, part of the effect of one risk factor may be mediated through another. For example, physical inactivity may have a direct effect on CHD but may also partly be mediated through its effects on BMI and blood pressure. It is recommended therefore that mortality benefits attributable to risk factors which may be causally related, or which overlap in population groups, should not be combined by simple addition. Ideally, their effects should instead be jointly estimated [12-16].

We do not currently have sources that allow joint estimation of relative risks for combinations of risk factors in this English population. However, several large cohort studies and meta-analyses have published independent risk reduction coefficients for each risk factor included in this study. One approach commonly used is to calculate the cumulative risk-reduction [17]. This approach accounts for risk factor prevalence overlap but assumes independence of effects [14-15]. The general equation for cumulative risk-reduction is stated as:

Combined $($ or cumulative $)$ effect $(\mathrm{CR})=$
$1-((1-a) \times(1-b) \times(1-c) \times \ldots \times(1-n))$
[1]

Thus for CHD risk factors, the specific equation is stated as:
$\mathbf{C R}=1-\left(\left(1-R_{\text {SBP }}\right) \times\left(1-\mathrm{R}_{\text {smoke }}\right) \times\left(1-\mathrm{R}_{\text {diabetes }}\right) \times \ldots \times\left(1-\mathrm{R}_{\mathrm{n}}\right)\right)$
where R denotes the mortality change attributable to a specific risk factor.

This is in contrast to additive risk-reduction (AR):
$\mathbf{A R}=\left(\mathrm{R}_{\text {SBP }}\right)+\left(\mathrm{R}_{\text {smoke }}\right)+\left(\mathrm{R}_{\text {diabetes }}\right)+\ldots . .+\left(\mathrm{R}_{\mathrm{n}}\right)$

### 2.5.2. 1.3.2 Implementation

For the purposes of this modelling study we first calculated the (additive) DPPs attributed to risk factor change. These were then adjusted down by using the ratio:

## Adjustment factor $=$ CR/AR

The adjustment factor would always be expected to be less than 1. In other words, cumulative risk factor reduction would be smaller than the mortality benefits arrived at by a simple summation of the benefits of each risk factor in turn.

The proportional change in the CHD mortality rate between two time points (denoted by R) was calculated using the following formulas [14-15]:

Continuous risk factors:
$\mathrm{R}_{\text {continuous }}=1-\exp ($ beta $\times$ absolute mean risk factor change $)$
and P denotes prevalence at the start-year; RR the relative risk in CHD mortality associated with risk factor presence; and $\Delta \mathrm{P}$ the change in prevalence between the start and final years.

Formulas [3] and [4] were used to calculate the proportional change in the CHD mortality rate ( R ) for each risk factor and the steps involved in their estimation are detailed below. However, we made two modifications to the methodology used in previous work [14-15]. First, we estimated aggregate change over a seven year period (2000-2007) rather than average annual change. Second, additive and cumulative risk-reduction was calculated by using the absolute values of $R$ (i.e. disregarding the direction of risk factor change). These are discussed in turn below.

### 2.5.3. Calculating aggregate change in risk factors over 2000 and 2007

Previous studies [14-15] estimating cumulative risk factor reduction calculated the average annual percentage change in CHD mortality attributable to annual falls in levels of smoking, blood pressure and cholesterol (where annual falls in CHD mortality and risk factor levels were estimated over a specified number of years). Rather than estimate the average annual change over a specific range of years, we were interested in calculating the R values between two fixed points in time (start and end years of the model), seven years apart, 2000 and 2007. We therefore adapted formulas [3] and [4], substituting change over the seven year study period for the estimation of annual average change. We checked our resulting estimates of cumulative risk reduction calculated over seven years against uprating the annual average by a factor of seven. The two sets of estimates were found to be virtually identical.

### 2.5.4. Regression models to estimate risk factor change, 2000-2007

Formula [3] requires estimates of absolute and relative change in risk factors, respectively. Regression modelling was used to estimate the magnitude of absolute and relative change. In order to smooth fluctuations in Health Survey for England data, we obtained estimates of risk factor change for each risk factor over 2000-2007 by using the predicted values from regression models. Separate models were fitted by sex and seven ten-year age-bands.

The dependent variable was the risk-factor level for each survey respondent; calendar year (i.e. year of interview) was the explanatory variable entered in the model as a continuous term.

Absolute change was measured as the difference between the predicted values for 2000 and 2007, by age and sex.

Estimates of risk factor change were not calculated separately by deprivation quintile owing to small sample sizes, especially in those risk factors covered by the survey in intermittent years. Data since 2003 were weighted for non-response at each stage of data collection. Although it was just beyond the time period covered by the IMPACT SEC model, the most recent survey data available (2008) was included in fitting the regressions to improve estimation of the underlying change. Analyses were conducted using Stata Version 11.1.

### 2.5.5. Adjustment factors by age-sex-IMD

The adjustment factors (section 3.10) fell within the range of 0.83 to 0.96 . The largest adjustment ( 0.83 ) was applied to the DPPs for women aged $65-74$ resident in the most deprived areas (IMDQ5). The adjustment factors for the deprivation quintiles were, on average, $\pm 0.01$ of the overall adjustment ratio for England across the 14 age and sex groups. The adjustments were on average, slightly higher for women ( 0.89 ) than men ( 0.92 ); and were higher in IMDQ5 than in IMDQ1 (mean values 0.8924 and 0.9089 , respectively). Hence the adjustment values indicated a larger downward adjustment to the additive DPPs in the most deprived areas relative to the most affluent.

### 2.6. Overlap between pharmacological and non-pharmacological contributions to risk factor DPPs

Risk factor improvements, such as lower blood pressure or lower total cholesterol, may be achieved through medications, lifestyle changes, or a combination. In order to separate the DPPs from pharmacological versus non-pharmacological contributions to CHD mortality, we subtracted the DPPs calculated in the treatment (primary prevention) component of the model from the DPPs calculated in the risk factor component. That is, to estimate the impact of population-wide reduction in total cholesterol due to non-pharmacological change, we subtracted the estimated effect of statins for the primary prevention of CHD from the overall number of DPPs due to change in mean total cholesterol. Similarly, to estimate the impact of the population-wide reduction in SBP we subtracted the estimated effect of anti-hypertensive medication for primary prevention from the overall number of DPPs due to change in mean SBP levels.

### 2.7. Net effects

As all treatments were in use in 2000, the net benefit of an intervention in 2007 was calculated by subtracting the expected number of deaths prevented if the uptake rates in 2000 remained constant from the estimated number of deaths prevented calculated using the 2007 uptake rates. This is illustrated in the example below.

## Example 5: Net effects for treatments

For example, in 2007, about 685,000 men aged $55-64$ were classified as the most affluent quintile. Uptake of statins in primary prevention was estimated to be approximately $15 \%$ with $50 \%$ assumed to comply. Statins in primary prevention reduces case fatality in patients by approximately $35 \%$. The underlying one-year case fatality rate in these men was approximately $0.6 \%$. The DPPs for at least a year were therefore calculated as:

Patient numbers $\times$ treatment uptake $\times$ compliance $\times$ relative mortality reduction $\times$ one year case fatality
$=685,000 \times 15 \% \times 50 \% \times 35 \% \times 0.6 \% \approx 108 \mathrm{DPPs}$

Applying the uptake rate in 2000 (2.7\%) gave a total of 19 DPPs:

The net DPPs were therefore:

$$
\text { Net DPPs = DPPs using uptake } 2007-\text { DPPs using uptake } 2000
$$

$=108-19=89$

The estimated changes in treatment uptake between 2000 and 2007 by deprivation quintile are shown in Table H .

### 2.8. Uncertainty analyses

We implemented uncertainty analysis in Excel using Ersatz (version 1.0 available at http://www.epigear.com). This is an add-on which allows probabilistic bootstrapping in Excel. Ersatz allows repeated random draws from specified distributions for input variables and then calculates the $95 \%$ uncertainty intervals from the realised values of the output variable (deaths prevented or postponed). For the IMPACT SEC model, we calculated the uncertainty intervals based on 1000 draws - taking the $95 \%$ uncertainty intervals from the $2.5^{\text {th }}$ and $97.5^{\text {th }}$ percentiles. The parameter distributions used for the input variables to the DPP calculations are shown in Table M. Worked examples using Ersatz are shown below Table M.

### 2.8.1. Allocating areas to socioeconomic quintiles using the Index of Multiple Deprivation, 2007

The Index of Multiple Deprivation (IMD) is a composite index of relative deprivation at small area level based on seven domains: income; employment; health deprivation and disability; education, skills and training; barriers to housing and services; crime and disorder; and living environment [19]. The IMD 2007 score of all small areas in England (average population 1,500) were ranked in ascending order and grouped into equal quintiles (about 6,500 areas in each), with quintile one (IMDQ1) including the most affluent and quintile five (IMDQ5) the most deprived areas. Based on their postcode of residence, patients treated in hospital (e.g. recorded in Hospital Episode Statistics) or in the community (e.g. in the General Practice Research Database) were matched via their area of residence to the corresponding deprivation quintile by the data providers to protect patient anonymity. Mortality counts were similarly aggregated into deprivation quintiles by the Office for National Statistics before being released to us for research purposes.

As the IMD 2007 includes rates of premature total mortality in the health deprivation and disability domain, its use to quantify health inequalities risks a tautology. However UK studies have shown that removing the health domain had little effect on either the assignment of areas into their deprivation quintile or the relationship between area-based deprivation and health [20].

Conceptually, the IMD 2007 is a measure of deprivation, not a measure of affluence. Hence, areas with the lowest scores are not necessarily the most affluent; rather they have the lowest concentration of deprived people. In this paper for clarity and to easily distinguish between the extreme ends of the deprivation spectrum, we have used the term 'most affluent' and 'most deprived' rather than 'least deprived' and 'most deprived'.

## 3 Data sources

### 3.1. Population and patient data sources used in the IMPACTSEC model

| Information | Source |
| :---: | :---: |
| Population data |  |
| Population counts and CHD deaths stratified by age, sex, and Index of Multiple Deprivation quintiles | Office for National Statistics (ONS): <br> (2000: ICD9 410-414) <br> (2007: ICD10 I20-I25) |
| Patients eligible for primary prevention therapies: |  |
| Lipid-lowering drugs | Prevalence of never having had angina or heart attack and currently taking lipid lowering drugs prescribed by a doctor from the Health Survey for England (HSfE 1998, 2003, and 2006) (http://www.ic.nhs.uk/statistics-and-data-collections/health-and-lifestyles-related-surveys/health-survey-for-england) |
| Hypertension treatment | Prevalence of never having had angina or heart attack and currently taking medication specifically prescribed to treat high blood pressure from the Health Survey for England (HSfE 1998, 2003, and 2006) |

Table A: Population and patient data sources used in the IMPACTSEC model

### 3.2. Demographic data 2000 and 2007 by sex and deprivation quintiles

|  | Year | England | IMDQ1 | IMDQ2 | IMDQ3 | IMDQ4 | IMDQ5 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Male |  |  |  |  |  |  |  |
| Population (000s) | 2000 | 16242 | 3353 | 3372 | 3321 | 3186 | 3011 |
|  | 2007 | 17002 | 3525 | 3542 | 3486 | 3335 | 3114 |
| Observed CHD deaths | 2000 | 56713 | 9146 | 10868 | 11671 | 12094 | 12934 |
|  | 2007 | 41713 | 6962 | 8129 | 8535 | 8723 | 9364 |
| Age-standardised rate (00,000) | 2000 | 310 | 238 | 270 | 301 | 349 | 415 |
|  | 2007 | 200 | 147 | 170 | 191 | 231 | 294 |
| Annual \% fall ${ }^{\dagger}$ |  | 6.0 | 6.6 | 6.4 | 6.3 | 5.7 | 4.8 |
| Expected deaths ${ }^{\dagger \dagger}$ | 2007 | 63685 | 11207 | 12856 | 13348 | 13098 | 13176 |
| Target DPPs ${ }^{\ddagger}$ | 2007 | 21972 | 4245 | 4727 | 4813 | 4375 | 3812 |
| \% of expected deaths prevented | 2007 | 34.5 | 37.9 | 36.8 | 36.1 | 33.4 | 28.9 |
| Female |  |  |  |  |  |  |  |
| Population (000s) | 2000 | 17710 | 3618 | 3663 | 3618 | 3493 | 3318 |
|  | 2007 | 18279 | 3803 | 3820 | 3747 | 3571 | 3337 |
| Observed CHD deaths | 2000 | 46530 | 7383 | 8959 | 9789 | 10093 | 10306 |
|  | 2007 | 32461 | 5350 | 6315 | 6812 | 6953 | 7031 |
| Age-standardised rate (00,000) | 2000 | 148 | 115 | 128 | 143 | 164 | 198 |
|  | 2007 | 94 | 70 | 79 | 90 | 107 | 136 |
| Annual \% fall ${ }^{\dagger}$ |  | 6.3 | 6.7 | 6.7 | 6.4 | 5.9 | 5.2 |
| Expected deaths ${ }^{\dagger \dagger}$ | 2007 | 48559 | 8458 | 9812 | 10348 | 10162 | 9778 |
| Target DPPs ${ }^{\text { }}$ | 2007 | 16098 | 3108 | 3497 | 3536 | 3209 | 2747 |
| \% of expected deaths prevented | 2007 | 33.2 | 36.7 | 35.6 | 34.2 | 31.6 | 28.1 |
| Total |  |  |  |  |  |  |  |
| Population (000s) | 2000 | 33952 | 6972 | 7035 | 6939 | 6678 | 6329 |
|  | 2007 | 35281 | 7328 | 7363 | 7233 | 6906 | 6451 |
| Observed CHD deaths | 2000 | 103243 | 16529 | 19827 | 21460 | 22187 | 23240 |
|  | 2007 | 74174 | 12312 | 14444 | 15347 | 15676 | 16395 |
| Age-standardised rate $(\mathbf{0 0 , 0 0 0})$ | 2000 | 229 | 177 | 199 | 222 | 257 | 306 |
|  | 2007 | 147 | 109 | 124 | 141 | 169 | 215 |
| Annual \% fall ${ }^{\dagger}$ |  | 6.1 | 6.7 | 6.5 | 6.3 | 5.8 | 4.9 |
| Expected deaths ${ }^{\dagger \dagger}$ | 2007 | 112244 | 19665 | 22669 | 23696 | 23260 | 22953 |
| Total DPPs ${ }^{\ddagger}$ | 2007 | 38070 | 7353 | 8225 | 8349 | 7584 | 6558 |
| \% of expected deaths prevented | 2007 | 33.9 | 37.4 | 36.3 | 35.2 | 32.6 | 28.6 |

Table B: Demographic data 2000 and 2007 by sex and deprivation quintiles
${ }^{\dagger}$ Annual \% fall $\left.=(\mathbf{1 - ( 2 0 0 7} \text { rate/2000 rate })^{\wedge}(\mathbf{1} / 7)\right)$
${ }^{\dagger}$ Expected deaths $=$ CHD deaths expected in 2007 had 2000 CHD rates remained.
${ }^{\text {\# DPP }}$ DPPs, deaths prevented or postponed. DPPs = expected - observed deaths in 2007

### 3.3. Data sources for treatment uptake levels

## Primary prevention therapies:

Lipid-lowering drugs Prevalence of never having had angina or heart attack and currently taking lipid lowering drugs prescribed by a doctor from the Health Survey for England (HSfE 1998, 2003, and 2006).

| Anti-hypertensive medication | Prevalence of never having had angina or heart attack and currently taking <br> medication specifically prescribed to treat high blood pressure from the Health |
| :--- | :--- |
|  | Survey for England (HSfE 1998, 2003, and 2006). | Survey for England (HSfE 1998, 2003, and 2006).

Table C: Data sources for treatment uptake levels

### 3.4. Treatment uptake in 2000 and 2007

|  | England |  |  | IMDQ1 |  |  | IMDQ2 |  |  | IMDQ3 |  |  | IMDQ4 |  |  | IMDQ5 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | N | Uptake (\%) |  | N | Uptake (\%) |  | N | Uptake (\%) |  | N | Uptake (\%) |  | $\mathbf{N}$ | Uptake (\%) |  | N | Uptake (\%) |  |
|  |  | 2000 | 2007 |  | 2000 | 2007 |  | 2000 | 2007 |  | 2000 | 2007 |  | 2000 | 2007 |  | 2000 | 2007 |
| Anti-hypertension | 35,280,843 | 8.3 | 13.5 | 7,328,217 | 8.3 | 14.0 | $\begin{aligned} & \hline 7,362, \\ & 561 \end{aligned}$ | 8.2 | 13.8 | $\begin{aligned} & 7,232,7 \\ & 79 \end{aligned}$ | 8.6 | 13.9 | 6,905,987 | 8.2 | 13.0 | 6,451,299 | 8.3 | 12.7 |
| Statins | 35,280,843 | 1.1 | 9.0 | 7,328,217 | 1.0 | 7.9 | $\begin{aligned} & 7,362, \\ & 561 \end{aligned}$ | 1.1 | 8.5 | $\begin{aligned} & 7,232,7 \\ & 79 \end{aligned}$ | 1.1 | 9.1 | 6,905,987 | 1.4 | 10.3 | 6,451,299 | 1.3 | 9.1 |

Table D: Treatment uptake in 2000 and 2007
$\dagger \dagger$ We assumed no change in community-based CPR between 2000 and 2007

### 3.5. Clinical efficacy of interventions: relative risk reductions obtained from meta-analyses, and randomised clinical trials

| Treatments | Relative reduction ${ }^{\dagger}$ | risk | Comments | Source paper: First author (year), notes |
| :---: | :---: | :---: | :---: | :---: |
| Primary prevention therapies: |  |  |  |  |
| Treatments for high blood pressure | $\begin{aligned} & 13 \% \quad(95 \% \\ & 6,19) \end{aligned}$ | CI: | $\mathrm{OR}=0.87 \quad$ (95\% CI: $0.81,0.94$ ); RRR=13\% (95\% CI: 6,19) in those with high blood pressure without disease at entry. [RRR=29\% (95\% CI: 17,37 ) those with average blood pressure and CHD, treated with ACE inhibitors] | Law (2003) [51] |
| Statins | $\begin{aligned} & 35 \% \quad(95 \% \\ & 11,52) \end{aligned}$ |  | $\mathrm{OR}=0.65 \quad(95 \% \quad \mathrm{CI}: \quad 0.48,0.89)$; RRR=35\% (95\% CI: 11,52) for CHD mortality (only trials using statins), Figure 3 on page 4 | Pignone (2000) [52] |

Table E; Relatives risk reductions used in the model
${ }^{\dagger}$ Relative risk reduction (RRR) calculated as $\mathbf{1}$ - odds ratio

### 3.6. Case fatality rates for each patient group

| Patient <br> group | Hypertension |  | Statins |  |
| :--- | :--- | :--- | :--- | :--- |
|  | Men | Women | Men | Women |
| $\mathbf{2 5 - 3 4}$ | 0.000 | 0.000 | 0.000 | 0.000 |
| $\mathbf{3 5 - 4 4}$ | 0.001 | 0.001 | 0.001 | 0.001 |
| $\mathbf{4 5 - 5 4}$ | 0.002 | 0.002 | 0.002 | 0.002 |
| $\mathbf{5 5 - 6 4}$ | 0.006 | 0.004 | 0.006 | 0.004 |
| $\mathbf{6 5 - 7 4}$ | 0.014 | 0.014 | 0.014 | 0.014 |
| $\mathbf{7 5 - 8 4}$ | 0.035 | 0.035 | 0.035 | 0.035 |
| $\mathbf{8 5 +}$ | 0.094 | 0.094 | 0.094 | 0.094 |

Table F: Case-fatality rates. Source Wijeysundera et.al (2010) [5]

### 3.7. Risk factors: variable definitions and source

The Health Survey for England (HSfE), an annual nationwide household survey of the English population, has been described in detail elsewhere [24]. Briefly, members of a stratified random sample (drawn from the Postcode Address File) that is socio-demographically representative of the English population were invited to participate. The annual household response rate was $75 \%$ in 2000 , falling steadily to $66 \%$ in 2007. Data were collected at two visits: an interviewer's visit, during which a questionnaire was administered, followed by a visit from a trained nurse for all those interviewed who agreed. The nurse visit, which did not take place in 2004 among the general population sample, includes measurements and collection of blood, as well as additional questioning including use of prescribed medication (1998, 2003, and 2006).

| Risk factor | HSfE survey <br> years | Description |
| :--- | :--- | ---: | :--- |
| SBP (mmHg) | All <br> between 2000-7 <br> except 2004 | Calculated as the mean of the 2 ${ }^{\text {rd }}$ and <br> eaten, consumed alcohol or smoked in <br> the 30 minutes prior to measurement. |
| Those reporting taking blood pressure |  |  |
| lowering drugs were included |  |  |

Table G: Definition of risk fators

### 3.8. Risk factor levels in $\mathbf{2 0 0 0}$ and $\mathbf{2 0 0 7}$ by sex and deprivation quintiles

The annual sample size of the Health Survey for England (HSE), roughly 14,000 adults aged 16 years and over, was not large enough to provide accurate and precise estimates of risk factor levels, and hence rates of change over time by age, sex, and deprivation quintiles. We considered a 'fixed gradient approach' for estimating risk factors changes.

The fixed gradient approach is based on the assumption that changes in pace and direction for each deprivation quintile were similar and therefore, most accurately measured by the overall national rates of change (across all age-sex groups). If this assumption holds, then relatively stable and plausible estimates for each quintile could be derived by scaling the national age-sex risk factor levels up or down using a fixed ratio/gradient.

The fixed gradient was derived by pooling together survey data for all available years from 2000 to 2007 to calculate risk factor estimates by age, sex, and deprivation quintiles. Then the pooled national estimate for 14 age-by-sex groups was set notionally to one, and the corresponding estimates for each deprivation quintile re-indexed to be below or above one (i.e. expressing the ratio of the deprivation quintile to national estimate). These index rates were then applied to the single year national estimates to derive the corresponding risk factor levels for that year. The fixed gradient was applied to both the start and end years of the model. The next table shows the risk factor levels in 2000 and 2007 by gender and deprivation quintiles using this approach.

|  | England |  | IMDQ1 |  | IMDQ2 |  | IMDQ3 |  |  | IMDQ4 |  | IMDQ5 |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :---: |
|  | $\mathbf{2 0 0 0}$ | $\mathbf{2 0 0 7}$ | $\mathbf{2 0 0 0}$ | $\mathbf{2 0 0 7}$ | $\mathbf{2 0 0 0}$ | $\mathbf{2 0 0 7}$ | $\mathbf{2 0 0 0}$ | $\mathbf{2 0 0 7}$ | $\mathbf{2 0 0 0}$ | $\mathbf{2 0 0 7}$ | $\mathbf{2 0 0 0}$ | $\mathbf{2 0 0 7}$ |  |
| Systolic blood pressure, $\mathbf{~ m m H g}$ |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Male | 133.1 | 130.6 | 133.1 | 130.5 | 133.4 | 130.8 | 133.3 | 130.7 | 133.0 | 130.6 | 133.0 | 130.6 |  |
| Female | 131.0 | 125.6 | 130.7 | 125.3 | 131.6 | 126.6 | 131.2 | 125.7 | 131.1 | 125.6 | 130.6 | 125.1 |  |
| Cholesterol, |  |  |  |  |  |  |  |  |  |  |  |  |  |
| mmol/L |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Male | 5.6 | 5.4 | 5.6 | 5.4 | 5.6 | 5.5 | 5.6 | 5.4 | 5.5 | 5.4 | 5.5 | 5.4 |  |
| Female | 5.7 | 5.5 | 5.7 | 5.6 | 5.8 | 5.6 | 5.7 | 5.5 | 5.6 | 5.4 | 5.6 | 5.5 |  |

Table H: Risk factor levels in 2000 and 2007

### 3.9. Beta coefficients for risk factors

Estimated $\beta$ coefficients from multiple regression analyses for the relationship between absolute changes in population mean risk factors and percentage changes in coronary heart disease mortality for men and women, stratified by age. Data sources, values and comments.

| Systolic blood pressure | Age group (years) |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 25-44 | 45-54 | 55-64 | 65-74 | 75+ |
| Men (hazard ratio per 20 mmHg ) | 0.49 | 0.49 | 0.52 | 0.58 | 0.65 |
| Men (log hazard ratio per 1 mmHg ) | -0.036 | -0.035 | -0.032 | -0.027 | -0.021 |
| Minimum | -0.029 | -0.028 | -0.026 | -0.022 | -0.017 |
| Maximum | -0.043 | -0.042 | -0.039 | -0.032 | -0.025 |
| Women (hazard ratio per 20 mmHg ) | 0.40 | 0.40 | 0.49 | 0.52 | 0.59 |
| Women (log hazard ratio per 1 mmHg ) | -0.046 | -0.046 | -0.035 | -0.032 | -0.026 |
| Minimum | -0.037 | -0.037 | -0.028 | -0.026 | -0.021 |
| Maximum | -0.055 | -0.055 | -0.042 | -0.039 | -0.031 |

Source: Prospective studies collaborative meta-analysis, Lancet 2002 [53]
Units: Percentage change in CHD mortality per 20 mmHg change in systolic blood pressure
Strengths: Large dataset, includes US data, adjusted for regression dilution bias, consistent with randomised controlled trials, results stratified by age and sex, with $95 \%$ confidence intervals
Limitations: Some publication bias still possible
Table I: Beta coefficients for SBP.
${ }^{\dagger}$ Risk reduction $=1 \mathbf{-}$ hazard ratio

| Cholesterol Age groups (years) |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 25-44 | 45-54 | 55-64 | 65-74 | 75-84 | 85+ |
| Mortality reduction per $1 \mathrm{mmol} / \mathrm{l}$ |  |  |  |  |  |  |
| Men | 0.55 | 0.53 | 0.36 | 0.21 | 0.21 | 0.21 |
| Women | 0.57 | 0.52 | 0.35 | 0.23 | 0.23 | 0.23 |
| Log coefficient |  |  |  |  |  |  |
| Men | -0.799 | -0.755 | -0.446 | -0.236 | -0.117 | -0.083 |
| Minimum | -0.639 | -0.604 | -0.357 | -0.189 | -0.093 | -0.067 |
| Maximum | -0.958 | -0.906 | -0.536 | -0.283 | -0.140 | -0.100 |
| Women | -0.844 | -0.734 | -0.431 | -0.261 | -0.174 | -0.051 |
| Minimum | -0.675 | -0.587 | -0.345 | -0.209 | -0.139 | -0.041 |
| Махітит | -1.013 | -0.881 | -0.517 | -0.314 | -0.209 | -0.062 |

Source: Prospective studies collaborative meta-analysis, Lancet 2007 [54]
Units: Percentage change in CHD mortality per $1 \mathrm{mmol} / \mathrm{l}$ change in total cholesterol
Strengths: Includes US data, adjusted for regression dilution bias, includes randomised controlled trials, RCT values consistent with observational data, results stratified by age and sex, with $95 \%$ confidence intervals
Limitations: $\quad$ Some publication bias still possible
Table J: Beta coefficients for cholesterol
${ }^{\dagger}$ Risk reduction $=1 \mathbf{-}$ hazard ratio

### 3.10. Cumulative benefit: Adjustment factors by age, sex and IMD quintile

In Section 1.2 we described how we adjusted down the DPPs calculated in an additive fashion over the risk factors by using the ratio of cumulative to additive risk-reduction. The 70 age-sex-IMD specific adjustment factors are shown below.

|  |  | Deprivation quintile |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | IMDQ1 | IMDQ2 | IMDQ3 | IMDQ4 | IMDQ5 | England |  |
| $\mathbf{2 5 - 3 4}$ | 0.9464 | 0.9449 | 0.9463 | 0.9462 | 0.9434 | 0.9453 |  |
| $\mathbf{3 5 - 4 4}$ | 0.9196 | 0.9169 | 0.9179 | 0.9126 | 0.9110 | 0.9153 |  |
| $\mathbf{4 5 - 5 4}$ | 0.9335 | 0.9278 | 0.9205 | 0.9193 | 0.9083 | 0.9219 |  |
| $\mathbf{5 5 - 6 4}$ | 0.8957 | 0.8957 | 0.8883 | 0.8851 | 0.8762 | 0.8886 |  |
| $\mathbf{6 5 - 7 4}$ | 0.8885 | 0.8843 | 0.8846 | 0.8817 | 0.8720 | 0.8827 |  |
| $\mathbf{7 5 - 8 4}$ | 0.9182 | 0.9146 | 0.9134 | 0.9214 | 0.9149 | 0.9162 |  |
| $\mathbf{8 5 +}$ | 0.9561 | 0.9569 | 0.9525 | 0.9520 | 0.9582 | 0.9547 |  |
|  |  |  | Women |  |  |  |  |
| $\mathbf{2 5 - 3 4}$ | 0.8799 | 0.8872 | 0.8846 | 0.8787 | 0.8782 | 0.8809 |  |
| $\mathbf{3 5 - 4 4}$ | 0.9148 | 0.9119 | 0.9014 | 0.9034 | 0.8892 | 0.9038 |  |
| $\mathbf{4 5 - 5 4}$ | 0.9038 | 0.9013 | 0.8937 | 0.8777 | 0.8546 | 0.8865 |  |
| $\mathbf{5 5 - 6 4}$ | 0.8862 | 0.8896 | 0.8842 | 0.8703 | 0.8560 | 0.8780 |  |
| $\mathbf{6 5 - 7 4}$ | 0.8620 | 0.8569 | 0.8523 | 0.8363 | 0.8307 | 0.8479 |  |
| $\mathbf{7 5 - 8 4}$ | 0.8803 | 0.8869 | 0.8824 | 0.8778 | 0.8622 | 0.8779 |  |
| $\mathbf{8 5 +}$ | 0.9394 | 0.9399 | 0.9409 | 0.9463 | 0.9386 | 0.9410 |  |
| Overall | 0.9089 | 0.9082 | 0.9045 | 0.9006 | 0.8924 | 0.9029 |  |

Table K: Adjustment factors by age, sex and IMD quintile

### 3.11. Uncertainty analysis: parameter distributions, functions and sources

Table M records the type of distribution and associated functions for each of the input variables in the IMPACT sec $_{\text {model }}$. We implemented stochastic uncertainty analysis in Excel using Ersatz (version 1.0 available at http://www.epigear.com), an add-in that allows probabilistic bootstrapping in Excel [62]. Ersatz allows repeated random draws from specified distributions for input variables that are used to recalculate iteratively the model. It then calculates the $95 \%$ uncertainty intervals from the realised values of the output variable (deaths prevented or postponed). For the $\mathrm{IMPACT}_{\text {SEC }}$ model, we calculated the uncertainty intervals based on 1000 draws taking the $95 \%$ uncertainty intervals as the $2.5^{\text {th }}$ and $97.5^{\text {th }}$ percentiles. Input variables taken from external sources (e.g. case fatality rates, beta coefficients and relative risk reductions) were randomly drawn from specified distributions but assumed constant across deprivation quintiles.

| Input parameters | Type of distribution and functions (Mean, Standard error) | Source |
| :---: | :---: | :---: |
| Population |  |  |
| Population counts and CHD deaths stratified by age, sex, and Index of Multiple Deprivation quintiles | - Population counts (no error) <br> - Deaths expected in 2007 had CHD mortality rates in 2000 persisted (Poisson distribution) | Office for National Statistics |
| Risk factors |  |  |
| Prevalence/mean estimates (pooled data; national estimates for 2000 and 2007) | - Continuous variables (Body Mass Index, SBP, total cholesterol, fruit and vegetable consumption): (Normal distribution: mean, SE of mean) | Health Survey for England |
| Beta coefficient: SBP | Normal distribution (mean, SE of mean): <br> $\mathrm{M}<45(-0.036,0.004)$; $\mathrm{M} 45-54(-0.035,0.004)$ <br> M 55-64 (-0.032,0.003); M 65-74 (-0.027,0.003) <br> M 75-84 (-0.021,0.002); M 85+ (-0.016,0.002) <br> F < $55(-0.046,0.005)$; F 55-64 (-0.035,0.004) <br> F 65-74 (-0.032,0.003); F 75-84 (-0.026,0.003) | Prospective studies collaborative meta-analysis (2002) [53]. Parameters on the log scale. |


|  | F 85+ (-0.019,0.002) |  |
| :---: | :---: | :---: |
| Beta coefficient: total cholesterol | Normal distribution (mean, SE of mean): <br> M < 45 (-0.799,0.081); M 45-54 (-0.755,0.077) <br> M 55-64 (-0.446,0.046); M 65-74 (-0.236,0.024) <br> M 75-84 (-0.117,0.012); M 85+ (-0.083,0.009) <br> F < 45 (-0.844,0.086); F 45-54 (-0.734,0.075) <br> F 55-64 (-0.431,0.044); F 65-74 (-0.261,0.027) <br> F 75-84 (-0.174,0.018); F 85+ (-0.051,0.005) | Prospective studies collaborative meta-analysis (2007) [54]. Parameters on the log-scale. |
| Aspirin <br> Beta blockers <br> ACE Inhibitors <br> Statins <br> Rehabilitation <br> Warfarin | M \& F (0.15,0.139) <br> M \& F $(0.23,0.185)$ <br> M \& F $(0.20,0.177)$ <br> M \& F $(0.24,0.245)$ <br> M \& F $(0.26,0.347)$ <br> M \& F $(0.22,0.305)$ | ATC (2002) [35] Freemantle (1999) [29] Flather (2000) [40] Hulten (2006) [41] Taylor (2004) [43] Anand and Yusuf (1999) [42] |
| Primary prevention therapies: Statins |  |  |
| Eligible patients: <br> Population | Population counts (no error) | Office for National Statistics |
| Treatment uptake | \% never having had angina or heart attack and currently taking lipid lowering drugs prescribed by a doctor: (Beta distribution: cases, sample-size minus cases) | Health Survey for England |
| Case fatality rate | Sample size ( $n$ ) = never having had angina or heart attack and currently taking lipid lowering drugs in 2006: <br> Beta distribution (cases $=n \times$ CFR estimate, non-cases $=$ $n$ - cases) | Wijeysundera et al (2010) [5] |
| Compliance | Beta distribution (cases $=n \times$ assumed compliance, noncases $=n$ - cases) | Health Survey for England |
| Relative risk reduction: Statins | Ersatz RR function (RRR, SE $\ln (R R R)$ ): <br> M \& F $(0.35,0.396)$ | Pignone (2000) [52] |
| Primary prevention therapies: Treatments for high blood pressure |  |  |
| Eligible patients: | Population counts (no error) | Office for National Statistics |


| Population |  |  |
| :--- | :--- | :--- |
| Treatment uptake | \% never having had angina or heart attack and currently <br> taking medication specifically prescribed to treat high <br> blood pressure: (Beta distribution: cases, sample-size <br> minus cases) | Health Survey for England |
| Case fatality rate | Sample size ( $n$ ) = never having had angina or heart attack <br> and currently taking medication to lower blood pressure <br> in 2006: <br> Beta distribution (cases $=n \times$ CFR estimate, non-cases $=$ <br> $n-$ cases $)$ | Wijeysundera et al (2010) [5] |
| Compliance | Beta distribution (cases $=n \times$ assumed compliance, non- <br> cases $=n-$ cases $)$ | Health Survey for England |
| Relative risk reduction: <br> Treatments for high blood <br> pressure | Ersatz RR function (RRR, SE ln(RRR)): <br> M \& F (0.13,0.294) | Law (2003) [51] |

Table L: Parameter distributions, functions and sources

## 4 Tables

| DPPS through changes in the population |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | England | $\begin{gathered} \text { IMD } \\ \text { quintile } 1 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 2 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 3 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 4 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 5 \end{gathered}$ |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 13253 | 2093 | 2666 | 2608 | 2742 | 3143 |
| 95\% LL | 8495 | 1187 | 1632 | 1577 | 1775 | 2302 |
| 95\% UL | 17371 | 2880 | 3551 | 3497 | 3590 | 3880 |
| DPPS through changes in the treatments uptakes |  |  |  |  |  |  |
|  | England | $\begin{gathered} \text { IMD } \\ \text { quintile } 1 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 2 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 3 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 4 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 5 \end{gathered}$ |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 7098 | 1411 | 1386 | 1713 | 1539 | 1049 |
| 95\% LL | 3479 | 656 | 665 | 800 | 716 | 500 |
| 95\% UL | 14195 | 3069 | 2811 | 3819 | 3141 | 2135 |

Table M: CHD deaths prevented or postponed through changes in population and treatment uptakes between 2000 and 2007 in England, stratified by deprivation quintiles.

| DPPS through SBP reduction |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Overall |  |  |  |  |  |  |
|  | England | $\begin{gathered} \text { IMD } \\ \text { quintile } 1 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 2 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 3 \end{gathered}$ | IMD quintile 4 | IMD <br> quintile 5 <br> Deprived |
|  |  | Affluent |  |  |  |  |
| Mean | 12960 | 2218 | 2579 | 2729 | 2736 | 2698 |
| 95\% LL | 8181 | 1295 | 1537 | 1690 | 1776 | 1868 |
| 95\% UL | 17463 | 3086 | 3560 | 3723 | 3649 | 3468 |
| Population wide changes |  |  |  |  |  |  |
|  | England | $\begin{gathered} \text { IMD } \\ \text { quintile } 1 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 2 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 3 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 4 \end{gathered}$ | IMD <br> quintile 5 <br> Deprived |
|  |  | Affluent |  |  |  |  |
| Mean | 11162 | 1861 | 2168 | 2321 | 2391 | 2421 |
| 95\% LL | 6500 | 978 | 1156 | 1322 | 1439 | 1612 |
| 95\% UL | 15093 | 2616 | 3024 | 3163 | 3190 | 3121 |
| Anti-hypertension treatment |  |  |  |  |  |  |


|  | England | IMD <br> quintile 1 | IMD <br> quintile 2 | IMD <br> quintile 3 | IMD <br> quintile 4 | IMD <br> quintile 5 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 1798 | 357 | 411 | 408 | 345 | 277 |
| $95 \%$ LL | 675 | 138 | 151 | 150 | 126 | 105 |
| $95 \%$ UL | 3860 | 784 | 907 | 898 | 780 | 606 |

Table N: CHD DPPs through medication and population changes in SBP between 2000 and 2007 in England, stratified by deprivation quintiles

| DPPS through Cholesterol reduction |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Overall |  |  |  |  |  |  |
|  | England | IMD quintile 1 | $\begin{gathered} \text { IMD } \\ \text { quintile } 2 \end{gathered}$ | IMD <br> quintile 3 | IMD quintile 4 | IMD <br> quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 7391 | 1286 | 1473 | 1592 | 1545 | 1494 |
| 95\% LL | 3851 | 551 | 794 | 700 | 725 | 930 |
| 95\% UL | 14493 | 2900 | 2819 | 3669 | 3161 | 2579 |
| Population wide changes |  |  |  |  |  |  |
|  | England | $\begin{gathered} \text { IMD } \\ \text { quintile } 1 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 2 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 3 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 4 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 5 \end{gathered}$ |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 2091 | 232 | 498 | 287 | 351 | 722 |
| 95\% LL | 1020 | 43 | 282 | 56 | 129 | 496 |
| 95\% UL | 3148 | 419 | 709 | 516 | 572 | 944 |
| Dyslipaedimia treatment |  |  |  |  |  |  |
|  | England | $\begin{gathered} \text { IMD } \\ \text { quintile } 1 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 2 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 3 \end{gathered}$ | $\begin{gathered} \text { IMD } \\ \text { quintile } 4 \end{gathered}$ | IMD quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 5300 | 1054 | 975 | 1305 | 1194 | 772 |
| 95\% LL | 2051 | 375 | 359 | 480 | 443 | 279 |
| 95\% UL | 12318 | 2679 | 2326 | 3369 | 2804 | 1869 |

Table O: CHD deaths prevented or postponed through medication and population changes in cholesterol between 2000 and 2007 in England, stratified by deprivation quintiles

## 5 Tables by gender

### 5.1. Men

| DPPS through changes in the population |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | England | IMD <br> quintile 1 | IMD <br> quintile 2 | IMD <br> quintile 3 | IMD <br> quintile 4 | IMD <br> quintile 5 |
|  |  |  | Deprived |  |  |  |
| Mean | 5872 | 1041 | 1015 | 1121 | 1247 | 1449 |
| $95 \%$ LL | 3029 | 495 | 411 | 510 | 675 | 912 |
| $95 \%$ UL | 8593 | 1557 | 1591 | 1709 | 1785 | 1960 |

DPPS through changes in the treatments uptakes

|  | England | IMD <br> quintile 1 | IMD <br> quintile 2 | IMD <br> quintile 3 | IMD <br> quintile 4 | IMD <br> quintile 5 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  |  | Deprived |  |  |  |
| Mean | 3017 | 474 | 763 | 751 | 596 | 434 |
| $95 \%$ LL | 1211 | 187 | 291 | 261 | 218 | 157 |
| $95 \%$ UL | 7005 | 1017 | 1867 | 2144 | 1470 | 1028 |

Table P: CHD deaths prevented or postponed through medication and population changes in SBP and Cholesterol between 2000 and 2007 in England, stratified by deprivation quintiles

| DPPS through SBP reduction |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Overall |  |  |  |  |  |  |
|  | England | IMD quintile 1 | IMD quintile 2 | IMD quintile 3 | IMD quintile 4 | IMD <br> quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 4812 | 806 | 941 | 996 | 1014 | 1054 |
| 95\% LL | 2011 | 265 | 320 | 390 | 463 | 540 |
| 95\% UL | 7625 | 1356 | 1573 | 1598 | 1557 | 1549 |
| Population wide changes |  |  |  |  |  |  |
|  | England | IMD quintile 1 | IMD quintile 2 | IMD quintile 3 | IMD <br> quintile 4 | IMD quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 4106 | 659 | 745 | 850 | 898 | 954 |
| 95\% LL | 1416 | 138 | 168 | 269 | 365 | 456 |
| 95\% UL | 6713 | 1165 | 1304 | 1414 | 1419 | 1442 |
| Anti-hypertension treatment |  |  |  |  |  |  |
|  | England | IMD <br> quintile 1 | IMD quintile 2 | IMD quintile 3 | IMD quintile 4 | IMD <br> quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 705 | 147 | 196 | 146 | 116 | 100 |
| 95\% LL | 198 | 45 | 46 | 39 | 31 | 30 |
| 95\% UL | 1808 | 370 | 528 | 386 | 312 | 247 |

Table Q: CHD deaths prevented or postponed through medication and population changes in SBP between 2000 and 2007 in England, stratified by deprivation quintiles

| DPPS through Cholesterol reduction |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Overall |  |  |  |  |  |  |
|  | England | IMD quintile 1 | IMD quintile 2 | IMD quintile 3 | IMD quintile 4 | IMD quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 4078 | 709 | 836 | 875 | 829 | 829 |
| 95\% LL | 2150 | 400 | 365 | 371 | 414 | 498 |
| 95\% UL | 8149 | 1246 | 1905 | 2242 | 1681 | 1407 |
| Population wide changes |  |  |  |  |  |  |
|  | England | IMD quintile 1 | IMD quintile 2 | $\begin{aligned} & \text { IMD } \\ & \text { quintile } 3 \end{aligned}$ | IMD quintile 4 | IMD quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 1766 | 381 | 270 | 271 | 349 | 495 |
| 95\% LL | 916 | 234 | 99 | 88 | 175 | 311 |
| 95\% UL | 2615 | 535 | 442 | 450 | 521 | 675 |
| Statins treatment |  |  |  |  |  |  |
|  | England | IMD quintile 1 | IMD quintile 2 | IMD quintile 3 | $\begin{aligned} & \text { IMD } \\ & \text { quintile } 4 \end{aligned}$ | IMD quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 2312 | 327 | 566 | 605 | 480 | 334 |
| 95\% LL | 684 | 85 | 155 | 159 | 130 | 83 |
| 95\% UL | 6184 | 861 | 1648 | 1992 | 1351 | 912 |

Table R: CHD deaths prevented or postponed through medication and population changes in cholesterol between 2000 and 2007 in England, stratified by deprivation quintiles

### 5.2. Women

| DPPS through changes in the population |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | England | IMD <br> quintile 1 | IMD <br> quintile 2 | IMD <br> quintile 3 | IMD <br> quintile 4 | IMD <br> quintile 5 |
|  |  |  | Deprived |  |  |  |
| Mean | 7380 | 1053 | 1652 | 1487 | 1495 | 1694 |
| $95 \%$ LL | 3673 | 341 | 834 | 682 | 730 | 1062 |
| 95\% UL | 10669 | 1679 | 2370 | 2197 | 2175 | 2264 |

DPPS through changes in the treatments uptakes

|  | England | IMD <br> quintile 1 | IMD <br> quintile 2 | IMD <br> quintile 3 | IMD <br> quintile 4 | IMD <br> quintile 5 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  |  | Deprived |  |  |  |
| Mean | 4081 | 937 | 623 | 962 | 944 | 615 |
| 95\% LL | 1692 | 342 | 261 | 383 | 365 | 246 |
| $95 \%$ UL | 8916 | 2402 | 1357 | 2250 | 2112 | 1494 |

Table S: CHD deaths prevented or postponed through medication and population changes in SBP and Cholesterol between 2000 and 2007 in England, stratified by deprivation quintiles

| DPPS through SBP reduction |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Overall |  |  |  |  |  |  |
|  | England | IMD quintile 1 | IMD quintile 2 | IMD quintile 3 | IMD quintile 4 | IMD quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 8149 | 1412 | 1638 | 1733 | 1722 | 1644 |
| 95\% LL | 4422 | 696 | 822 | 917 | 955 | 1011 |
| 95\% UL | 11540 | 2064 | 2366 | 2475 | 2420 | 2218 |
| Population wide changes |  |  |  |  |  |  |
|  | England | IMD quintile 1 | IMD quintile 2 | IMD quintile 3 | IMD <br> quintile 4 | IMD quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 7056 | 1202 | 1424 | 1471 | 1492 | 1467 |
| 95\% LL | 3446 | 513 | 628 | 701 | 745 | 854 |
| 95\% UL | 10329 | 1816 | 2136 | 2176 | 2161 | 2018 |
| Anti-hypertension treatment |  |  |  |  |  |  |
|  | England | IMD <br> quintile 1 | IMD quintile 2 | IMD quintile 3 | IMD quintile 4 | IMD quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 1093 | 210 | 215 | 262 | 229 | 177 |
| 95\% LL | 319 | 63 | 64 | 75 | 65 | 53 |
| 95\% UL | 2624 | 510 | 520 | 641 | 575 | 433 |

Table T: CHD deaths prevented or postponed through medication and population changes in SBP between 2000 and 2007 in England, stratified by deprivation quintiles

| DPPS through Cholesterol reduction |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Overall |  |  |  |  |  |  |
|  | England | IMD quintile 1 | IMD quintile 2 | IMD <br> quintile 3 | IMD <br> quintile 4 | IMD <br> quintile 5 <br> Deprived |
|  |  | Affluent |  |  |  |  |
| Mean | 3313 | 577 | 637 | 717 | 717 | 665 |
| 95\% LL | 1069 | 18 | 298 | 179 | 171 | 304 |
| 95\% UL | 8202 | 2065 | 1335 | 2005 | 1904 | 1562 |
| Population wide changes |  |  |  |  |  |  |
|  | England | IMD quintile 1 | IMD <br> quintile 2 | IMD <br> quintile 3 | IMD quintile 4 | IMD quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 325 | -149 | 228 | 16 | 2 | 227 |
| 95\% LL | -315 | -264 | 99 | -123 | -134 | 97 |
| 95\% UL | 996 | -31 | 364 | 161 | 144 | 365 |
| Statins treatment |  |  |  |  |  |  |
|  | England | IMD quintile 1 | IMD <br> quintile 2 | IMD quintile 3 | IMD quintile 4 | IMD quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 2988 | 727 | 409 | 700 | 714 | 438 |
| 95\% LL | 922 | 190 | 115 | 197 | 199 | 123 |
| 95\% UL | 7822 | 2203 | 1095 | 2009 | 1905 | 1323 |

Table U: CHD deaths prevented or postponed through medication and population changes in cholesterol between 2000 and 2007 in England, stratified by deprivation quintiles

### 5.3. Percentage difference in men relative to women

## DPPS through changes in the population

|  | England | IMD <br> quintile 1 | IMD <br> quintile 2 | IMD <br> quintile 3 | IMD <br> quintile 4 | IMD <br> quintile 5 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  | Affluent |  |  |  | Deprived |
| Mean | $13 \%$ | $-17 \%$ | $33 \%$ | $16 \%$ | $8 \%$ | $11 \%$ |
| $95 \% \mathrm{LL}$ | $-74 \%$ | $-222 \%$ | $-35 \%$ | $-80 \%$ | $-87 \%$ | $-49 \%$ |


| 95\% UL | 61\% | 58\% | 75\% | 67\% | 58\% | 49\% |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| DPPS through changes in the treatments uptakes |  |  |  |  |  |  |
|  | England | IMD quintile 1 | IMD quintile 2 | IMD quintile 3 | IMD quintile 4 | IMD quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 15\% | 40\% | -43\% | 6\% | 26\% | 19\% |
| 95\% LL | -101\% | -49\% | -262\% | -165\% | -91\% | -105\% |
| 95\% UL | 74\% | 85\% | 56\% | 76\% | 80\% | 79\% |

Table V: Percentage difference of DPPs for men relative to women through medication and population changes in SBP and Cholesterol between 2000 and 2007 in England

| DPPS through SBP reduction |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Overall |  |  |  |  |  |  |
|  | England | IMD quintile 1 | IMD quintile 2 | IMD quintile 3 | IMD quintile 4 | IMD quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 36\% | 37\% | 40\% | 38\% | 37\% | 33\% |
| 95\% LL | -24\% | -32\% | -33\% | -25\% | -20\% | -19\% |
| 95\% UL | 75\% | 80\% | 79\% | 77\% | 74\% | 68\% |
| Population wide changes |  |  |  |  |  |  |
|  | England | IMD quintile 1 | IMD quintile 2 | IMD <br> quintile 3 | IMD <br> quintile 4 | IMD quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 34\% | 36\% | 44\% | 37\% | 34\% | 31\% |
| 95\% LL | -38\% | -47\% | -37\% | -41\% | -37\% | -29\% |
| 95\% UL | 80\% | 89\% | 88\% | 81\% | 76\% | 70\% |
| Anti-hypertension treatment |  |  |  |  |  |  |
|  | England | IMD quintile 1 | IMD quintile 2 | IMD quintile 3 | IMD <br> quintile 4 | IMD <br> quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 19\% | 17\% | -15\% | 29\% | 35\% | 32\% |
| 95\% LL | -128\% | -145\% | -243\% | -103\% | -90\% | -96\% |
| 95\% UL | 81\% | 79\% | 77\% | 85\% | 87\% | 84\% |

Table W: Percentage difference of DPPs for men relative to women through medication and population changes in SBP between 2000 and 2007 in England

| DPPS through Cholesterol reduction |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Overall |  |  |  |  |  |  |
|  | England | IMD quintile 1 | IMD quintile 2 | IMD quintile 3 | IMD quintile 4 | IMD quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | -52\% | -458\% | -48\% | -69\% | -57\% | -40\% |
| 95\% LL | -273\% | -1102\% | -246\% | -448\% | -367\% | -180\% |
| 95\% UL | 53\% | 78\% | 49\% | 61\% | 61\% | 48\% |
| Population wide changes |  |  |  |  |  |  |
|  | England | IMD quintile 1 | IMD quintile 2 | IMD <br> quintile 3 | IMD <br> quintile 4 | IMD quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | -25\% | 402\% | -35\% | 1065\% | 680\% | -148\% |
| 95\% LL | -6102\% | 215\% | -211\% | -6121\% | -7144\% | -436\% |
| 95\% UL | 6040\% | 1190\% | 59\% | 5990\% | 6646\% | -15\% |
| Statins treatment |  |  |  |  |  |  |
|  | England | IMD quintile 1 | IMD quintile 2 | IMD quintile 3 | IMD quintile 4 | IMD quintile 5 |
|  |  | Affluent |  |  |  | Deprived |
| Mean | 5\% | 42\% | -80\% | -15\% | 14\% | 4\% |
| 95\% LL | -173\% | -84\% | -465\% | -288\% | -160\% | -184\% |
| 95\% UL | 79\% | 91\% | 65\% | 79\% | 84\% | 84\% |

Table X: Percentage difference of DPPs for men relative to women through medication and population changes in cholesterol between 2000 and 2007 in England

## 6 Appendix reference list

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## Falls in blood pressure and cholesterol have saved 20,000+ lives in England

## Impact of statins greatest among most affluent but drugs only accounted for $14 \%$ of total fall in deaths

[The contribution of primary prevention medication and dietary change in coronary mortality reduction in England between 2000 and 2007: a modelling study doi10.1136/bmjopen-2014-006070]

Falls in blood pressure and total cholesterol staved off more than 20,000 deaths from coronary heart disease in England between 2000 and 2007, shows a mathematical analysis published in the online journal BMJ Open.

The impact of statins was greatest among the most affluent in the population, suggesting that these drugs have helped maintain health inequalities between rich and poor, say the researchers.

The researchers wanted to quantify the contributions made by drug treatment (primary prevention) and changes in population risk factors (blood pressure and total cholesterol) to the falling rates of coronary heart disease deaths, stratified by socioeconomic background.

They used trial data, analyses of published evidence, national surveys, and official statistics to calculate the number of deaths postponed or prevented across the population of England.

The analysis showed that between 2000 and 2007 deaths from coronary heart disease fell by 38,000 , of which 20,400 lives were saved as a direct result of reductions in blood pressure and total cholesterol.

In absolute terms, a higher proportion of lives were saved among the least affluent sectors of the population, which is to be expected given their much higher prevalence of risk factors, say the researchers.

The substantial fall in blood pressure accounted for well over half of the total, the calculations indicated, with around 13,000 deaths prevented or postponed.

But only a small proportion (1800) of these were attributable to drug treatment, with the rest accounted for by changes in risk factors at the population level.

Falls in blood pressure prevented almost twice as many deaths among the population's poorest as among the richest.

Falls in total cholesterol accounted for some 7400 deaths prevented or postponed, of which ( 5300 or $14 \%$ of the total) were attributable to statins, with the remainder attributable to changes in risk factors at the population level.

Statins prevented almost $50 \%$ more deaths among the richest compared with the poorest, whereas changes at the population level prevented three times as many deaths among the poorest as among the richest.

The researchers were not able to account for 14\% of the total fall in coronary heart disease deaths between 2000 and 2007 (17,600 lives saved). These might be attributable to other risk factors for heart disease, such as stress, they suggest.

They conclude that population-wide approaches, focusing on prevention, such as public health initiatives to curb salt and trans fat levels in processed and take-away foods may have more of an impact than prescribing drugs to individuals.
"Targeting high-risk individuals with medication appears less effective and may also widen socioeconomic inequalities in [coronary heart disease] mortality," they write.
"Any intervention that requires people to mobilise their own resources (material and psychological) will understandably favour those who have greater resources, and thus widen social inequalities," they add.

When healthcare budgets are stretched, as now, preventive approaches are a better way to get results, they suggest.

